# Sta 205 : Home Work #6

### Due: 2009 Mar 04

## 1. Fubini and Tonelli.

(a) Let X be a positive random variable (i.e,  $X \ge 0$  a.s) on  $(\Omega, \mathcal{F}, \mathsf{P})$ . Show that

$$\mathsf{E}(X) = \int_0^\infty \mathsf{P}(X > t) \ dt$$

(note X need not have an absolutely-continuous distribution). Also verify that for any  $\alpha > 0$ 

$$\mathsf{E}(X^{\alpha}) = \int_0^\infty \alpha t^{\alpha - 1} \mathsf{P}(X > t) dt$$

(b) Define probability spaces  $(\Omega_i, \mathcal{B}_i, \mu_i)$ , for i = 1, 2 as follows. Let each  $\Omega_i := (0, 1]$  be the unit interval, with  $\sigma$ -algebras

$$\mathcal{B}_1 =$$
 Borel sets of  $(0,1]$   $\mathcal{B}_2 =$  All subsets of  $(0,1]$ 

and let  $\mu_1$  be Lebesgue measure and  $\mu_2$  counting measure—so that  $\mu_1(A)$  is the length of any set  $A \in \mathcal{B}_1$  and  $\mu_2(A)$  is the number of points in  $A \in \mathcal{B}_2$ . Define

$$f(x,y) := \mathbf{1}_{x=y}(x,y)$$

Set

$$I_1 := \int_{\Omega_1} \left[ \int_{\Omega_2} f(x, y) \mu_2(dy) \right] \mu_1(dx) \qquad I_2 := \int_{\Omega_2} \left[ \int_{\Omega_1} f(x, y) \mu_1(dx) \right] \mu_2(dy)$$

Compute  $I_1$  and  $I_2$ . Is  $I_1 = I_2$ ? Are the measures  $\mu_1$  and  $\mu_2$   $\sigma$ -finite? Why doesn't Fubini's theorem hold here?

(c) This problem is a probabilistic version of the familiar integration-by-parts formula from calculus. Suppose F and G are two distribution functions with no common points of discontinuity in an interval (a, b]. Show that

$$\int_{(a,b]} G(x)F(dx) = F(b)G(b) - F(a)G(a) - \int_{(a,b]} F(x)G(dx)$$

Show that the above formula need not hold true if F and G have common discontinuities.

# 2. Uniform Integrability (UI)

A family  $\{X_{\alpha}\}$  of random variables is called *uniformly integrable* (or UI) if

$$\mathsf{E}\Big[|X_{\alpha}|\,\mathbf{1}_{\{|X_{\alpha}|>t\}}\Big]\to 0$$

as  $t \to \infty$ , uniformly in  $\alpha$ .

- (a) Let  $\{X_n\}$  be a sequence of iid,  $L_1$  random variables. Set  $S_n \equiv \sum_{i=1}^n X_i$ . Show that the sequence of random variables  $\{Y_n\}$  defined by  $Y_n \equiv S_n/n$  is UI.
- (b) Let  $X_n \sim No(0, \sigma_n^2)$ . Find a simple condition on  $\{\sigma_n^2\}$  such that  $\{X_n\}$  is UI.
- (c) If  $\{X_n\}$  and  $\{Y_n\}$  are UI, show that so is  $\{X_n + Y_n\}$ .
- (d) Suppose  $\{X_n, n \ge 1\}$  is an **arbitrary** sequence of non-negative random variables, and set  $M_n \equiv \bigvee_{i=1}^n X_i$ . If  $\{X_n\}$  is UI, show that  $\mathsf{E}(M_n)/n \to 0$ .
- (e) Let  $\phi(x)$  be a function which grows faster than x at infinity, i.e.,  $\phi(x)/x \nearrow \infty$ (monotonically) as  $x \to \infty$ . Let  $\mathcal{C}$  be a collection of random variables such that, for some fixed  $B < \infty$  and all  $Z \in \mathcal{C}$ ,

$$\mathsf{E}\Big(\phi(|Z|)\Big) \le B.$$

Show that C is UI. What is the implication for  $\phi(x) = x^2$ ?

## 3. Convergence Theorems Revisited.

(a) Let X be a non-negative real valued random variable. Show that:

i.  $\lim_{n\to\infty} n\mathsf{E}(\frac{1}{X}\mathbf{1}_{[X>n]}) = 0.$ 

ii. 
$$\lim_{n \to \infty} n^{-1} \mathsf{E}(\frac{1}{X} \mathbf{1}_{[X > n^{-1}]}) = 0.$$

(b) Suppose  $\{p_k, k \ge 0\}$  is a probability mass function on  $\{0, 1, ...\}$  and define the generating function

$$P(z) = \sum_{k=0}^{\infty} p_k z^k \qquad 0 \le z \le 1$$

Prove using Dominated Convergence theorem that

$$\frac{d}{dz}P(z) = \sum_{k=1}^{\infty} p_k k z^{k-1} \qquad 0 \le z \le 1.$$

Note you may wish to consider the cases z < 1 and z = 1 separately. What is P'(1)? P'(0)? Can you express the variance of a random variable X, if it exists, in terms of P(z) and its derivatives? How about  $p_k = \mathsf{P}[X = k]$ ?