

Random Variables

Let Ω be any set, \mathcal{F} any Sigma Field on Ω , and \mathbb{P} any probability measure defined for each element of \mathcal{F} ; such a triple $(\Omega, \mathcal{F}, \mathbb{P})$ is called a *probability space*. Let \mathbb{R} denote the real numbers $(-\infty, \infty)$ and \mathcal{B} the Borel sets on \mathbb{R} generated by (for example) the half-open sets $(a, b]$.

Definition. A real-valued Random Variable is a function $X : \Omega \rightarrow \mathbb{R}$ that is “ $\mathcal{F} \setminus \mathcal{B}$ -measurable,” i.e., that satisfies $X^{-1}(B) = \{\omega : X(\omega) \in B\} \in \mathcal{F}$ for each Borel set $B \in \mathcal{B}$ (or, equivalently, simply for each set B of the form $(-\infty, b]$ for some rational $-\infty < b < \infty$).

This is sometimes denoted simply “ $X^{-1}(\mathcal{B}) \subset \mathcal{F}$.” Since the probability measure \mathbb{P} is only defined on sets $F \in \mathcal{F}$, a random variable *must* satisfy this condition if we are to be able to find the probability $\Pr[X \in B]$ for each Borel set B , or even if we want to find the distribution function (DF) $F_X(b) \equiv \Pr[X \leq b]$ for each rational number b . Note that set-inverses are rather well-behaved functions from one class of sets to another; specifically, for any collection $\{A_\alpha\} \subset \mathcal{B}$,

$$[X^{-1}(A_\alpha)]^c = X^{-1}((A_\alpha)^c) \quad \bigcap_{\alpha} X^{-1}(A_\alpha) = X^{-1}\left(\bigcap_{\alpha} A_\alpha\right) \quad \bigcup_{\alpha} X^{-1}(A_\alpha) = X^{-1}\left(\bigcup_{\alpha} A_\alpha\right)$$

and thus, measurable or not, $X^{-1}(\mathcal{B})$ is a Sigma Field if \mathcal{B} is; it is denoted \mathcal{F}_X (or $\sigma(X)$), called the “sigma field generated by X ,” and is the smallest sigma field \mathcal{G} such that X is $(\mathcal{G} \setminus \mathcal{B})$ -measurable. In particular, X is $(\mathcal{F} \setminus \mathcal{B})$ -measurable if and only if $\sigma(X) \subset \mathcal{F}$.

In probability and statistics, sigma field’s represent *information*: a random variable Y is measurable over \mathcal{F}_X if and only if the value of Y can be found from that of X , i.e., if there exists some function φ such that $Y = \varphi(X)$. Note the difference in perspective between real analysis, on the one hand, and probability/statistics, on the other; in analysis it is only *Lebesgue* measurability that mathematicians worry about, and only to avoid paradoxes and pathologies. In probability and statistics we study measurability for a variety of sigma field’s, and the (technical) concept of measurability corresponds to the (empirical) notion of *observability*.

DISTRIBUTIONS.

A random variable X on a probability space $(\Omega, \mathcal{F}, \mathbb{P})$ induces a measure μ_X on $(\mathbb{R}, \mathcal{B})$, called the *distribution measure* (or simply the *distribution*), via the relation

$$\mu(B) = \mathbb{P}[X \in B],$$

sometimes written more succinctly as $\mu_X = \mathbb{P} \circ X^{-1}$ or even $\mathbb{P}X^{-1}$.

Functions of Random Variables

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space, X a (real-valued) random variable, and $f : \mathbb{R} \rightarrow \mathbb{R}$ a (real-valued $\mathcal{B} \setminus \mathcal{B}$) measurable function. Then $Y = f(X)$ is a random variable, i.e.,

$$Y^{-1}(B) = X^{-1}(f^{-1}(B)) \in \mathcal{F}$$

for any $B \in \mathcal{B}$. Also every continuous or piecewise-continuous real-valued function on \mathbb{R} is $\mathcal{B} \setminus \mathcal{B}$ -measurable.

Random Vectors

Denote by \mathbb{R}^2 the set of points (x, y) in the plane, and by \mathcal{B}^2 the sigma field generated by rectangles of the form $\{(x, y) : a < x \leq b, c < y \leq d\} = (a, b] \times (c, d]$. Note that finite unions of those rectangles form a field \mathcal{F}_0^2 , so the minimal sigma field and minimal λ system containing \mathcal{F}_0^2 coincide, and the assignment $\lambda_0^2((a, b] \times (c, d]) = (b - a) \times (d - c)$ has a unique extension to a measure on all of \mathcal{B}^2 , called two-dimensional Lebesgue measure (and denoted λ^2). Of course, it's just the area of sets in the plane.

A $\mathcal{F} \setminus \mathbb{R}^2$ -measurable mapping $X : \Omega \rightarrow \mathbb{R}^2$ is called a (two-dimensional) *random vector*, or simply an \mathbb{R}^2 -valued random variable, or (a bit ambiguously) an \mathbb{R}^2 -RV. It's easy to show that the components X_1, X_2 of a \mathbb{R}^2 -RV X are each RV's, and conversely that for any two random variables X_1 and X_2 the two-dimensional RV $(X, Y) : \Omega \rightarrow \mathbb{R}^2$ is $\mathcal{F} \setminus \mathbb{R}^2$ -measurable, *i.e.*, is a \mathbb{R}^2 -RV.

Also, any measurable (and in particular, any piecewise-continuous) function $f : \mathbb{R}^2 \rightarrow \mathbb{R}$ induces a random variable $f(X, Y)$: this shows that such combinations as $X + Y, X/Y, X \wedge Y, X \vee Y$, *etc.* are all random variables if X and Y are.

The same ideas work in any finite number of dimensions, so without any special notice we will regard n -tuples (X_1, \dots, X_n) as \mathbb{R}^n -valued RV's, or $\mathcal{F} \setminus \mathcal{B}^n$ -measurable functions, and will use Lebesgue n -dimensional measure λ^n on \mathcal{B}^n . Again $\sum_i X_i, \prod_i X_i, \min_i X_i$, and $\max_i X_i$ are all random variables.

Even if we have *infinitely many* random variables we can verify the measurability of $\sum_i X_i, \inf_i X_i$, and $\sup_i X_i$, and of $\liminf_i X_i$, and $\limsup_i X_i$ as well: for example,

$$[\omega : \sup_i X_i(\omega) \leq r] = \bigcap_{i=1}^{\infty} [\omega : X_i(\omega) \leq r]$$

$$[\omega : \limsup_i X_i(\omega) \leq r] = \bigcup_{i=1}^{\infty} \bigcap_{j=i}^{\infty} [\omega : X_j(\omega) \leq r].$$

The event " X_i converges" is the same as

$$[\omega : \limsup_i X_i(\omega) - \liminf_i X_i(\omega) = 0],$$

and so is \mathcal{F} -measurable and has a well defined probability $P[\limsup_i X_i = \liminf_i X_i]$. This is one point where countable additivity (and not just finite additivity) of P is crucial, and where \mathcal{F} needs to be a sigma field (and not just a field).

Example: Discrete RV's

If an RV X can take on only a finite or countable set of values, say b_i , then each set $\Lambda_i = [\omega : X(\omega) = b_i]$ must be in \mathcal{F} , the Λ_i are disjoint, and X can be represented in the form

$$X(\omega) = \sum_i b_i 1_{\Lambda_i}(\omega), \quad \text{where} \quad (*)$$

$$1_{\Lambda}(\omega) = \begin{cases} 1 & \text{if } \omega \in \Lambda \\ 0 & \text{if } \omega \notin \Lambda \end{cases}$$

is the so-called *indicator function* of Λ . By including a term with $b_i = 0$, if necessary, we can assume that $\Omega = \cup \Lambda_i$ so the $\{\Lambda_i\}$ form a "countable partition" of Ω . Any RV can be approximated as well as we like by a simple RV of the form (*).

EXPLICIT CONSTRUCTION OF SIGMA FIELDS (OMIT ON FIRST READING)**Ordinals and Transfinite Induction**

Every finite set S (say, with $n < \infty$ elements) can be *totally ordered* $a_1 \prec a_2 \prec a_3 \prec \dots$ in $n!$ ways, but in some sense every one of these is the same— if \prec_1 and \prec_2 are two orderings, there exists a 1–1 order-preserving isomorphism $\varphi : (S, \prec_1) \longleftrightarrow (S, \prec_2)$. Thus *up to isomorphism* there is only one ordering for any finite set.

For countably infinite sets there are many different orderings. The obvious one is $a_1 \prec a_2 \prec a_3 \prec \dots$, ordered just like the positive integers \mathbb{N} ; this ordering is called ω , the first *limit ordinal*. But we could pick any element (say, $b_1 \in S$) and order the remainder of S in the usual way, but declare $a_n \prec b_1$ for every $n \in \mathbb{N}$; one element is “bigger” (in the ordering) than all the others. This is *not* isomorphic to ω , and it is called $\omega+1$, the *successor* to ω . If we set aside two elements (say, $b_1 \prec b_2$) to follow all the others we have $\omega+2$, and similarly we have $\omega+n$ for each $n \in \mathbb{N}$. The limit of all these is $\omega+\omega$, or $2\omega\dots$ it is the ordering we would get if we lexicographically ordered the set $\{(i, j) : i = 1, 2, j \in \mathbb{N}\}$ of the first two rows of integers in the first quadrant, declaring $(1, i) \prec (2, j)$ for every i, j and otherwise $(i, j) \prec (i, k)$ if $j < k$.

We would get the successor to this, $2\omega+1$, by extending the lexicographical ordering as we add $(3, 1)$ to S ; in an obvious way we get $2\omega+n$ and eventually the limit ordinals $3\omega, 4\omega, \text{etc.}$, and the successor ordinals $m\omega+n$. The limit of all these is $\omega\omega$ or ω^2 , the lexicographical ordering of the entire first quadrant of integers (i, j) . It too has successors ω^2+n (graphically you can think about integer triplets (i, j, k)), and limits like $\omega^2+\omega$ and ω^3 and ω^ω (which turns out to be the same as 2^ω).

In general an ordinal is a *successor* ordinal if it has a maximal element, and otherwise is a *limit* ordinal. Every ordinal α has a successor $\alpha+1$, and every set of ordinals $\{\alpha_n\}$ has a limit (least upper bound) λ . Let Ω be the first *uncountable* ordinal.

Proofs and constructions by transfinite induction usually have one step at each successor ordinal, and another at each limit ordinal. The *Borel sets* can be defined by transfinite construction as follows. Let \mathcal{F}_1 be any class of subsets of some probability space \mathcal{X} (perhaps \mathcal{F}_1 is the open sets in $\mathcal{X} = \mathbb{R}$, for example).

Succ: For any ordinal α , let $\mathcal{F}_{\alpha+1}$ be the class of countable unions of sets $E_n \in \mathcal{F}_\alpha$ and their complements $E_n^c : E_n^c \in \mathcal{F}_\alpha$.

Lim: For any limit ordinal λ , let $\mathcal{F}_\lambda = \cup_{\alpha < \lambda} \mathcal{F}_\alpha$.

Together these define \mathcal{F}_α for all ordinals, limit and successor; the sigma field *generated by* \mathcal{F}_1 is just \mathcal{F}_Ω . It remains to prove that:

1. $\mathcal{F}_1 \subset \mathcal{F}_\Omega$, *i.e.*, \mathcal{F}_Ω contains the open sets;
2. $E \in \mathcal{F}_\Omega \implies E^c \in \mathcal{F}_\Omega$, *i.e.*, \mathcal{F}_Ω is closed under complements;
3. $E_n \in \mathcal{F}_\Omega \implies \cup_{n=1}^\infty E_n \in \mathcal{F}_\Omega$, *i.e.*, \mathcal{F}_Ω is closed under countable unions;
4. $\mathcal{F}_\Omega \subset \mathcal{G}$ for any sigma field \mathcal{G} containing \mathcal{F}_1 .

Item 1. is trivial since $\mathcal{F}_\Omega = \cup_{\alpha < \Omega} \mathcal{F}_\alpha$, and in particular contains \mathcal{F}_1 . Item 2. follows by transfinite induction upon noting that $E \in \mathcal{F}_\alpha \implies E^c \in \mathcal{F}_{\alpha+1}$. Item 3 follows by noting that $E_n \in \mathcal{F}_\Omega \implies E_n \in \mathcal{F}_{\alpha_n}$ for some $\alpha_n < \Omega$, and $\beta = \sup_{n < \infty} \alpha_n$ is an ordinal satisfying $\alpha_n \preceq \beta < \Omega$ and hence $E_n \in \mathcal{F}_\beta$ for all n and $\cup_{n=1}^\infty E_n \in \mathcal{F}_{\beta+1}$. Verifying the minimality condition Item 4 is left as an exercise.

It isn't immediately obvious from the construction that we couldn't have stopped earlier—for example, that \mathcal{F}_2 or \mathcal{F}_ω isn't already the Borel sets, unchanging as we allow successively more intersections and unions. In fact that happens if the original space \mathcal{X} is countable or finite; in the case of \mathbb{R} , however, one can show that $\mathcal{F}_\alpha \neq \mathcal{F}_{\alpha+1}$ for every $\alpha < \Omega$.

Do you think this explicit construction is clearer or more complicated than the completion argument used in Billingsly's book?

INFINITE COIN TOSS

For each $\omega \in \Omega = (0, 1]$ and integer $n \in \mathbb{N}$ let $\delta_n(\omega)$ be the n^{th} bit in the nonterminating binary expansion of ω . There's some ambiguity in the dyadic expansion of rationals... for example, one-half can be written either as $0.10b$ or as the infinitely repeating $0.0111111...b$. If we had used the convention that the dyadic rationals have only finitely many 1's in their expansion (so $1/2 = 0.10b$) then $\delta_n(\omega) = \lfloor 2^n \omega \rfloor \pmod{2}$; with our convention that all expansions must have infinitely many ones, we have

$$\delta_n(\omega) = (\lceil 2^n \omega \rceil + 1) \pmod{2}.$$

We can think of $\{\delta_n\}$ as an infinite sequence of *random variables*, all defined on the same measurable space (Ω, \mathcal{B}^1) , with the random variable δ_1 equal to zero on $(0, 1/2]$ and one on $(1/2, 1]$; δ_2 equal to zero on $(0, 1/4] \cup (1/2, 3/4]$ and one on $(1/4, 1/2] \cup (3/4, 1]$; and, in general, δ_n equal to one on a union of 2^{n-1} intervals, each of length 2^{-n} (for a total length of $1/2$), and equal to zero on the complementary set, also of length $1/2$. For the Lebesgue probability measure \mathbb{P} on Ω that just assigns to each event $E \in \mathcal{B}^1$ its length $\mathbb{P}(E)$, we have $\mathbb{P}[X_n = 0] = \mathbb{P}[X_n = 1] = 1/2$.

Question 1: If we had used the other convention that every binary expansion must have infinitely many zero's (instead of one's), so e.g. $1/2 = 0.10b$, then what would the event $E_1 \equiv \{\omega : \delta_1(\omega) = 1\}$ have been?

The sigma field “generated by” any family of random variables $\{X_\alpha\}$ (whether countable or not) is defined to be the smallest sigma field for which each X_α is measurable, *i.e.*, the smallest one containing each $X_\alpha^{-1}(B)$ for every Borel set $B \subset \mathbb{R}$. For each fixed n the σ -algebra \mathcal{F}_n generated by $\delta_1, \dots, \delta_n$ is just the field $\mathcal{F}_n = \{\cup_i (a_i/2^n, b_i/2^n)\}$ consisting of all (finite) unions of left-open intervals with both endpoints an integer over 2^n . Each set in \mathcal{F}_n can be specified by listing which of the 2^n intervals $(\frac{i}{2^n}, \frac{i+1}{2^n}]$ ($0 \leq i < 2^n$) it contains, so there are 2^{2^n} sets in \mathcal{F}_n altogether. The union $\cup \mathcal{F}_n$ consists of all finite unions of left-open intervals with dyadic rational endpoints. It is closed under taking complements but it still isn't a sigma field, since it isn't closed under taking *countable* unions and intersections; for example, it contains the set $E_n = \{\omega : \delta_n = 1\}$ for each $n \in \mathbb{N}$ and finite intersections like $E_1 \cap \dots \cap E_n = (1 - 2^{-n}, 1]$, but not their countable intersection $\cap_{n=1}^\infty E_n = \{1\}$. By definition the “join” $\mathcal{F} = \bigvee_n \mathcal{F}_n \equiv \sigma(\cup_n \mathcal{F}_n)$ is just the smallest sigma field that contains each \mathcal{F}_n (and so contains their union); this is just the familiar Borel sets in $(0, 1]$.

Lebesgue measure \mathbb{P} , which assigns to any interval $(a, b]$ its length, is determined on each \mathcal{F}_n by the rule $\mathbb{P}[\cup_i (a_i/2^n, b_i/2^n)] = \sum (b_i - a_i)2^{-n}$ or, equivalently, by the joint distribution of the random variables $\delta_1, \dots, \delta_n$: independent Bernoulli's, each with $\mathbb{P}[\delta_i = 1] = 1/2$. For any number $0 < p < 1$ we can make a similar measure \mathbb{P}_p on (Ω, \mathcal{F}_n) by requiring $\mathbb{P}_p[\delta_n = 1] = p$ and, more generally,

$$\mathbb{P}[\delta_i = d_i, 1 \leq i \leq n] = p^{\sum d_i} (1-p)^{n-\sum d_i},$$

the four intervals in \mathcal{F}_2 would have probabilities $[(1-p)^2, p(1-p), p(1-p), \text{ and } p^2]$, for example, instead of $[1/4, 1/4, 1/4, 1/4]$. This determines a measure on each \mathcal{F}_n , which extends uniquely to a measure \mathbb{P}_p on $\mathcal{F} = \bigvee_n \mathcal{F}_n$. For $p = 1/2$ this is Lebesgue Measure, characterized by the property that $\mathbb{P}[(a, b]] = b - a$ for each $0 \leq a \leq b \leq 1$, but the other \mathbb{P}_p 's are new. This example (the family δ_n of random variables on the spaces $(\Omega, \mathcal{F}, \mathbb{P}_p)$) is an important one, and lets us build other important examples.

Under each of these probability distributions all the δ_n are both identically distributed and independent, *i.e.*,

$$\mathbb{P}[\delta_1 \in A_1, \dots, \delta_n \in A_n] = \prod_{i=1}^n \mathbb{P}[\delta_1 \in A_i].$$

Any probability assignment to intervals $(a, b] \subset \Omega$ determines *some* joint probability distribution for all the $\{\delta_n\}$, but typically the δ_n will be neither independent nor identically distributed. For any DF (*i.e.*, non-decreasing right-continuous function $F(x)$ satisfying $F(0) = 0$ and $F(1) = 1$), the prescription $\mathbf{P}_F((a, b]) \equiv F(b) - F(a)$ determines a probability distribution on every \mathcal{F}_n that extends uniquely to \mathcal{F} , determining the joint distribution of all the $\{\delta_n\}$.

Question 2: For $F(x) = x^2$, are δ_1 and δ_2 identically distributed? Independent? Find the marginal probability distribution for each δ_n under \mathbf{P}_F .

MEASURABILITY AND OBSERVABILITY

Fix any measure \mathbf{P}_p on (Ω, \mathcal{F}) (say, Lebesgue measure $\mathbf{P} = \mathbf{P}_{.5}$), and define a new sequence of random variables Y_n on $(\Omega, \mathcal{F}, \mathbf{P})$ by

$$Y_n(\omega) = \sum_{i=1}^n (-1)^{\delta_n(\omega)} = \sum_{i=1}^n (2\delta_n(\omega) - 1),$$

the sum of n independent terms, each ± 1 with probability $1/2$ each. This is the “symmetric random walk” (it would be asymmetric with \mathbf{P}_p for $p \neq .5$), starting at the origin and moving left or right with equal probability at each step; each Y_n is $2S_n - n$ for the binomial $\text{Bi}(n, .5)$ random variable $S_n = \sum_{i=1}^n \delta_i$, the partial sums of the δ_n 's.

The sigma field generated by the first n Y_i 's, that generated by the first n S_i 's, and that generated by the first n δ_i 's are all the same, the finite field \mathcal{F}_n of all unions of half-open intervals with endpoints of the form $j2^{-n}$, and a random variable Z on $(\Omega, \mathcal{F}, \mathbf{P})$ is \mathcal{F}_n -measurable if *and only if* Z can be written as a function $Z = \varphi_n(\delta_1, \dots, \delta_n)$ of the first n δ 's. Thus “measurability” *means* something for us— Z is **measurable** over \mathcal{F}_n if and only if you can tell its value by **observing** the first n values of δ_i (or, equivalently, of Y_i or S_i). We'll see that a function Z on Ω is \mathcal{F} -measurable (*i.e.*, is a random variable) if and only if you can approximate it arbitrarily well by a function of the first n δ_i 's, as $n \rightarrow \infty$.

UNIFORMS, NORMALS, AND MORE

From the infinite sequence of independent random bits $\{\delta_n\}$ we can construct as many random variables as we like of *any* distribution, all on the same space $(\Omega, \mathcal{F}, \mathbf{P})$, the unit interval with Lebesgue measure (length). For example, set:

$$\begin{aligned} U_1(\omega) &= \sum_{i=1}^{\infty} 2^{-i} \delta_{2^i}(\omega) & U_3(\omega) &= \sum_{i=1}^{\infty} 2^{-i} \delta_{5^i}(\omega) \\ U_2(\omega) &= \sum_{i=1}^{\infty} 2^{-i} \delta_{3^i}(\omega) & U_4(\omega) &= \sum_{i=1}^{\infty} 2^{-i} \delta_{7^i}(\omega), \end{aligned}$$

each the sum of *different* (and therefore independent) random bits; it is easy to see that $\{U_n\}$ will be independent, uniformly distributed random variables for $n = 1, 2, 3, 4$, and that we could construct as many of them as we like using successive primes $\{2, 3, 5, 7, 11, 13, \dots\}$.

Question 3: Why did I use δ_{2^i} , δ_{3^i} , δ_{5^i} , δ_{7^i} ? Give another choice that would have worked.

Let $F(x)$ be any DF (right-continuous, non-decreasing function on \mathbb{R} with limits 0 and 1 $x \rightarrow -\infty$ and $x \rightarrow +\infty$, respectively) and define:

$$\begin{aligned} X_1(\omega) &= \inf\{x \in \mathbb{R} : F(x) \geq U_1(\omega)\} & X_3(\omega) &= \inf\{x \in \mathbb{R} : F(x) \geq U_3(\omega)\} \\ X_2(\omega) &= \inf\{x \in \mathbb{R} : F(x) \geq U_2(\omega)\} & X_4(\omega) &= \inf\{x \in \mathbb{R} : F(x) \geq U_4(\omega)\}; \end{aligned}$$

it's not hard to see or show (we'll do it in a week or so) that the $\{X_n\}$ are independent, each with DF $F(x) = P[X_n \leq x]$. For example, we could take $X_n = \Phi^{-1}(U_n)$ to get independent random variables with the standard normal distribution or $X_n = -\log(1 - U_n)$ for the exponential distribution.

Independent normal random variables can be constructed even more efficiently via:

$$\begin{aligned} Z_1(\omega) &= \cos(2\pi U_1)\sqrt{-2 \ln U_2} & Z_3(\omega) &= \cos(2\pi U_3)\sqrt{-2 \ln U_4} \\ Z_2(\omega) &= \sin(2\pi U_1)\sqrt{-2 \ln U_2} & Z_4(\omega) &= \sin(2\pi U_3)\sqrt{-2 \ln U_4}; \end{aligned}$$

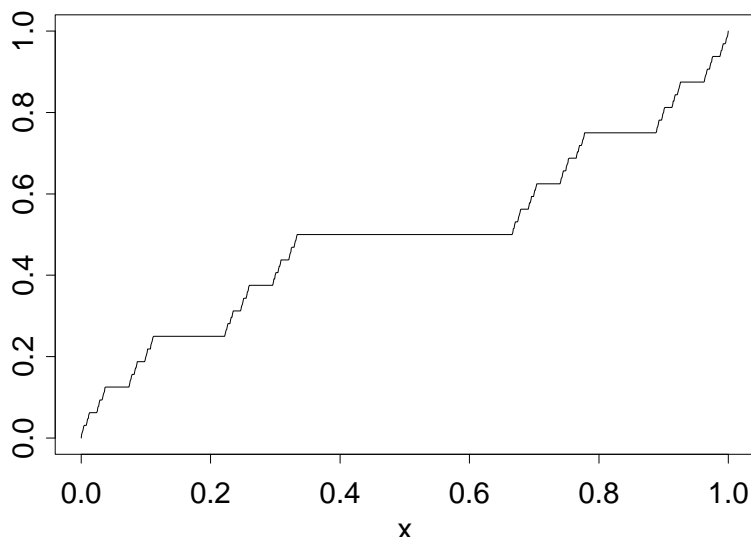
We've seen that from ordinary length measure on the unit interval (or, equivalently, from a single uniformly-distributed random variable ω) we can construct first an infinite sequence of independent 0 – 1 bits δ_n ; then an infinite sequence of independent uniform random variables U_n ; then an infinite sequence of independent normal random variables Z_n or, more generally, random variables X_n with any distribution(s) we choose.

The Cantor Distribution

Set $Y \equiv \sum_{n=1}^{\infty} 2\delta_n 3^{-n}$; then the ternery expansion of $y = Y(\omega)$ includes only zero's (where $\delta_n = 0$) and two's (where $\delta_n = 1$), and so lies in the Cantor set. Since Y takes on uncountably many different values, it cannot have a discrete random variable. Its CDF can be given analytically by the expression

$$F(y) = \sum_{n=1}^{\infty} \{2^{-n} : t_n > 0, t_m \neq 1, 1 \leq m < n\},$$

in terms of the ternery expansion $t_n \equiv \lfloor 3^n y \rfloor \pmod{3}$ of $y = \sum_{n=1}^{\infty} t_n 3^{-n}$ or graphically as



Evidently $F(x)$ has derivative $F' = 0$ wherever it is differentiable; this distribution is an example of a *singular* distribution, one that is neither absolutely continuous nor discrete.

Theorem. Let $F(x)$ be any distribution function. Then there exist unique numbers $p_d \geq 0$, $p_c \geq 0$, $p_s \geq 0$ with $p_d + p_c + p_s = 1$ and distribution functions $F_d(x)$, $F_c(x)$, $F_s(x)$ with the properties that F_d is discrete with some probability mass function $f_d(x)$, F_c is absolutely

continuous with some probability density function $f_c(x)$, and F_s is singular, satisfying $F(x) = p_d F_d(x) + p_c F_c(x) + p_s F_s(x)$ and

$$F_d(x) = \sum_{t \leq x} f_d(t), \quad F_c(x) = \int_{t \leq x} f_c(t) dt, \quad F'_s(x) = 0.$$

EXPECTATION AND INTEGRAL INEQUALITIES

Discrete RV's

If a random variable Y can take on only a finite or countably infinite set of values, say b_i , then each set $\Lambda_i = [\omega : Y(\omega) = b_i]$ must be in \mathcal{F} ; the Λ_i are disjoint, and Y can be represented in the form

$$Y(\omega) = \sum_i b_i 1_{\Lambda_i}(\omega), \quad \text{where } 1_{\Lambda_i}(\omega) = \begin{cases} 1 & \text{if } \omega \in \Lambda_i \\ 0 & \text{if } \omega \notin \Lambda_i \end{cases} \quad (\star)$$

is the so-called *indicator function* of Λ_i . By adding a term with $b_i = 0$, if necessary, we can assume that $\Omega = \cup \Lambda_i$ so the $\{\Lambda_i\}$ form a “countable partition” of Ω . Any RV X can be approximated as well as we like by a simple RV of the form (\star) by choosing $\epsilon > 0$, setting $b_i \equiv i\epsilon$, and

$$\Lambda_\epsilon \equiv \{\omega : b_i \leq X(\omega) < b_i + \epsilon\} \quad X_\epsilon(\omega) \equiv \sum_{-\infty}^{\infty} b_i 1_{\Lambda_i}(\omega) = \epsilon \lfloor X(\omega)/\epsilon \rfloor$$

It is easy to define the *expectation* of such a simple RV, or (equivalently) the *integral* of X_ϵ over $(\Omega, \mathcal{F}, \mathbf{P})$, if X is bounded below or above (to avoid indeterminate sums):

$$\mathbf{E}X_\epsilon = \int_{\Omega} X_\epsilon(\omega) \mathbf{P}(d\omega) = \int_{\Omega} X_\epsilon(\omega) d\mathbf{P}(\omega) = \int_{\Omega} X_\epsilon d\mathbf{P} = \sum_i b_i \mathbf{P}(\Lambda_i)$$

Since $X_\epsilon(\omega) \leq X(\omega) < X_\epsilon(\omega) + \epsilon$, we have $\mathbf{E}X_\epsilon \leq \mathbf{E}X < \mathbf{E}X_\epsilon + \epsilon$, *i.e.*,

$$\sum_i i\epsilon \mathbf{P}[i\epsilon \leq X < (i+1)\epsilon] \leq \mathbf{E}X < \sum_i i\epsilon \mathbf{P}[i\epsilon \leq X < (i+1)\epsilon] + \epsilon. \quad (\star\star)$$

This determines the value of $\mathbf{E}X = \int_{\Omega} X d\mathbf{P}$ for each random variable X . If we take $\epsilon = 2^{-n}$ above, and simplify the notation by writing X_n for $X_{2^{-n}} = 2^{-n} \lfloor 2^n X \rfloor$, the sequence X_n increases monotonically to X and we can define $\mathbf{E}X = \lim_n \mathbf{E}X_n$.

Note that even for $\Omega = (0, 1]$, $\mathbf{P} = \lambda(dx)$ (Lebesgue measure), and X continuous, the passage to the limit suggested in $(\star\star)$ is *not* the same as the limit of Riemann sums that is used to introduce integration in undergraduate calculus courses; for the Riemann sum it is the x -axis that is broken up into integral multiples of some ϵ , determining the integral of *continuous* functions, while here it is the y axis that is broken up, determining the integral of all *measurable* functions. The two definitions of integral agree for continuous functions where they are both defined, of course, but the present one is much more general.

If X is *not* bounded below or above, we can set $X^+ \equiv 0 \vee X$ and $X^- \equiv 0 \vee -X$, so that $X = X^+ - X^-$ with both X^+ and X^- bounded below (by zero), so their expectations are well-defined; if either $\mathbf{E}X^+ < \infty$ or $\mathbf{E}X^- < \infty$, we can unambiguously define $\mathbf{E}X \equiv \mathbf{E}X^+ - \mathbf{E}X^-$, while if $\mathbf{E}X^+ = \mathbf{E}X^- = \infty$ we regard $\mathbf{E}X$ as undefined.

For any measurable set $\Lambda \in \mathcal{F}$ we write $\int_{\Lambda} X d\mathbf{P}$ for $\mathbf{E}X 1_{\Lambda}$. For $\Omega \subset \mathbb{R}$, if \mathbf{P} gives positive probability to either $\{a\}$ or $\{b\}$ then the integrals over the sets (a, b) , $(a, b]$, $[a, b)$, and $[a, b]$ may all be different; the notation $\int_a^b X d\mathbf{P}$ isn't expressive enough to distinguish them.

Frequently in Probability and Statistics we need to calculate or estimate integrals and expectations; usually this is done through limiting arguments in which a sequence of integrals is shown to converge to the one whose value we need. Here are some important properties of integrals for any measurable set $\Lambda \in \mathcal{F}$ and random variables $\{X_n\}$, X , Y , useful for bounding or estimating the integral of a random variable X (they're only listed here for reference and so we can talk about them— don't worry, you won't have to remember them all or know how to prove them!):

1. $\int_{\Lambda} X dP$ is well-defined and finite if and only if $\int_{\Lambda} |X| dP < \infty$, and $\left| \int_{\Lambda} X dP \right| \leq \int_{\Lambda} |X| dP$. We can also define $\int_{\Lambda} X dP \leq \infty$ for any X bounded below by some $b > -\infty$.
2. **Lebesgue's Monotone Convergence Thm:** If $0 \leq X_n \nearrow X$, then $\int_{\Lambda} X_n dP \nearrow \int_{\Lambda} X dP \leq \infty$. In particular, the sequence of integrals converges (possibly to $+\infty$).
3. **Lebesgue's Dominated Convergence Thm:** If $X_n \rightarrow X$, and if $|X_n| \leq Y$ for some RV $Y \geq 0$ with $EY < \infty$, then $\int_{\Lambda} X_n dP \rightarrow \int_{\Lambda} X dP$ and $\int_{\Lambda} |X| dP \leq \int_{\Lambda} Y dP < \infty$. In particular, the sequence of integrals converges to a finite limit.

4. **Fatou's Lemma:** If $X_n \geq 0$ on Λ , then $\int_{\Lambda} (\liminf X_n) dP \leq \liminf (\int_{\Lambda} X_n dP)$. The two sides may be unequal (example?), and the result is false for \limsup .
5. **Fubini's Thm:** If *either* each $X_n \geq 0$, *or* $\sum_n \int_{\Lambda} |X_n| dP < \infty$, then the order of integration and summation can be exchanged: $\sum_n \int_{\Lambda} X_n dP = \int_{\Lambda} \sum_n X_n dP$. If both these conditions fail, the orders may not be exchangeable (example?).
6. For any $p > 0$, $E|X|^p = \int_0^{\infty} p x^{p-1} P[|X| > x] dx$ and $E|X|^p < \infty \Leftrightarrow \sum_{n=1}^{\infty} n^{p-1} P[|X| \geq n] < \infty$. The case $p = 1$ is easiest and most important: if $S \equiv \sum_{n=1}^{\infty} P[|X| \geq n] < \infty$, then $S \leq E|X| < S+1$. If X takes on only nonnegative integer values, $EX = S$.

7. If μ_X is the distribution of X , and if f is a measurable real-valued function on \mathbb{R} , then $Ef(X) = \int_{\Omega} f(X(\omega)) dP = \int_{\mathbb{R}} f(x) \mu_X(dx)$ if either side exists. In particular, $\mu = EX = \int x \mu_X(dx)$ and $\sigma^2 = E(X - \mu)^2 = \int (x - \mu)^2 \mu_X(dx)$.
8. **Hölder's Inequality:** Let $p > 1$ and $q = \frac{p}{p-1}$ (e.g., $p = q = 2$ or $p = 1.01$, $q = 101$). Then $EXY \leq E|XY| \leq [E|X|^p]^{\frac{1}{p}} [E|Y|^q]^{\frac{1}{q}}$. In particular, for $p = q = 2$,

$$\text{Cauchy-Schwartz Inequality: } EXY \leq E|XY| \leq \sqrt{EX^2 EY^2}.$$

9. **Minkowski's Inequality:** Let $1 \leq p \leq \infty$ and let $X, Y \in L_p(\Omega, \mathcal{F}, P)$. Then

$$(E|X + Y|^p)^{\frac{1}{p}} \leq (E|X|^p)^{\frac{1}{p}} + (E|Y|^p)^{\frac{1}{p}}$$

Thus the norm $\|X\|_p \equiv (E|X|^p)^{\frac{1}{p}}$ obeys the triangle inequality on $L_p(\Omega, \mathcal{F}, P)$.

10. **Jensen's Inequality:** Let $\varphi(x)$ be a convex function on \mathbb{R} , X an integrable RV. Then $\varphi(E[X]) \leq E[\varphi(X)]$. Examples: $\varphi(x) = |x|^p$, $p \geq 1$; $\varphi(x) = e^x$; $\varphi(x) = [0 \vee x]$.
11. **Markov's & Chebychev's Inequalities:** If φ is positive and increasing, then $P[|X| \geq u] \leq E[\varphi(|X|)]/\varphi(u)$. In particular $P[|X - \mu| > u] \leq \frac{\sigma^2}{u^2}$ and $P[|X| > u] \leq \frac{\sigma^2 + \mu^2}{u^2}$.
One-Sided Version: $P[X > u] \leq \frac{\sigma^2}{\sigma^2 + (u - \mu)^2}$.
12. **Hoeffding's Inequality:** If $\{X_j\}$ are independent and $(\exists \{a_j, b_j\})$ s.t. $P[a_j \leq X_j \leq b_j] = 1$, then $(\forall c > 0)$, $S_n := \sum_{j=1}^n X_j$ satisfies $P[S_n - ES_n \geq c] \leq \exp(-2c^2 / \sum_{j=1}^n |b_j - a_j|^2)$. Hoeffding proved this improvement on Chebychev's inequality (at UNC) in 1963. See also related **Azuma's** inequality (1967), **Bernstein's** inequality (1937), and **Chernoff** bounds (1952).