

# Canonical Correlations

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Let  $X$  be an  $n$ -dimensional random vector with finite covariance, and  $Y$  an  $m$ -dimensional one. Are  $X$  and  $Y$  independent? If not, how can one quantify the *degree* of dependence between them? Here's one form of answer, borrowed in large part from the Wikipedia.

## 1 Mutual Information

One classical approach from information theory is to measure the *mutual information*. If the joint distribution of  $X$  and  $Y$  has a density function  $f(x, y)$  with marginals  $f_x(x)$  and  $f_y(y)$ , this is defined to be

$$I(X : Y) = \iint \lg \left[ \frac{f(x, y)}{f_x(x) f_y(y)} \right] f(x, y) dx dy$$

(here  $\lg x \equiv \log_2 x = \frac{\log x}{\log 2}$ ; traditionally logarithms in Information Theory are taken to base two, so  $I(X : Y)$  is measured in ‘bits’). A similar expression applies to discrete random variables, with a sum replacing the integral— or, more generally, to any joint distribution absolutely-continuous with respect to the product of the two marginal distributions (otherwise  $I(X : Y) = \infty$ ). It's easy to show that  $I(X : Y) \geq 0$  for all  $X, Y$  and that  $I(X : Y) = 0$  if and only if  $X$  and  $Y$  are independent. The quantity  $I(X : Y)$  is invariant under affine changes of variables for  $X$  and  $Y$ .

For univariate normal random variables  $X \sim \text{No}(\mu_x, \sigma_x^2)$  and  $Y \sim \text{No}(\mu_y, \sigma_y^2)$  with correlation  $\rho$ , hence covariance and precision matrices

$$\Sigma = \begin{bmatrix} \sigma_x^2 & \rho\sigma_x\sigma_y \\ \rho\sigma_x\sigma_y & \sigma_y^2 \end{bmatrix} \quad \Lambda = \frac{1}{\sigma_x^2\sigma_y^2(1-\rho^2)} \begin{bmatrix} \sigma_y^2 & -\rho\sigma_x\sigma_y \\ -\rho\sigma_x\sigma_y & \sigma_x^2 \end{bmatrix}$$

we have (setting  $\mu_x = \mu_y = 0$  without loss of generality)

$$\begin{aligned}
I(X : Y) &= \iint \lg \left[ \frac{[2\pi\sigma_x^2\sigma_y^2(1-\rho^2)]^{-1/2} \exp \left\{ -\frac{\sigma_y^2x^2 - 2\rho\sigma_x\sigma_yxy + \sigma_x^2y^2}{2\sigma_x^2\sigma_y^2(1-\rho^2)} \right\}}{[2\pi\sigma_x^2]^{-1/2} \exp(-x^2/2\sigma_x^2) [2\pi\sigma_y^2]^{-1/2} \exp(-y^2/2\sigma_y^2)} \right] f(x, y) dx dy \\
&= \iint \lg \left[ \frac{1}{\sqrt{1-\rho^2}} \exp \left\{ \frac{x^2}{2\sigma_x^2} + \frac{y^2}{2\sigma_y^2} - \frac{\sigma_y^2x^2 - 2\rho\sigma_x\sigma_yxy + \sigma_x^2y^2}{2\sigma_x^2\sigma_y^2(1-\rho^2)} \right\} \right] f(x, y) dx dy \\
&= -\frac{1}{2} \lg(1-\rho^2) + \frac{1}{\log 2} \left\{ \frac{\sigma_x^2}{2\sigma_x^2} + \frac{\sigma_y^2}{2\sigma_y^2} - \frac{\sigma_y^2\sigma_x^2 - 2\rho\sigma_x\sigma_y\rho\sigma_x\sigma_y + \sigma_x^2\sigma_y^2}{2\sigma_x^2\sigma_y^2(1-\rho^2)} \right\} \\
&= -\log(1-\rho^2)/(2\log 2). \tag{1}
\end{aligned}$$

As expected,  $I(X : Y) = 0$  if  $X \perp\!\!\!\perp Y$  (i.e. if  $\rho = 0$ ), while  $I(X : Y) \rightarrow \infty$  as  $\rho \rightarrow \pm 1$ . We'll return to this below.

There are close links between  $I(X : Y)$ , the Kullback-Leibler divergence from  $p(x, y)$  to  $p_x(x)p_y(y)$ , and the mutual (Shannon, not Fisher) information  $H(X, Y)$  between  $X$  and  $Y$ .

## 2 Canonical Correlations

For  $a \in \mathbb{R}^n$  and  $b \in \mathbb{R}^m$ , what is the highest possible correlation  $\rho$  between  $U = a'X$  and  $V = b'Y$ ? First let's standardize by setting  $c = \Sigma_{XX}^{-\frac{1}{2}}a$  and  $d = \Sigma_{YY}^{-\frac{1}{2}}b$ ; then

$$\begin{aligned}
\rho &= \frac{a'\Sigma_{XY}b}{\sqrt{a'\Sigma_{XX}a}\sqrt{b'\Sigma_{YY}b}} \\
&= \frac{c'\Sigma_{XX}^{-\frac{1}{2}}\Sigma_{XY}\Sigma_{YY}^{-\frac{1}{2}}d}{\sqrt{c'c}\sqrt{d'd}}
\end{aligned}$$

so, by the Cauchy-Schwartz inequality,

$$\rho^2 \leq \frac{(c'AA'c)(d'd)}{c'c \ d'd} = \frac{c'AA'c}{c'c}$$

where

$$A = \Sigma_{XX}^{-\frac{1}{2}}\Sigma_{XY}\Sigma_{YY}^{-\frac{1}{2}}.$$

The largest this can possibly be is the maximum eigenvalue  $\lambda_1$  of the positive-definite matrix  $AA'$ ; it's attained when  $c$  is the (say, unit) eigenvector  $c_1$  for  $AA'$  and where  $d = d_1$  is proportional to  $A'c_1$  and (say) of unit length.

We call  $\rho_1$  the “first canonical correlation” and the vectors

$$U_1 = c'_1 \Sigma_X^{-\frac{1}{2}} X = a'_1 X \quad V_1 = d'_1 \Sigma_Y^{-\frac{1}{2}} Y = b'_1 Y$$

the first Canonical Components.

We can define similarly the second, third, etc. canonical correlations and components, up to the rank of  $AA'$ ,  $(n \wedge m)$ .

These may be used to quantify the degree of dependence between  $X$  and  $Y$ . For example, if all  $\rho_i = 0$  for  $i > r$ , the dependence has rank (at most)  $r$ ; a classical test of that null hypothesis may be based on the asymptotic (as the number of replicates  $p \rightarrow \infty$ )  $\chi^2_\nu$  distribution of the test statistic

$$Q_r = - \left( p - 1 - \frac{1}{2}(m + n + 1) \right) \sum_{j=r+1}^{n \wedge m} \log(1 - \hat{\rho}_j^2), \quad (2)$$

with degrees of freedom  $\nu = (m - r)(n - r)$ . As an exploratory tool, a plot of  $Q_r$  *vs.*  $r$  may suggest the dimensionality of the linear interdependence of  $X$  and  $Y$ .

## 2.1 Back to Information

The sum in Equation (2) is just  $(-2 \log 2)$  times the mutual information  $I(X_r : Y_r)$  between the projections of  $X$  and  $Y$  onto the orthogonal complements of the spaces spanned by the first  $r$  canonical components (see Equation (1)).