Analyzing Stochastic Diffusion Processes

Introduction

- Many interesting ecological diffusions
- Emerging diseases avian flu, H1N1 flu
- Exotic organisms invasive plants, gypsy moths
- Size and age distributions
- Transformation of landscape, deforestation, land use classifications, urban growth

The objective

- Our objective: forecast likely spread in space and time with associated uncertainty
- Nonlinear, nonhomogeneous in space and time
- Explanatory covariates
- Start with deterministic integro-differential equations or with partial differential equations
- How to add uncertainty?

cont.

- Theoretical models, "varying" around them
- Too much simplification required to obtain analytical solutions
- Discretization to fit models
- Do we care of the deterministic equation? Should we just work with the discrete time version we want? Dynamic spatial models?

Hierarchical modeling

The hierarchical paradigm:

[data|process, parameters][process|parameters][parameters]

- A paradigm shift designed experiments to observational studies; controlled experiments to integrated (big picture) investigation
- Prior information from: empirical studies, mechanistic knowledge, ecological theory, etc.
- Multiple information sources
- Conditional uncertainty in components (a natural way to specify models)
- Different resolutions in space and time
- Structured dependence in space and time
- Complex dependence structure through latent variables

Computation

- Fit within a Bayesian framework; enables full inference, exact inference
- Model fitting, associated computation is challenging
- High dimension, sparsity, dimension reduction
- MCMC/Gibbs sampler model fitting
- Model validation? model comparison?

Dynamics

- **•** Continuous space, discrete time, i.e., $w_t(s)$
- Without loss of generality $\mathbf{t} = (1, 2, ..., T)$
- Envision $w_t(s)$ as a *dynamical* process
- In fact, simplify to first order Markov, i.e., for locations $s_1, s_2, ..., s_n$, let $\mathbf{w}_t = (w_t(s_1), w_t(s_2), ..., w_t(s_n))^T$. Then

$$[\mathbf{w}_t | \mathbf{w}_0, \mathbf{w}_1, \dots \mathbf{w}_{t-1}] = [\mathbf{w}_t | \mathbf{w}_{t-1}]$$

- For example, w_t = Hw_{t-1} + η_t where η_t(s) incorporates spatial structure
- A vector AR(1) model and H is called the propagator matrix
- ??Specifying H??

Specifying H

- I = I not stationary (explosive), no interaction across space and time, not realistic for most dynamic processes of interest
- H = Diag(h) where Diag(h) has diagonal elements 0 < h_i < 1 - Still no interactions</p>
- Integro-difference equation (IDE) dynamics:

$$w_t(s) = \int h(s, r; \phi) w_{t-1}(r) dr + \eta_t(s)$$

h is a "redistribution kernel" that determines the rate of diffusion and the advection

cont

• If require w > 0, work with

$$\mathsf{og} w_t(s) = \mathsf{log}(\int h(s,r;\phi) w_{t-1}(r) dr) + \eta_t(s)$$

Alternatively,

$$v_t(s) = \int h(s, r; \phi) v_{t-1}(r) dr$$

and

$$\log w_t(s) = \log v_t(s) + \eta_t(s)$$

- Discretization to obtain H
- **•** Forms for $h(s,r;\phi)$; $h(s,r;\phi(r))$?, $h_t(s,r;\phi)$?

cont

- **P** Recall linear PDE, $\frac{dw(s,t)}{dt} = h(s)w(s,t)$
- Finite differencing yields $w(s,t+\Delta t) - w(s,t) = h(s)w(s,t)\Delta t$, i.e., $w(s,t+1) \approx \tilde{h}(s)w(s,t)$. Same limitations as above.
- Need more general PDE's
- PDE can motivate IDE, can clarify H
- "forward" vs. "backward" perspective
- IDE's can be specified directly without using PDE's,
 e.g., h(s,r) can be a sum of a survival/spread term + a birth/replenishment term

Diffusion PDE's

- Diffusion in one dimension Fick's Law: diffusive flux from *high* concentration to *low* is $-\delta \frac{\partial w(x,t)}{\partial x}$ with δ , the diffusion coefficient. Location varying diffusion $\delta(x)$
- And, diffusion equation is ∂w/∂t = −∂flux/∂x, i.e., $\frac{∂w(x,t)}{∂t} = \frac{∂}{∂x} (\delta(x) \frac{∂w(x,t)}{∂x})$
- That is, the 1-dim diffusion equation is

$$\frac{\partial w(x,t)}{\partial t} = \delta'(x)\frac{\partial w(x,t)}{\partial x} + \delta(x)\frac{\partial^2 w(x,t)}{\partial x^2}$$

- In 2-dim, diffusive flux is $-\delta(x, y)\nabla w(x, y, t)$ ($\nabla w(x, y, t)$ is the concentration gradient at time *t*)
- The resulting diffusion PDE is

$$\frac{\partial w(x,y,t)}{\partial t} = \frac{\partial}{\partial x} (\delta(x,y) \frac{\partial w(x,y,t)}{\partial x}) + \frac{\partial}{\partial y} (\delta(x,y) \frac{\partial w(x,y,t)}{\partial y})_{\text{-1.157}}$$

Discretizing the diffusion equation

- Complete the differentiation of the diffusion equation
- Yields second order partial derivatives, $\frac{\partial^2 w}{\partial x^2}$ and $\frac{\partial^2 w}{\partial y^2}$
- **9** Introduce Δt , Δx , Δy
- Replace ∂'s with finite differences (first forward and second order centered) careful detail, ugly expression!
- After the smoke clears, we obtain $w_{t+\Delta t} = Hw_t$
- $onumber \,$ Again, add η_t
- We are back to our earlier redistribution form

Add growth rate

- Previous dynamics simply redistribute existing population spatially over time
- In many situations, there is also growth of the population
- Population growth can be captured by a logistic differential equation

$$\frac{\partial w(s,t)}{\partial t} = rw(s,t)(1 - w(s,t)/K)$$

- \bullet r is the growth rate, K is the carrying capacity
- Add growth to the diffusion PDE for $\frac{\partial w(s,t)}{\partial t}$
- r(s)?, K(s)?

Eurasian collared dove data

- An example from Wikle et al., using data from the Breeding Bird Survey (BBS)
- Escaped to U.S. from Bahamas, introduced in Florida, expanding dramatically across North America
- 4000+ routes in the survey (some sampled more than once per year, others not sampled in a given year), length of route is ≈ 40 kms, 50 stops per route, count birds by sight for 3 minutes item 18 years: 1986-2003
- Route is a "point", response at a point is a count
- Aggregate to grid boxes
- Z_{it} is count in box i in year t, n_{it} is number of visits to cell i in year t.
- λ_{it} is *intensity* for box *i* in year *t*

Modeling specifics

- ϵ_{it} are i.i.d. (pure error or micro-scale variation)
- The focus is on the w_t . They tell the diffusion story, i.e., $w_t = H(\delta)w_{t-1} + \eta_t$
- Model for η_t ?
- $\mathbf{w}_0 \sim N(0, 10I)$
- δ is the vector of local diffusion coefficients, one for each grid cell
- A dimension reduction for δ; many possibilities here basis functions, EOF's, predictive processes

Enriching the modeling

- Again, we focus on w(s,t)
- w(s,t) can arise as a mean model for a geostatistical model or in a space-time GLM (as in Wikle) or as a cumulative intensity Λ(s,t) for a space-time point pattern (which drives the cumulative diffusion)
- A general diffusion PDE (nonstochastic) looks like $\frac{\partial w(s,t)}{\partial t} = a(w(s,t), z(s,t), \theta)$ where z(s,t) are other potential variables (z(s,t) = t for example)
- How to make the PDE stochastic?
- For the remainder, we use the logistic DE, i.e., a diffusion driven by a logistic growth model
- Discretization as proposed above

First a DE

Ignoring location s for the moment, we have:

$$dw(t) = a(w(t), t, \theta)dt \quad \text{with} \quad w(0) = w_0$$

Simplest way to add stochasticity is to make θ random.

Instead:

$$dw(t) = a(w(t), t, \theta)dt + b(w(t), t, \theta)dZ(t)$$

where Z(t) is Brownian motion over R^1 with a and b the "drift" and "volatility" respectively. Now a *stochastic* differential equation (SDE)

• θ would still be random

Next: dw(t) = a(w(t), t, θ(t))dt where (with Z(t) is variance 1 Brownian motion)

 $d\theta(t) = g(\theta(t), t, \beta)dt + h(\theta(t), \sigma)dZ(t)$

- This includes the previous example
- For the logistic equation:

$$dw(t) = \theta(t)w(t)\left[1 - \frac{w(t)}{K}\right]dt$$

If $\theta(t) = \mu + \zeta(t)$ with $d\zeta(t) = -a\zeta(t) dt + \sigma_{\zeta} dZ(t)$ equivalently

$$d\theta (t) = -\alpha (\mu - \theta (t)) dt + \sigma_{\zeta} dZ (t) ,$$

a self-reverting Ornstein-Uhlenbeck (OU) process and $\theta(t)$ is a stationary GP with $cov(\theta(t), \theta(t') = (\sigma^2/\alpha)exp(-\alpha|t - t'|).$

Add space

Now, we add space. First,

 $dw(s,t)=a(w(s,t),t,\theta(s))dt \quad \text{with} \quad w(s,0))=w_0(s),$

a PDE. Randomness through $\theta(s)$, a process realization, so $\theta(s)$ provides the spatial dependence. Hence, a stochastic process of differential equations.

Next,

 $dw(s,t) = a(w(s,t),t,\theta(s))dt + b(w(s,t),t,\theta(s))dZ(s,t)$

- Modeling Z(s,t)? For a fixed finite set of spatial locations assume independent Brownian motion at each location.
- Or a discrete space approximation to spatial Brownian motion employing a Gaussian process (GP)on R²

cont.

Next,

$$dw(s,t) = a(w(s,t),t,\theta(s,t))dt$$

where say

$$d\theta(s,t) = \gamma(\theta(s,t) - \theta(s))dt + bdZ(s,t)$$

- **9** Again, $\theta(s)$ is process realization
- Now, $\theta(s,t)$ given through an infinite dimensional SDE
- This version produces a covariance function that is separable in space and time

Important points

- A differential equation in time at every spatial location, i.e., parameters indexed by location
- The parameters vary spatially as realizations of a spatial process
- Instead, the differential equation is a stochastic differential equation (SDE), e.g., a spatial Ornstein-Uhlenbeck (O-U) process
- For the logistic PDE, the rate parameter in the differential equation is assumed to change over time. It can be modeled as a realization of a spatio-temporal process
- It can be modeled using a SDE, yielding an SDE embedded within the differential equation

- Spatio-temporal modeling settings for the above?
- The usual "geostatistics" setting with observations at locations and times:

$$Y(s,t) = \Lambda(s,t) + \epsilon(s,t)$$

with $\Lambda(s,t)$ modeled through a differential equation as above, i.e., process model is a stochastic PDE.

• Space-time point pattern setting with data of the form $(s_i, t_i), i = 1, 2, ...n$, i.e., random locations at random times. Now, we model the space-time intensity associated with the point pattern, again, say $\Lambda(s, t)$, through a differential equation as above resulting in a space-time Cox process.

A Geostatistical Example

The logistic PDE in space and time:

$$\frac{\partial \Lambda\left(t,s\right)}{\partial t} = r(t,s)\Lambda\left(t,s\right)\left[1 - \frac{\Lambda\left(t,s\right)}{K\left(s\right)}\right]$$

- Time discretized to intervals Δt , indexed as $t_j, j = 0, 1, 2, ...J$. At location s_i , data $Y(t_j, s_i)$
- **Dynamic model:** $Y(t_j, s_i) = \Lambda(t_j, s_i) + \varepsilon(t_j, s_i)$
- Using Euler's approximation yields difference equation: $\Delta\Lambda(t_j,s) = r(t_{j-1},s)\Lambda(t_{j-1},s)\left[1 - \frac{\Lambda(t_{j-1},s)}{K(s)}\right]\Delta t,$ $\Lambda(t_j,s) \approx \Lambda(0,s) + \sum_{l=1}^{j} \Delta\Lambda(t_l,s)$
- No spatial flux, no Δx , no Δy , $\Lambda_t = Diag(h)\Lambda_{t-1}$

cont.

- Can not add scaled Brownian motion to the logistic PDE.
- Instead a time-varying growth rate at each location
- An O-U process for r(t,s):

$$\frac{\partial r\left(t,s\right)}{\partial t} = \alpha_r\left(\mu_r\left(s\right) - r(t,s)\right) + \frac{\partial B\left(t,s\right)}{\partial t}$$

- Model the initial A (0, s) and K (s) as log-Gaussian spatial processes with regression forms for the means
- Below we assume K (s) known and set to 1, i.e., 100% is the capacity.
- Similar modeling for $\mu_r(s)$

A Simulation Example

- \blacksquare 10 × 10 study region
- 44 locations over 30 time periods
- 4 sites for holdout
- Matérn covariance function used for $\Lambda_0(s)$, for r(s), $(\nu = 3/2 \text{ used in the example})$
- E.g., the space time covariance function for r(t,s)becomes $\varrho(t_{j_1} - t_{j_2}, s_{i_1} - s_{i_2}) = \sigma_r^2 \exp\left(-\alpha_r |t_{j_1} - t_{j_2}|\right) (\phi_r |s_{i_1} - s_{i_2}|)^{\nu} \kappa_{\nu} (\phi_r |s_{i_1} - s_{i_2}|)$
- weak priors on σ^2 's, weak log normal prior on α_r , discrete prior on ϕ 's

Space-time point patterns

- Spatio-temporal Cox process models using SDEŠs
- Motivation: Urban development using spatio-temporal point processes.
- Differential equation models for cumulative intensity
- We have 21 years of urban development data for Irving, TX but we just show a simulation example with five years

Spatio-temporal Cox Process

In a study region D during a period of [0,T], N_T events:

Point pattern:
$$X_T = \left\{ x_{1,t_1}, \dots, x_{N_T,t_{N_T}} \right\}$$

where $x_{i,t_i} = (x_i^1, x_i^2, t_i)$

 X_T is a Poisson process with inhomogeneous intensity

$$\Omega\left(t,s\right), s \in D, t \in [0,T]$$

Specifying the intensity?

$$\Omega(t,s) = f(t,\theta_l(t,s); l = 1,\ldots,p)$$

 $\theta_l(t,s), s \in D, l = 1, \dots, p$ are processes for parameters of interest.

The cumulative intensity

Discretize the spatio-temporal Cox process in time:

Spatial point pattern: $X_{[t_1,t_2]}$ during $t \in [t_1,t_2)$ $x_i = (x_i^1, x_i^2), x_i \in X_{[t_1,t_2]}$

The cumulative intensity for $X_{[t_1,t_2]}$ is

$$\int_{t_1}^{t_2} \Omega(t,s) \, dt = \int_{t_1}^{t_2} f(t,\theta_l(t,s); l = 1,\dots,p) \, dt$$

We consider models for the cumulative intensity

$$\Lambda\left(t,s\right) = \int_{0}^{t} \Omega\left(\tau,s\right) d\tau$$

Comments

- So house locations and times over $(0,T] \times D$
- Need a ∆t and an area A in order to observe a point pattern
- If Ω(t, s) ≥ 0 then Λ(t, s) increases in t; we do not allow house removal
- Work with cumulative intensity Λ(t, s) easier to think about mechanistically. In fact, Λ(t₂, s) - Λ(t₁, s) provides the intensity for the interval (t₁, t₂].
- Dynamics in Λ(t, s) provide dynamics for the discretized spatial point process

Illustrative growth models (each of which has an explicit solution)

• Exponential growth

$$\frac{d\Lambda\left(t,s\right)}{dt} = r\left(s\right)\Lambda\left(t,s\right)$$

• Gompertz growth

$$\frac{d\Lambda\left(t,s\right)}{dt} = r(s)e^{-\alpha(s)t}\Lambda\left(t,s\right)$$

Logistic growth

$$\frac{d\Lambda(t,s)}{dt} = r(s)\Lambda(t,s) \left[1 - \frac{\Lambda(t,s)}{K(s)}\right]$$

local growth rate local carrying capacity

Process Models for the Parameters

r(s), K(s) and initial intensity

$$\Lambda_0(s) = \int_{-\infty}^0 \Omega(\tau, s) \, d\tau$$

are parameter processes which are modeled on log scale as

$$\log \Lambda_0(s) = \mu_{\Lambda}(s; \beta_{\Lambda}) + \theta_{\Lambda}(s), \quad \theta_{\Lambda}(s) \sim GP(0, C_{\Lambda}(\phi_{\Lambda}))$$
$$\log r(s) = \mu_r(s; \beta_r) + \theta_r(s), \quad \theta_r(s) \sim GP(0, C_r(\phi_r))$$
$$\log K(s) = \mu_K(s; \beta_K) + \theta_K(s), \quad \theta_K(s) \sim GP(0, C_K(\phi_K))$$

Hence, given r (s), K (s) and Λ₀ (s) the growth curve is fixed. Also, the μ's are trend surfaces.

The problem

- We want the differential equations to be dependent at every location BUT
- We would not insist that the process exactly follows a logistic differential equation at every location
- We want to introduce some noise so convert to an SDE
- Again a time varying rate at each location
- Again $Z(t(s)) = \log r(t(s))$ with

 $dZ(t(s)) = \xi(Z(t(s)) - \log r(s)) + \sigma_Z dW(t, s)$

 Induces a stationary process with a separable space-time covariance function

Discretizing Time (Euler Approximation)

Back to the original model, the intensity for the spatial point pattern in a time interval:

$$\Delta\Lambda_{j}(s) = \int_{t_{j,1}}^{t_{j,2}} \Omega(\tau, s) \, d\tau \approx \Omega(t_{j,1}, s) \, \Delta t$$

• Difference equation model:

$$\Delta \Lambda_{j}(s) = r(s)\Lambda_{j-1}(s) \left[1 - \frac{\Lambda_{j-1}(s)}{K(s)}\right] \Delta t$$

$$\Lambda_{j}(s) = \Lambda_{0}(s) + \sum_{l=1}^{j-1} \Delta \Lambda_{l}(s)$$
explicit transition
a recursion

Discrete-time Model

Model parameters and latent processes:

 $\begin{aligned} \theta_r \left(s \right), \theta_K \left(s \right) \text{ and } \theta_\Lambda \left(s \right) \\ \beta_\Lambda, \beta_r, \beta_r \\ \phi_\Lambda, \phi_r, \phi_K \end{aligned}$



Discretizing Space

Divide region D into M cells. Rescaling and assuming homogeneous intensity in each cell. We obtain (with r(m), k(m) average growth rate and cumulative carrying capacity):

$$\frac{d\Lambda(t,m)}{dt} = r(m)\Lambda(t,m)[1 - \frac{\Lambda(t,m)}{K(m)}]$$

with induced transition

$$\Delta\Lambda_{j}(m) = r(m)\Lambda_{j-1}(m) \left[1 - \frac{\Lambda_{j-1}(m)}{K(m)}\right] \Delta t.$$

The joint likelihood (product Poisson):

$$\prod_{j=1}^{J} \left[\exp\left(-\sum_{m=1}^{M} \Delta \Lambda_{j}(m) A(m)\right) \prod_{m=1}^{M} \Delta \Lambda_{j}(m)^{n_{jm}} \right]$$
$$\cdot \exp\left(-\sum_{m=1}^{M} \Lambda_{0}(m) A(m)\right) \prod_{m=1}^{M} \Lambda_{0}(m)^{n_{0m}},$$