Poverty vs. region (east, west)

Lecture 20 - Introduction to Multiple Regression

Sta102 / BME102

Colin Rundel

November 16, 2015

str(poverty)

##	da	ata.frame': 51 obs. of	8	3 varia	ables	3:										
##	\$	State	:	Factor	cw/	51 le	vels	"Alal	bama"	,"Ala	ska",	:	123	34	56	5 7
##	\$	Metropolitan.Residence	::	num 5	55.4	65.6	88.2	52.5	94.4	84.5	87.7	80.	1 100	89	.3	
##	\$	Caucasian	:	num 7	71.3	70.8	87.7	81 77	7.5 9	0.2 8	5.4 7	6.3	36.2	80.	6.	
##	\$	Graduates	:	num 7	79.9	90.6	83.8	80.9	81.1	88.7	87.5	88.	7 86	84.	7.	
##	\$	Poverty	:	num 1	14.6	8.3 1	3.3 1	L8 12	.8 9.	4 7.8	8.1	16.8	12.1			
##	\$	PercFemaleHH	:	num 1	14.2	10.8	11.1	12.1	12.6	9.6	12.1	13.1	18.9	9 12		
##	\$	region2	:	Factor	rw/	2 lev	'els	'east'	","we	st":	122	2 2 2	2 1	1 1	1	
##	\$	region4	:	Factor	rw/	4 lev	rels '	'nortl	heast	","mi	dwest	;",	: 4 3	334	43	3 3

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Linear regression with categorical predictors

Poverty vs. region (east, west)

by	<pre>(poverty\$Poverty, pov function(x) c(mean=m</pre>	<i>v</i> 0	-	,sd=sd(x),iqr=IQR(x)))
## ## ##	poverty\$region2: eas mean med 11.170370 10.300000	sd 3.085427		
##	poverty\$region2: wes	t		
## ##	mean med 11.550000 10.700000	sd 3.168459	iqr 4.000000	

Linear regression with categorical predictors

Poverty vs. region (east, west)

summary(lm(Poverty ~ region2, data=poverty))

```
##
## Call:
## lm(formula = Poverty ~ region2, data = poverty)
##
## Residuals:
## Min 1Q Median 3Q
                                   Max
## -5.5704 -2.2000 -0.8704 2.0398 6.4500
##
## Coefficients:
## Estimate Std. Error t value Pr(>|t|)
## (Intercept) 11.1704 0.6013 18.576 <2e-16 ***
## region2west 0.3796 0.8766 0.433 0.667
## ---
## Signif. codes:
## 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.125 on 49 degrees of freedom
## Multiple R-squared: 0.003813, Adjusted R-squared: -0.01652
## F-statistic: 0.1875 on 1 and 49 DF, p-value: 0.6669
```

Linear regression with categorical predictors

Poverty vs. region (east, west)

% poverty =
$$11.17 + 0.38 \times \mathbb{1}_{west}$$

- Explanatory variable: region
- Reference level: east
- Intercept: estimated average % poverty in eastern states is 11.17%
 - This is the value we get if we plug in 0 for the explanatory variable
- *Slope:* estimated average % poverty in western states is 0.38% higher than eastern states.
 - Estimated average % poverty in western states is 11.17 + 0.38 = 11.55%.
 - This is the value we get if we plug in 1 for the explanatory variable

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Linear regression with categorical predictors

Poverty vs. Region (Northeast, Midwest, West, South)

summary(lm(Poverty ~ region4, data=poverty))

```
##
## Call:
## lm(formula = Poverty ~ region4, data = poverty)
##
## Residuals:
##
     Min
             1Q Median
                           ЗQ
                                 Max
## -6.359 -1.559 -0.025 1.574 6.508
##
## Coefficients:
##
                 Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                   9.5000
                            0.8682 10.943 1.62e-14 ***
## region4midwest 0.0250
                             1.1485 0.022 0.982725
## region4west
                   1.7923
                              1.1294 1.587 0.119220
## region4south
                   4.1588
                              1.0736 3.874 0.000331 ***
## ---
## Signif. codes:
## 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 2.604 on 47 degrees of freedom
## Multiple R-squared: 0.3361, Adjusted R-squared: 0.2938
## F-statistic: 7.933 on 3 and 47 DF, p-value: 0.0002205
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```

Linear regression with categorical predictors

Poverty vs. Region (Northeast, Midwest, West, South)

Which region (Northeast, Midwest, West, South) is the reference level?

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	9.50	0.87	10.94	0.00
region4midwest	0.03	1.15	0.02	0.98
region4west	1.79	1.13	1.59	0.12
region4south	4.16	1.07	3.87	0.00

Interpretation:

- Predict 9.50% poverty in Northeast
- Predict 9.53% poverty in Midwest
- Predict 11.29% poverty in West
- Predict 13.66% poverty in South

Linear regression with categorical predictors

Poverty vs. Region (Northeast, Midwest, West, South)

by(poverty\$Poverty, poverty\$region4, function(x) admonstration(x) admonstration(x)

function(x) c(mean=mean(x),med=median(x),sd=sd(x),iqr=IQR(x)))

##	poverty\$r	region4: no:	rtheast		
##	mean	med	sd	iqr	
##	9.500000	9.600000 2	.381701 2	.500000	
##					
##	poverty\$r	region4: mi	dwest		
##	mean	med	sd	iqr	
##	9.525000	9.550000 1	.415579 1	.550000	
##					
##	poverty\$r	region4: we	st		
##	mear	n med	S	d iqr	
##	11.292308	3 10.800000	2.64747	1 3.400000)
##					
##	poverty\$r	region4: so	uth		
##	mear	n med	S	d iqr	
##	13.658824	4 14.200000	3.23343	1 3.900000)

Poverty vs. Region (Northeast, Midwest, West, South)

summary(aov(poverty\$Poverty~poverty\$region4))

##	Df	Sum Sq	Mean Sq F	value	Pr(>F)
<pre>## poverty\$region4</pre>	3	161.4	53.81	7.933	0.00022 ***
## Residuals	47	318.8	6.78		
##					
## Signif. codes:					
## 0 '***' 0.001 '	**'	0.01 '*	' 0.05 '.	' 0.1	' ' 1

Poverty vs. Region (Northeast, Midwest, West, South)

summary(lm(Poverty ~ region4, data=poverty))

	Call: lm(formula = Poverty ~	region4, dat	a = poverty)			
##			1			
##	Residuals:					
##	Min 1Q Median	3Q Max				
##	-6.359 -1.559 -0.025	.574 6.508				
##						
##	Coefficients:					
##	Estimate	e Std. Error	t value Pr(> t)			
	(Intercept) 9.5000			***		
	region4midwest 0.0250					
	region4west 1.7923					
##	region4south 4.1588	1.0736	3.874 0.000331	***		
##						
	Signif. codes:					
##	0 '***' 0.001 '**' 0.01	'*' 0.05 '.	' 0.1 ' ' 1			
##						
	Residual standard error		0			
	Multiple R-squared: 0	- 0	-			
##	F-statistic: 7.933 on 3	and 47 DF,	p-value: 0.00022	205		
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Multiple predictors in a linear model

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Weights of books

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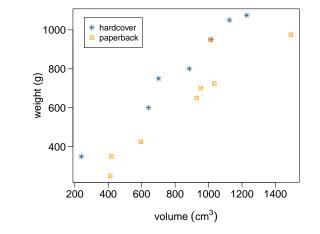
	weight (g)	volume (cm ³)	cover
1	800	885	hc
2	950	1016	hc
3	1050	1125	hc
4	350	239	hc
5	750	701	hc
6	600	641	hc
7	1075	1228	hc
8	250	412	pb
9	700	953	pb
10	650	929	pb
11	975	1492	pb
12	350	419	pb
13	950	1010	pb
14	425	595	pb
15	725	1034	pb



Multiple predictors in a linear model

Weights of hard cover and paperback books

Can you identify a trend in the relationship between volume and weight of hardcover and paperback books?



Modeling weights of books using volume and cover type

book_mlr = lm(weight ~ volume + cover, data = allbacks)
summary(book_mlr)

Coefficients:

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##		Estimate	Std.	Error	t	value	Pr(> t)	
##	(Intercept)	197.96284	59.	19274		3.344	0.005841	**
##	volume	0.71795	0.	06153	1	1.669	6.6e-08	***
##	cover:pb	-184.04727	40.	49420	-	4.545	0.000672	***
##	-							
##								

Residual standard error: 78.2 on 12 degrees of freedom
Multiple R-squared: 0.9275, Adjusted R-squared: 0.9154
F-statistic: 76.73 on 2 and 12 DF, p-value: 1.455e-07

Linear model

	Estimate	Std. Error	t value	$\Pr(> t)$
(Intercept)	197.96	59.19	3.34	0.01
volume	0.72	0.06	11.67	0.00
cover:pb	-184.05	40.49	-4.55	0.00

 $\widetilde{\texttt{weight}} = 197.96 + 0.72 \text{ volume} - 184.05 \text{ cover:pb}$

For hardcover books: plug in 0 for cover

For paperback books: plug in 1 for cover

 $\widehat{\text{weight}}$ = 197.96 + 0.72 volume - 184.05 × 1 = 13.91 + 0.72 volume

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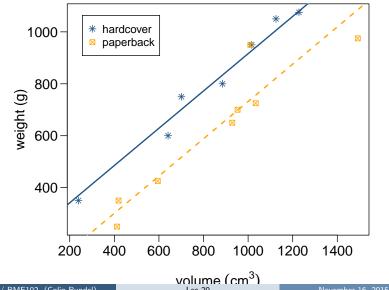
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Multiple predictors in a linear model

Visualising the linear model



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Multiple predictors in a linear model

Interpretation of the regression coefficients

	Estimate	Std. Error	t value	$\Pr(> t)$
(Intercept)	197.96	59.19	3.34	0.01
volume	0.72	0.06	11.67	0.00
cover:pb	-184.05	40.49	-4.55	0.00

- *Slope of volume:* <u>All else held constant</u>, for each 1 cm³ increase in volume we would expect weight to increase on average by 0.72 grams.
- *Slope of cover:* <u>All else held constant</u>, the model predicts that paperback books weigh 184 grams less than hardcover books, on average.
- *Intercept:* Hardcover books with no volume are expected on average to weigh 198 grams.
 - Obviously, the intercept does not make sense in context. It only serves to adjust the height of the line.

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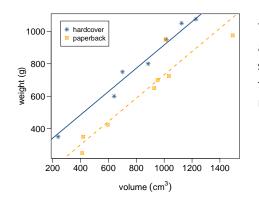
Prediction

What is the correct calculation for the predicted weight of a paperback book that has a volume of 600 cm^3 ?

	Estimate	Std. Error	t value	$\Pr(> t)$
(Intercept)	197.96	59.19	3.34	0.01
volume	0.72	0.06	11.67	0.00
cover:pb	-184.05	40.49	-4.55	0.00

A note on interactions

$$\widetilde{\texttt{weight}} = 197.96 + 0.72 \text{ volume} - 184.05 \text{ cover:pb}$$



This model assumes that hardcover and paperback books have the same slope for the relationship between their volume and weight. If this isn't reasonable, then we would include an "interaction" variable in the model.

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Multiple predictors in a linear model

Example of an interaction

summary(lm(weight ~ volume + cover + volume:cover, data = allbacks))

```
## Coefficients:
```

```
##
                   Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                   161.58654
                               86.51918
                                         1.868
                                                  0.0887
## volume
                     0.76159
                                0.09718
                                         7.837 7.94e-06 ***
## coverpb
                  -120.21407 115.65899 -1.039
                                                  0.3209
                                0.12802 -0.592
## volume:coverpb
                   -0.07573
                                                  0.5661
##
## Residual standard error: 80.41 on 11 degrees of freedom
## Multiple R-squared: 0.9297, Adjusted R-squared: 0.9105
```

F-statistic: 48.5 on 3 and 11 DF, p-value: 1.245e-06

 $\widehat{\texttt{weight}} = 161.58 + 0.76 \text{ volume} - 120.21 \text{ cover:pb} - 0.076 \text{ volume} \times \texttt{cover:pb}$

Multiple predictors in a linear model

Example of an interaction - interpretation

	Estimate	Std. Error	t value	$\Pr(> t)$
(Intercept)	161.5865	86.5192	1.87	0.0887
volume	0.7616	0.0972	7.84	0.0000
coverpb	-120.2141	115.6590	-1.04	0.3209
volume:coverpb	-0.0757	0.1280	-0.59	0.5661

Regression equations for hardbacks:

$$\widehat{\text{weight}} = 161.58 + 0.76 \text{ volume} - 120.21 \times \text{O} - 0.076 \text{ volume} \times \text{O}$$

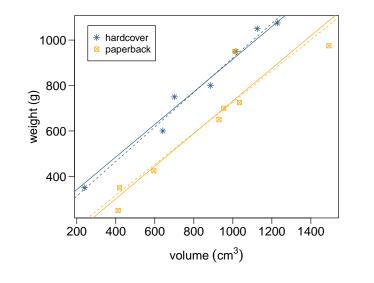
= 161.58 + 0.76 volume

Regression equations for paperbacks:

$$\widetilde{\texttt{weight}} = 161.58 + 0.76 \text{ volume} - 120.21 imes 1 - 0.076 \text{ volume} imes 1$$

= 41.37 + 0.686 volume

Example of an interaction - Results



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Another look at R

For a linear regression we have defined the correlation coefficient to be

$$R = \operatorname{Cor}(X, Y) = \frac{\operatorname{Cov}(X, Y)}{\sigma_X \sigma_Y}$$

This definition works fine for the simple linear regression case where X and Y are numeric variables, but does not work for regression with a categorical predictor or for multiple regression.

A more useful, and equivalent, definition is $R = \text{Cor}(Y, \hat{Y})$, which will work for all regression examples we will see in this class.

 R^2 and Adjusted R^2

Another look at R^2

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So how can we claim that R^2 is a measure of variability "explained" by the model?

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Remember, in an ANOVA we can partition total uncertainty into model (group) uncertainty and residual (error) uncertainty.

$$SST = SSG + SSE$$
$$\sum_{i=1}^{k} \sum_{j=1}^{n_i} (y_{ij} - \bar{y})^2 = \sum_{i=1}^{k} \sum_{j=1}^{n_i} (\bar{y}_i - \bar{y})^2 + \sum_{i=1}^{k} \sum_{j=1}^{n_i} (y_{ij} - \bar{y}_i)^2$$

For a regression we can do the same thing, just replacing \bar{y}_i with \hat{y}_i

$$SST = SSR + SSE$$

 $\sum_{i=1}^{n} (y_i - \bar{y})^2 = \sum_{i=1}^{n} (\hat{y}_i - \bar{y})^2 + \sum_{i=1}^{n} (y_i - \hat{y}_i)^2$

 R^2 and Adjusted R^2

Another look at R, cont.

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Claim: $Cor(X, Y) = Cor(Y, \hat{Y})$ Remember: $Cor(X, Y) = \frac{Cov(X, Y)}{\sigma_X \sigma_Y}$, $\hat{Y} = b_0 + b_1 X$, $Var(aX + b) = a^2 Var(X)$, Cov(aX + b, Y) = a Cov(X, Y)

$$\operatorname{Cor}(Y, \hat{Y}) = \frac{\operatorname{Cov}(Y, \hat{Y})}{\sqrt{\operatorname{Var}(Y)\operatorname{Var}(\hat{Y})}}$$
$$= \frac{\operatorname{Cov}(Y, b_0 + b_1 X)}{\sqrt{\sigma_Y^2 \operatorname{Var}(b_0 + b_1 X)}}$$
$$= \frac{b_1 \operatorname{Cov}(Y, X)}{\sigma_Y \sqrt{b_1^2 \operatorname{Var}(X)}}$$
$$= \frac{b_1 \operatorname{Cov}(Y, X)}{b_1 \sigma_Y \sigma_X}$$
$$= \operatorname{Cor}(X, Y)$$

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R^2 and Adjusted R^2

Another look at R^2

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After a fair bit of algebra we can show that,

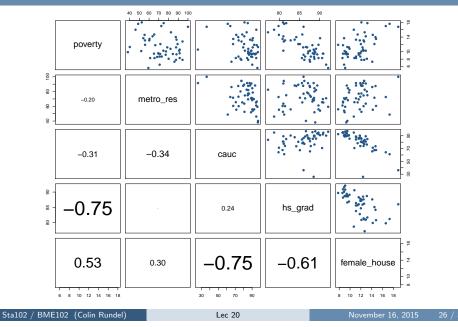
$$R^{2} = \operatorname{Cor}(Y, \hat{Y})^{2} = \frac{\operatorname{Cov}(Y, \hat{Y})^{2}}{\operatorname{Var}(Y)\operatorname{Var}(\hat{Y})}$$

$$=\frac{\sum_{i=1}^{n}(\hat{Y}_{i}-\bar{Y})^{2}}{\sum_{i=1}^{n}(Y_{i}-\bar{Y})^{2}}=\frac{SSR}{SST}$$

$$=\frac{SST-SSE}{SST}=1-\frac{SSE}{SST}$$

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Revisit: Modeling poverty

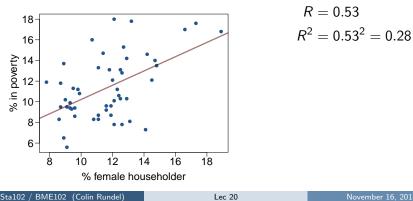


 R^2 and Adjusted R^2

Predicting poverty using % female householder

summary(lm(poverty ~ female_house, data = poverty))

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	3.31	1.90	1.74	0.09
emale_house	0.69	0.16	4.32	0.00
emale_nouse	0.09	0.10	4.32	



R^2 and Adjusted R^2

Another look at R^2 - from last week

anova(lm(poverty ~ female_house, data = poverty))

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
female_house	1	132.57	132.57	18.68	0.00
Residuals	49	347.68	7.10		
Total	50	480.25			

$$SS_{Tot} = \sum (y - \bar{y})^2 = 480.25 \rightarrow \text{total variability}$$

$$SS_{Err} = \sum e_i^2 = 347.68 \rightarrow \text{unexplained variability}$$

$$SS_{Reg} = SS_{Total} - SS_{Error} \rightarrow \text{explained variability}$$

$$= 480.25 - 347.68 = 132.57$$

$$R^{2} = \frac{\text{explained variability}}{\text{total variability}} = \frac{132.57}{480.25} = 0.28 \text{ v}$$

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R^2 and Adjusted R^2

Predicting poverty using % female hh + % cauc

pov_mlr = lm(poverty ~ female_house + cauc, data = poverty)
summary(pov_mlr)

	Estimate	Std. Error	t value	$\Pr(> t)$
(Intercept)	-2.58	5.78	-0.45	0.66
female_house	0.89	0.24	3.67	0.00
cauc	0.04	0.04	1.08	0.29

anova(pov_mlr)

		Df	Sum Sq	Mean Sq	F value	Pr(>F)	
	female_house	1	132.57	132.57	18.74	0.00	
	cauc	1	8.21	8.21	1.16	0.29	
	Residuals	48	339.47	7.07			
	Total	50	480.25				
			d variability variability	$=\frac{132.57}{480}$	$\frac{+8.21}{.25} =$	0.29	
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 R^2 and Adjusted R^2

Adjusted R^2

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Adjusted R^2

$$R_{adj}^2 = 1 - \left(rac{SS_{Error}}{SS_{Total}} imes rac{n-1}{n-k-1}
ight)$$

where n is the number of cases and k is the number of predictors (explanatory variables excluding the intercept) in the model.

- Because k is never negative, R_{adj}^2 will always be less than or equal to R^2 .
- R²_{adj} applies a penalty for the number of predictors included in the model.

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• Therefore, we prefer models with higher R_{adi}^2

R^2 vs. adjusted R^2

	R^2	Adjusted R^2
Model 1 (poverty vs. female_house)	0.2760	0.2613
Model 2 (poverty vs. female_house + cauc)	0.2931	0.2637

- We would like to have some criteria to evaluate if adding an additional variable makes a difference in the explanatory power of the model.
- When any variable is added to the model R^2 increases.
- Adjusted R^2 is based on R^2 but it penalizes the addition of variables.

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R^2 and Adjusted R^2

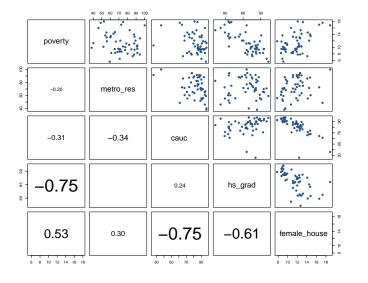
Calculate adjusted R^2

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
female_house	1	132.57	132.57	18.74	0.0001
cauc	1	8.21	8.21	1.16	0.2868
Residuals	48	339.47	7.07		
Total	50	480.25			

$$R_{adj}^{2} = 1 - \left(\frac{SS_{Error}}{SS_{Total}} \times \frac{n-1}{n-k-1}\right)$$

= $1 - \left(\frac{339.47}{480.25} \times \frac{51-1}{51-2-1}\right)$
= $1 - \left(\frac{339.47}{480.25} \times \frac{50}{48}\right)$
= $1 - 0.74$
= 0.26

We saw that adding the variable cauc to the model only marginally increased adjusted R^2 , i.e. did not add much useful information to the model. Why?



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Collinearity between explanatory variables (cont.)

• Two predictor variables are said to be collinear when they are correlated, and this *collinearity* (also called *multicollinearity*) complicates model estimation.

Remember: Predictors are also called explanatory or <u>independent</u> variables, so ideally they should be independent of each other.

- We don't like adding predictors that are associated with each other to the model, because often times the addition of such variable brings nothing to the table. Instead, we prefer the simplest model that explains as much as possible the most *parsimonious* model.
- In addition, inclusion of collinear variables can result in biased estimates of the slope parameters.
- While it's impossible to avoid all collinearity, often experiments are designed to control for correlated predictors.

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