Lecture 10 - Inference for One Mean

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Confidence Intervals in the Real World

Lets assume we are interested in understanding the blood pressure of high school athletes, we collect a large sample (n=200) from the population and record each student's blood pressure in mmHg. Here we do not know either the population mean (μ) or variance (σ^2).

We want to construct a 95% confidence interval based on the observed sample average, which we do by calculating:

$$CI_{95\%} = \bar{X} \pm Z^* SE$$
$$= \bar{X} \pm 1.96 \frac{\sigma}{\sqrt{n}}$$

Missing σ

When working with real samples the population standard deviation (σ) is almost never known, we address this by plugging in the sample standard deviation when calculating the standard error. However, when we do this it changes the sampling distribution.

• We estimate the standard error using the sample standard deviation, this adds uncertainty to inference process.

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- We estimate the standard error using the sample standard deviation, this adds uncertainty to inference process.
- Our new sampling distribution is still symmetric and roughly bell shaped, but its tails are *thicker* than the normal distribution.
- Observations are more likely to fall beyond two SDs from the mean than with the normal distribution.

t distribution



First described by William Gosset ...

- Oxford Graduate with a degree in Chemistry and Mathematics
- Hired by the Guinness Brewery in 1899
- Spent 1906 1907 studying with Karl Pearson
- Published "The probable error of a mean" in 1908 under the pseudonym "A. Student"
- Much of his work was promoted by R.A. Fisher



Properties of the *t* distribution

- it is centered at zero*, like the standard normal (*Z*) distribution.
- it has a single parameter, *df* (*degrees of freedom*), which determines the thickness of the tails.



• as *df* increases the *t* distribution converges to the *Z* distribution.

Finding probabilities

As before we can find any probability we are interested by knowing how to calculate the area under the tail of the *t* distribution. For example, if we want to know $P(T_{df=19} > 1.16)$ then we can use:

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- R: 1-pt(1.16,df=19) ## [1] 0.1302092
- App: (https://gallery.shinyapps.io/dist_calc/):



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Locate the *T* value on the appropriate *df* row, obtain the probability from the corresponding column heading (one or two tail).

one tail	0.100	0.050	0.025	0.010	0.005
two tails	0.200	0.100	0.050	0.020	0.010
df 1	3.08	6.31	12.71	31.82	63.66
2	1.89	2.92	4.30	6.96	9.92
3	1.64	2.35	3.18	4.54	5.84
:		:	÷	:	
17	1.33	1.74	2.11	2.57	2.90
18	1.33	1.73	2.10	2.55	2.88
19	1.33	1.73	2.09	2.54	2.86
20	1.33	1.72	2.09	2.53	2.85
÷	:	:	:	÷	
400	1.28	1.65	1.97	2.34	2.59
500	1.28	1.65	1.96	2.33	2.59
∞	1.28	1.64	1.96	2.33	2.58

Finding probabilities - upper tail

Using the table below find:

 $P(T_{df=19} > 1.16)$

one tail	0.100	0.050	0.025	0.010	0.005
two tails	0.200	0.100	0.050	0.020	0.010
df 17	1.33	1.74	2.11	2.57	2.90
18	1.33	1.73	2.10	2.55	2.88
19	1.33	1.73	2.09	2.54	2.86
20	1.33	1.72	2.09	2.53	2.85

Finding probabilities - upper tail

Using the table below find:

 $P(T_{df=19} < -2)$

1.1						
	one tail	0.100	0.050	0.025	0.010	0.005
	two tails	0.200	0.100	0.050	0.020	0.010
	df 17	1.33	1.74	2.11	2.57	2.90
	18	1.33	1.73	2.10	2.55	2.88
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CLT vs. t

From the Central Limit Distribution we have,

$$\bar{X} \sim N(\mu, \sigma^2/n)$$

 $\frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1)$

Since σ is unknown we must use s which results in the following

$$\frac{\bar{X} - \mu}{s/\sqrt{n}} \sim t_{df=n-1}$$

Implications of t distribution for Confidence intervals

Confidence intervals are always of the form

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If our point estimate is a sample mean and σ is unknown, then our sample mean follows a *t* distribution (and not a *Z* distribution), the critical value is then given by t_{df}^{\star} (as opposed to a *Z*^{*}) and the *SE* is s/\sqrt{n} (and not σ/\sqrt{n}).

$$\bar{X} \pm t_{df}^{\star} imes rac{\mathsf{S}}{\sqrt{n}}$$

Finding the critical t (t*)



$$n = 10, df = 10 - 1 = 9$$

 t^* is at the intersection of row df = 9 and two tails column 0.05.

one tail	0.100	0.050	0.025	0.010	0.005
two tails	0.200	0.100	0.050	0.020	0.010
df 6	1.44	1.94	2.45	3.14	3.71
7	1.41	1.89	2.36	3.00	3.50
8	1.40	1.86	2.31	2.90	3.36
9	1.38	1.83	2.26	2.82	3.25
10	1.37	1.81	2.23	2.76	3.17

We would like to calculate a 95% confidence interval for the average rental price of an apartment in Durham. We sample craigslist and find

 $\mathsf{Rent} = \{625, 733, 895, 929, 775, 1349, 599, 749, 1020, 799, \\705, 665, 1282, 1143, 1209, 500, 1495, 1076, 975, 879\}$

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 $\bar{X} = 920.1$ s = 271 n = 20 SE = s/ \sqrt{n} = 60.6

Confidence Intervals as Inference

In 2001 the average GPA of students at Duke University was 3.37. Last semester 63 introductory statistics students reported their GPA on an in class survey. The mean was 3.58, and the standard deviation 0.53. A histogram of the data is shown below.



Assuming that this sample is random and representative of all Duke students, do these data provide convincing evidence that the average GPA of Duke students has changed over the last decade and a half? Your friend has collected some data as part of a summer REU they collected tadpoles from a local different stream and measured their lengths. From the stream they were able collect 50 tadpoles which had an average length 2.3 cm and a standard deviation of 0.2 cm.

They argue that since it is well known that the distribution of tadpole lengths is normal they should be able to use the Z distribution when constructing their confidence intervals for the average lengths. Are they correct? If not, how serious a mistake are they making? (Construct the CIs both ways for both steams and compare)

Recap: Inference using CIs for sample means

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- independence of observations (often verified by a random sample, and if sampling without replacement, n < 10% of population)
- sample size is large or population not overly skewed or heavy/light tailed

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Confidence interval:

$$\bar{X} \pm t_{df}^{\star} \frac{s}{\sqrt{n}}$$
, where $df = n - 1$

Hypothesis Tests for one mean

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- We examine how likely our data (or something more extreme) is under this assumption, and use that as evidence against the null hypothesis (and hence for the alternative).

In 2001 the average GPA of students at Duke University was 3.37. Last semester Duke students in a Stats class were surveyed and ask for their current GPA. This survey had 63 respondents and yielded an average GPA of 3.56 with a standard deviation of 0.31.

Assuming that this sample is random and representative of all Duke students, do these data provide convincing evidence that the average GPA of Duke students has *changed* over the last decade?

Setting the hypotheses

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$$H_0: \mu = 3.37$$

• We test the claim that average GPA has changed.

$$H_{\rm A}: \mu \neq 3.37 \tag{24}$$

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- Large p-value (> α) we claim it is likely to observe these data if the null hypothesis were true, and therefore *do not reject* H₀.
- We never accept H_0 since we're not in the business of trying to prove it. We just want to know if the data provide *convincing* evidence against H_0 .

What is a p-value:

- The probability of the observed data (sample statistic) or something more extreme in favor of the null hypothesis given the null hypothesis is true.
- Indirect evidence against H_0 .

What a p-value *isn't*:

- A p-value is not the probably H_0 is true
- A p-value is not the probably H_A is false

Back to the GPA example, in order to perform inference on these data we need to use the CLT, and therefore we need to check the conditions: Back to the GPA example, in order to perform inference on these data we need to use the CLT, and therefore we need to check the conditions:

- 1. Independence:
 - We have already assumed this sample is random.
 - Assume sampling without replacement, but 63 < 10% of all current Duke students.

 \Rightarrow it appears reasonable to assume that GPA of one student in this sample is independent of another.

Conditions for inference - GPA

2. Sample size / skew: The distribution appears to be slightly left skewed (but not extremely) and n = 63 so we will assume that the sampling distribution of the sample means should be nearly normal by the CLT.



p-value - probability of observing data at least as favorable to H_A as our current data set, if in fact H_0 is true (the true population mean $\mu = 3.37$).

In this case because we are not making any claims about GPAs going up or down, we need to consider GPA changes in both directions. E.g. a sample average GPA of 3.18 is just as much in favor of H_A as a sample average GPA of 3.56.

Drawing a Conclusion / Inference

$$p-value = 4.2 \times 10^{-6}$$

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If the true average GPA of Duke students is 3.37, there is approximately a 4.2×10^{-6} chance of observing a random sample of 63 Duke students with an average GPA of 3.56 and above or 3.18 and below.

• This is a very small probability, it seems very unlikely that a 3.56 sample average GPA could have happened by chance.

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- This is a very small probability, it seems very unlikely that a 3.56 sample average GPA could have happened by chance.
- Since the p-value is small (lower than 5%) we reject H_0 .
- Claim the data provide convincing evidence that Duke students' average GPA has changed since 2001. E.g. the difference between the null value of a 3.37 GPA and observed sample mean of 3.56 GPA is *not due to chance / sampling variability*.

> $H_0: \mu = 8$ $H_A: \mu > 8$

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Yes - Independence \checkmark , Nearly Normal \checkmark

p-value - probability of observing data at least as favorable to H_A as our current data set (a sample mean greater than 9.7), if in fact H_0 was true (the true population mean is 8).

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College Applications - Making a decision

p – *value* < 0.005

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- This is a very small probability, it seems very unlikely that a sample mean of 9.7 or more schools is likely to happen simply by chance.
- Since p-value is *low* (lower than 5%) we *reject* H_0 .
- The data provide convincing evidence that Duke students apply on average to more than 8 schools.

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We construct a 95% confidence interval using $t^*_{df=205} \approx t^*_{df=200} =$ 1.97,

Regardless of the sample statistic of interest, all null value hypothesis testing takes exactly the same form.

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- 2. Check assumptions and conditions
- 3. Calculate a *test statistic* and a p-value (draw a picture!)
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- 4. Make a decision, and interpret it in context of the research question
 - If p-value $< \alpha$, reject H_0
 - If p-value > α , do not reject H_0