## Lecture 14 - Inference for proportions

Sta 102 June 7th, 2016

Colin Rundel & Mine Çetinkaya-Rundel

## Inference

## Independent Variable

	None	Categorical (2 levels)	Categorical (>2 levels)	Numerical
Numerical	Test of One Mean	Test of Two Means	ANOVA	Regression
Categorical (2 levels)	Test of One Proportion	Test of Two Proportions	$\chi^2$ - Test of Independence	Logistic Regression
Categorical (>2 levels)	$\chi^2$ - GoF	χ <sup>2</sup> - Test of Independence	$\chi^2$ - Test of Independence	Multinomial Regression

## Inference for a single proportion

Two scientists want to know if a certain drug is effective against high blood pressure. The first scientist wants to give the drug to 1000 people with high blood pressure and see how many of them experience lower blood pressure levels. The second scientist wants to give the drug to 500 people with high blood pressure, and not give the drug to another 500 people with high blood pressure, and see how many in both groups experience lower blood pressure levels. Which is the better way to test this drug?

(a) All 1000 get the drug

(b) 500 get the drug, 500 don't

The GSS asks the same question, below is the distribution of responses from the 2010 survey:

All 1000 get the drug	99
500 get the drug 500 don't	571
Total	670

We would like to estimate the proportion of all Americans who have good intuition about experimental design, i.e. would answer "500 get the drug 500 don't".

What are the parameter of interest and the point estimate?

We would like to estimate the proportion of all Americans who have good intuition about experimental design, i.e. would answer "500 get the drug 500 don't".

What are the parameter of interest and the point estimate?

 Parameter of interest: Proportion of all Americans who have good intuition about experimental design.

p (a population proportion)

We would like to estimate the proportion of all Americans who have good intuition about experimental design, i.e. would answer "500 get the drug 500 don't".

What are the parameter of interest and the point estimate?

 Parameter of interest: Proportion of all Americans who have good intuition about experimental design.

p (a population proportion)

 Point estimate: Proportion of sampled Americans who have good intuition about experimental design.

 $\hat{p}$  (a sample proportion)

What percent of all Americans have a good intuition about experimental design, i.e. would answer "500 get the drug 500 don't"?

What percent of all Americans have a good intuition about experimental design, i.e. would answer "500 get the drug 500 don't"?

 We can answer this research question using a confidence interval, which we know is has the form

point estimate  $\pm$  critical value  $\times$  standard error

What percent of all Americans have a good intuition about experimental design, i.e. would answer "500 get the drug 500 don't"?

 We can answer this research question using a confidence interval, which we know is has the form

point estimate  $\pm$  critical value  $\times$  standard error

What we need to know then is

 $SE_{\hat{p}} = ?$  CV = ?

It may be useful to instead think about  $K = n\hat{p}$ , what distribution will that have?

It may be useful to instead think about  $K = n\hat{p}$ , what distribution will that have?

 $K \sim Binom(n, p)$ 

It may be useful to instead think about  $K = n\hat{p}$ , what distribution will that have?

 $K \sim Binom(n, p)$ 

$$n\hat{p} \approx X \sim N\left(\mu = np, \, \sigma = \sqrt{np(1-p)}\right)$$

It may be useful to instead think about  $K = n\hat{p}$ , what distribution will that have?

 $K \sim Binom(n, p)$ 

$$n\hat{p} \approx X \sim N\left(\mu = np, \, \sigma = \sqrt{np(1-p)}\right)$$

We can then find the distribution of  $\hat{p}$  by dividing by n,

$$\hat{p} \approx \frac{X}{n} \sim N\left(\mu = p, \, \sigma = \sqrt{\frac{p(1-p)}{n}}\right)$$

A sample proportion will have a sampling distribution that is approximately normal with,

$$\hat{p} \sim N\left(\mu = p, \ \sigma = SE = \sqrt{\frac{p(1-p)}{n}}\right)$$

A sample proportion will have a sampling distribution that is approximately normal with,

$$\hat{p} \sim N\left(\mu = p, \ \sigma = SE = \sqrt{\frac{p(1-p)}{n}}\right)$$

But of course this is true only under certain conditions ... any guesses?

A sample proportion will have a sampling distribution that is approximately normal with,

$$\hat{p} \sim N\left(\mu = p, \ \sigma = SE = \sqrt{\frac{p(1-p)}{n}}\right)$$

But of course this is true only under certain conditions ... any guesses?

Assumptions/conditions:

- 1. Independence:
  - Random sample
  - 10% condition: If sampling without replacement, n < 10% of the population.
- 2. Normality: At least 10 successes  $(np \ge 10)$  and 10 failures  $(n(1-p) \ge 10)$ .

Given: n = 670,  $\hat{p} = \frac{571}{670} = 0.85$ .

Given: n = 670,  $\hat{p} = \frac{571}{670} = 0.85$ .

Are CLT conditions met?

Given: n = 670,  $\hat{p} = \frac{571}{670} = 0.85$ .

Are CLT conditions met?

 Independence: The sample is random, and 670 < 10% of all Americans, therefore we can assume that one respondent's response is independent of another.

Given: n = 670,  $\hat{p} = \frac{571}{670} = 0.85$ .

Are CLT conditions met?

- Independence: The sample is random, and 670 < 10% of all Americans, therefore we can assume that one respondent's response is independent of another.
- 2. *Success-failure*: 571 people answered correctly (successes) and 99 answered incorrectly (failures), both are greater than 10.

CI =point estimate  $\pm$  margin of error

 $CI = \text{point estimate} \pm \text{margin of error}$ = point estimate  $\pm$  critical value  $\times SE$ 

 $CI = \text{point estimate} \pm \text{margin of error}$ 

= point estimate  $\pm$  critical value  $\times$  *SE* 

 $= \hat{p} \pm Z^{\star} \times SE$ 

 $\begin{aligned} CI &= \text{point estimate} \pm \text{margin of error} \\ &= \text{point estimate} \pm \text{critical value} \times SE \\ &= \hat{p} \pm Z^* \times SE \\ &= 0.85 \pm 1.96 \times \sqrt{\frac{0.85 \times 0.15}{670}} = (0.82, 0.88) \end{aligned}$ 

$$0.01 \geq 1.96 \times \sqrt{\frac{p \times (1-p)}{n}}$$

$$\begin{array}{rcl} 0.01 & \geq & 1.96 \times \sqrt{\frac{p \times (1-p)}{n}} \\ 0.01 & \geq & 1.96 \times \sqrt{\frac{0.85 \times 0.15}{n}} \rightarrow \textit{Using } \hat{p} \textit{ from previous study} \end{array}$$

$$\begin{array}{rcl} 0.01 & \geq & 1.96 \times \sqrt{\frac{p \times (1-p)}{n}} \\ 0.01 & \geq & 1.96 \times \sqrt{\frac{0.85 \times 0.15}{n}} \rightarrow \textit{Using } \hat{p} \textit{ from previous study} \\ 0.01^2 & \geq & 1.96^2 \times \frac{0.85 \times 0.15}{n} \end{array}$$

$$\begin{array}{rcl} 0.01 &\geq& 1.96 \times \sqrt{\frac{p \times (1-p)}{n}} \\ 0.01 &\geq& 1.96 \times \sqrt{\frac{0.85 \times 0.15}{n}} \rightarrow \textit{Using } \hat{p} \textit{ from previous study} \\ 0.01^2 &\geq& 1.96^2 \times \frac{0.85 \times 0.15}{n} \\ n &\geq& \frac{1.96^2 \times 0.85 \times 0.15}{0.01^2} \end{array}$$

$$\begin{array}{rcl} 0.01 &\geq& 1.96 \times \sqrt{\frac{p \times (1-p)}{n}} \\ 0.01 &\geq& 1.96 \times \sqrt{\frac{0.85 \times 0.15}{n}} \rightarrow \textit{Using } \hat{p} \textit{ from previous study} \\ 0.01^2 &\geq& 1.96^2 \times \frac{0.85 \times 0.15}{n} \\ n &\geq& \frac{1.96^2 \times 0.85 \times 0.15}{0.01^2} \\ n &\geq& 4898.04 \end{array}$$
How many people should you sample in order to reduce the margin of error of a 95% confidence interval down to 1%.

 $ME = Z^* \times SE$ 

$$\begin{array}{rcl} 0.01 & \geq & 1.96 \times \sqrt{\frac{p \times (1-p)}{n}} \\ 0.01 & \geq & 1.96 \times \sqrt{\frac{0.85 \times 0.15}{n}} \rightarrow \textit{Using } \hat{p} \textit{ from previous study} \\ 0.01^2 & \geq & 1.96^2 \times \frac{0.85 \times 0.15}{n} \\ n & \geq & \frac{1.96^2 \times 0.85 \times 0.15}{0.01^2} \\ n & \geq & 4898.04 \rightarrow n \textit{ should be at least } 4,899 \end{array}$$

... use  $\hat{p} = 0.5$ . Why?

... use  $\hat{p} = 0.5$ . Why?

• if you don't know any better, 50-50 is a good guess

... use  $\hat{p} = 0.5$ . Why?

- if you don't know any better, 50-50 is a good guess
- p̂ = 0.5 gives the most conservative estimate largest standard error and thus the largest possible sample size.



 $H_A$  is what we are interested in and  $H_0$  represents the status quo, both *must* be about the population parameter of interest.

 $H_A$  is what we are interested in and  $H_0$  represents the status quo, both *must* be about the population parameter of interest.

Parameter of interest: p, point estimate:  $\hat{p}$ 

 $H_A$  is what we are interested in and  $H_0$  represents the status quo, both *must* be about the population parameter of interest.

Parameter of interest: p, point estimate:  $\hat{p}$ 

Hypotheses:

$$H_0: p = 0.8$$
  
 $H_A: p > 0.8$ 

For a test of one proportion our null and alternative hypotheses are about p, therefore when we assume  $H_0$  is true we fix  $p = p_0$ . Hence,

- Standard error:
  - CI: calculate using observed sample proportion:

$$SE = \sqrt{rac{p(1-p)}{n}} \approx \sqrt{rac{\hat{p}(1-\hat{p})}{n}}$$

• HT: calculate using the null value:

$$SE = \sqrt{\frac{p_0(1-p_0)}{n}}$$

For a test of one proportion our null and alternative hypotheses are about p, therefore when we assume  $H_0$  is true we fix  $p = p_0$ . Hence,

- Standard error:
  - CI: calculate using observed sample proportion:

$$SE = \sqrt{rac{p(1-p)}{n}} \approx \sqrt{rac{\hat{p}(1-\hat{p})}{n}}$$

• HT: calculate using the null value:

$$SE = \sqrt{\frac{p_0(1-p_0)}{n}}$$

- Success-failure condition:
  - CI: At least 10 *observed* successes and failures, calculated using the sample proportion, p̂
  - HT: At least 10 *expected* successes and failures, calculated using the null value, *p*<sub>0</sub>

 $H_0: p = 0.80$   $H_A: p > 0.80$ 

$$H_0: p = 0.80$$
  $H_A: p > 0.80$ 

$$SE = \sqrt{\frac{0.80 \times 0.20}{670}} = 0.0154$$

$$H_0: p = 0.80$$
  $H_A: p > 0.80$ 

$$SE = \sqrt{\frac{0.80 \times 0.20}{670}} = 0.0154$$
$$Z = \frac{0.85 - 0.80}{0.0154} = 3.25$$

$$H_0: p = 0.80$$
  $H_A: p > 0.80$ 

$$SE = \sqrt{\frac{0.80 \times 0.20}{670}} = 0.0154$$
$$Z = \frac{0.85 - 0.80}{0.0154} = 3.25$$
$$p - value = 1 - 0.9994 = 0.0006$$



$$H_0: p = 0.80$$
  $H_A: p > 0.80$ 

$$SE = \sqrt{\frac{0.80 \times 0.20}{670}} = 0.0154$$
$$Z = \frac{0.85 - 0.80}{0.0154} = 3.25$$
$$p - value = 1 - 0.9994 = 0.0006$$



Since p-value is small we reject  $H_0$ .

11% of 1,001 Americans responding to a 2006 Gallup survey stated that they have objections to celebrating Halloween on religious grounds. At 95% confidence level, the margin of error for this survey a is  $\pm 3\%$ . A news piece on this study's findings states: "More than 10% of all Americans have objections on religious grounds to celebrating Halloween."

Is this statement justified?

# Inference for difference of two proportion

Scientists predict that global warming may have big effects on the polar regions within the next 100 years. One of the possible effects is that the northern ice cap may completely melt. Would this bother you a great deal, some, a little, or not at all if it actually happened?

- (a) A great deal
- (b) Some
- (c) A little
- (d) Not at all

The GSS asks this question, below is the distribution of responses from the 2010 survey:

A great deal	454
Not a great deal	226
Total	680

The GSS asks this question, below is the distribution of responses from the 2010 survey:

A great deal	454
Not a great deal	226
Total	680

The same question was asked of 88 Duke students, of which 56 said it would bother them a great deal.

We will collapse the data such that we consider only the responses of a great deal and its compliment, not a great deal.

	US	Duke	Total
A great deal	454	56	510
Not a great deal	226	32	258
Total	680	88	768

This is an example of a contingency table (specifically a  $2 \times 2$  contingency table).

	US	Duke	Total
A great deal	454	56	510
Not a great deal	226	32	258
Total	680	88	768

This is an example of a contingency table (specifically a  $2 \times 2$  contingency table).

We are interested in comparing proportion of Duke students who say it would both them a gread deal ( $p_{GD|Duke} = 56/88$ ) to the proportion of all Americans who say it would both them a gread deal ( $p_{GD|US} = 454/680$ ).

Knowing which of the two variables to condition on can be tricky some times.

Ask yourself - which of the two variables is most likely the dependent variable (y) and which is most likely the independent variable (x). In other words, changes in x should *cause* changes in y (not the other way around).

Once we know this then the two proportions of interest are:

$$p_{y_1|x_1}$$
 and  $p_{y_1|x_2}$ 

## Parameter and point estimate

 Parameter of interest: Difference between the proportions of all Duke students and all Americans who would be bothered a great deal by the northern ice cap melting.

 $p_{GD|Duke} - p_{GD|US}$ 

## Parameter and point estimate

 Parameter of interest: Difference between the proportions of all Duke students and all Americans who would be bothered a great deal by the northern ice cap melting.

#### $p_{GD|Duke} - p_{GD|US}$

 Point estimate: Difference between the proportions of sampled Duke students and sampled Americans who would be bothered a great deal by the northern ice cap melting.

 $\hat{p}_{GD|Duke} - \hat{p}_{GD|US}$ 

# Inference for comparing proportions

The details for inference are the same as what we've seen previously,

# Inference for comparing proportions

The details for inference are the same as what we've seen previously,

• CI: point estimate ± critical value × std error

The details for inference are the same as what we've seen previously,

- CI: point estimate  $\pm$  critical value  $\times$  std error
- HT: Test Statistic = point estimate-null value std error
  p-value using sampling distribution.

The details for inference are the same as what we've seen previously,

- CI: point estimate  $\pm$  critical value  $\times$  std error
- HT: Test Statistic = point estimate-null value std error
  p-value using sampling distribution.
- We just need to figure out the appropriate sampling distribution and its parameters..

We can combine that result with the approach we used for the test of two means to find the distribution of  $\hat{p}_1 - \hat{p}_2$ 

$$(\hat{p}_1 - \hat{p}_2) \sim N(\mu = E(\hat{p}_1 - \hat{p}_2), \ \sigma^2 = Var(\hat{p}_1 - \hat{p}_2))$$

We can combine that result with the approach we used for the test of two means to find the distribution of  $\hat{p}_1 - \hat{p}_2$ 

$$(\hat{p}_1 - \hat{p}_2) \sim N(\mu = E(\hat{p}_1 - \hat{p}_2), \ \sigma^2 = Var(\hat{p}_1 - \hat{p}_2))$$

$$E(\hat{p}_1 - \hat{p}_2) = E(\hat{p}_1) - E(\hat{p}_2) \qquad Var(\hat{p}_1 - \hat{p}_2) = Var(\hat{p}_1) + Var(\hat{p}_2)$$
$$= p_1 - p_2 \qquad \qquad = \frac{p_1(1 - p_1)}{n_1} + \frac{p_2(1 - p_2)}{n_1}$$

We can combine that result with the approach we used for the test of two means to find the distribution of  $\hat{p}_1 - \hat{p}_2$ 

$$(\hat{p}_1 - \hat{p}_2) \sim N(\mu = E(\hat{p}_1 - \hat{p}_2), \ \sigma^2 = Var(\hat{p}_1 - \hat{p}_2))$$

$$E(\hat{p}_1 - \hat{p}_2) = E(\hat{p}_1) - E(\hat{p}_2) \qquad Var(\hat{p}_1 - \hat{p}_2) = Var(\hat{p}_1) + Var(\hat{p}_2)$$
$$= p_1 - p_2 \qquad \qquad = \frac{p_1(1 - p_1)}{n_1} + \frac{p_2(1 - p_2)}{n_1}$$

Note - as with the test of two means, this result requires that  $\hat{p}_1$  and  $\hat{p}_2$  are independent.

### 1. Independence within groups:

• The US group is sampled randomly and we're assuming that the Duke group represents a random sample as well.
- The US group is sampled randomly and we're assuming that the Duke group represents a random sample as well.
- *n*<sub>Duke</sub> < 10% of all Duke students and 680 < 10% of all Americans.</li>

- The US group is sampled randomly and we're assuming that the Duke group represents a random sample as well.
- *n*<sub>Duke</sub> < 10% of all Duke students and 680 < 10% of all Americans.</li>

We can assume that the attitudes of Duke students in the sample are independent of each other, and attitudes of US residents in the sample are independent of each other as well.

- The US group is sampled randomly and we're assuming that the Duke group represents a random sample as well.
- *n*<sub>Duke</sub> < 10% of all Duke students and 680 < 10% of all Americans.</li>

We can assume that the attitudes of Duke students in the sample are independent of each other, and attitudes of US residents in the sample are independent of each other as well.

2. *Independence between groups:* The sampled Duke students and the US residents are independent of each other.

- The US group is sampled randomly and we're assuming that the Duke group represents a random sample as well.
- *n*<sub>Duke</sub> < 10% of all Duke students and 680 < 10% of all Americans.</li>

We can assume that the attitudes of Duke students in the sample are independent of each other, and attitudes of US residents in the sample are independent of each other as well.

- 2. *Independence between groups:* The sampled Duke students and the US residents are independent of each other.
- 3. Success-failure:

At least 10 observed successes and 10 observed failures in *both* groups.

Construct a 95% confidence interval for the difference between the proportions of Duke students and Americans who would be bothered a great deal by the melting of the northern ice cap  $(p_{GD|Duke} - p_{GD|US}).$ 

	Duke	US
A great deal	56	454
Not a great deal	32	226
Total	88	680

Construct a 95% confidence interval for the difference between the proportions of Duke students and Americans who would be bothered a great deal by the melting of the northern ice cap  $(p_{GD|Duke} - p_{GD|US}).$ 

	Duke	US
A great deal	56	454
Not a great deal	32	226
Total	88	680

 $\hat{p}_{GD|Duke} = 56/88 = 0.636$  $\hat{p}_{GD|US} = 454/680 = 0.668$  Construct a 95% confidence interval for the difference between the proportions of Duke students and Americans who would be bothered a great deal by the melting of the northern ice cap  $(p_{GD|Duke} - p_{GD|US}).$ 

	Duke	US
A great deal	56	454
Not a great deal	32	226
Total	88	680

 $\hat{p}_{GD|Duke} = 56/88 = 0.636$  $\hat{p}_{GD|US} = 454/680 = 0.668$ 

$$SE \approx \sqrt{\frac{\hat{p}_{GD|Duke}(1-\hat{p}_{GD|Duke})}{n_{Duke}} + \frac{\hat{p}_{GD|US}(1-\hat{p}_{GD|US})}{n_{US}}}$$
$$= \sqrt{\frac{0.636(1-0.636)}{88} + \frac{0.668(1-0.668)}{680}} = 0.0537$$

## CI for difference of proportions, cont.

 $\hat{p}_{GD|Duke} = 0.636$  $\hat{p}_{GD|US} = 0.668$ SE = 0.0537

## CI for difference of proportions, cont.

 $\hat{p}_{GD|Duke} = 0.636$  $\hat{p}_{GD|US} = 0.668$ SE = 0.0537

$$\begin{aligned} CI &= PE \pm CV \times SE \\ &= (\hat{p}_{GD|Duke} - \hat{p}_{GD|US}) \pm Z^* \times \sqrt{\frac{\hat{p}_{GD|Duke}(1 - \hat{p}_{GD|Duke})}{n_{Duke}}} \\ &= (0.636 - 0.668) \pm 1.96 \times 0.0537 \\ &= (-0.138, 0.074) \end{aligned}$$

## CI for difference of proportions, cont.

 $\hat{p}_{GD|Duke} = 0.636$  $\hat{p}_{GD|US} = 0.668$ SE = 0.0537

$$\begin{aligned} CI &= PE \pm CV \times SE \\ &= (\hat{p}_{GD|Duke} - \hat{p}_{GD|US}) \pm Z^* \times \sqrt{\frac{\hat{p}_{GD|Duke}(1 - \hat{p}_{GD|Duke})}{n_{Duke}}} \\ &= (0.636 - 0.668) \pm 1.96 \times 0.0537 \\ &= (-0.138, 0.074) \end{aligned}$$

What conclusion should we draw here?

Just like the other hypothesis tests we have seen thus far, we formulate our null and alternative hypotheses for testing if the proportion of all Duke students who would be bothered a great deal by the melting of the northern ice cap differs from the proportion of all Americans who do as follows,

$$H_0: p_{GD|Duke} = p_{GD|US} \implies p_{GD|Duke} - p_{GD|US} = 0$$
$$H_A: p_{GD|Duke} \neq p_{GD|US} \implies p_{GD|Duke} - p_{GD|US} \neq 0$$

When constructing a confidence interval for a population proportion, we check if the *observed* number of successes and failures are at least 10.

$$n\hat{p} \ge 10$$
  $n(1-\hat{p}) \ge 10$ 

When constructing a confidence interval for a population proportion, we check if the *observed* number of successes and failures are at least 10.

$$n\hat{p} \ge 10$$
  $n(1-\hat{p}) \ge 10$ 

When conducting a hypothesis test for a population proportion, we check if the *expected* number of successes and failures are at least 10.

$$np_0 \ge 10$$
  $n(1-p_0) \ge 10$ 

As such, we don't have a specific null value we can use to calculated the *expected* number of successes and failures in each group *or* the standard error. So, we know the following

As such, we don't have a specific null value we can use to calculated the *expected* number of successes and failures in each group *or* the standard error. So, we know the following

$$P_{GD|Duke} = P_{GD|US}$$
  
 $p_{GD|Duke} = ?$   
 $p_{GD|US} = ?$ 

As such, we don't have a specific null value we can use to calculated the *expected* number of successes and failures in each group *or* the standard error. So, we know the following

$$p_{GD|Duke} = p_{GD|US}$$
  
 $p_{GD|Duke} = ?$   
 $p_{GD|US} = ?$ 

Does this null give us any additional useful information?

Think about the sample proportions as probabilities, what does it mean if

P(GD|Duke) = P(GD|US)

Think about the sample proportions as probabilities, what does it mean if

$$P(GD|Duke) = P(GD|US)$$

If these two probabilities are equal then global warming concern is *independent* of the Duke vs. US grouping. Which means that,

P(GD|Duke) = P(GD|US) = P(GD)

# Pooling

As such, our null hypothesis is equivalent to claiming that our two categorical variables are independent. So when conducting the hypothesis test we assume the null hypothesis to be true, which means we must also assume that the two variables are independent.

Under the assumption of independence our best guess for both  $p_{GD|Duke}$ and  $p_{GD|US}$  will be  $\hat{p}_{GD}$ , which is the sample proportion of *all* respondents (from Duke or US) who answered "A great deal".

We call this value  $\hat{p}_{pooled}$ ,

$$\hat{p}_{pooled} = \frac{\text{\# of successes in } 1 + \text{\# of successes in } 2}{n_1 + n_2} = \frac{n_1 \hat{p}_1 + n_2 \hat{p}_2}{n_1 + n_2}$$

Calculate the estimated pooled proportion of Duke students and Americans who would be bothered a great deal by the melting of the northern ice cap.

	Duke	US	Total
A great deal	56	454	510
Not a great deal	32	226	258
Total	88	680	788

$$\hat{p}_{pooled} = \frac{56 + 454}{88 + 680} = \frac{510}{788} = 0.664$$

Calculate the estimated pooled proportion of Duke students and Americans who would be bothered a great deal by the melting of the northern ice cap.

	Duke	US	Total
A great deal	56	454	510
Not a great deal	32	226	258
Total	88	680	788

$$\hat{p}_{pooled} = \frac{56 + 454}{88 + 680} = \frac{510}{788} = 0.664$$

Which sample proportion  $(\hat{p}_{GD|Duke} \text{ or } \hat{p}_{GD|US})$  is closer to the pooled estimate? Why?

Under the null hypothesis we are stating that  $p_1 = p_2$  which does not uniquely identify either  $p_1$  or  $p_2$ . Therefore we are using the pooled proportion ( $\hat{p}$ ) as our best guess for  $p_1$  and  $p_2$  under the null hypothesis.

For a *confidence interval* we use  $\hat{p}_1$  and  $\hat{p}_2$  to approximate for  $p_1$  and  $p_2$ 

$$SE = \sqrt{\frac{p_1(1-p_1)}{n_1} + \frac{p_2(1-p_2)}{n_2}} \approx \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}$$

While for a *hypothesis test* we use  $\hat{p}_{pooled}$  to approximate for  $p_1$  and  $p_2$ 

$$SE = \sqrt{\frac{p_1(1-p_1)}{n_1} + \frac{p_2(1-p_2)}{n_2}} \approx \sqrt{\frac{\hat{p}_p(1-\hat{p}_p)}{n_1} + \frac{\hat{p}_p(1-\hat{p}_p)}{n_2}}$$
$$= \sqrt{\hat{p}_p(1-\hat{p}_p)\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}$$

$$\hat{p}_{pooled} = 0.664, \quad n_1 = 88, \quad n_2 = 680$$

$$\hat{p}_{pooled} = 0.664, \quad n_1 = 88, \quad n_2 = 680$$

$$SE = \sqrt{\hat{p}_{\rho}(1-\hat{p}_{\rho})\left(\frac{1}{n_{1}}+\frac{1}{n_{2}}\right)} = \sqrt{0.664(1-0.664)\left(\frac{1}{88}+\frac{1}{680}\right)} = 0.0535$$

$$\hat{p}_{pooled} = 0.664, \quad n_1 = 88, \quad n_2 = 680$$

$$SE = \sqrt{\hat{p}_p (1 - \hat{p}_p) \left(\frac{1}{n_1} + \frac{1}{n_2}\right)} = \sqrt{0.664(1 - 0.664) \left(\frac{1}{88} + \frac{1}{680}\right)} = 0.0535$$
$$Z = \frac{\left(\frac{56}{88} - \frac{454}{680}\right) - 0}{0.0535} = -0.59$$

$$\hat{p}_{pooled} = 0.664, \quad n_1 = 88, \quad n_2 = 680$$

$$SE = \sqrt{\hat{p}_{\rho}(1 - \hat{p}_{\rho})\left(\frac{1}{n_{1}} + \frac{1}{n_{2}}\right)} = \sqrt{0.664(1 - 0.664)\left(\frac{1}{88} + \frac{1}{680}\right)} = 0.0535$$
$$Z = \frac{(56/88 - 454/680) - 0}{0.0535} = -0.59$$
$$p\text{-value} = P(Z < -0.59 \text{ or } Z > 0.59)$$
$$= 0.277 + 0.277 = 0.555$$

Confidence interval:

CI = (-0.138, 0.074)

Hypothesis test:

 $H_0: p_{GD|Duke} = p_{GD|US} \qquad Z = -0.59$  $H_A: p_{GD|Duke} \neq p_{GD|US} \qquad p-value = 0.555$ 

Do the results of the Confidence interval and hypothesis test agree? Do the necessarily have to agree?

	Duke	US	Total
A great deal	56	454	510
Not a great deal	32	226	258
Total	88	680	788

	Duke	US	Total
A great deal	56	454	510
Not a great deal	32	226	258
Total	88	680	788

$$H_0: p_{NGD|Duke} = p_{NGD|US} H_0: p_{NGD|Duke} \neq p_{NGD|US}$$
$$\hat{p}_{pooled} = \frac{32 + 226}{88 + 680} = \frac{258}{788} = 0.336$$

	Duke	US	Total
A great deal	56	454	510
Not a great deal	32	226	258
Total	88	680	788

 $H_{0}: p_{NGD|Duke} = p_{NGD|US}$  $H_{0}: p_{NGD|Duke} \neq p_{NGD|US}$  $\hat{p}_{pooled} = \frac{32 + 226}{88 + 680} = \frac{258}{788} = 0.336$ 

$$SE = \sqrt{0.336(1 - 0.336)\left(\frac{1}{88} + \frac{1}{680}\right)} = 0.0535$$

	Duke	US	Total
A great deal	56	454	510
Not a great deal	32	226	258
Total	88	680	788

 $H_{0}: p_{NGD|Duke} = p_{NGD|US}$  $H_{0}: p_{NGD|Duke} \neq p_{NGD|US}$  $\hat{p}_{pooled} = \frac{32 + 226}{88 + 680} = \frac{258}{788} = 0.336$ 

$$SE = \sqrt{0.336(1 - 0.336)\left(\frac{1}{88} + \frac{1}{680}\right)} = 0.0535$$

$$Z = \frac{(32/88 - 226/680) - 0}{0.0535} = 0.585$$

p-value = P(Z < -0.59 or Z > 0.59)= 0.277 + 0.277 = 0.555<sup>39</sup>

	Duke	US	Total
A great deal	56	454	510
Not a great deal	32	226	258
Total	88	680	788

	Duke	US	Total
A great deal	56	454	510
Not a great deal	32	226	258
Total	88	680	788

$$\begin{array}{l} H_0: \ p_{Duke|GD} = p_{Duke|NGD} \\ H_0: \ p_{Duke|GD} \neq p_{Duke|NGD} \end{array} \qquad \qquad \hat{p}_{pooled} = \frac{56 + 32}{510 + 258} = \frac{88}{788} = 0.115 \end{array}$$

	Duke	US	Total
A great deal	56	454	510
Not a great deal	32	226	258
Total	88	680	788

 $H_0: p_{Duke|GD} = p_{Duke|NGD}$  $H_0: p_{Duke|GD} \neq p_{Duke|NGD} \qquad \hat{p}_{pooled} = \frac{56 + 32}{510 + 258} = \frac{88}{788} = 0.115$ 

$$SE = \sqrt{0.115(1 - 0.115)\left(\frac{1}{510} + \frac{1}{258}\right)} = 0.0241$$

	Duke	US	Total
A great deal	56	454	510
Not a great deal	32	226	258
Total	88	680	788

 $\begin{array}{l} H_0: \ p_{Duke|GD} = p_{Duke|NGD} \\ H_0: \ p_{Duke|GD} \neq p_{Duke|NGD} \end{array} \qquad \qquad \hat{p}_{pooled} = \frac{56 + 32}{510 + 258} = \frac{88}{788} = 0.115 \end{array}$ 

$$SE = \sqrt{0.115(1 - 0.115)\left(\frac{1}{510} + \frac{1}{258}\right)} = 0.0241$$

$$Z = \frac{(56/510 - 32/258) - 0}{0.0241} = 0.59$$

p-value = P(Z < -0.59 or Z > 0.59)= 0.2775 + 0.2775 = 0.555 40
# Recap

### Recap - inference for one proportion

• Population parameter: p, point estimate:  $\hat{p}$ 

## Recap - inference for one proportion

- Population parameter: p, point estimate:  $\hat{p}$
- Conditions:
  - independence
    - random sample and 10% condition
  - at least 10 successes and failures
    - observed for CI
    - expected for HT

## Recap - inference for one proportion

- Population parameter: p, point estimate:  $\hat{p}$
- Conditions:
  - independence
    - random sample and 10% condition
  - at least 10 successes and failures
    - observed for CI
    - expected for HT
- Standard error:  $SE = \sqrt{\frac{p(1-p)}{n}}$ 
  - for CI: use p̂
  - for HT: use p<sub>0</sub>
  - for Power:
    - Step 1 use p<sub>0</sub>
    - Step 2 use  $p_A = p_0 + \delta$

• Population parameter:  $(p_1 - p_2)$ , point estimate:  $(\hat{p}_1 - \hat{p}_2)$ 

- Population parameter:  $(p_1 p_2)$ , point estimate:  $(\hat{p}_1 \hat{p}_2)$
- Conditions:

- Population parameter:  $(p_1 p_2)$ , point estimate:  $(\hat{p}_1 \hat{p}_2)$
- Conditions:
  - independence within groups
    - random sample and 10% condition met for both groups
  - independence between groups
  - at least 10 successes and failures in each group
    - observed for CI
    - expected for HT

- Population parameter:  $(p_1 p_2)$ , point estimate:  $(\hat{p}_1 \hat{p}_2)$
- Conditions:
  - independence within groups
    - random sample and 10% condition met for both groups
  - independence between groups
  - at least 10 successes and failures in each group
    - observed for CI
    - expected for HT

• 
$$SE = \sqrt{\frac{p_1(1-p_1)}{p_1} + \frac{p_2(1-p_2)}{p_2}}$$

- for CI: use  $\hat{p}_1$  and  $\hat{p}_2$
- for HT:
  - when  $H_0: p_1 = p_2$ : use  $\hat{p}_{pool} = \frac{\#suc_1 + \#suc_2}{n_1 + n_2}$
  - when H<sub>0</sub>: p<sub>1</sub> − p<sub>2</sub> = (some value other than 0): use p̂<sub>1</sub> and p̂<sub>2</sub>
    this is pretty rare
- for Power:
  - Step 1 use p
     p
     pool
  - Step 2 use  $\hat{p}_1$  and  $\hat{p}_2$

## Reference - standard error calculations

	one sample	two samples
mean	$SE = \frac{\sigma}{\sqrt{n}}$	$SE = \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$
proportion	$SE = \sqrt{\frac{p(1-p)}{n}}$	$SE = \sqrt{\frac{p_1(1-p_1)}{n_1} + \frac{p_2(1-p_2)}{n_2}}$

#### **Reference - standard error calculations**

	one sample	two samples
mean	$SE = \frac{\sigma}{\sqrt{n}}$	$SE = \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$
proportion	$SE = \sqrt{\frac{p(1-p)}{n}}$	$SE = \sqrt{\frac{p_1(1-p_1)}{n_1} + \frac{p_2(1-p_2)}{n_2}}$

 When working with means, it's very rare that σ is known, so we usually use s as an approximation.

#### Reference - standard error calculations

	one sample	two samples
mean	$SE = \frac{\sigma}{\sqrt{n}}$	$SE = \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$
proportion	$SE = \sqrt{\frac{p(1-p)}{n}}$	$SE = \sqrt{\frac{p_1(1-p_1)}{n_1} + \frac{p_2(1-p_2)}{n_2}}$

- When working with means, it's very rare that σ is known, so we usually use s as an approximation.
- When working with proportions, we will not know *p* therefore
  - if doing a hypothesis test, p comes from the null hypothesis
  - if constructing a confidence interval, use  $\hat{p}$  instead