

# Lecture 5

## Random Effects Models

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9/19/2018

## Random Effects Models

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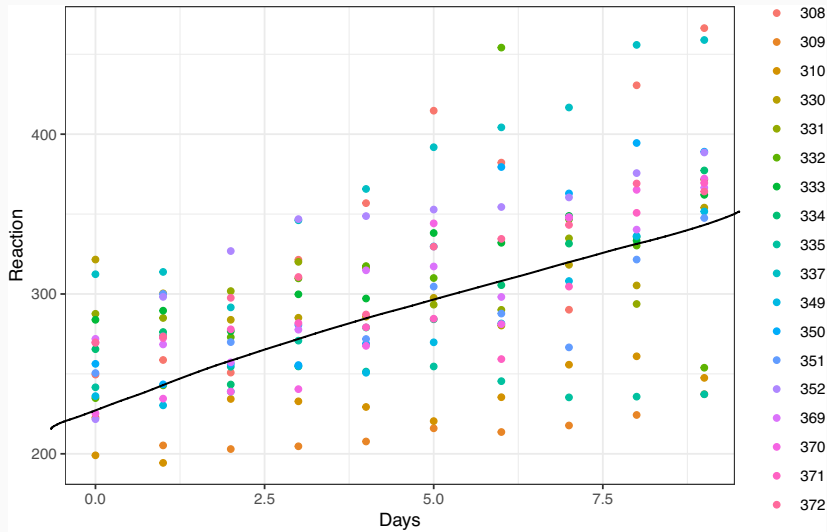
## Sleep Study Data

The average reaction time per day for subjects in a sleep deprivation study. On day 0 the subjects had their normal amount of sleep . Starting that night they were restricted to 3 hours of sleep per night. The observations represent the average reaction time on a series of tests given each day to each subject.

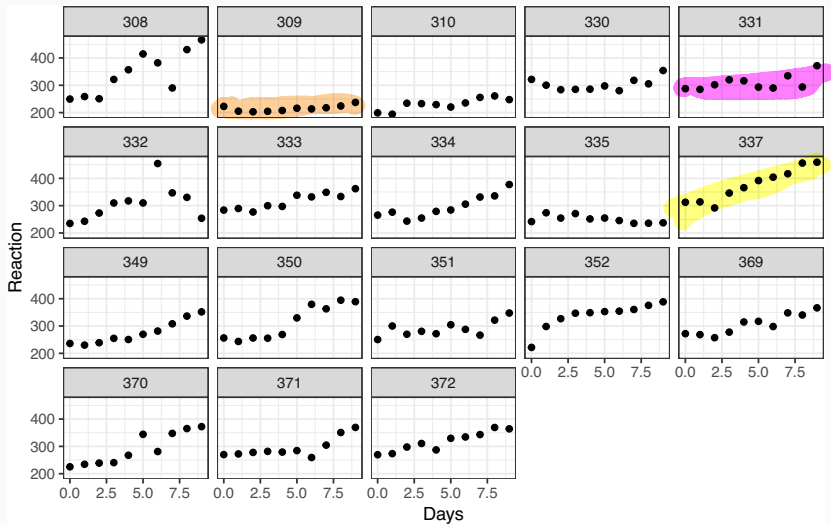
```
sleep = lme4::sleepstudy %>% tbl_df()
```

```
sleep
```

```
## # A tibble: 180 x 3
##   Reaction Days Subject
##   <dbl> <dbl> <fct>
## 1    250.     0 308
## 2    259.     1 308
## 3    251.     2 308
## 4    321.     3 308
## 5    357.     4 308
## 6    415.     5 308
## 7    382.     6 308
## 8    290.     7 308
## 9    431.     8 308
## 10   466.     9 308
## # ... with 170 more rows
```



# EDA (small multiples)



## Bayesian Linear Model

```
sleep_lm = "model{
  # Likelihood
  for(i in 1:length(y)){
    y[i] ~ dnorm(mu[i], tau)
    mu[i] = beta[1] + beta[2]*x[i]

    y_pred[i] ~ dnorm(mu[i],tau)
  }

  # Prior for beta
  beta[1] ~ dnorm(0, 1/10000)
  beta[2] ~ dnorm(0, 1/10000)

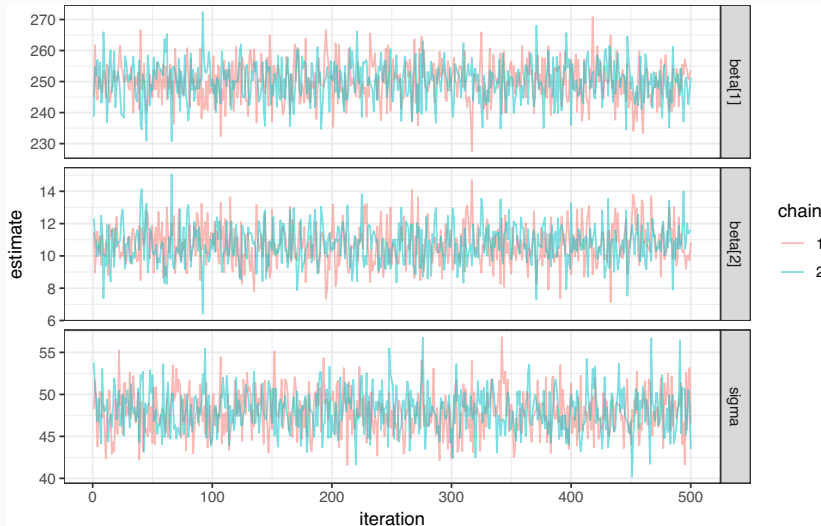
  # Prior for sigma / tau
  sigma ~ dunif(0, 100)
  tau = 1/(sigma*sigma)
}"
```

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$

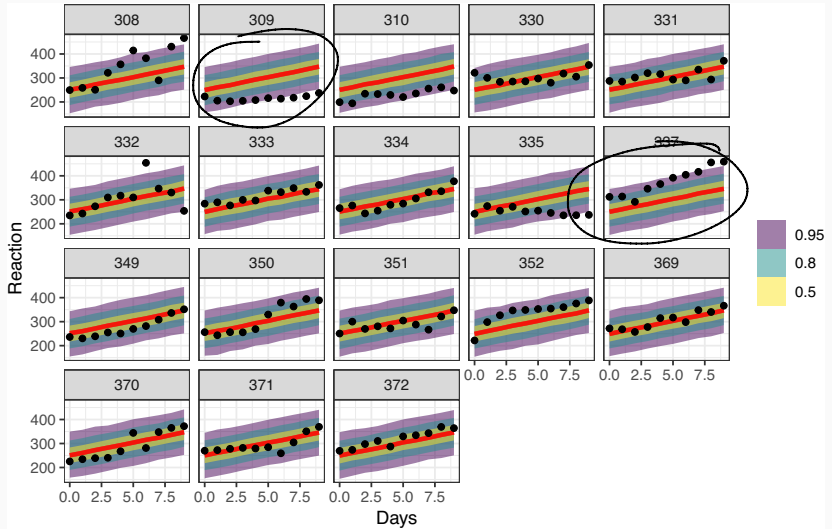
↑  
↳ km time

↑  
days

# MCMC Diagnostics

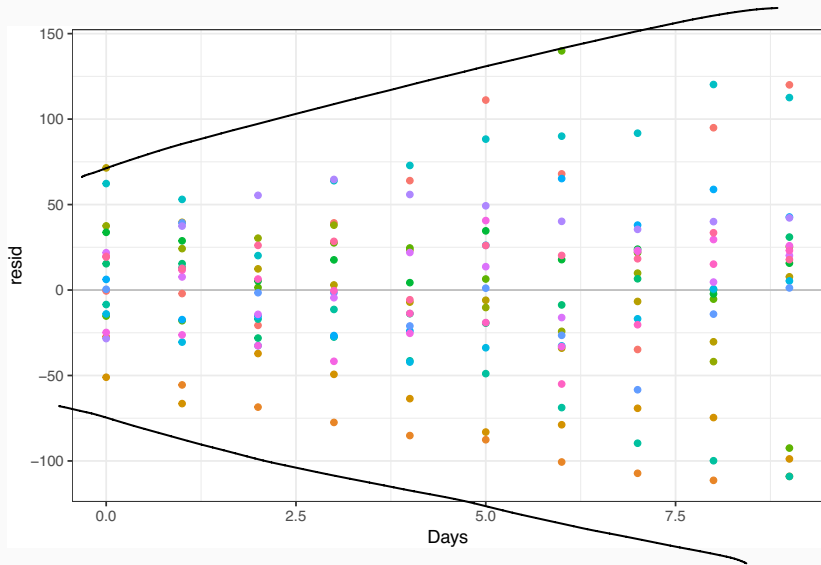


# Model fit

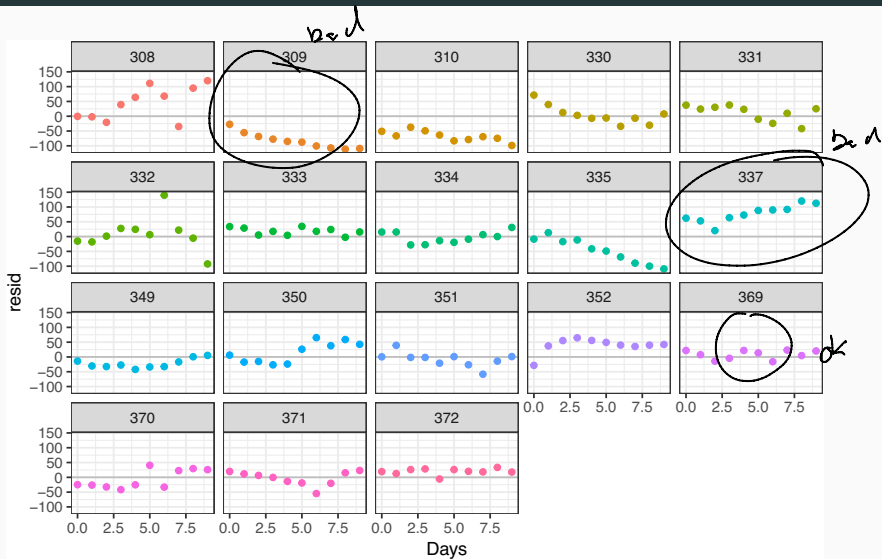




# Residuals



# Residuals by subject



## Random Intercept Model

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```
sleep = sleep %>%  
  mutate(Subject_index = as.integer(Subject))
```

```
sleep[c(1:2,11:12,21:22,31:32),]
```

```
## # A tibble: 8 x 4
```

```
##   Reaction  Days Subject Subject_index
```

```
##   <dbl> <dbl> <fct>         <int>
```

```
## 1    250.     0 308             1
```

```
## 2    259.     1 308             1
```

```
## 3    223.     0 309             2
```

```
## 4    205.     1 309             2
```

```
## 5    199.     0 310             3
```

```
## 6    194.     1 310             3
```

```
## 7    322.     0 330             4
```

```
## 8    300.     1 330             4
```

## Random Intercept Model

Let  $i$  represent each observation and  $j(i)$  be subject in observation  $i$  then

$$y_i = \alpha_{j(i)} + \beta \times \text{Day}_i + \epsilon_i$$

$$\alpha_j \sim \mathcal{N}(\beta_\alpha, \sigma_\alpha^2)$$

$$\epsilon_i \sim \mathcal{N}(0, \sigma^2)$$

$$\beta_\alpha \sim \mathcal{N}(0, 10^4)$$

$$\beta \sim \mathcal{N}(0, 10^4)$$

$$\sigma, \sigma_\alpha \sim \text{Unif}(0, 10^2)$$

## Random Intercept Model - JAGS

```
sleep_ri = "model{
  for(i in 1:length(Reaction)) {
    Reaction[i] ~ dnorm(mu[i],tau)
    mu[i] = alpha[Subject_index[i]] + beta*Days[i]

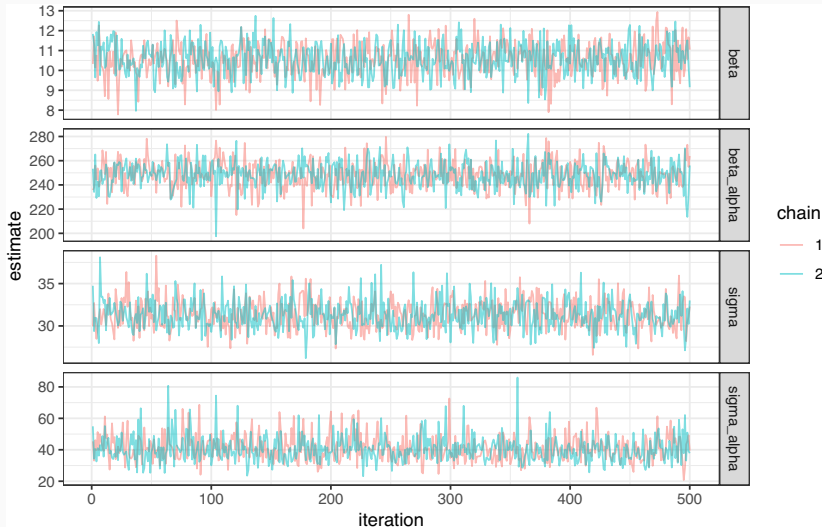
    y_pred[i] ~ dnorm(mu[i],tau)
  }

  for(j in 1:18) {
    alpha[j] ~ dnorm(beta_alpha, tau_alpha)
  }

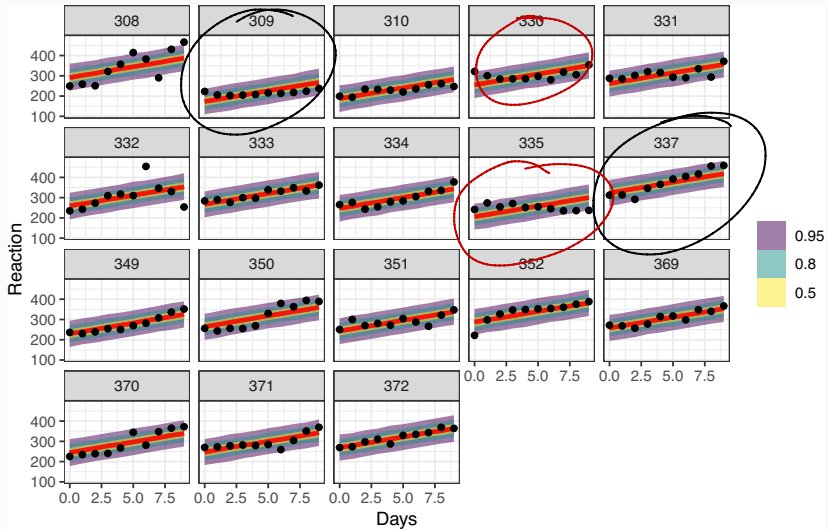
  beta_alpha ~ dnorm(0,1/10000)
  sigma_alpha ~ dunif(0, 100)
  tau_alpha = 1/(sigma_alpha*sigma_alpha)

  beta ~ dnorm(0,1/10000)
  sigma ~ dunif(0, 100)
  tau = 1/(sigma*sigma)
}"
```

# MCMC Diagnostics

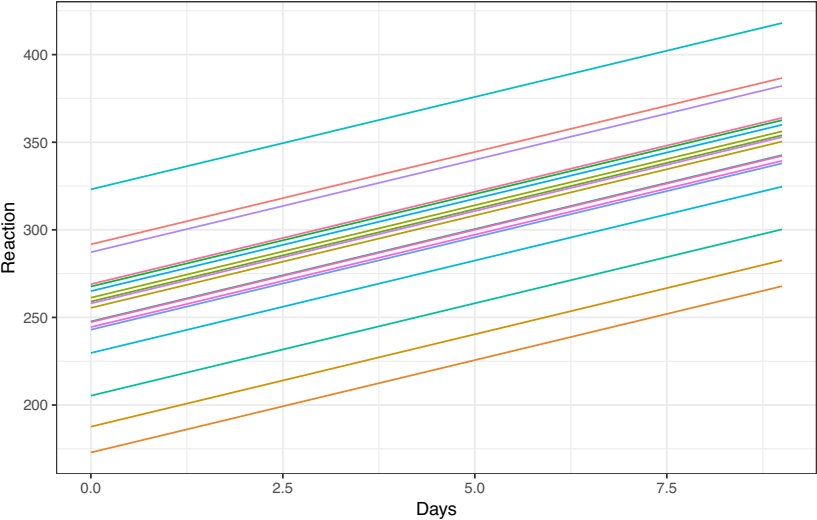


# Model fit

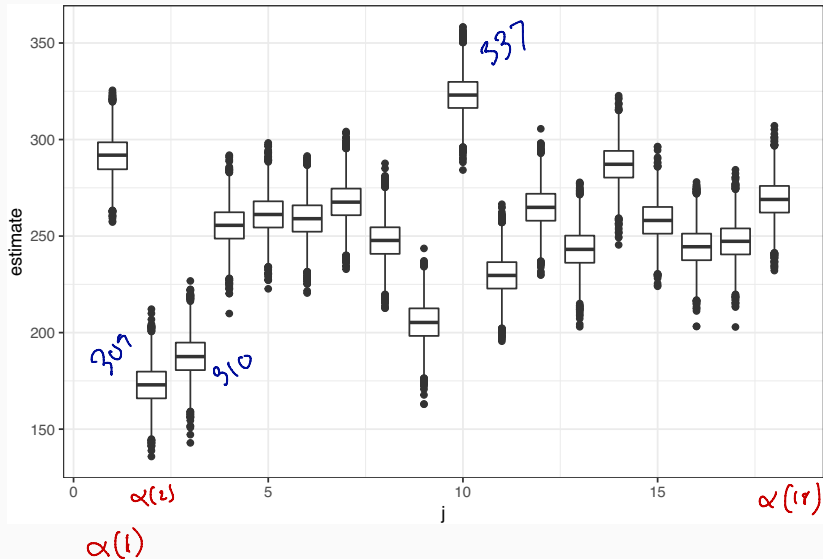




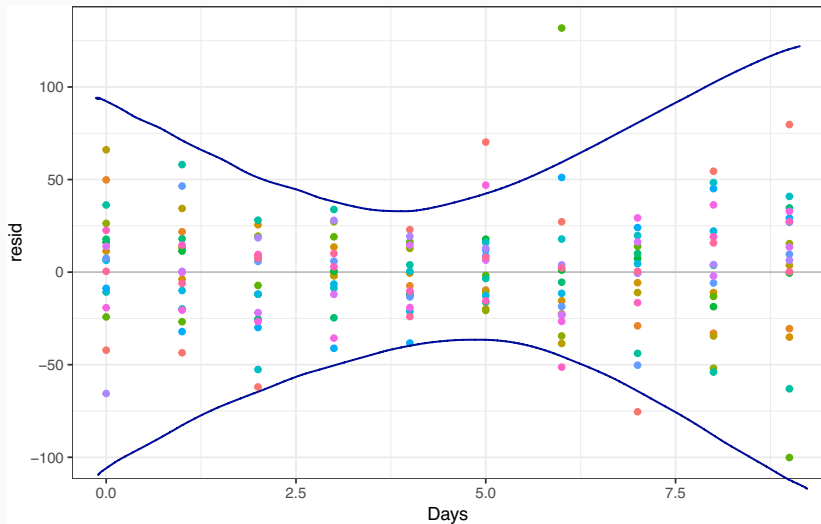
# Model fit - lines



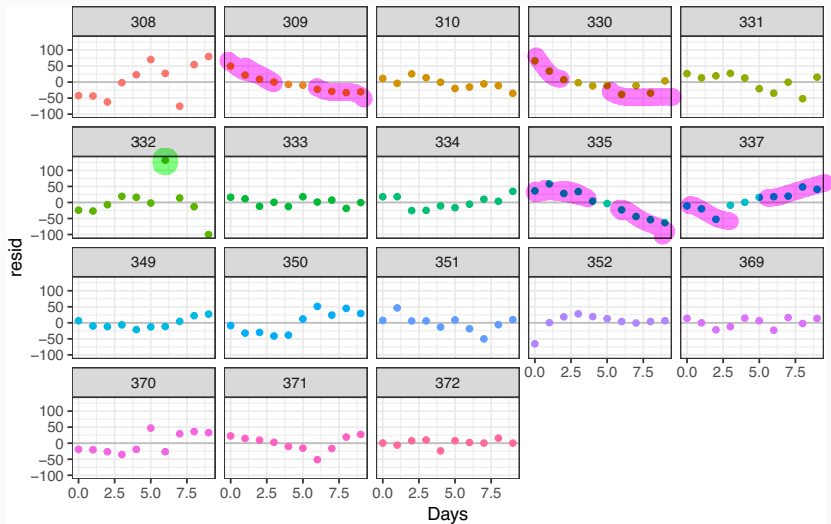
# Random Effects



# Residuals



# Residuals by subject



## Why not a fixed effect for Subject?

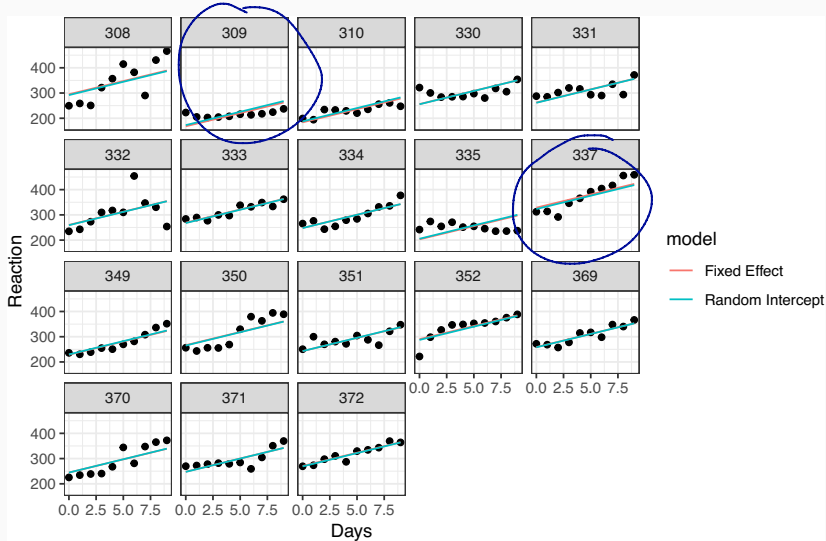
We are not going to bother with the Bayesian model here to avoid the headache of dummy coding and the resulting  $\beta$ s.

## Why not a fixed effect for Subject?

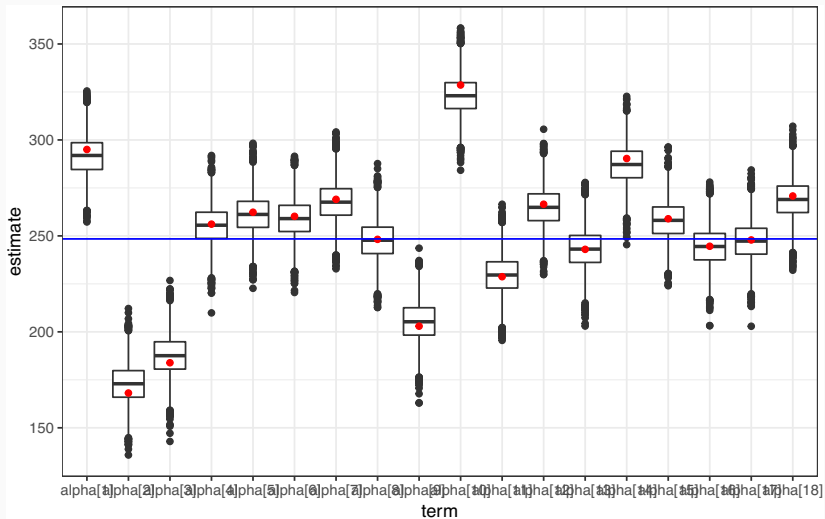
We are not going to bother with the Bayesian model here to avoid the headache of dummy coding and the resulting  $\beta$ s.

```
l = lm(Reaction ~ Days + Subject - 1, data=sleep)
summary(l)
##
## Call:
## lm(formula = Reaction ~ Days + Subject - 1, data = sleep)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -100.540  -16.389   -0.341   15.215  131.159
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## Days              10.4673     0.88042   13.02 <2e-16 ***
## Subject308        295.0310    10.4471  28.24 <2e-16 ***
## Subject309        168.1302    10.4471  16.09 <2e-16 ***
## Subject310        183.8985    10.4471  17.60 <2e-16 ***
## Subject330        256.1186    10.4471  24.52 <2e-16 ***
## Subject331        262.3333    10.4471  25.11 <2e-16 ***
## Subject332        260.1993    10.4471  24.91 <2e-16 ***
## Subject333        269.0555    10.4471  25.75 <2e-16 ***
## Subject334        248.1993    10.4471  23.76 <2e-16 ***
## Subject335        202.9673    10.4471  19.43 <2e-16 ***
## Subject337        328.6182    10.4471  31.45 <2e-16 ***
## Subject349        228.7317    10.4471  21.89 <2e-16 ***
## Subject350        266.4999    10.4471  25.51 <2e-16 ***
## Subject351        242.9950    10.4471  23.26 <2e-16 ***
## Subject352        290.3188    10.4471  27.79 <2e-16 ***
## Subject369        258.9319    10.4471  24.79 <2e-16 ***
## Subject370        244.5990    10.4471  23.41 <2e-16 ***
## Subject371        247.8813    10.4471  23.73 <2e-16 ***
## Subject372        270.7833    10.4471  25.92 <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

# Comparing Model fit



# Random effects vs fixed effects





## Random Intercept Model (strong prior for $\sigma_\alpha$ )

```
sleep_ri2 = "model{
  for(i in 1:length(Reaction)) {
    Reaction[i] ~ dnorm(mu[i],tau)
    mu[i] = alpha[Subject_index[i]] + beta*Days[i]

    y_pred[i] ~ dnorm(mu[i],tau)
  }

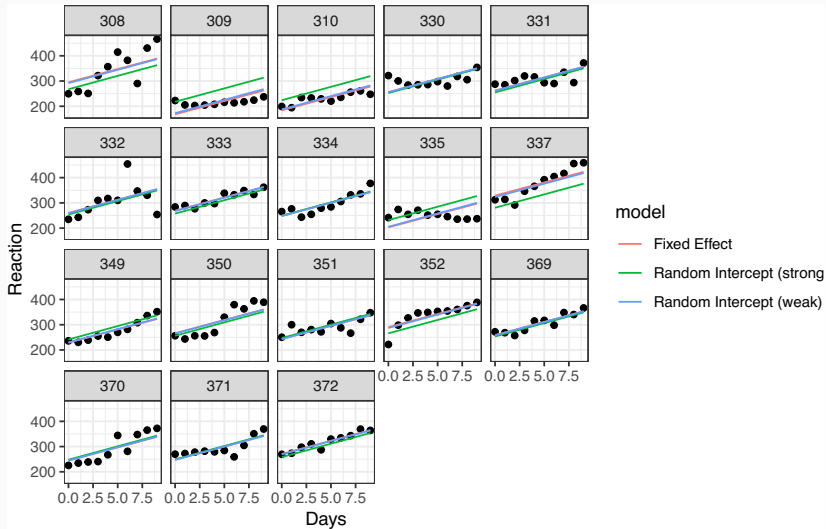
  for(j in 1:18) {
    alpha[j] ~ dnorm(beta_alpha, tau_alpha)
  }

  beta_alpha ~ dnorm(0,1/10000)
  sigma_alpha ~ dunif(0, 10)
  tau_alpha = 1/(sigma_alpha*sigma_alpha)

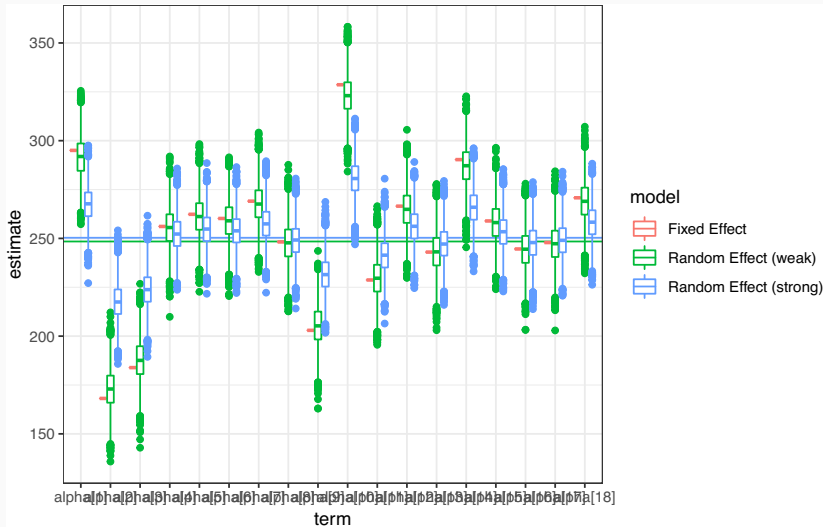
  beta ~ dnorm(0,1/10000)
  sigma ~ dunif(0, 100)
  tau = 1/(sigma*sigma)
}"
```

*vs dunif( $\sigma$ , 100)*

# Comparing Model fit



# Prior Effect on $\alpha$



## Some Distribution Theory

$$Y_i = \alpha_{j(i)} + \beta_1 X_i + \varepsilon_i$$

$$\alpha_j \sim N(\rho_\alpha, \sigma_\alpha^2)$$

$$\varepsilon_i \sim N(0, \sigma^2)$$

$$\underline{Y} \sim N(\underline{\mu}, \underline{\Sigma})$$

$$\underline{\mu} = ?$$

$$\underline{\Sigma} = ?$$

$$E(Y_i | \underline{\theta}) = E(\alpha_{j(i)} + \beta_1 X_i + \varepsilon_i | \underline{\theta})$$

$$= E(\alpha_{j(i)} | \underline{\theta}) + E(\beta_1 X_i | \underline{\theta}) + E(\varepsilon_i | \underline{\theta})$$

$$= \boxed{\rho_\alpha + \beta_1 X_i} + 0$$

$$\underline{\mu} = \beta_\alpha \frac{\mathbf{1}}{n \times 1} + \beta_1 \underline{X}$$

$$\underline{\theta} = \{\beta_1, \rho_\alpha, \sigma_\alpha^2, \sigma^2\}$$

## Some Distribution Theory

$$\{ \underline{\xi} \}_{i,k} = \text{Cov}(Y_i, Y_k | \underline{\theta})$$

$$= E \left( (Y_i - E(Y_i | \underline{\theta})) (Y_k - E(Y_k | \underline{\theta})) | \underline{\theta} \right)$$

$$= E \left( (\alpha_{j(i)} + \cancel{\beta_1 X_i} + \xi_i - (\beta_\alpha + \cancel{\beta_1 X_i})) (\alpha_{j(k)} + \cancel{\beta_1 X_k} + \xi_k - (\beta_\alpha + \cancel{\beta_1 X_k})) \right)$$

$$= E \left( (\alpha_{j(i)} + \xi_i - \beta_\alpha) (\alpha_{j(k)} + \xi_k - \beta_\alpha) | \underline{\theta} \right)$$

$$= E \left( \alpha_{j(i)} \alpha_{j(k)} + \xi_i \xi_k - \alpha_{j(i)} \beta_\alpha - \alpha_{j(k)} \beta_\alpha + \beta_\alpha^2 | \underline{\theta} \right)$$

$$= E \left( \alpha_{j(i)} \alpha_{j(k)} | \underline{\theta} \right) + E \left( \xi_i \xi_k | \underline{\theta} \right) - \beta_\alpha^2$$

$$= E(\alpha_{j(i)}\alpha_{j(k)} | \underline{\theta}) + E(\varepsilon_i \varepsilon_k | \underline{\theta}) - \beta_\alpha^2$$

$i=k$

$$\text{cov}(-) = E(\alpha_j(i)^2 | \underline{\theta}) + E(\varepsilon_i^2 | \underline{\theta}) - \beta_\alpha^2$$

$$= \text{Var}(\alpha_j(i)^2 | \underline{\theta}) + E(\alpha_j(i)^2)$$

$$+ \text{Var}(\varepsilon_i) + E(\varepsilon_i^2)$$

$$- \beta_\alpha^2$$

$$= \sigma_\alpha^2 + \sigma^2$$

# Some Distribution Theory

$$= E(\alpha_{j(i)} \alpha_{j(k)} | \underline{\theta}) + E(\epsilon_i \epsilon_k | \underline{\theta}) - \beta_\alpha^2$$

$$\alpha_{j(i)} = \alpha_{j(k)}$$

$$i \neq k$$

Diff obs same subj

$$\text{Cov}(\cdot) = E(\alpha_{j(i)}^2) + E(\epsilon_i \epsilon_k) - \beta_\alpha^2$$

$$= \text{Var}(\alpha_{j(i)}) + E(\alpha_{j(i)})^2 - \beta_\alpha^2$$

$$= \sigma_\alpha^2$$

$$i \neq k \quad \alpha_{j(i)} \neq \alpha_{j(k)}$$

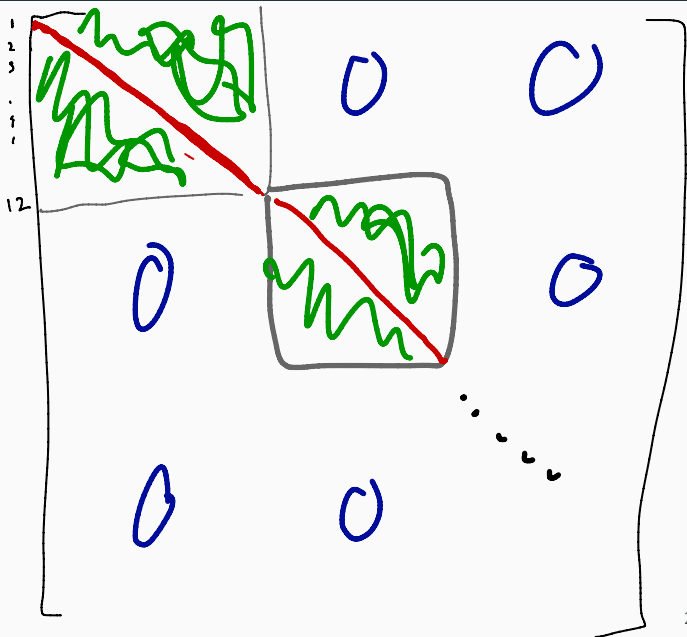
$$\text{Cov}(\cdot) = E(\alpha_{j(i)}) E(\alpha_{j(k)}) + E(\epsilon_i) E(\epsilon_k) - \beta_\alpha^2$$

$$= \beta_\alpha \beta_\alpha + 0 \cdot 0 - \beta_\alpha^2 = 0$$

# Some Distribution Theory

$$\sigma_x^2 + \sigma^2$$
$$\sigma_x^2$$
$$0$$

$$\xi =$$
$$\frac{\quad}{n \times n}$$





## Random intercept and slope model

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Let  $i$  represent each observation and  $j(i)$  be the subject in observation  $i$  then

$$Y_i = \alpha_{j(i)} + \beta_{j(i)} \times \text{Days} + \epsilon_i$$

$$\alpha_j \sim \mathcal{N}(\beta_0, \sigma_\alpha^2)$$

$$\beta_j \sim \mathcal{N}(\beta_1, \sigma_\beta^2)$$

$$\epsilon_i \sim \mathcal{N}(0, \sigma^2)$$

$$\beta_\alpha, \beta_\beta \sim \mathcal{N}(0, 10000)$$

$$\sigma, \sigma_\alpha, \sigma_\beta \sim \text{Unif}(0, 100)$$

## Model - JAGS

```
sleep_ris = "model{
  for(i in 1:length(Reaction)) {
    Reaction[i] ~ dnorm(mu[i],tau)
    mu[i] = alpha[Subject_index[i]] + beta[Subject_index[i]] * Days[i]
    y_pred[i] ~ dnorm(mu[i], tau)
  }

  sigma ~ dunif(0, 100)
  tau = 1/(sigma*sigma)

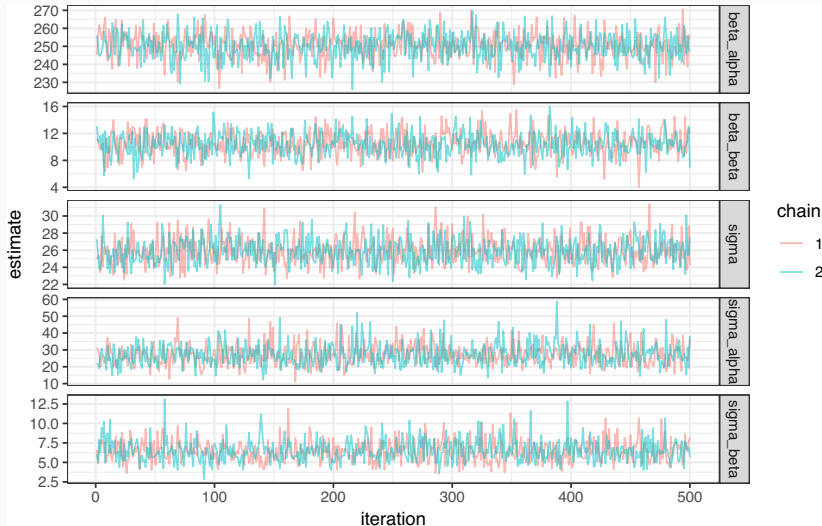
  for(j in 1:18) {
    alpha[j] ~ dnorm(beta_alpha, tau_alpha)
    beta[j] ~ dnorm(beta_beta, tau_beta)
  }

  beta_alpha ~ dnorm(0,1/10000)
  beta_beta ~ dnorm(0,1/10000)

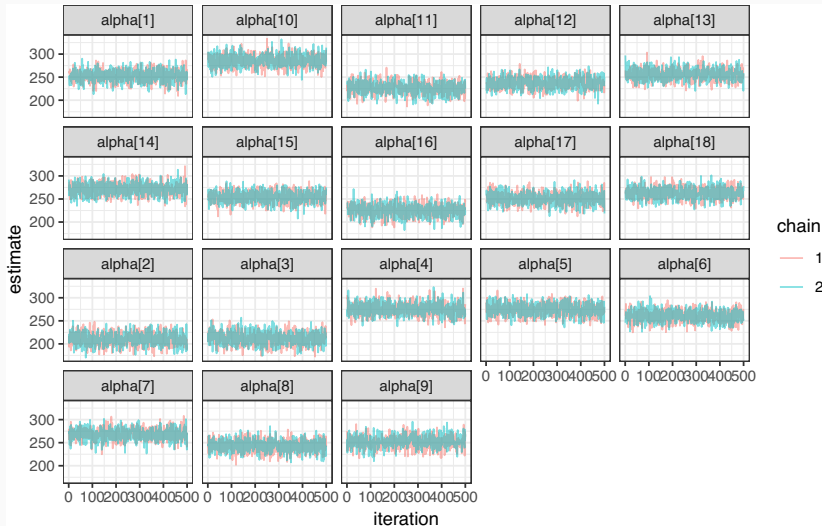
  sigma_alpha ~ dunif(0, 100)
  tau_alpha = 1/(sigma_alpha*sigma_alpha)

  sigma_beta ~ dunif(0, 100)
  tau_beta = 1/(sigma_beta*sigma_beta)
}"
```

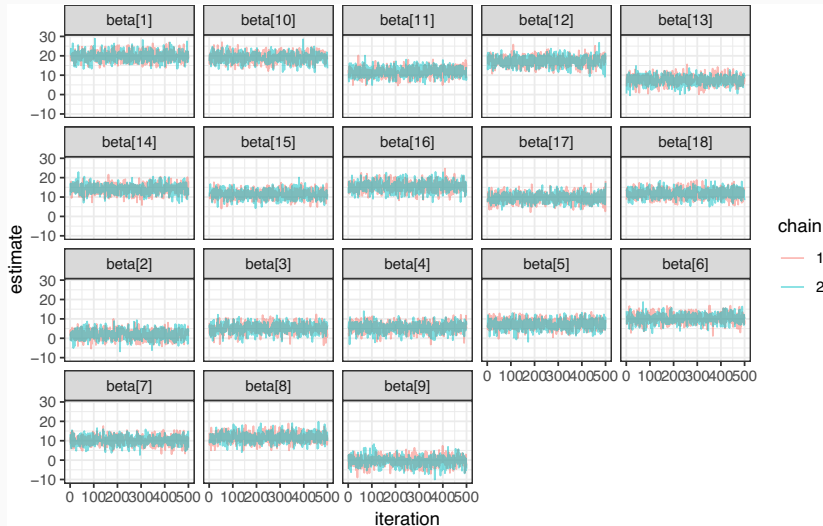
# MCMC Diagnostics



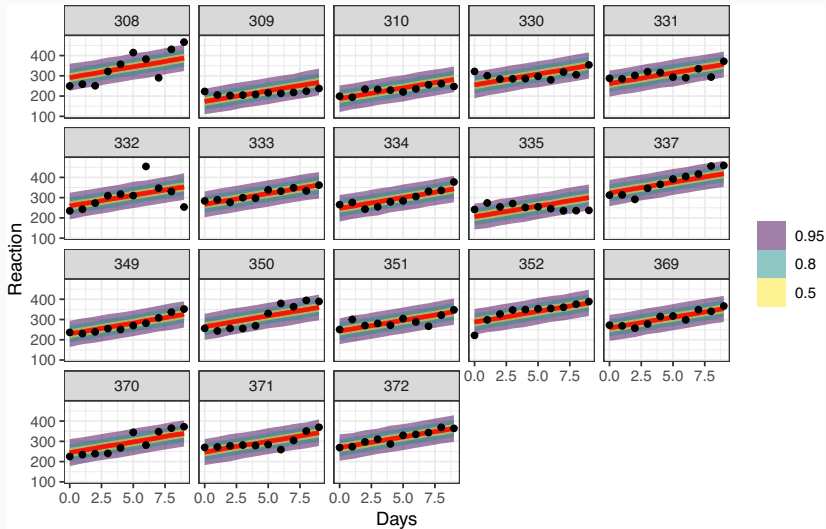
# MCMC Diagnostics - $\alpha$



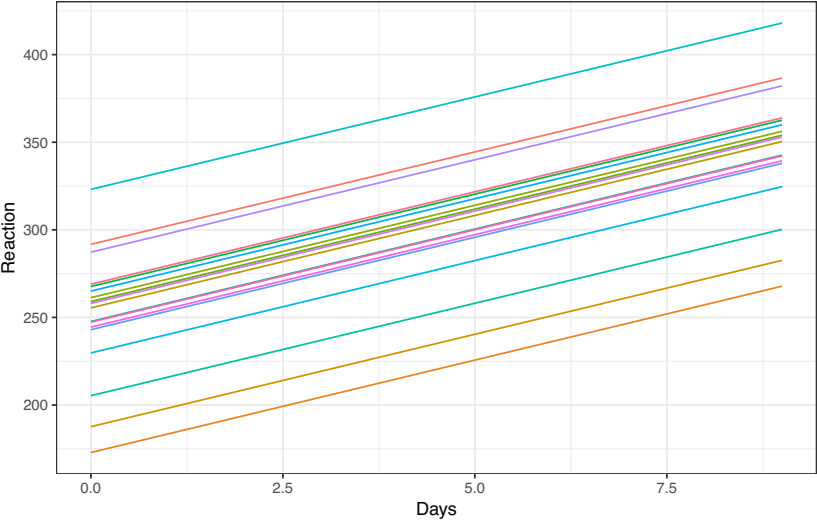
# MCMC Diagnostics - $\beta$



# Model fit

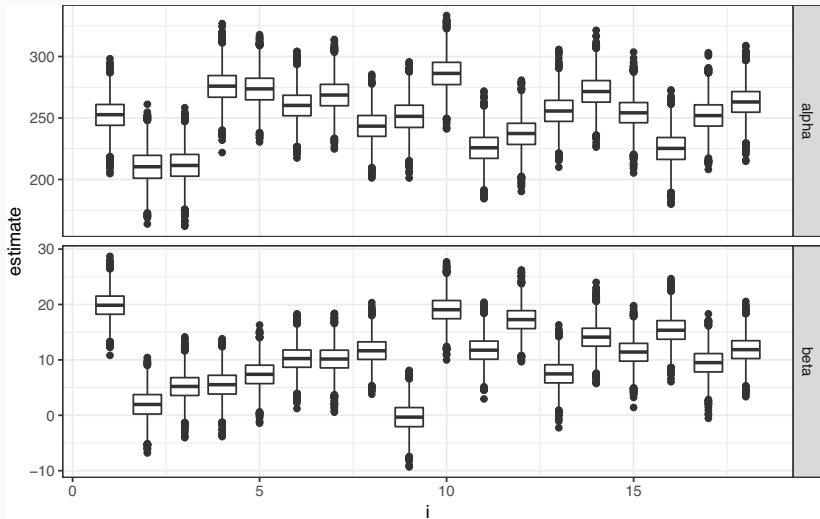


# Model fit - lines





# Random Effects



# Residuals by subject

