

# Lecture 14

## Full Posterior Pred & Covariance Functions

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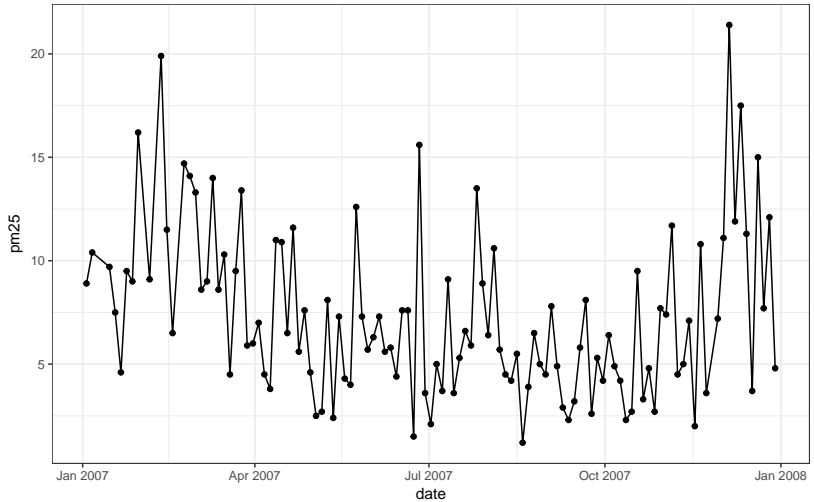
Colin Rundel

03/06/2017

## Full Posterior Predictive Distribution

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# FRN Data



## JAGS Model

```
## model{
##   y ~ dnorm(mu, inverse(Sigma))
##
##   for (i in 1:N) {
##     mu[i] = beta[1]+ beta[2] * x[i] + beta[3] * x[i]^2
##   }
##
##   for (i in 1:(N-1)) {
##     for (j in (i+1):N) {
##       Sigma[i,j] = sigma2 * exp(- pow(l+d[i,j],2))
##       Sigma[j,i] = Sigma[i,j]
##     }
##   }
##
##   for (k in 1:N) {
##     Sigma[k,k] = sigma2 + sigma2_w
##   }
##
##   for (i in 1:3) {
##     beta[i] ~ dt(0, 2.5, 1)
##   }
##   sigma2_w ~ dnorm(10, 1/25) T(0,)
##   sigma2   ~ dnorm(10, 1/25) T(0,)
##   l       ~ dt(0, 2.5, 1) T(0,)
## }
```

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param	post_mean	post_med	post_lower	post_upper
beta[1]	9.2136151	11.4359371	-0.4309078	15.2615892
beta[2]	-0.0361357	-0.0551308	-0.1012205	0.0849476
beta[3]	0.0001007	0.0001367	-0.0001924	0.0002552
l	0.8787410	0.0698553	0.0065124	7.0905582
sigma2	8.4807746	7.8609848	1.5342164	18.6524860
sigma2_w	9.7527513	10.4646243	2.2091857	14.8425142

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# Predicting

```
l = post %>% filter(param == 'l') %>% select(post_med) %>% unlist()
sigma2 = post %>% filter(param == 'sigma2') %>% select(post_med) %>% unlist()
sigma2_w = post %>% filter(param == 'sigma2_w') %>% select(post_med) %>% unlist()

beta0 = post %>% filter(param == 'beta[1]') %>% select(post_med) %>% unlist()
beta1 = post %>% filter(param == 'beta[2]') %>% select(post_med) %>% unlist()
beta2 = post %>% filter(param == 'beta[3]') %>% select(post_med) %>% unlist()

reps=1000

x = pm25$day
y = pm25$pm25
x_pred = 1:365 + rnorm(365, 0.01)

mu = beta0 + beta1*x + beta2*x^2
mu_pred = beta0 + beta1*x_pred + beta2*x_pred^2

dist_o = rdist(x)
dist_p = rdist(x_pred)
dist_op = rdist(x, x_pred)
dist_po = t(dist_op)

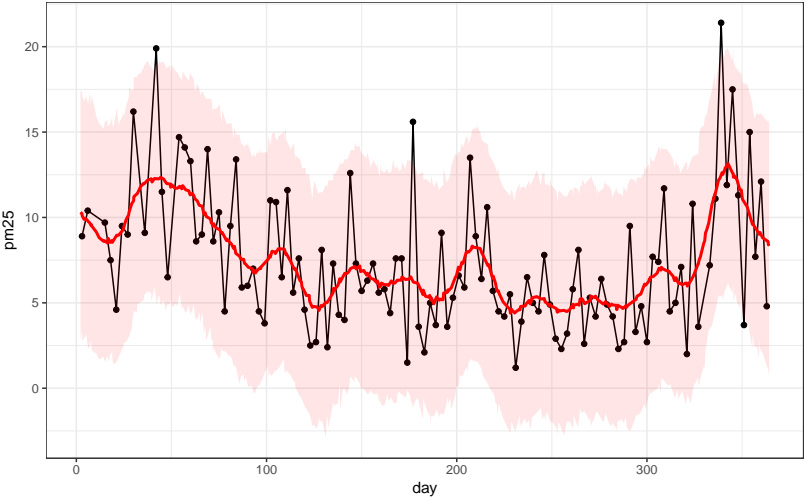
cov_o = sq_exp_cov(dist_o, sigma2 = sigma2, l = l) + nugget_cov(dist_o, sigma2 = sigma2_w)
cov_p = sq_exp_cov(dist_p, sigma2 = sigma2, l = l) + nugget_cov(dist_p, sigma2 = sigma2_w)
cov_op = sq_exp_cov(dist_op, sigma2 = sigma2, l = l) + nugget_cov(dist_op, sigma2 = sigma2_w)
cov_po = sq_exp_cov(dist_po, sigma2 = sigma2, l = l) + nugget_cov(dist_po, sigma2 = sigma2_w)

cond_cov = cov_p - cov_po %>% solve(cov_o) %>% cov_op
cond_mu = mu_pred + cov_po %>% solve(cov_o) %>% (y - mu)

pred = cond_mu %>% matrix(1, ncol=reps) + t(chol(cond_cov)) %>% matrix(rnorm(length(x_pred)*reps), ncol=reps)

pred_df = pred %>% t() %>% post_summary() %>% mutate(day=x_pred)
```

# Predictions



## Full Posterior Predictive Distribution

Our posterior consists of samples from

$$l, \sigma^2, \sigma_w^2, \beta_0, \beta_1, \beta_2 \mid \mathbf{y}$$

and for the purposes of generating the posterior predictions we sampled

$$\mathbf{y}_{pred} \mid l^{(m)}, \sigma^{2(m)}, \sigma_w^{2(m)}, \beta_0^{(m)}, \beta_1^{(m)}, \beta_2^{(m)}, \mathbf{y}$$

where  $l^{(m)}$ , etc. are the posterior median of that parameter.



## Full Posterior Predictive Distribution

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and for the purposes of generating the posterior predictions we sampled

$$\mathbf{y}_{pred} \mid l^{(m)}, \sigma^{2(m)}, \sigma_w^{2(m)}, \beta_0^{(m)}, \beta_1^{(m)}, \beta_2^{(m)}, \mathbf{y}$$

where  $l^{(m)}$ , etc. are the posterior median of that parameter.

In practice we should instead be sampling

$$\mathbf{y}_{pred}^{(i)} \mid l^{(i)}, \sigma^{2(i)}, \sigma_w^{2(i)}, \beta_0^{(i)}, \beta_1^{(i)}, \beta_2^{(i)}, \mathbf{y}$$

since this takes into account the additional uncertainty in the model parameters.

# Full Posterior Predictive Distribution

```
if (!file.exists("gp_pred.Rdata"))
{
  x = pm25$day; y = pm25$pm25

  n_post_samp = nrow(param)

  x_pred = 1:365 + rnorm(365, 0.01)
  y_pred = matrix(NA, nrow=n_post_samp, ncol=length(x_pred))
  colnames(y_pred) = paste0("Y_pred[", round(x_pred,0), "]")

  for(i in 1:n_post_samp)
  {
    l = param[i,'l']
    sigma2 = param[i,'sigma2']
    sigma2_w = param[i,'sigma2_w']
    beta0 = betas[i,"beta[1]"]
    beta1 = betas[i,"beta[2]"]
    beta2 = betas[i,"beta[3]"]

    mu = beta0 + beta1*x + beta2*x^2
    mu_pred = beta0 + beta1*x_pred + beta2*x_pred^2

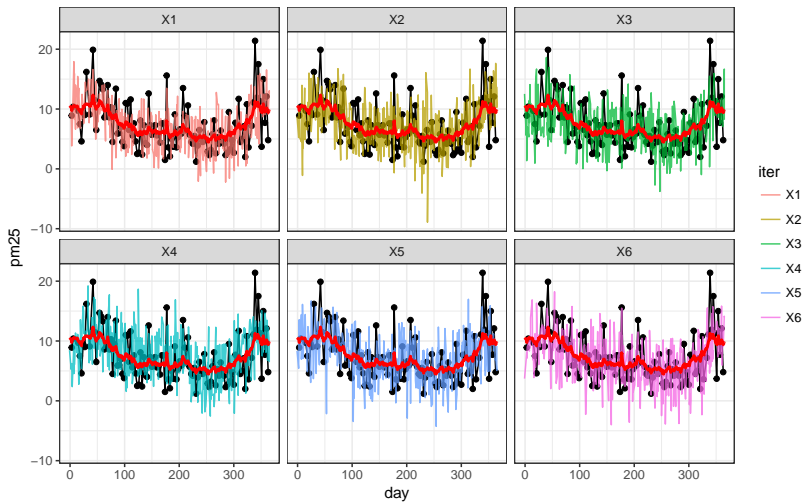
    dist_o = rdist(x)
    dist_p = rdist(x_pred)
    dist_op = rdist(x, x_pred)
    dist_po = t(dist_op)

    cov_o = sq_exp_cov(dist_o, sigma2 = sigma2, l = l) + nugget_cov(dist_o, sigma2 = sigma2_w)
    cov_p = sq_exp_cov(dist_p, sigma2 = sigma2, l = l) + nugget_cov(dist_p, sigma2 = sigma2_w)
    cov_op = sq_exp_cov(dist_op, sigma2 = sigma2, l = l) + nugget_cov(dist_op, sigma2 = sigma2_w)
    cov_po = sq_exp_cov(dist_po, sigma2 = sigma2, l = l) + nugget_cov(dist_po, sigma2 = sigma2_w)

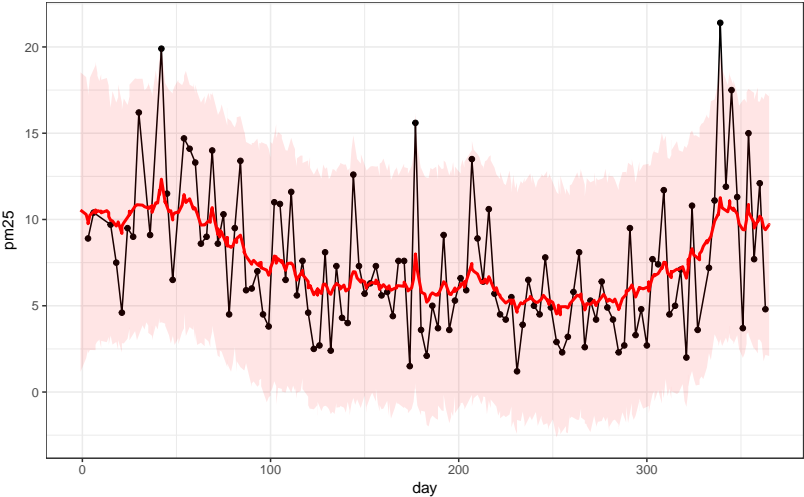
    cond_cov = cov_p - cov_po %>% solve(cov_o) %>% cov_op
    cond_mu = mu_pred + cov_po %>% solve(cov_o) %>% (y - mu)

    y_pred[i,] = cond_mu + t(chol(cond_cov)) %>% matrix(rnorm(length(x_pred)), ncol=1)
  }
}
```

# Full Posterior Predictive Distribution - Plots



# Full Posterior Predictive Distribution - Mean + CI

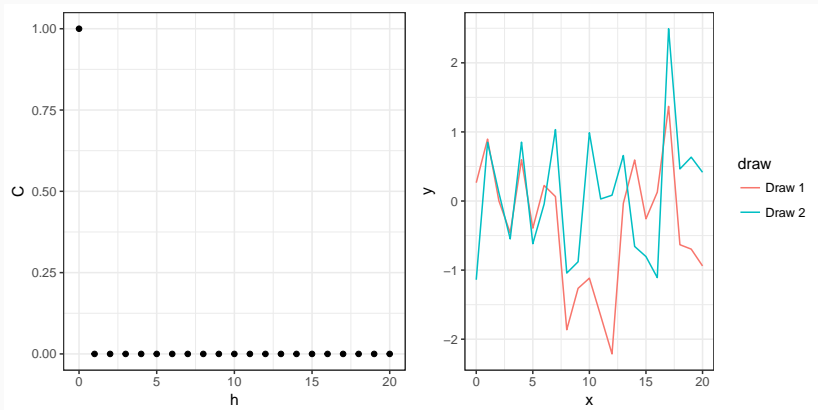


## More on Covariance Functions

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# Nugget Covariance

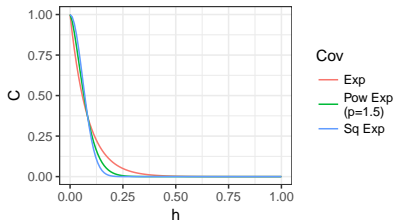
$$\text{Cov}(y_{t_i}, y_{t_j}) = \text{Cov}(h = |t_i - t_j|) = \sigma^2 \mathbb{1}_{\{h=0\}}$$



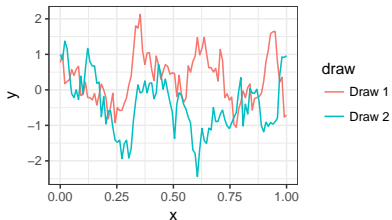
# (- / Power / Square) Exponential Covariance

$$\text{Cov}(y_{t_i}, y_{t_j}) = \text{Cov}(h = |t_i - t_j|) = \sigma^2 \exp(-(h l)^p)$$

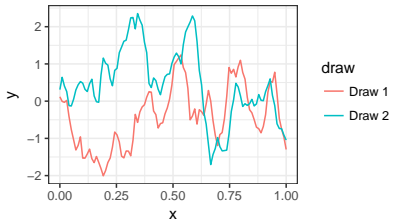
Covariance - l=12, sigma2=1



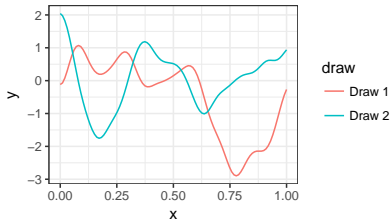
Exponential



Powered Exponential (p=1.5)

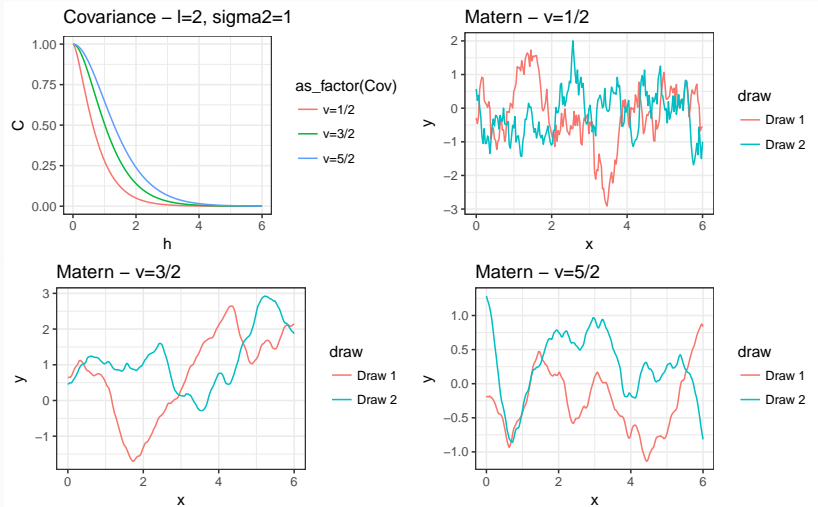


Square Exponential



# Matern Covariance

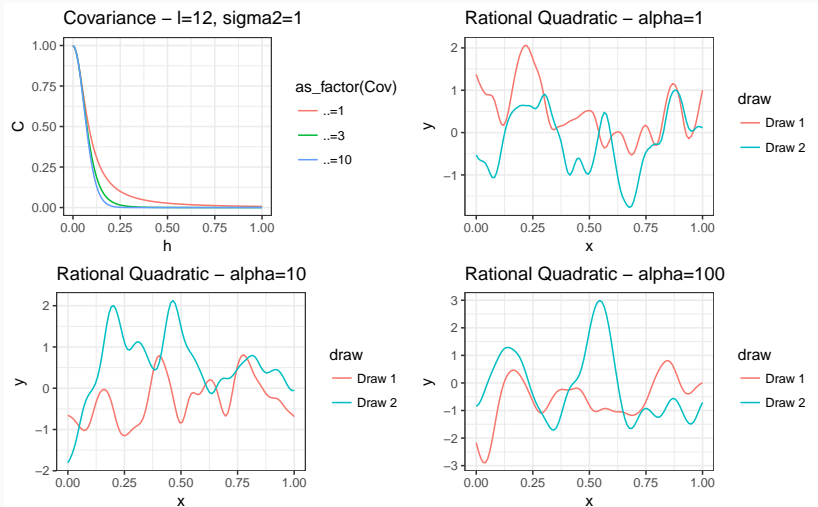
$$\text{Cov}(y_{t_i}, y_{t_j}) = \text{Cov}(h = |t_i - t_j|) = \sigma^2 \frac{2^{1-\nu}}{\Gamma(\nu)} \left(\sqrt{2\nu} h \cdot l\right)^\nu K_\nu \left(\sqrt{2\nu} h \cdot l\right)$$





# Rational Quadratic Covariance

$$\text{Cov}(y_{t_i}, y_{t_j}) = \text{Cov}(h = |t_i - t_j|) = \sigma^2 \left( 1 + \frac{h^2 l^2}{\alpha} \right)^{-\alpha}$$



- **Matern Covariance**

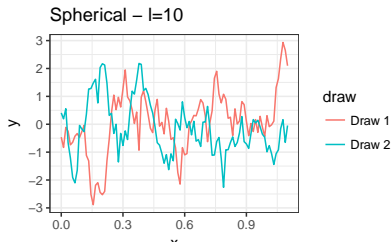
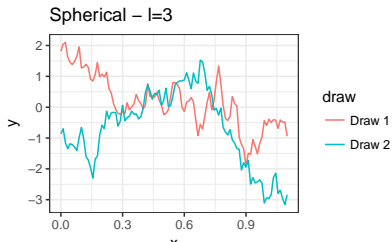
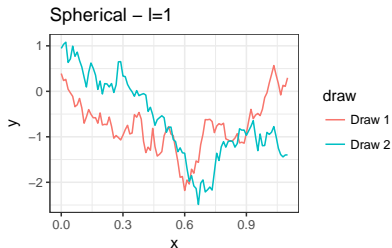
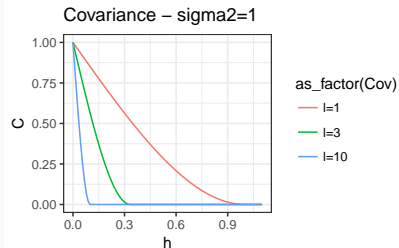
- A Gaussian process with Matérn covariance has sample functions that are  $\lceil \nu - 1 \rceil$  times differentiable.
- When  $\nu = 1/2 + p$  for  $p \in \mathbb{N}^+$  then the Matern has a simplified form (product of an exponential and a polynomial of order  $p$ ).
- When  $\nu = 1/2$  the Matern is equivalent to the exponential covariance.
- As  $\nu \rightarrow \infty$  the Matern converges to the square exponential covariance.

- **Rational Quadratic Covariance**

- is a scale mixture (infinite sum) of squared exponential covariance functions with different characteristic length-scales ( $l$ ).
- As  $\alpha \rightarrow \infty$  the rational quadratic converges to the square exponential covariance.
- Has sample functions that are infinitely differentiable for any value of  $\alpha$

# Spherical Covariance

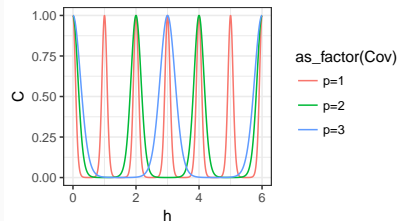
$$\text{Cov}(y_{t_i}, y_{t_j}) = \text{Cov}(h = |t_i - t_j|) = \begin{cases} \sigma^2 \left(1 - \frac{3}{2}h \cdot l + \frac{1}{2}(h \cdot l)^3\right) & \text{if } 0 < h < 1/l \\ 0 & \text{otherwise} \end{cases}$$



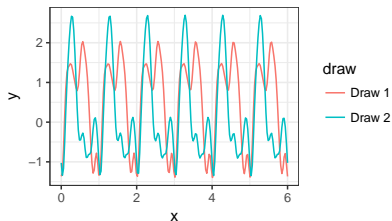
# Periodic Covariance

$$\text{Cov}(y_{t_i}, y_{t_j}) = \text{Cov}(h = |t_i - t_j|) = \sigma^2 \exp\left(-2l^2 \sin^2\left(\pi \frac{h}{p}\right)\right)$$

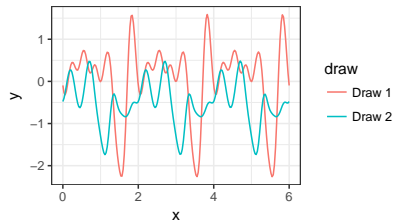
Covariance - l=2, sigma2=1



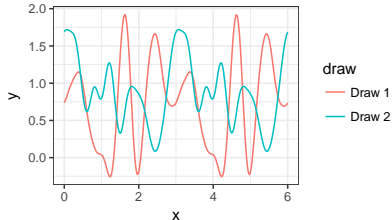
Periodic - p=1



Periodic - p=2

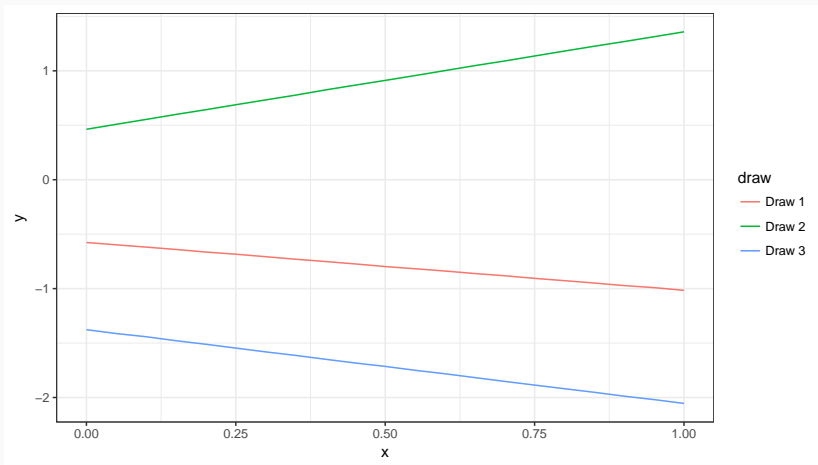


Periodic - p=3



# Linear Covariance

$$\text{Cov}(y_{t_i}, y_{t_j}) = \sigma_b^2 + \sigma_v^2 (t_i - c)(t_j - c)$$



## Combining Covariances

If we define two valid covariance functions,  $Cov_a(y_{t_i}, y_{t_j})$  and  $Cov_b(y_{t_i}, y_{t_j})$  then the following are also valid covariance functions,

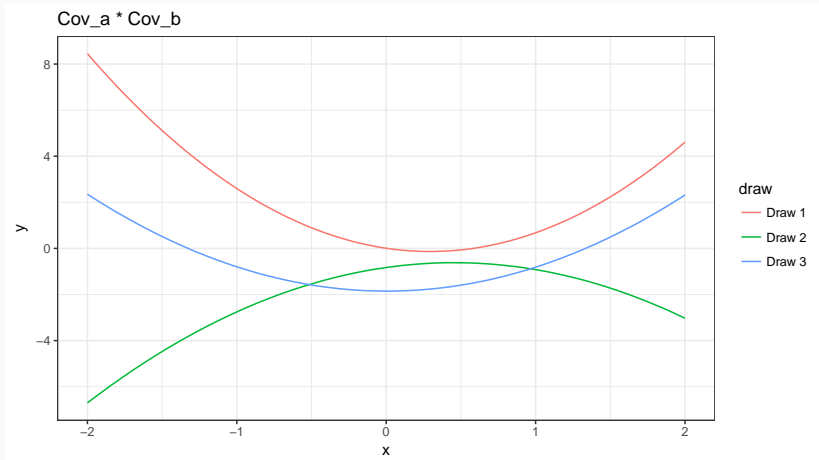
$$Cov_a(y_{t_i}, y_{t_j}) + Cov_b(y_{t_i}, y_{t_j})$$

$$Cov_a(y_{t_i}, y_{t_j}) \times Cov_b(y_{t_i}, y_{t_j})$$

## Linear $\times$ Linear $\rightarrow$ Quadratic

$$\text{Cov}_a(y_{t_i}, y_{t_j}) = 1 + 2 (t_i \times t_j)$$

$$\text{Cov}_b(y_{t_i}, y_{t_j}) = 2 + 1 (t_i \times t_j)$$

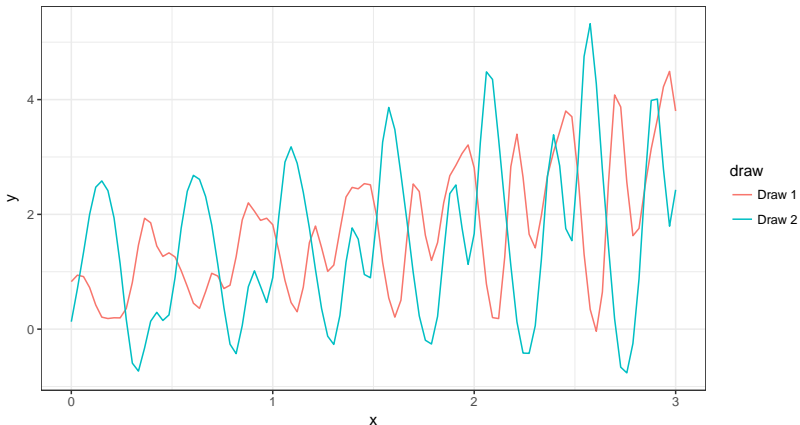


# Linear $\times$ Periodic

$$\text{Cov}_a(y_{t_i}, y_{t_j}) = 1 + 1 (t_i \times t_j)$$

$$\text{Cov}_b(y_{t_i}, y_{t_j}) = \text{Cov}(h = |t_i - t_j|) = \exp(-2 \sin^2(2\pi h))$$

Cov\_a \* Cov\_b

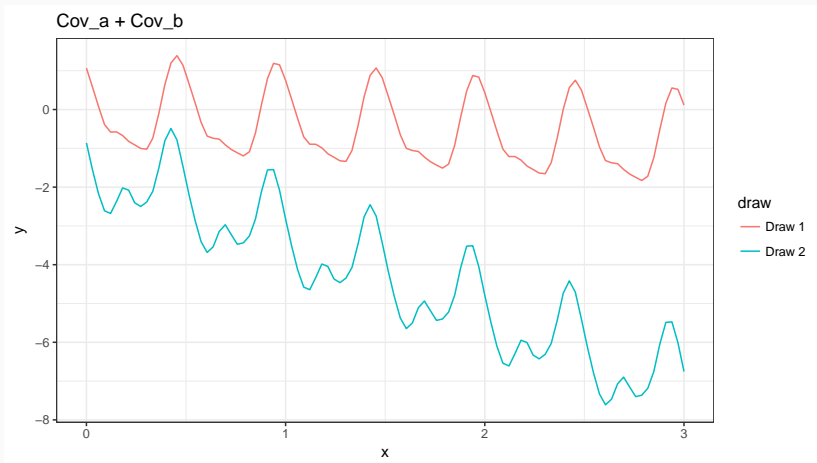




# Linear + Periodic

$$\text{Cov}_a(y_{t_i}, y_{t_j}) = 1 + 1 (t_i \times t_j)$$

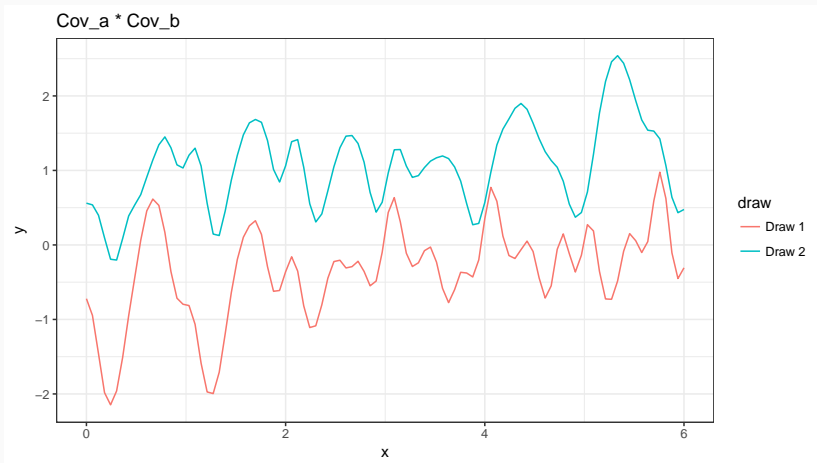
$$\text{Cov}_b(h = |t_i - t_j|) = \exp(-2 \sin^2(2\pi h))$$



## Sq Exp $\times$ Periodic $\rightarrow$ Locally Periodic

$$\text{Cov}_a(h = |t_i - t_j|) = \exp(-(1/3)h^2)$$

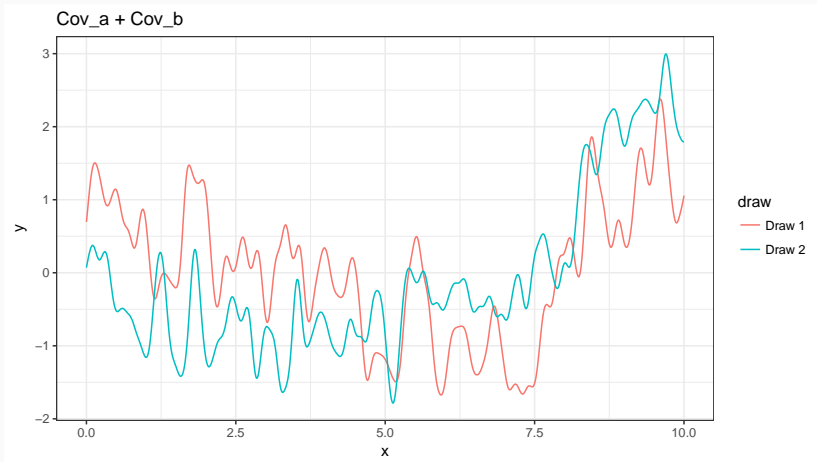
$$\text{Cov}_b(h = |t_i - t_j|) = \exp(-2 \sin^2(\pi h))$$



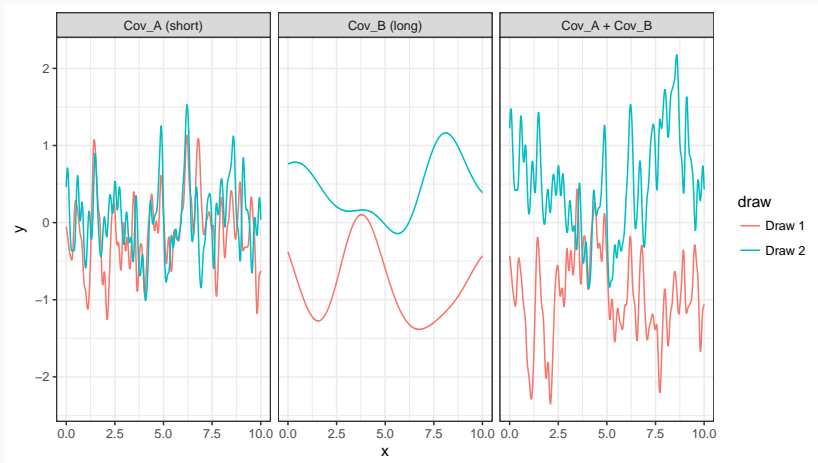
## Sq Exp (short) + Sq Exp (long)

$$\text{Cov}_a(h = |t_i - t_j|) = (1/4) \exp(-4\sqrt{3}h^2)$$

$$\text{Cov}_b(h = |t_i - t_j|) = \exp(-(\sqrt{3}/2)h^2)$$



## Sq Exp (short) + Sq Exp (long) (Seen another way)

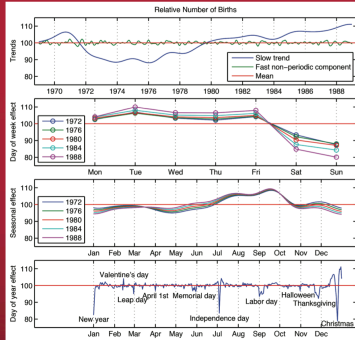


## BDA3 example

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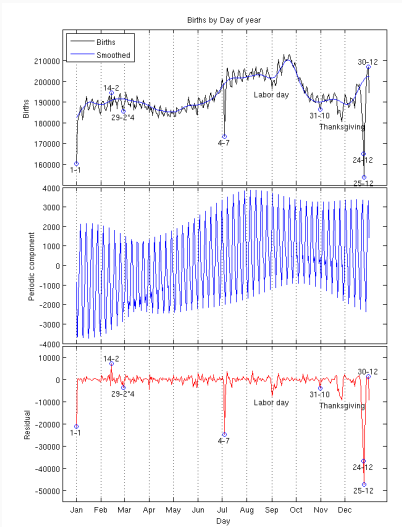
# Bayesian Data Analysis

## Third Edition



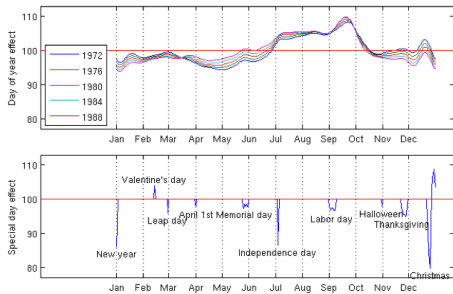
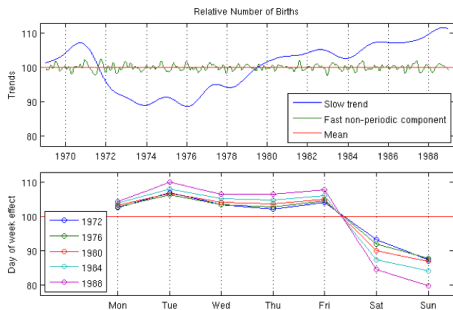
Andrew Gelman, John B. Carlin, Hal S. Stern,  
David B. Dunson, Aki Vehtari, and Donald B. Rubin

# Births (one year)



1. Smooth long term trend  
( $sq \exp cov$ )
2. Seven day periodic trend with decay  
( $periodic \times sq \exp cov$ )
3. Constant mean
4. Student t observation model

# Births (multiple years)



1. slowly changing trend ( $sq \exp cov$ )
2. small time scale correlating noise ( $sq \exp cov$ )
3. 7 day periodical component capturing day of week effect ( $periodic \times sq \exp cov$ )
4. 365.25 day periodical component capturing day of year effect ( $periodic \times sq \exp cov$ )
5. component to take into account the special days and interaction with weekends ( $linear cov$ )
6. independent Gaussian noise ( $nugget cov$ )
7. constant mean