

# Lecture 19

## Fitting CAR and SAR Models

---

Colin Rundel

03/29/2017

## Fitting areal models

---

- Simultaneous Autoregressive (SAR)

$$y(s_i) = \phi \sum_{j=1}^n W_{ij} y(s_j) + \epsilon$$

$$\mathbf{y} \sim \mathcal{N}(0, \sigma^2 ((I - \phi\mathbf{W})^{-1})(I - \phi\mathbf{W})^{-1})^t)$$

- Conditional Autoregressive (CAR)

$$y(s_i) | \mathbf{y}_{-s_i} \sim \mathcal{N} \left( \phi \sum_{j=1}^n W_{ij} y(s_j), \sigma^2 \right)$$

$$\mathbf{y} \sim \mathcal{N}(0, \sigma^2 (I - \phi\mathbf{W})^{-1})$$

## Some specific generalizations

Generally speaking we will want to work with a scaled / normalized version of the weight matrix,

$$\frac{W_{ij}}{W_i}$$

When  $W$  is an adjacency matrix we can express this as

$$D^{-1}W$$

where  $D = \text{diag}(m_i)$  and  $m_i = |N(s_i)|$ .

We can also allow  $\sigma^2$  to vary between locations, we can define this as  $D_\tau = \text{diag}(1/\sigma_i^2)$  and most often we use

$$D_\tau = \text{diag} \left( \frac{1}{\sigma^2/|N(s_i)|} \right) = D/\sigma^2$$

- Conditional Model

$$y(s_i) | \mathbf{y}_{-s_i} \sim \mathcal{N} \left( x_i \cdot \beta + \phi \sum_{j=1}^n \frac{W_{ij}}{D_{ii}} (y(s_j) - x_j \cdot \beta), \sigma^2 D_{ii}^{-1} \right)$$

- Joint Model

$$\mathbf{y} \sim \mathcal{N}(\mathbf{X}\beta, \Sigma_{CAR})$$

$$\begin{aligned} \Sigma_{CAR} &= (D_\sigma (I - \phi D^{-1} W))^{-1} \\ &= (1/\sigma^2 D (I - \phi D^{-1} W))^{-1} \\ &= (1/\sigma^2 (D - \phi W))^{-1} \\ &= \sigma^2 (D - \phi W)^{-1} \end{aligned}$$

- Formula Model

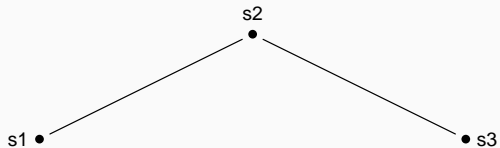
$$y(s_i) = x_i \cdot \beta + \phi \sum_{j=1}^n D_{ij}^{-1} W_{ij} (y(s_j) - x_j \cdot \beta) + \epsilon_i$$

- Joint Model

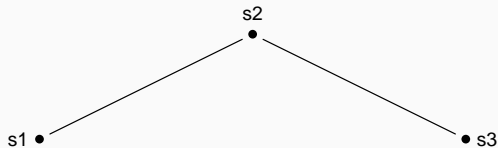
$$\begin{aligned} y &= X\beta + \phi D^{-1}W (y - X\beta) + \epsilon \\ (y - X\beta) &= \phi D^{-1}W (y - X\beta) + \epsilon \\ (y - X\beta)(I - \phi D^{-1}W)^{-1} &= \epsilon \\ y &= X\beta + (I - \phi D^{-1}W)^{-1}\epsilon \end{aligned}$$

$$y \sim \mathcal{N} \left( X\beta, (I - \phi D^{-1}W)^{-1} \sigma^2 D^{-1} ((I - \phi D^{-1}W)^{-1})^t \right)$$

## Toy CAR Example



## Toy CAR Example



$$W = \begin{pmatrix} 0 & 1 & 0 \\ 1 & 0 & 1 \\ 0 & 1 & 0 \end{pmatrix} \quad D = \begin{pmatrix} 1 & 0 & 0 \\ 0 & 2 & 0 \\ 0 & 0 & 1 \end{pmatrix}$$

$$\Sigma = \sigma^2 (D - \phi W) = \sigma^2 \begin{pmatrix} 1 & -\phi & 0 \\ -\phi & 2 & -\phi \\ 0 & -\phi & 1 \end{pmatrix}^{-1}$$



## When does $\Sigma$ exist?

```
check_sigma = function(phi) {  
  Sigma_inv = matrix(c(1,-phi,0,-phi,2,-phi,0,-phi,1), ncol=3, byrow=TRUE)  
  solve(Sigma_inv)  
}
```

```
check_sigma(phi=0)  
##      [,1] [,2] [,3]  
## [1,]    1  0.0  0  
## [2,]    0  0.5  0  
## [3,]    0  0.0  1
```

```
check_sigma(phi=0.5)  
##      [,1]      [,2]      [,3]  
## [1,] 1.1666667 0.3333333 0.1666667  
## [2,] 0.3333333 0.6666667 0.3333333  
## [3,] 0.1666667 0.3333333 1.1666667
```

```
check_sigma(phi=-0.6)  
##      [,1]      [,2]      [,3]  
## [1,] 1.28125 -0.46875 0.28125  
## [2,] -0.46875 0.78125 -0.46875  
## [3,] 0.28125 -0.46875 1.28125
```

```
check_sigma(phi=1)
```

```
## Error in solve.default(Sigma_inv): Lapack routine dgesv: system is exactl
```

```
check_sigma(phi=-1)
```

```
## Error in solve.default(Sigma_inv): Lapack routine dgesv: system is exactl
```

```
check_sigma(phi=1.2)
```

```
##           [,1]      [,2]      [,3]
```

```
## [1,] -0.6363636 -1.363636 -1.6363636
```

```
## [2,] -1.3636364 -1.136364 -1.3636364
```

```
## [3,] -1.6363636 -1.363636 -0.6363636
```

```
check_sigma(phi=-1.2)
```

```
##           [,1]      [,2]      [,3]
```

```
## [1,] -0.6363636  1.363636 -1.6363636
```

```
## [2,]  1.3636364 -1.136364  1.3636364
```

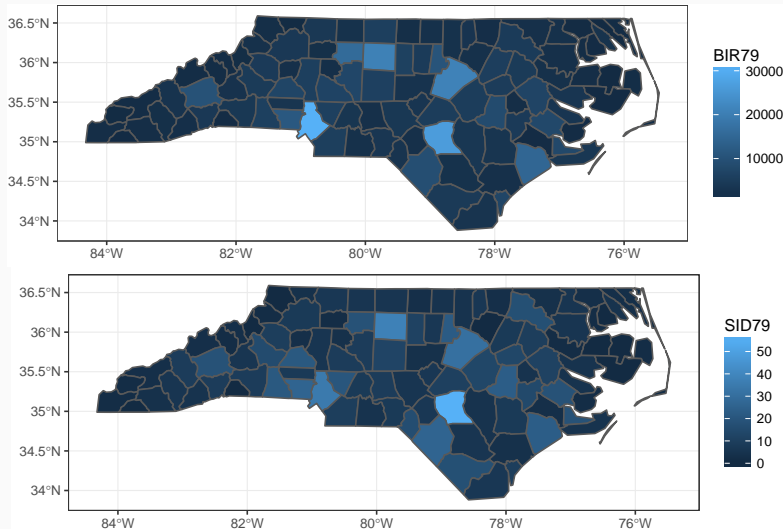
```
## [3,] -1.6363636  1.363636 -0.6363636
```

Generally speaking just like the AR(1) model for time series we require that  $|\phi| < 1$  for the CAR model to be proper.

These results for  $\phi$  also apply in the context where  $\sigma_i^2$  is constant across locations (i.e.  $\Sigma = (\sigma^2 (I - \phi D^{-1} W))^{-1}$ )

As a side note, the special case where  $\phi = 1$  is known as an intrinsic autoregressive (IAR) model and they are popular as an *improper* prior for spatial random effects. An additional sum constraint is necessary for identifiability ( $\sum_i y(s_i) = 0$ ).

# Example - NC SIDS



## Using `spautolm` from `spdep`

```
library(spdep)

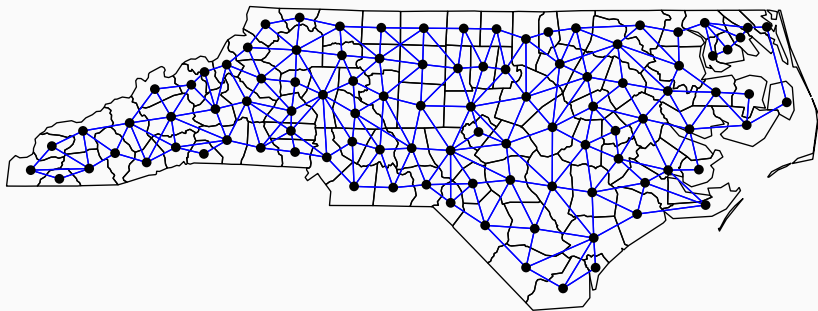
W = st_touches(nc, sparse=FALSE)
listW = mat2listw(W)

listW
## Characteristics of weights list object:
## Neighbour list object:
## Number of regions: 100
## Number of nonzero links: 490
## Percentage nonzero weights: 4.9
## Average number of links: 4.9
##
## Weights style: M
## Weights constants summary:
##      n   nn  S0  S1   S2
## M 100 10000 490 980 10696
```

```
nc_coords = nc %>% st_centroid() %>% st_coordinates()
```

```
plot(st_geometry(nc))
```

```
plot(listW, nc_coords, add=TRUE, col="blue", pch=16)
```



# CAR Model

```
nc_car = spautolm(formula = SID74 ~ BIR74, data = nc,
                  listw = listW, family = "CAR")

summary(nc_car)
##
## Call:
## spautolm(formula = SID74 ~ BIR74, data = nc, listw = listW, family = "CAR")
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -10.38934  -1.58600  -0.52154   1.14729  13.54059
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  1.06911902  0.67501301   1.5838   0.1132
## BIR74         0.00175249  0.00010107  17.3401  <2e-16
##
## Lambda: 0.13222 LR test value: 8.8654 p-value: 0.0029062
## Numerical Hessian standard error of lambda: 0.030094
##
## Log likelihood: -275.7655
## ML residual variance (sigma squared): 13.695, (sigma: 3.7007)
## Number of observations: 100
## Number of parameters estimated: 4
## AIC: 559.53
```

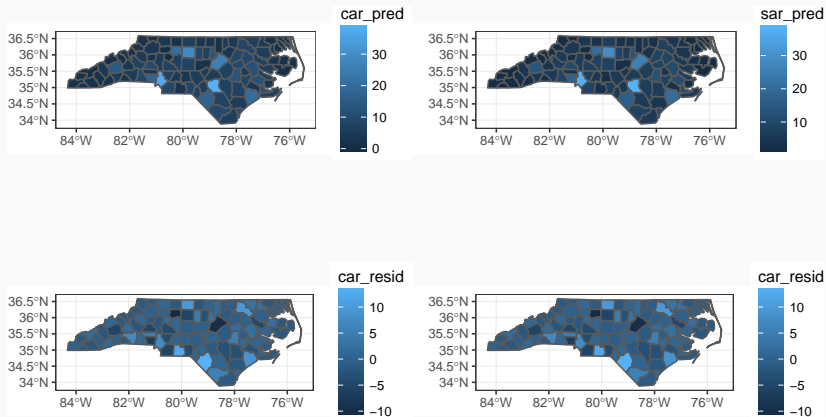
# SAR Model

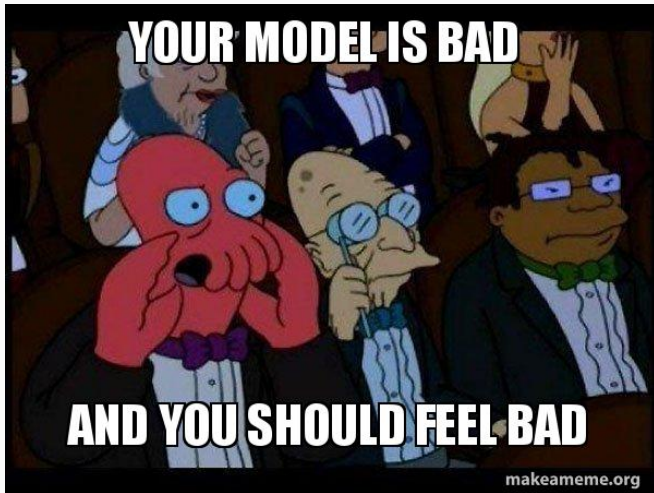
```
nc_sar = spautolm(formula = SID74 ~ BIR74, data = nc,
                  listw = listW, family = "SAR")

summary(nc_sar)
##
## Call:
## spautolm(formula = SID74 ~ BIR74, data = nc, listw = listW, family = "SAR")
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -10.94771  -1.72354  -0.56866   1.23273  14.70511
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  1.01971585  0.64910408   1.571   0.1162
## BIR74         0.00174741  0.00010105  17.292  <2e-16
##
## Lambda: 0.075265 LR test value: 8.4013 p-value: 0.0037495
## Numerical Hessian standard error of lambda: 0.024085
##
## Log likelihood: -275.9975
## ML residual variance (sigma squared): 14.158, (sigma: 3.7627)
## Number of observations: 100
## Number of parameters estimated: 4
## AIC: 560
```

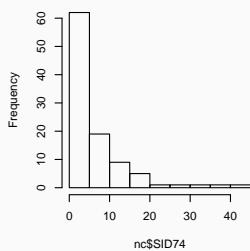


# Residuals

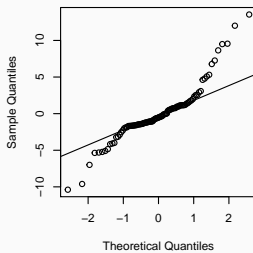




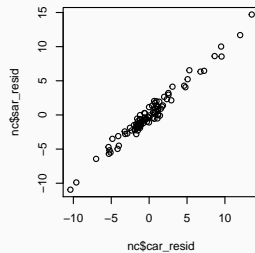
### Histogram of nc\$SID74



### CAR Residuals



### CAR vs SAR Residuals



## Jags CAR Model

```
## model{  
##   y ~ dnorm(beta0 + beta1*x, tau * (D - phi*W))  
##   y_pred ~ dnorm(beta0 + beta1*x, tau * (D - phi*W))  
##  
##   beta0 ~ dnorm(0, 1/100)  
##   beta1 ~ dnorm(0, 1/100)  
##  
##   tau <- 1 / sigma2  
##   sigma2 ~ dnorm(0, 1/100) T(0,)  
##   phi ~ dunif(-0.99, 0.99)  
## }
```

```
y = nc$SID74
```

```
x = nc$BIR74
```

```
W = W * 1L
```

```
D = diag(rowSums(W))
```

## Jags CAR Model

```
## model{  
##   y ~ dnorm(beta0 + beta1*x, tau * (D - phi*W))  
##   y_pred ~ dnorm(beta0 + beta1*x, tau * (D - phi*W))  
##  
##   beta0 ~ dnorm(0, 1/100)  
##   beta1 ~ dnorm(0, 1/100)  
##  
##   tau <- 1 / sigma2  
##   sigma2 ~ dnorm(0, 1/100) T(0,)  
##   phi ~ dunif(-0.99, 0.99)  
## }
```

```
y = nc$SID74
```

```
x = nc$BIR74
```

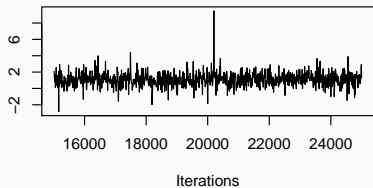
```
W = W * 1L
```

```
D = diag(rowSums(W))
```

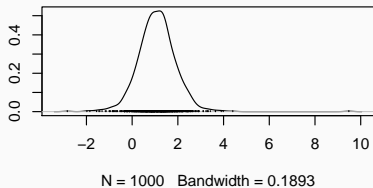
Why don't we use the conditional definition for the  $y$ 's?

# Model Results

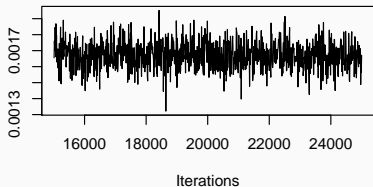
### Trace of beta0



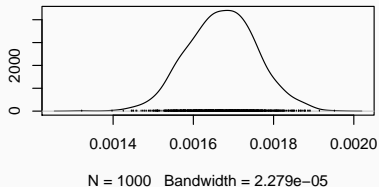
### Density of beta0



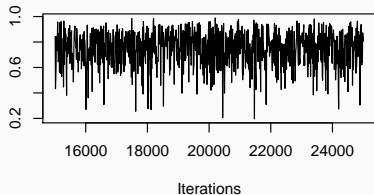
### Trace of beta1



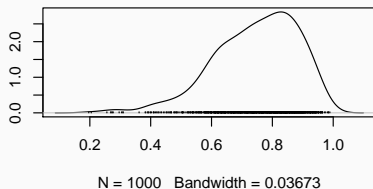
### Density of beta1



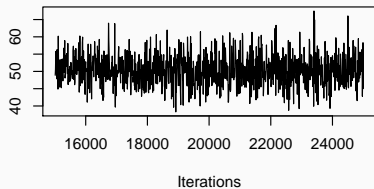
**Trace of phi**



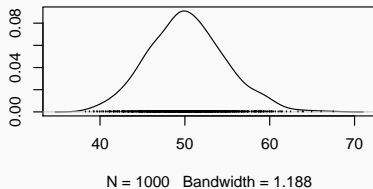
**Density of phi**



**Trace of sigma2**



**Density of sigma2**



# Predictions

```
nc$jags_pred = y_pred$post_mean  
nc$jags_resid = nc$SID74 - y_pred$post_mean
```

```
sqrt(mean(nc$jags_resid^2))
```

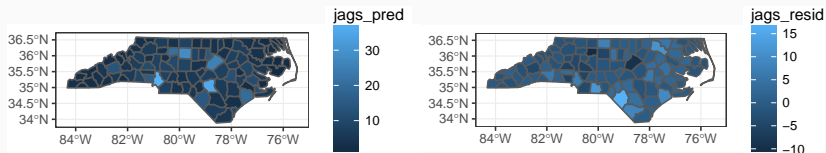
```
## [1] 3.987985
```

```
sqrt(mean(nc$car_resid^2))
```

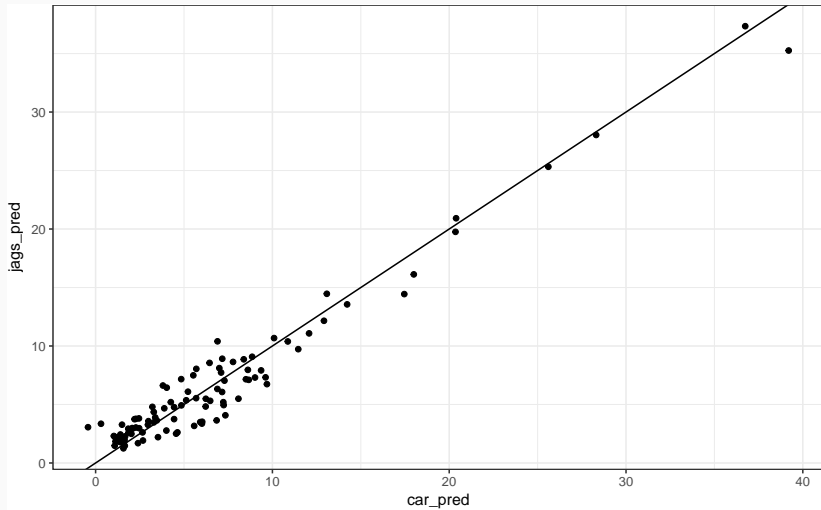
```
## [1] 3.72107
```

```
sqrt(mean(nc$sar_resid^2))
```

```
## [1] 3.762664
```







$$\Sigma_{SAR} = (I - \phi D^{-1} W)^{-1} \sigma^2 D^{-1} ((I - \phi D^{-1} W)^{-1})^t$$

$$\begin{aligned}\Sigma_{SAR}^{-1} &= \left( (I - \phi D^{-1} W)^{-1} \sigma^2 D^{-1} ((I - \phi D^{-1} W)^{-1})^t \right)^{-1} \\ &= \left( ((I - \phi D^{-1} W)^{-1})^t \right)^{-1} \frac{1}{\sigma^2} D (I - \phi D^{-1} W) \\ &= \frac{1}{\sigma^2} (I - \phi D^{-1} W)^t D (I - \phi D^{-1} W)\end{aligned}$$

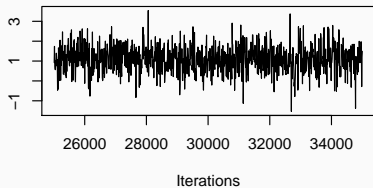
## Jags SAR Model

```
## model{  
##   y ~ dnorm(beta0 + beta1*x, tau * (D - phi*W))  
##   y_pred ~ dnorm(beta0 + beta1*x, tau * (D - phi*W))  
##  
##   beta0 ~ dnorm(0, 1/100)  
##   beta1 ~ dnorm(0, 1/100)  
##  
##   tau <- 1 / sigma2  
##   sigma2 ~ dnorm(0, 1/100) T(0,)  
##   phi ~ dunif(-0.99, 0.99)  
## }
```

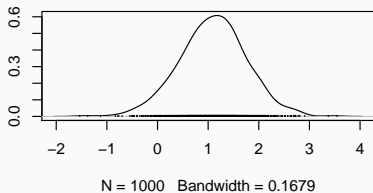
```
D_inv = diag(1/diag(D))  
W_tilde = D_inv %**% W  
I = diag(1, ncol=length(y), nrow=length(y))
```

# Model Results

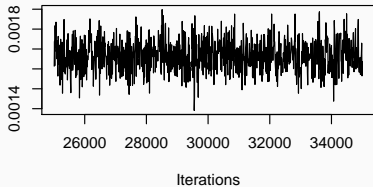
### Trace of beta0



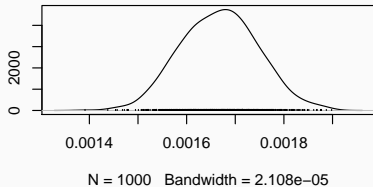
### Density of beta0



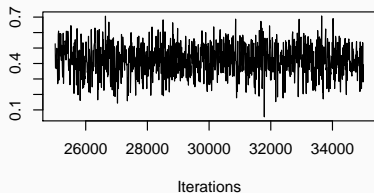
### Trace of beta1



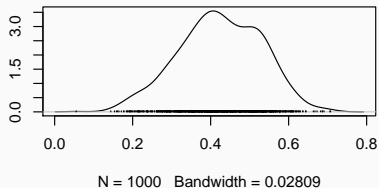
### Density of beta1



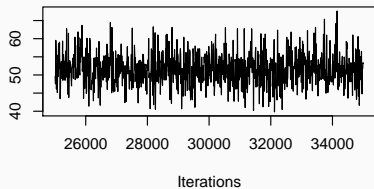
**Trace of phi**



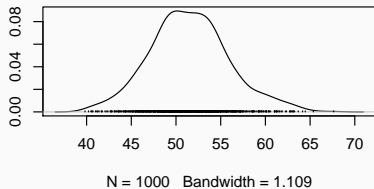
**Density of phi**



**Trace of sigma2**



**Density of sigma2**



# Comparing Model Results

