

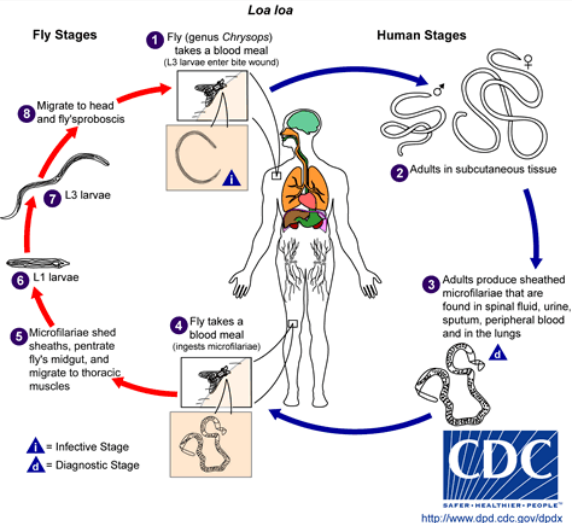
Lecture 21

More Spatial Random Effects Models

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04/10/2017

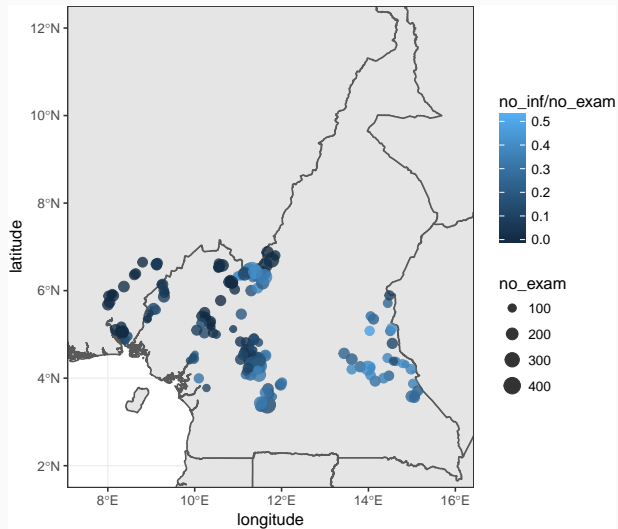
Loa Loa Example



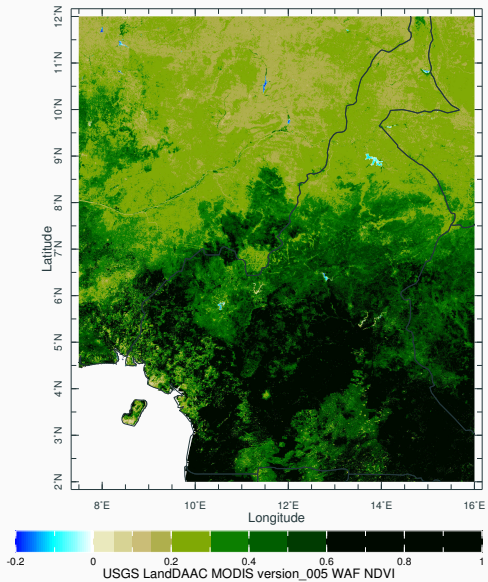
```
library(PrevMap)
loaloa = tbl_df(loaloa) %>% setNames(., tolower(names(.)))

loaloa
## # A tibble: 197 × 11
##   row villcode longitude latitude no_exam no_inf elevation mean9901 m
##   <int> <int> <dbl> <dbl> <int> <int> <int> <dbl>
## 1     1     214  8.041860  5.736750     162     0     108 0.4389815
## 2     2     215  8.004330  5.680280     167     1     99 0.4258333
## 3     3     118  8.905556  5.347222     88     5     783 0.4914815
## 4     4     219  8.100720  5.917420     62     5     104 0.4324074
## 5     5     212  8.182510  5.104540     167     3     109 0.4150000
## 6     6     116  8.929167  5.355556     66     3     909 0.4363889
## 7     7      16 11.360000  4.885000     163    11     503 0.5019444
## 8     8     217  8.067490  5.897800     83     0     103 0.3731481
## 9     9     112  9.018056  5.593056     30     4     751 0.4808333
## 10    10     104  9.312500  6.004167     57     4     268 0.4865741
## # ... with 187 more rows, and 2 more variables: min9901 <dbl>, stdev9901
```

Spatial Distribution



Normalized Difference Vegetation Index (NDVI)



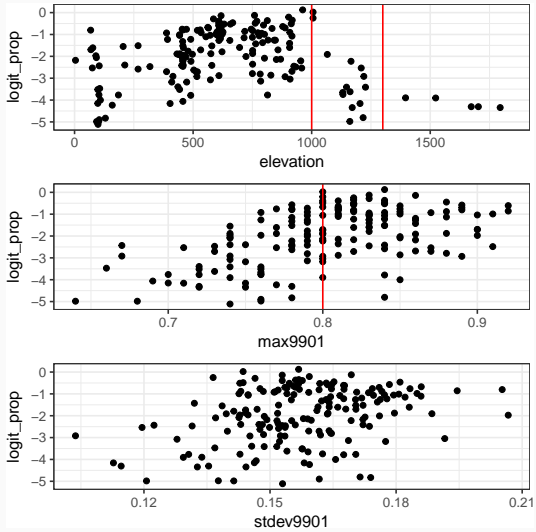
Original paper - Diggle, et. al. (2007). *Spatial modelling and prediction of Loa loa risk: decision making under uncertainty*. Annals of Tropical Medicine and Parasitology, 101, 499-509.

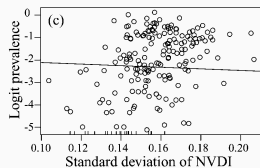
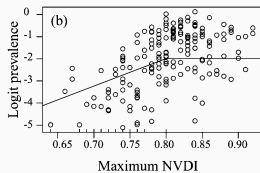
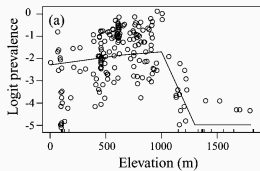
- `no_exam` and `no_inf` - Collected between 1991 and 2001 by NGOs (original paper mentions 168 villages and 21,938 observations)
- `elevation` - USGS gtopo30 (1km resolution)
- `mean9901` to `stdev9901` - aggregated data from 1999 to 2001 from the Flemish Institute for Technological Research (1 km resolution)

$$\log\left(\frac{p(x)}{1-p(x)}\right) = \alpha + f_1(\text{ELEVATION}) + f_2(\text{max (NDVI)}) + f_3(\text{sd (NDVI)}) + S(X)$$

where

$$s(x) \sim \mathcal{N}(0, \Sigma)$$
$$\{\Sigma\}_{ij} = \sigma^2 \exp(-d \phi)$$





Model EDA

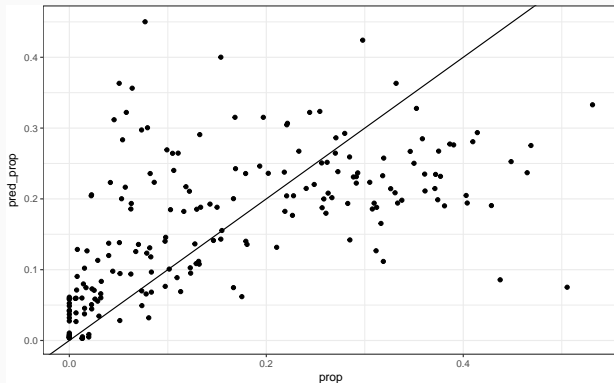
```
loaloe = loaloe %>%
  mutate(elev_factor = cut(elevation, breaks=c(0,1000,1300,2000), dig.lab=5),
         max_factor = cut(max9901, breaks=c(0,0.8,1)))

g = glm(no_inf/no_exam ~ elevation:elev_factor + max9901:max_factor + stdev9901,
       data=loaloe, family=binomial, weights=loaloe$no_exam)

summary(g)
##
## Call:
## glm(formula = no_inf/no_exam ~ elevation:elev_factor + max9901:max_factor +
##      stdev9901, family = binomial, data = loaloe, weights = loaloe$no_exam)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -7.1434  -2.5887  -0.8993   1.6375  10.9052
##
## Coefficients:
##
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -8.343e+00  4.825e-01 -17.291 < 2e-16 ***
## stdev9901    8.781e+00  1.205e+00   7.288 3.14e-13 ***
## elevation:elev_factor(0,1000]  1.606e-03  8.749e-05  18.358 < 2e-16 ***
## elevation:elev_factor(1000,1300]  1.631e-04  8.792e-05   1.855  0.0636 .
## elevation:elev_factor(1300,2000] -1.432e-03  1.887e-04  -7.588 3.25e-14 ***
## max9901:max_factor(0,0.8]    5.511e+00  6.299e-01   8.749 < 2e-16 ***
## max9901:max_factor(0.8,1]    5.626e+00  5.793e-01   9.711 < 2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

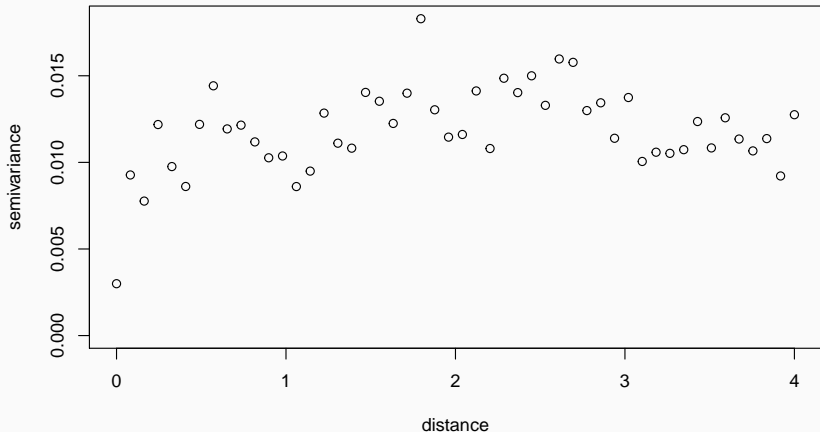
Residuals

```
loaloa = loaloa %>%  
  mutate(pred_prop = predict(g, type="response"),  
         resid = prop - pred_prop)  
  
ggplot(loaloa, aes(x=prop, y=pred_prop)) +  
  geom_point() +  
  geom_abline(slope = 1, intercept = 0)
```



```
library(geoR)
```

```
variog(coords = cbind(loaloo$longitude, loaloo$latitude),  
        data = loaloo$resid,  
        uvec = seq(0, 4, length.out = 50)) %>% plot()  
## variog: computing omnidirectional variogram
```



spBayes GLM Model

```
library(spBayes)

spg = spGLM(no_inf/no_exam ~ elevation:elev_factor + max9901:max_factor + stdev9901,
            data=loaloe, family="binomial", weights=loaloe$no_exam,
            coords=cbind(loaloe$longitude, loaloe$latitude),
            cov.model="exponential", n.samples=20000,
            #starting=list(beta=coefficients(g), phi=9, sigma.sq=1, w=0),
            starting=list(beta=rep(0,7), phi=3, sigma.sq=1, w=0),
            priors=list(phi.unif=c(0.1, 10), sigma.sq.ig=c(2, 2)),
            amcmc=list(n.batch=1000, batch.length=20, accept.rate=0.43))

save(spg, loaloe, file="loaloe.Rdata")
```

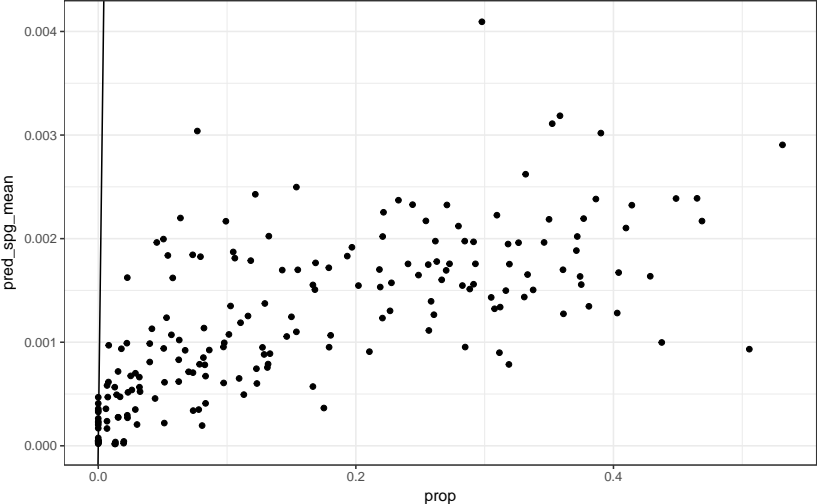
```

spg$p.beta.theta.samples %>%
  post_summary() %>%
  knitr::kable(digits=5)

```

param	post_mean	post_med	post_lower	post_upper
(Intercept)	-12.69885	-11.61326	-21.65388	-6.96361
stdev9901	9.24231	9.15244	-14.48649	29.76058
elevation:elev_factor(0,1000]	0.00048	0.00077	-0.00474	0.00291
elevation:elev_factor(1000,1300]	-0.00048	-0.00032	-0.00359	0.00169
elevation:elev_factor(1300,2000]	-0.00814	-0.00581	-0.02900	0.00004
max9901:max_factor(0,0.8]	4.87762	3.99492	-2.93030	15.63246
max9901:max_factor(0.8,1]	5.08690	4.44632	-2.18626	14.89011
sigma.sq	0.38088	0.34626	0.12793	0.88673
phi	6.22996	5.18205	0.69584	18.67107

Prediction



spBayes GLM Model - Fixed?

```
library(spBayes)

spg_good = spGLM(no_inf ~ elevation:elev_factor + max9901:max_factor + stdev9901,
  data=loaloe, family="binomial", weights=loaloe$no_exam,
  coords=cbind(loaloe$longitude, loaloe$latitude),
  cov.model="exponential", n.samples=20000,
  #starting=list(beta=coefficients(g), phi=9, sigma.sq=1, w=0),
  starting=list(beta=rep(0,7), phi=3, sigma.sq=1, w=0),
  priors=list(phi.unif=c(0.1, 10), sigma.sq.ig=c(2, 2)),
  amcmc=list(n.batch=1000, batch.length=20, accept.rate=0.43))

save(spg_good, loaloe, file="loaloe_good.Rdata")
```

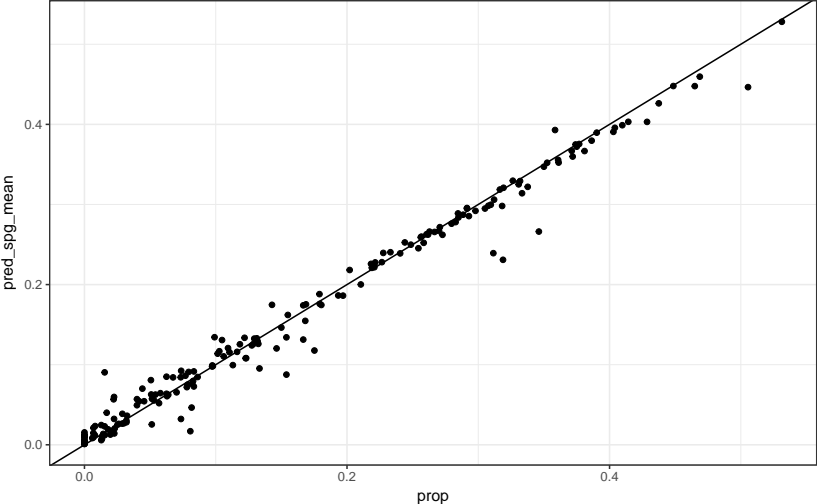
```

spg_good$p.beta.theta.samples %>%
  post_summary() %>%
  knitr::kable(digits=5)

```

param	post_mean	post_med	post_lower	post_upper
(Intercept)	-2.66090	-2.13138	-6.31576	-0.80487
stdev9901	-0.12840	-0.41947	-5.86766	8.58835
elevation:elev_factor(0,1000]	0.00023	0.00024	-0.00051	0.00086
elevation:elev_factor(1000,1300]	-0.00054	-0.00055	-0.00128	0.00020
elevation:elev_factor(1300,2000]	-0.00204	-0.00200	-0.00285	-0.00127
max9901:max_factor(0,0.8]	0.88041	0.90550	-1.03795	3.63477
max9901:max_factor(0.8,1]	1.28673	1.13796	-0.26884	3.83860
sigma.sq	1.47552	1.39146	0.43359	3.05883
phi	2.22372	2.09524	0.86456	4.14663

Prediction



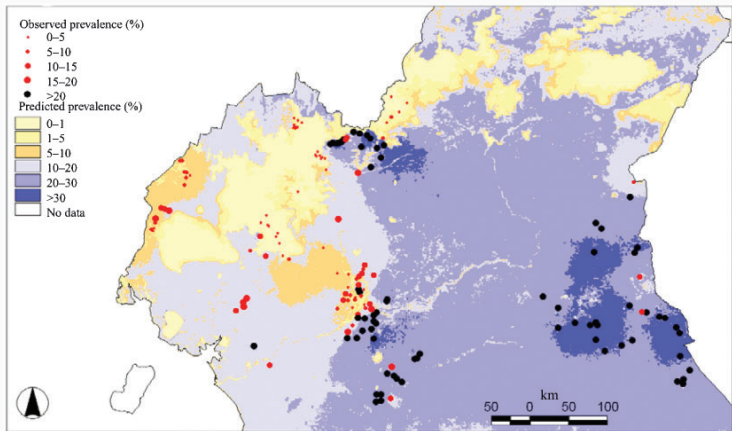
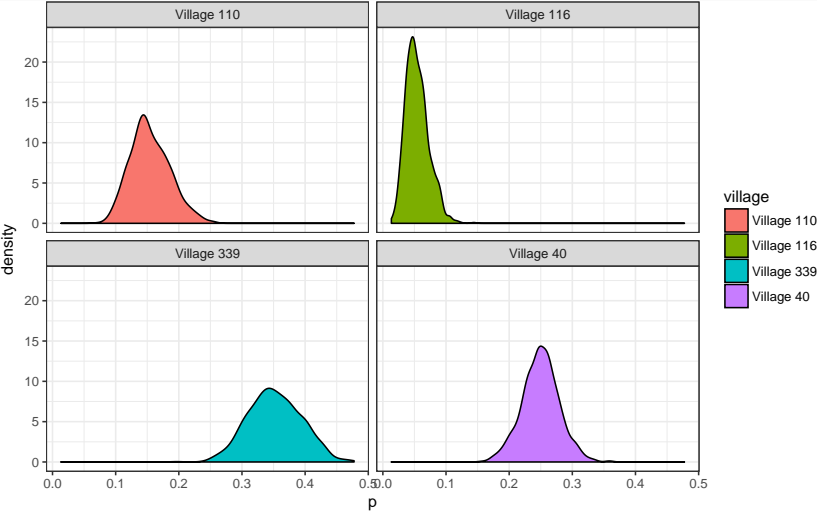


FIG. 2. Point estimates of the prevalence of *Loa loa* microfilaraemia, over-laid with the prevalences observed in field studies.

Exceedance Probability - Posterior Summary



Exceedance Probability Predictive Surface

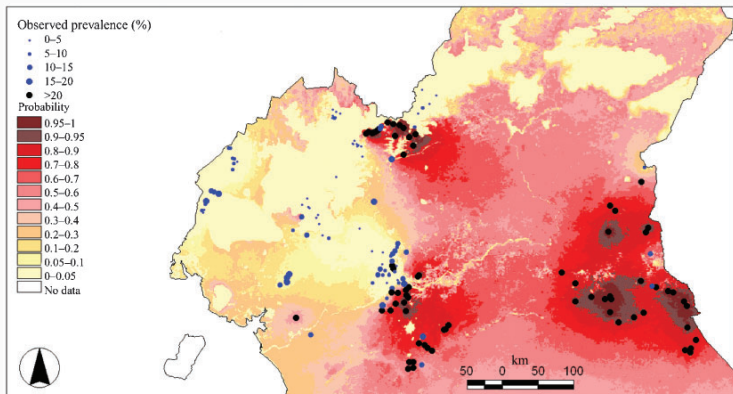


FIG. 4. A probability contour map, indicating the probability that the prevalence of *Loa loa* microfilaraemia in each area exceeds 20%, over-laid with the prevalences observed in field studies.

Spatial Assignment of Migratory Birds

Using intrinsic markers (genetic and isotopic signals) for the purpose of inferring migratory connectivity.

- Existing methods are too coarse for most applications
- Large amounts of data are available (>150,000 feather samples from >500 species)
- Genetic assignment methods are based on Wasser, et al. (2004)
- Isotopic assignment methods are based on Wunder, et al. (2005)

Hermit Thrush (*Catharus guttatus*)

- 138 individuals
- 14 locations
- 6 loci
- 9-27 alleles / locus

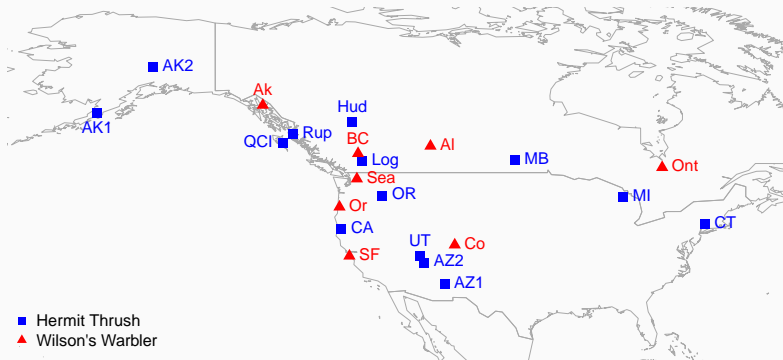


Wilson's Warbler (*Wilsonia pusilla*)

- 163 individuals
- 8 locations
- 9 loci
- 15-31 alleles / locus



Sampling Locations



Allele Frequency Model

For the allele i , from locus l , at location k

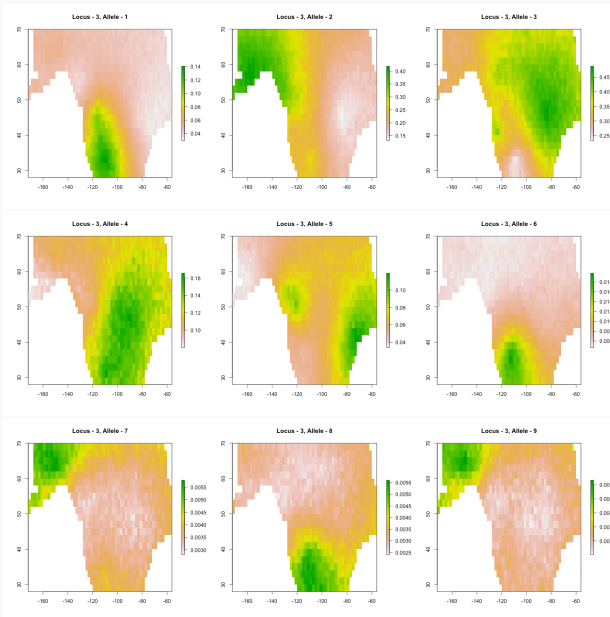
$$y_{\cdot lk} | \Theta \sim \mathcal{N}(\sum_i y_{ilk}, f_{\cdot lk})$$

$$f_{ilk} = \frac{\exp(\Theta_{ilk})}{\sum_i \exp(\Theta_{ilk})}$$

$$\Theta_{il} | \alpha, \mu \sim \mathcal{N}(\mu_{il}, \Sigma)$$

$$\{\Sigma\}_{ij} = \sigma^2 \exp\left(-(\{d\}_{ij} r)^\psi\right) + \sigma_n^2 1_{i=j}$$

Predictions by Allele (Locus 3)



Assignment model assuming Hardy-Weinberg equilibrium and allowing for genotyping (δ) and single amplification (γ) errors.

$$P(S_G|\mathbf{f}, k) = \prod_l P(i_l, j_l|\mathbf{f}, k)$$

$$P(i_l, j_l|\mathbf{f}, k) = \begin{cases} \gamma P(i_l|\mathbf{f}, k) + (1 - \gamma)P(i_l|\tilde{\mathbf{f}}, k)^2 & \text{if } i = j \\ (1 - \gamma)P(i_l|\mathbf{f}, k)P(j_l|\mathbf{f}, k) & \text{if } i \neq j \end{cases}$$

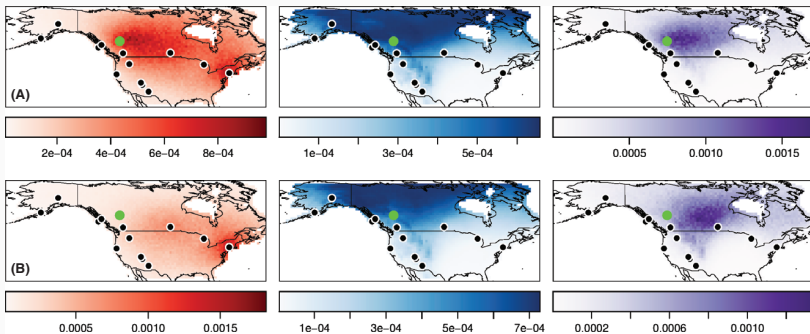
$$P(i_l|\mathbf{f}, k) = (1 - \delta)f_{lik} + \delta/m_l$$

Combined Model

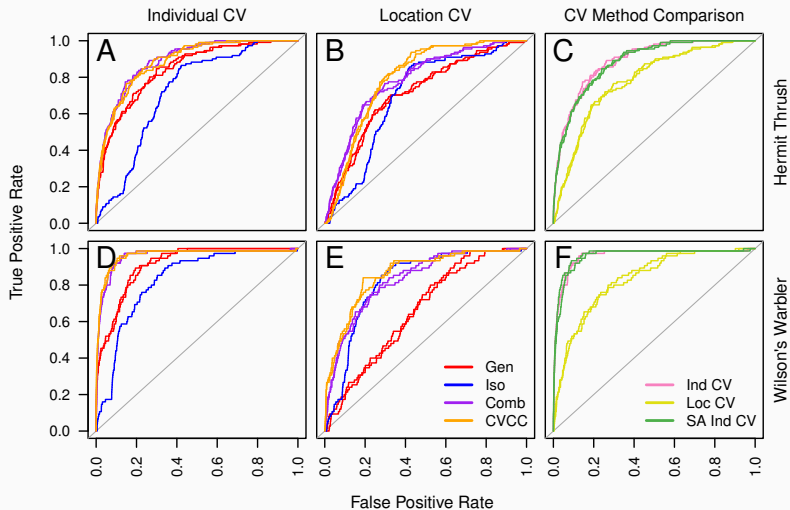
Genetic

Isotopic

Combined



Model Assessment



Migratory Connectivity

