

# Lecture 23

## Spatio-temporal Models

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## Spatial Models with AR time dependence

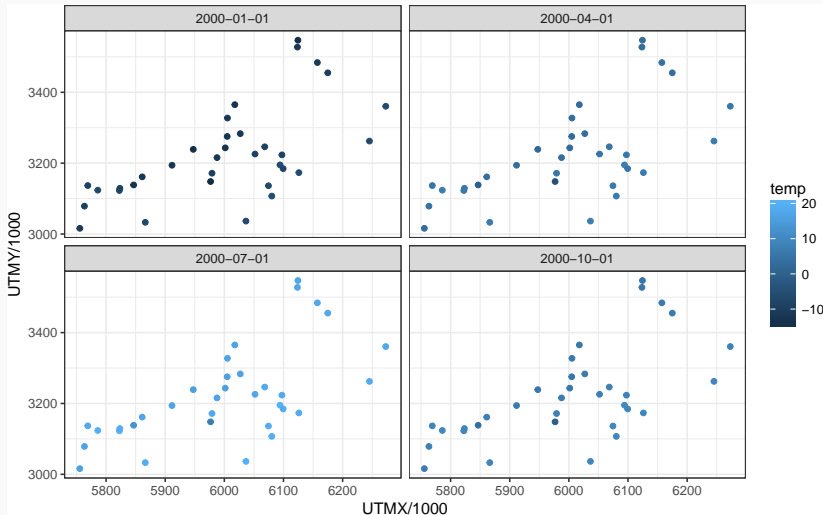
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## Example - Weather station data

Based on Andrew Finley and Sudipto Banerjee's notes from National Ecological Observatory Network (NEON) Applied Bayesian Regression Workshop, March 7 - 8, 2013 Module 6

NETemp.dat - Monthly temperature data (Celsius) recorded across the Northeastern US starting in January 2000.

```
library(spBayes)
data("NETemp.dat")
ne_temp = NETemp.dat %>%
  filter(UTMX > 5.5e6, UTMX > 3e6) %>%
  select(1:27) %>%
  tbl_df()
ne_temp
## # A tibble: 34 × 27
##   elev   UTMX   UTMY     y.1     y.2     y.3     y.4     y.5
##   <int> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl> <dbl>
## 1    102 6094162 3195181 -6.388889 -3.611111  3.722222  6.777778 12.555556
## 2     1 6245390 3262354 -6.277778 -4.111111  2.611111  6.555556 11.388889
## 3    157 6157302 3484043 -11.111111 -9.444444 -0.388889  3.944444  9.888889
## 4    176 6123610 3527665 -11.611111 -9.722222 -1.166667  2.888889  9.666667
## 5    400 6004871 3275456 -12.611111 -9.055556 -1.611111  2.555556  8.555556
## 6    133 6051946 3225830 -9.111111 -6.388889  1.222222  4.944444 10.888889
## 7     56 6099462 3184587 -7.944444 -6.055556  2.055556  5.555556 11.111111
## 8     59 6074601 3136288 -6.555556 -3.500000  3.166667  6.166667 11.500000
## 9    160 6174891 3455064 -9.944444 -8.944444 -0.277778  3.555556  9.611111
## 10   360 6005282 3327413 -12.277778 -9.444444 -1.500000  2.944444  9.000000
## # ... with 24 more rows, and 19 more variables: y.6 <dbl>, y.7 <dbl>,
## #   y.8 <dbl>, y.9 <dbl>, y.10 <dbl>, y.11 <dbl>, y.12 <dbl>, y.13 <dbl>
```



$$y_t = \underset{1 \times p}{F_t'} \underset{p \times 1}{\theta_t} + v_t \quad \text{observation equation}$$

$$\underset{p \times 1}{\theta_t} = \underset{p \times p}{G_t} \underset{p \times 1}{\theta_{t-1}} + \underset{p \times 1}{\omega_t} \quad \text{evolution equation}$$

$$v_t \sim \mathcal{N}(0, V_t)$$

$$\omega_t \sim \mathcal{N}(0, W_t)$$

ARMA / ARIMA are a special case of a dynamic linear model, for example an  $AR(p)$  can be written as

$$F_t' = (1, 0, \dots, 0)$$
$$G_t = \begin{pmatrix} \phi_1 & \phi_2 & \cdots & \phi_{p-1} & \phi_p \\ 1 & 0 & \cdots & 0 & 0 \\ 0 & 1 & \cdots & 0 & 0 \\ \vdots & \vdots & \ddots & \vdots & 0 \\ 0 & 0 & \cdots & 1 & 0 \end{pmatrix}$$

$$\omega_t = (\omega_1, 0, \dots, 0), \quad \omega_1 \sim \mathcal{N}(0, \sigma^2)$$

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$$\omega_t = (\omega_1, 0, \dots, 0), \quad \omega_1 \sim \mathcal{N}(0, \sigma^2)$$

$$y_t = \theta_t + v_t, \quad v_t \sim \mathcal{N}(0, \sigma_v^2)$$

$$\theta_t = \sum_{i=1}^p \phi_i \theta_{t-i} + \omega_1, \quad \omega_1 \sim \mathcal{N}(0, \sigma_\omega^2)$$

The observed temperature at time  $t$  and location  $s$  is given by  $y_t(s)$  where,

$$y_t(s) = x_t(s)\beta_t + u_t(s) + \epsilon_t(s)$$

$$\epsilon_t(s) \stackrel{ind.}{\sim} \mathcal{N}(0, \tau_t^2)$$

$$\beta_t = \beta_{t-1} + \eta_t$$

$$\eta_t \stackrel{i.i.d.}{\sim} \mathcal{N}(0, \Sigma_\eta)$$

$$u_t(s) = u_{t-1}(s) + w_t(s)$$

$$w_t(s) \stackrel{ind.}{\sim} \mathcal{N}(\mathbf{0}, \Sigma_t(\phi_t, \sigma_t^2))$$



## Dynamic spatio-temporal models

The observed temperature at time  $t$  and location  $s$  is given by  $y_t(s)$  where,

$$y_t(s) = x_t(s)\beta_t + u_t(s) + \epsilon_t(s)$$

$$\epsilon_t(s) \stackrel{\text{ind.}}{\sim} \mathcal{N}(0, \tau_t^2)$$

$$\beta_t = \beta_{t-1} + \eta_t$$

$$\eta_t \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(0, \Sigma_\eta)$$

$$u_t(s) = u_{t-1}(s) + w_t(s)$$

$$w_t(s) \stackrel{\text{ind.}}{\sim} \mathcal{N}(\mathbf{0}, \Sigma_t(\phi_t, \sigma_t^2))$$

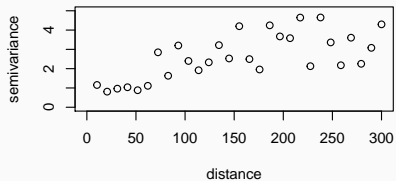
Additional assumptions for  $t = 0$ ,

$$\beta_0 \sim \mathcal{N}(\boldsymbol{\mu}_0, \boldsymbol{\Sigma}_0)$$

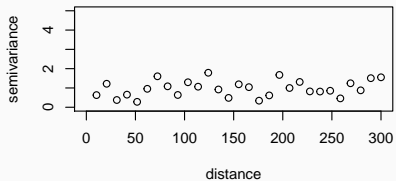
$$u_0(s) = 0$$

# Variograms by time

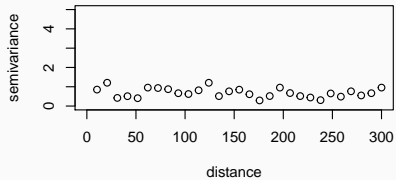
**Jan 2000**



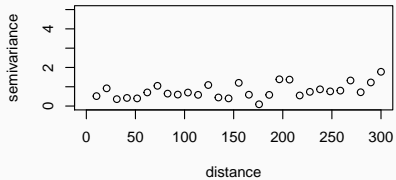
**Apr 2000**



**Jul 2000**



**Oct 2000**



## Data and Model Parameters

**\*\*Data\*:**

```
coords = ne_temp %>% select(UTMX, UTM Y) %>% as.matrix() / 1000  
y_t = ne_temp %>% select(starts_with("y.")) %>% as.matrix()
```

```
max_d = coords %>% dist() %>% max()  
n_t = ncol(y_t)  
n_s = nrow(y_t)
```

**\*\*Parameters\*:**

```
n_beta = 2  
starting = list(  
  beta = rep(0, n_t * n_beta), phi = rep(3/(max_d/2), n_t),  
  sigma.sq = rep(1, n_t), tau.sq = rep(1, n_t),  
  sigma.eta = diag(0.01, n_beta)  
)  
  
tuning = list(phi = rep(1, n_t))  
  
priors = list(  
  beta.0.Norm = list(rep(0, n_beta), diag(1000, n_beta)),  
  phi.Unif = list(rep(3/(0.9 * max_d), n_t), rep(3/(0.05 * max_d), n_t)),  
  sigma.sq.IG = list(rep(2, n_t), rep(2, n_t)),  
  tau.sq.IG = list(rep(2, n_t), rep(2, n_t)),  
  sigma.eta.IW = list(2, diag(0.001, n_beta))  
)
```

## Fitting with spDynLM from spBayes

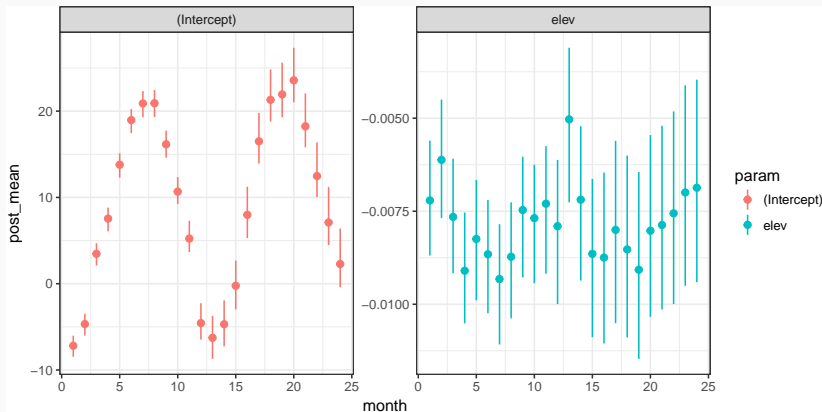
```
n_samples = 10000
models = lapply(paste0("y.",1:24, "~elev"), as.formula)

m = spDynLM(
  models, data = ne_temp, coords = coords, get.fitted = TRUE,
  starting = starting, tuning = tuning, priors = priors,
  cov.model = "exponential", n.samples = n_samples, n.report = 1000)

save(m, ne_temp, models, coords, starting, tuning, priors, n_samples,
      file="dynlm.Rdata")

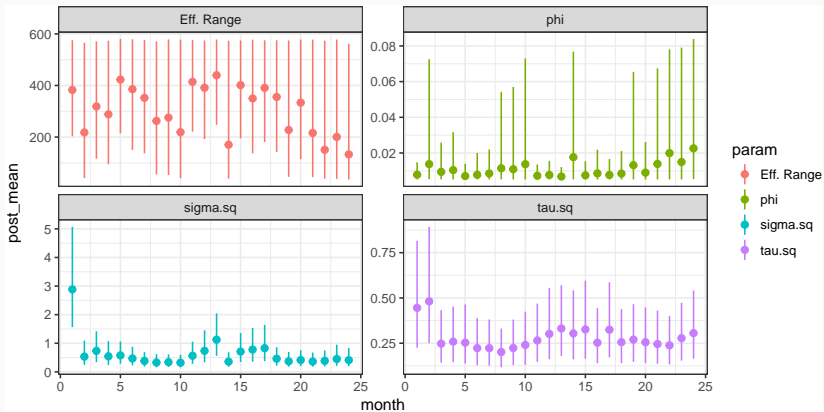
## -----
##      General model description
## -----
## Model fit with 34 observations in 24 time steps.
##
## Number of missing observations 0.
##
## Number of covariates 2 (including intercept if specified).
##
## Using the exponential spatial correlation model.
##
## Number of MCMC samples 10000.
##
## ...
```

# Posterior Inference - $\beta$ s

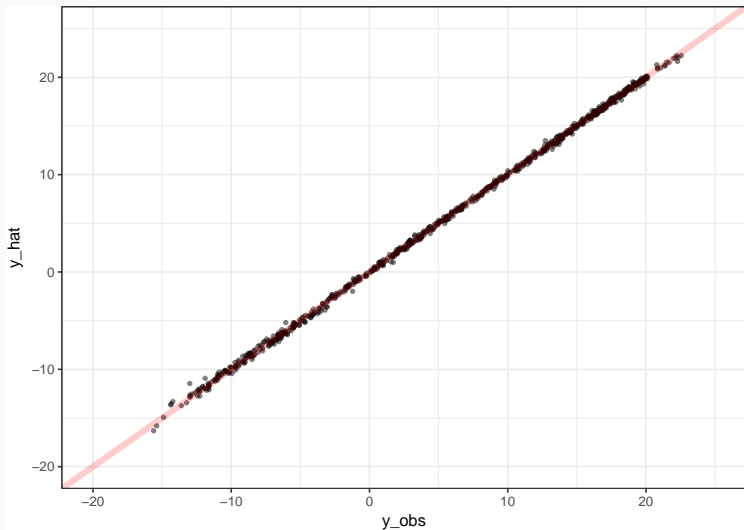


Lapse Rate

# Posterior Inference - $\theta$



## Posterior Inference - Observed vs. Predicted



## Prediction

spPredict does not support spDynLM objects.

```
r = raster(xmn=575e4, xmx=630e4, ymn=300e4, ymx=355e4, nrow=20, ncol=20)
```

```
pred = xyFromCell(r, 1:length(r)) %>%  
  cbind(elev=0, ., matrix(NA, nrow=length(r), ncol=24)) %>%  
  as.data.frame() %>%  
  setNames(names(ne_temp)) %>%  
  rbind(ne_temp, .) %>%  
  select(1:15) %>%  
  select(-elev)
```

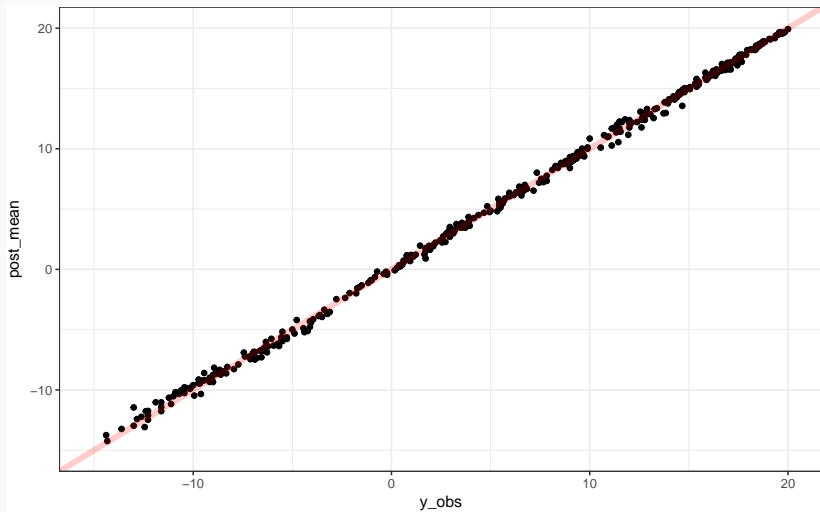
```
models_pred = lapply(paste0("y.",1:n_t, "~1"), as.formula)
```

```
n_samples = 5000
```

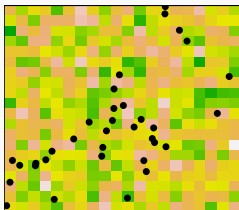
```
m_pred = spDynLM(  
  models_pred, data = pred, coords = coords_pred, get.fitted = TRUE,  
  starting = starting, tuning = tuning, priors = priors,  
  cov.model = "exponential", n.samples = n_samples, n.report = 1000)
```

```
save(m_pred, pred, models_pred, coords_pred, y_t_pred, n_samples,  
     file="dynlm_pred.Rdata")
```

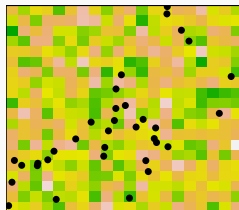




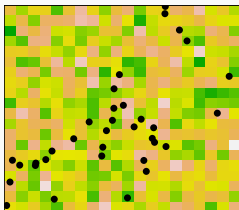
Jan 2000



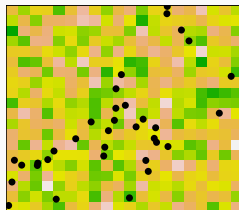
Apr 2000



Jul 2000



Oct 2000



## Spatio-temporal models for continuous time

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In general, spatiotemporal models will have a form like the following,

$$\begin{aligned} y(\mathbf{s}, t) &= \underbrace{\mu(\mathbf{s}, t)}_{\text{mean structure}} + \underbrace{e(\mathbf{s}, t)}_{\text{error structure}} \\ &= \underbrace{x(\mathbf{s}, t) \boldsymbol{\beta}(\mathbf{s}, t)}_{\text{Regression}} + \underbrace{w(\mathbf{s}, t)}_{\text{Spatiotemporal RE}} + \underbrace{\epsilon(\mathbf{s}, t)}_{\text{Error}} \end{aligned}$$

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The simplest possible spatiotemporal model is one where we assume there is no dependence between observations in space and time,

$$w(\mathbf{s}, t) = \alpha(t) + \omega(\mathbf{s})$$

these are straight forward to fit and interpret but are quite limiting (no shared information between space and time).

## Spatiotemporal Covariance

Lets assume that we want to define our spatiotemporal random effect to be a single stationary Gaussian Process (in 3 dimensions\*),

$$\mathbf{w}(\mathbf{s}, t) \sim \mathcal{N}(\mathbf{0}, \Sigma(\mathbf{s}, t))$$

where our covariance function depends on both  $\|\mathbf{s} - \mathbf{s}'\|$  and  $|t - t'|$ ,

$$\text{cov}(\mathbf{w}(\mathbf{s}, t), \mathbf{w}(\mathbf{s}', t')) = c(\|\mathbf{s} - \mathbf{s}'\|, |t - t'|)$$

- Note that the resulting covariance matrix  $\Sigma$  will be of size  $n_s \cdot n_t \times n_s \cdot n_t$ .
  - Even for modest problems this gets very large (past the point of direct computability).
  - If  $n_t = 52$  and  $n_s = 100$  we have to work with a  $5200 \times 5200$  covariance matrix

## Separable Models

One solution is to use a separable form, where the covariance is the product of a valid 2d spatial and a valid 1d temporal covariance / correlation function,

$$\text{cov}(w(\mathbf{s}, t), w(\mathbf{s}', t')) = \sigma^2 \rho_1(\|\mathbf{s} - \mathbf{s}'\|; \boldsymbol{\theta}) \rho_2(|t - t'|; \phi)$$

## Separable Models

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If we define our observations as follows (stacking time locations within spatial locations)

$$w_s^t = (w(s_1, t_1), \dots, w(s_1, t_{n_t}), w(s_2, t_1), \dots, w(s_2, t_{n_t}), \dots, \dots, w(s_{n_s}, t_1), \dots, w(s_{n_s}, t_{n_t}))$$

then the covariance can be written as

$$\boldsymbol{\Sigma}_w(\sigma^2, \boldsymbol{\theta}, \boldsymbol{\phi}) = \sigma^2 \mathbf{H}_s(\boldsymbol{\theta}) \otimes \mathbf{H}_t(\boldsymbol{\phi})$$

where  $\mathbf{H}_s(\boldsymbol{\theta})$  and  $\mathbf{H}_t(\boldsymbol{\phi})$  are  $n_s \times n_s$  and  $n_t \times n_t$  sized correlation matrices respectively and their elements are defined by

$$\{\mathbf{H}_s(\boldsymbol{\theta})\}_{ij} = \rho_1(\|\mathbf{s}_i - \mathbf{s}_j\|; \boldsymbol{\theta})$$

$$\{\mathbf{H}_t(\boldsymbol{\phi})\}_{ij} = \rho_2(|t_i - t_j|; \boldsymbol{\phi})$$



Definition:

$$\underset{[m \times n]}{\mathbf{A}} \otimes \underset{[p \times q]}{\mathbf{B}} = \underset{[m \cdot p \times n \cdot q]}{\begin{pmatrix} a_{11}\mathbf{B} & \cdots & a_{1n}\mathbf{B} \\ \vdots & \ddots & \vdots \\ a_{m1}\mathbf{B} & \cdots & a_{mn}\mathbf{B} \end{pmatrix}}$$

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Properties:

$$\mathbf{A} \otimes \mathbf{B} \neq \mathbf{B} \otimes \mathbf{A} \quad (\text{usually})$$

$$(\mathbf{A} \otimes \mathbf{B})^t = \mathbf{A}^t \otimes \mathbf{B}^t$$

$$\det(\mathbf{A} \otimes \mathbf{B}) = \det(\mathbf{B} \otimes \mathbf{A})$$

$$= \det(\mathbf{A})^{\text{rank}(\mathbf{B})} \det(\mathbf{B})^{\text{rank}(\mathbf{A})}$$

$$(\mathbf{A} \otimes \mathbf{B})^{-1} = \mathbf{A}^{-1} \otimes \mathbf{B}^{-1}$$

If we have a spatiotemporal random effect with a separable form,

$$\mathbf{w}(s, t) \sim \mathcal{N}(\mathbf{0}, \Sigma_w)$$

$$\Sigma_w = \sigma^2 H_s \otimes H_t$$

then the likelihood for  $\mathbf{w}$  is given by

$$\begin{aligned} & -\frac{n}{2} \log 2\pi - \frac{1}{2} \log |\Sigma_w| - \frac{1}{2} \mathbf{w}^t \Sigma_w^{-1} \mathbf{w} \\ = & -\frac{n}{2} \log 2\pi - \frac{1}{2} \log [(\sigma^2)^{n_t \cdot n_s} |H_s|^{n_t} |H_t|^{n_s}] - \frac{1}{2} \mathbf{w}^t \frac{1}{\sigma^2} (H_s^{-1} \otimes H_t^{-1}) \mathbf{w} \end{aligned}$$

- Additive and separable models are still somewhat limiting
- Cannot treat spatiotemporal covariances as 3d observations
- Possible alternatives:
  - Specialized spatiotemporal covariance functions, i.e.

$$c(\mathbf{s} - \mathbf{s}', t - t') = \sigma^2 (|t - t'| + 1)^{-1} \exp(-\|\mathbf{s} - \mathbf{s}'\| (|t - t'| + 1)^{-\beta/2})$$

- Mixtures, i.e.  $w(\mathbf{s}, t) = w_1(\mathbf{s}, t) + w_2(\mathbf{s}, t)$ , where  $w_1(\mathbf{s}, t)$  and  $w_2(\mathbf{s}, t)$  have seperable forms.