

Lecture 12

Gaussian Process Models

3/01/2018

Multivariate Normal

Multivariate Normal Distribution

For an n -dimension multivariate normal distribution with covariance Σ (positive semidefinite) can be written as

$$\mathbf{Y}_{n \times 1} \sim N(\boldsymbol{\mu}_{n \times 1}, \boldsymbol{\Sigma}_{n \times n}) \text{ where } \{\boldsymbol{\Sigma}\}_{ij} = \sigma_{ij}^2 = \rho_{ij} \sigma_i \sigma_j$$

$$\begin{pmatrix} Y_1 \\ \vdots \\ Y_n \end{pmatrix} \sim N \left(\begin{pmatrix} \mu_1 \\ \vdots \\ \mu_n \end{pmatrix}, \begin{pmatrix} \rho_{11} \sigma_1 \sigma_1 & \cdots & \rho_{1n} \sigma_1 \sigma_n \\ \vdots & \ddots & \vdots \\ \rho_{n1} \sigma_n \sigma_1 & \cdots & \rho_{nn} \sigma_n \sigma_n \end{pmatrix} \right)$$

For the n dimensional multivariate normal given on the last slide, its density is given by

$$(2\pi)^{-n/2} \det(\Sigma)^{-1/2} \exp\left(-\frac{1}{2} \begin{pmatrix} \mathbf{Y} - \boldsymbol{\mu} \\ 1 \times n \end{pmatrix}' \begin{pmatrix} \Sigma^{-1} \\ n \times n \end{pmatrix} \begin{pmatrix} \mathbf{Y} - \boldsymbol{\mu} \\ n \times 1 \end{pmatrix}\right)$$

and its log density is given by

$$-\frac{n}{2} \log 2\pi - \frac{1}{2} \log \det(\Sigma) - \frac{1}{2} \begin{pmatrix} \mathbf{Y} - \boldsymbol{\mu} \\ 1 \times n \end{pmatrix}' \begin{pmatrix} \Sigma^{-1} \\ n \times n \end{pmatrix} \begin{pmatrix} \mathbf{Y} - \boldsymbol{\mu} \\ n \times 1 \end{pmatrix}$$

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- Draw n iid unit normals ($\mathcal{N}(0, 1)$) as \mathbf{z}

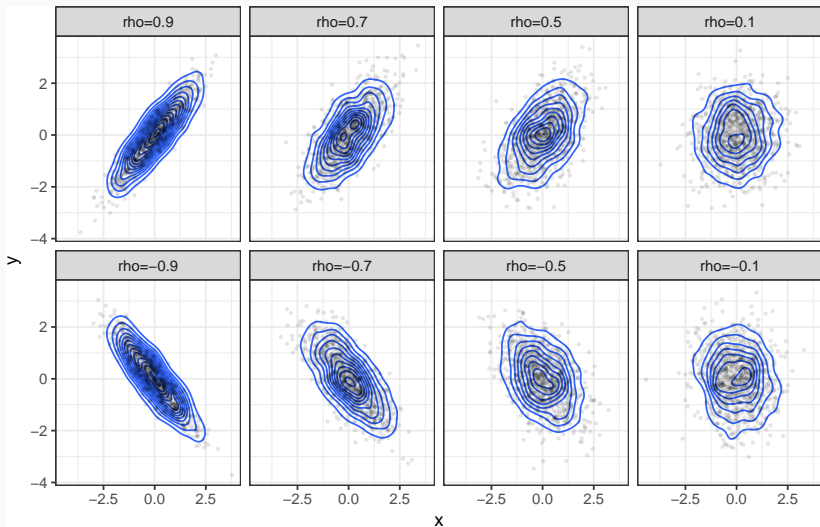
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- Construct multivariate normal draws using

$$\mathbf{Y} = \boldsymbol{\mu} + \mathbf{A} \mathbf{z}$$

Bivariate Example

$$\mu = \begin{pmatrix} 0 \\ 0 \end{pmatrix} \quad \Sigma = \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix}$$



Marginal distributions

Proposition - For an n -dimensional multivariate normal with mean $\boldsymbol{\mu}$ and covariance matrix $\boldsymbol{\Sigma}$, any marginal or conditional distribution of the \boldsymbol{y} 's will also (multivariate) normal.

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For a k -dimensional marginal distribution,

$$\mathbf{y}_{i_1, \dots, i_k} = \mathcal{N} \left(\begin{pmatrix} \boldsymbol{\mu}_{i_1} \\ \vdots \\ \boldsymbol{\mu}_{i_k} \end{pmatrix}, \begin{pmatrix} \gamma_{i_1 i_1} & \cdots & \gamma_{i_1 i_k} \\ \vdots & \ddots & \vdots \\ \gamma_{i_k i_1} & \cdots & \gamma_{i_k i_k} \end{pmatrix} \right)$$

Conditional Distributions

If we partition the n -dimensions into two pieces such that $\mathbf{Y} = (\mathbf{Y}_1, \mathbf{Y}_2)^t$ then

$$\mathbf{Y}_{n \times 1} \sim \mathcal{N} \left(\begin{pmatrix} \boldsymbol{\mu}_1 \\ \boldsymbol{\mu}_2 \end{pmatrix}, \begin{pmatrix} \boldsymbol{\Sigma}_{11} & \boldsymbol{\Sigma}_{12} \\ \boldsymbol{\Sigma}_{21} & \boldsymbol{\Sigma}_{22} \end{pmatrix} \right)$$

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then the conditional distributions are given by

$$\mathbf{Y}_1 | \mathbf{Y}_2 = \mathbf{a} \sim \mathcal{N}(\boldsymbol{\mu}_1 + \boldsymbol{\Sigma}_{12} \boldsymbol{\Sigma}_{22}^{-1} (\mathbf{a} - \boldsymbol{\mu}_2), \boldsymbol{\Sigma}_{11} - \boldsymbol{\Sigma}_{12} \boldsymbol{\Sigma}_{22}^{-1} \boldsymbol{\Sigma}_{21})$$

$$\mathbf{Y}_2 | \mathbf{Y}_1 = \mathbf{b} \sim \mathcal{N}(\boldsymbol{\mu}_2 + \boldsymbol{\Sigma}_{21} \boldsymbol{\Sigma}_{11}^{-1} (\mathbf{b} - \boldsymbol{\mu}_1), \boldsymbol{\Sigma}_{22} - \boldsymbol{\Sigma}_{21} \boldsymbol{\Sigma}_{11}^{-1} \boldsymbol{\Sigma}_{21})$$

From Shumway,

A process, $\mathbf{Y} = \{Y(t) : t \in T\}$, is said to be a Gaussian process if all possible finite dimensional vectors $\mathbf{y} = (y_{t_1}, y_{t_2}, \dots, y_{t_n})^t$, for every collection of time points t_1, t_2, \dots, t_n , and every positive integer n , have a multivariate normal distribution.

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So far we have only looked at examples of time series where T is discrete (and evenly spaced & contiguous), it turns out things get a lot more interesting when we explore the case where T is defined on a *continuous* space (e.g. \mathbb{R} or some subset of \mathbb{R}).

Gaussian Process Regression

Imagine we have a Gaussian process defined such that

$$\mathbf{Y} = \{Y(t) : t \in [0, 1]\},$$

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- The unconstrained covariance matrix for the observed data can have up to $n(n + 1)/2$ unique values*
- Necessary to make some simplifying assumptions:
 - Stationarity
 - Simple parameterization of Σ

More on these next week, but for now some simple / common examples

Covariance Functions

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Exponential Covariance:

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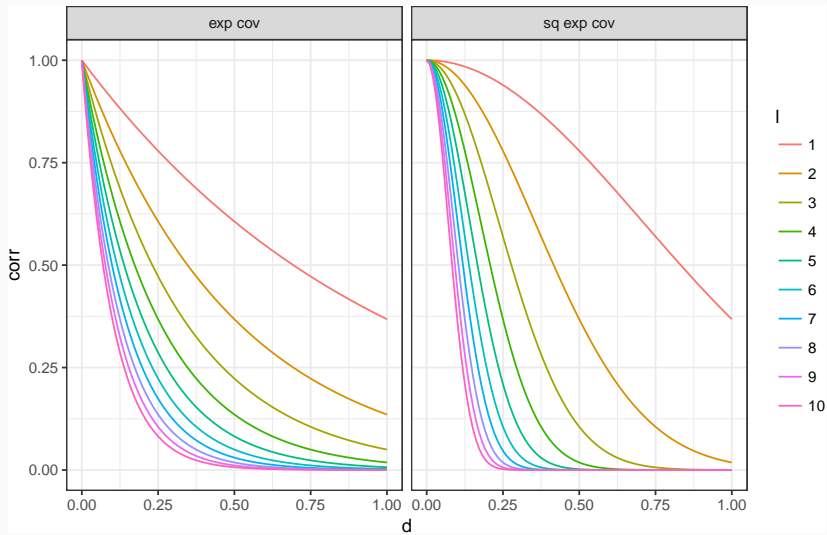
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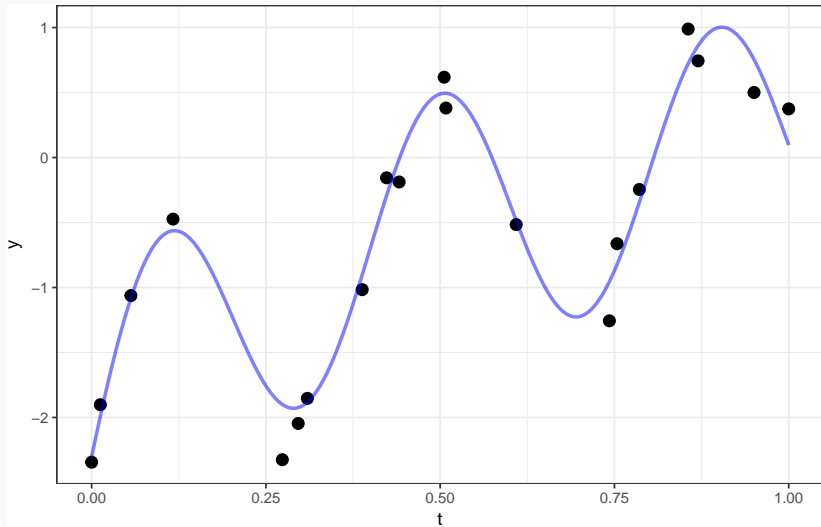
Powered Exponential Covariance ($p \in (0, 2]$):

$$\Sigma(y_t, y_{t'}) = \sigma^2 \exp(-(|t - t'| / l)^p)$$

Covariance Function - Correlation Decay



Example



Prediction

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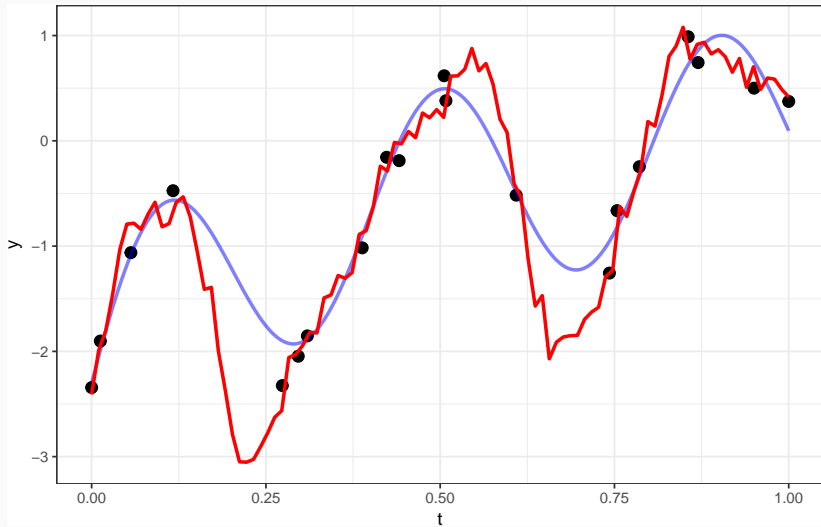
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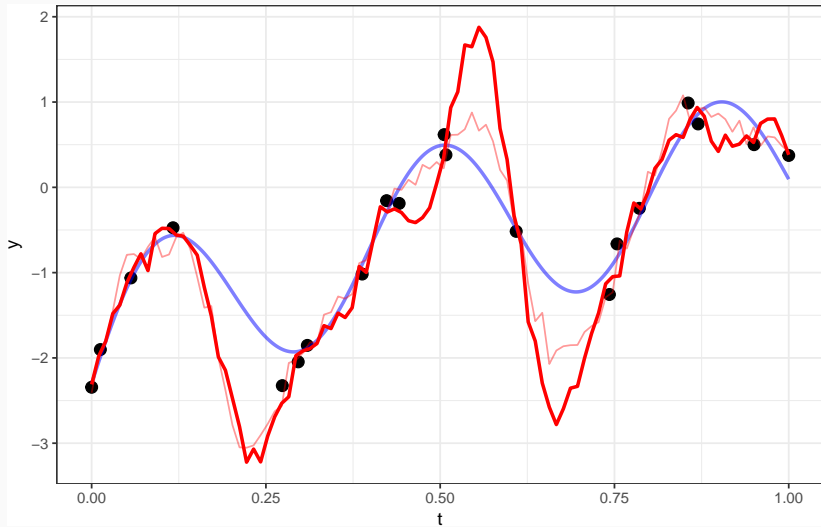
We therefore want to sample from $\mathbf{Y}_{pred} | \mathbf{Y}_{obs}$.

$$\mathbf{Y}_{pred} | \mathbf{Y}_{obs} = \mathbf{y} \sim \mathcal{N}(\Sigma_{po} \Sigma_{obs}^{-1} \mathbf{y}, \Sigma_{pred} - \Sigma_{po} \Sigma_{pred}^{-1} \Sigma_{op})$$

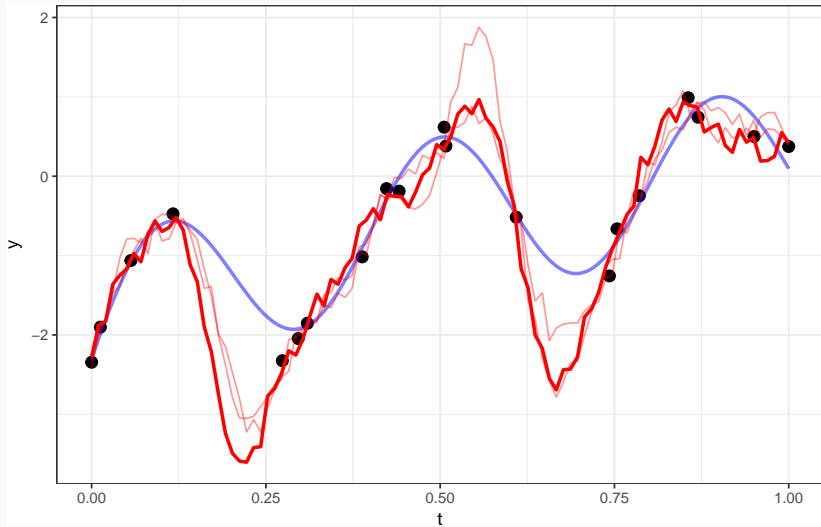
Draw 1



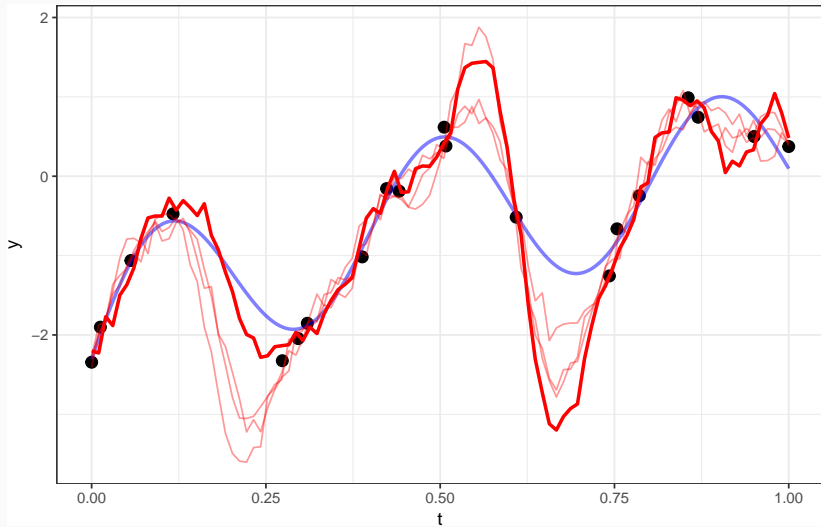
Draw 2



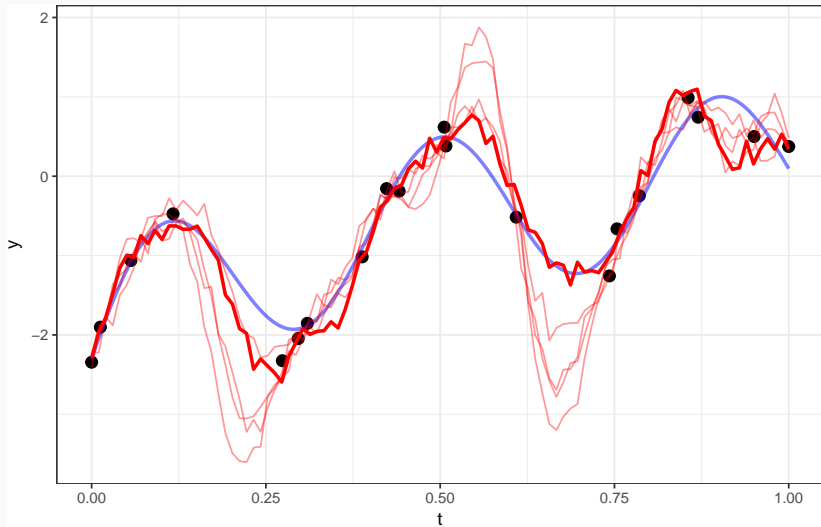
Draw 3



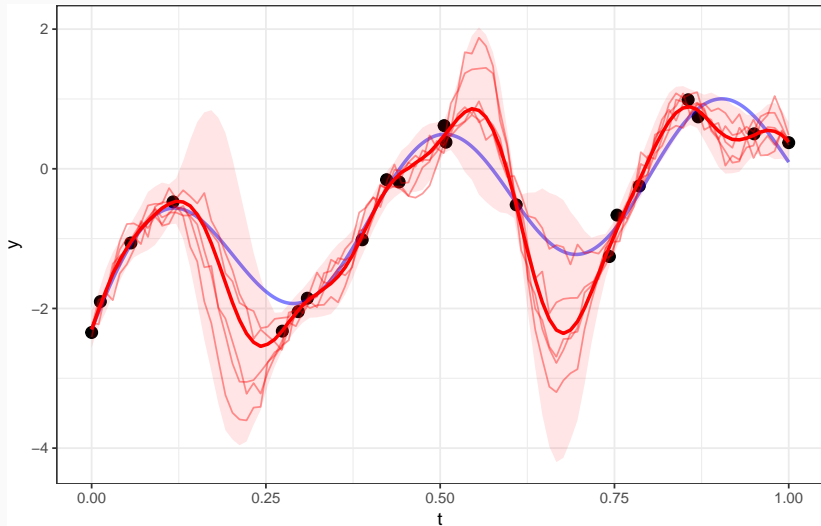
Draw 4



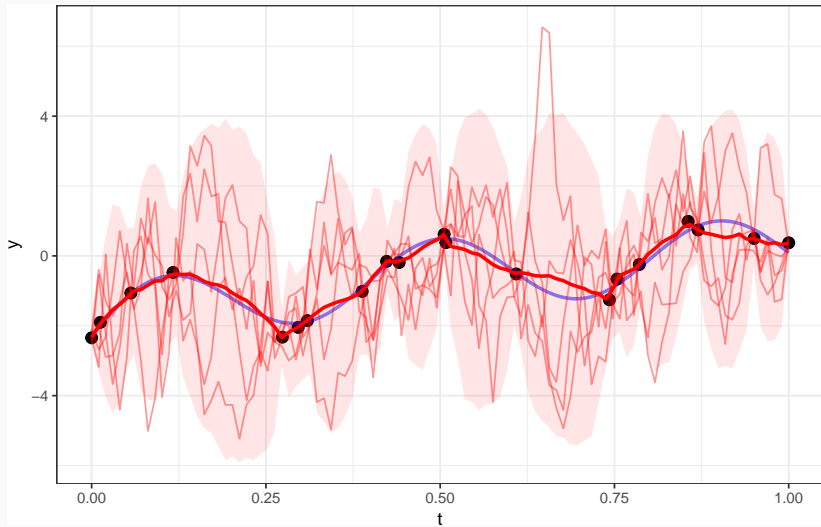
Draw 5



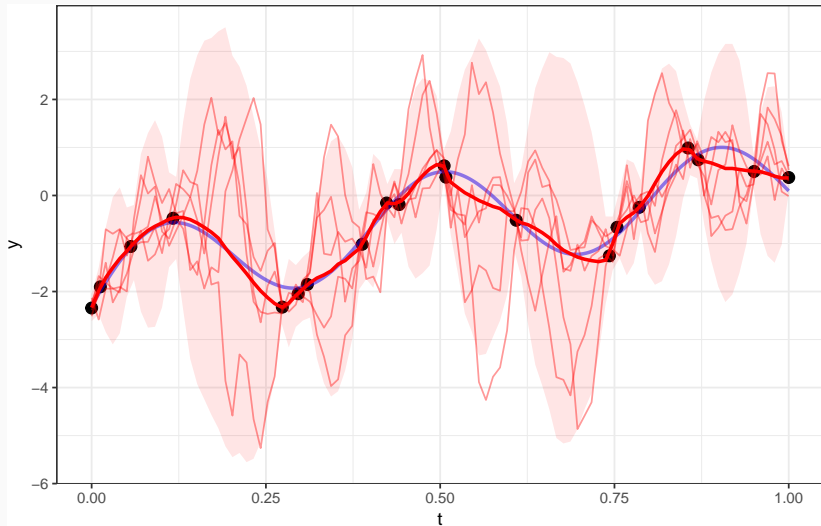
Many draws later



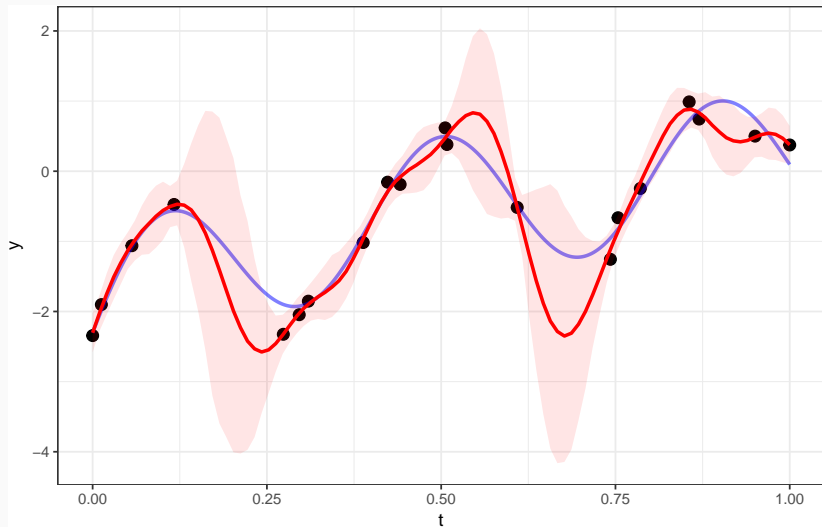
Exponential Covariance



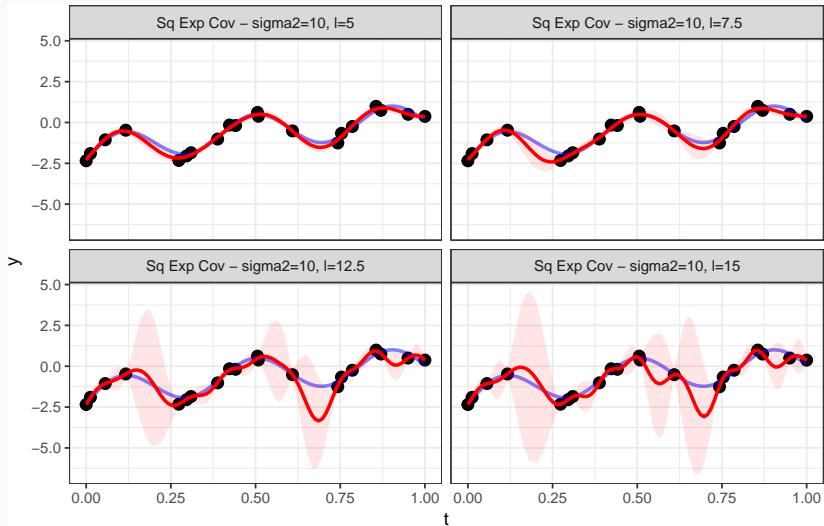
Powered Exponential Covariance ($p = 1.5$)



Back to the square exponential



Changing the range (l)

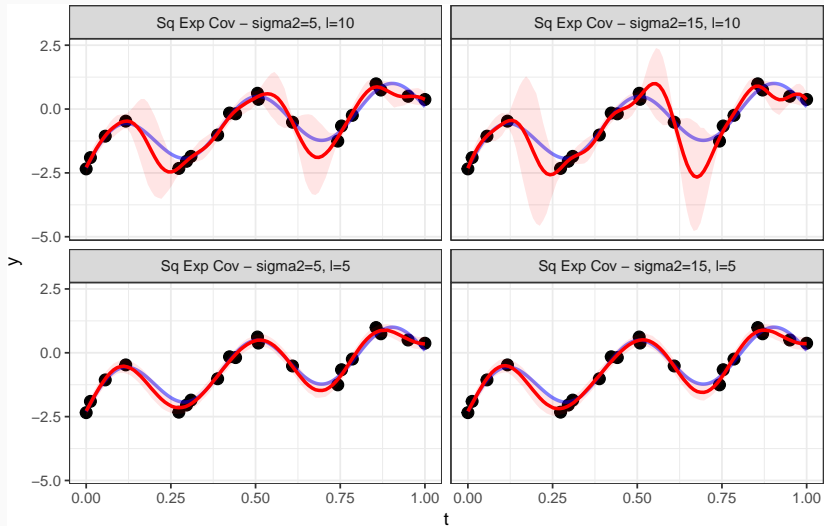


For the square exponential covariance

$$\begin{aligned}Cov(d) &= \sigma^2 \exp(-(l \cdot d)^2) \\Corr(d) &= \exp(-(l \cdot d)^2)\end{aligned}$$

we would like to know, for a given value of l , beyond what distance apart must observations be to have a correlation less than 0.05?

Changing the scale (σ^2)



```
gp_sq_exp_model = "model{
  y ~ dnorm(mu, inverse(Sigma))

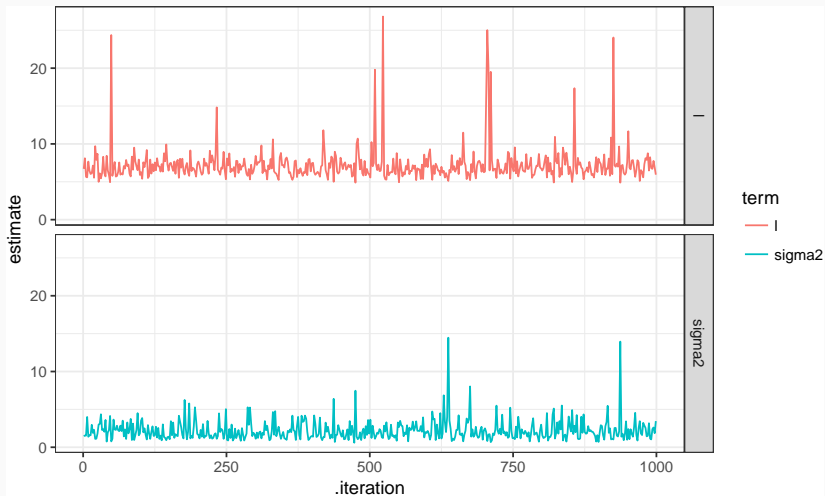
  for (i in 1:N) {
    mu[i] <- 0
  }

  for (i in 1:(N-1)) {
    for (j in (i+1):N) {
      Sigma[i,j] <- sigma2 * exp(- pow(l*d[i,j],2))
      Sigma[j,i] <- Sigma[i,j]
    }
  }

  for (k in 1:N) {
    Sigma[k,k] <- sigma2 + 0.01
  }

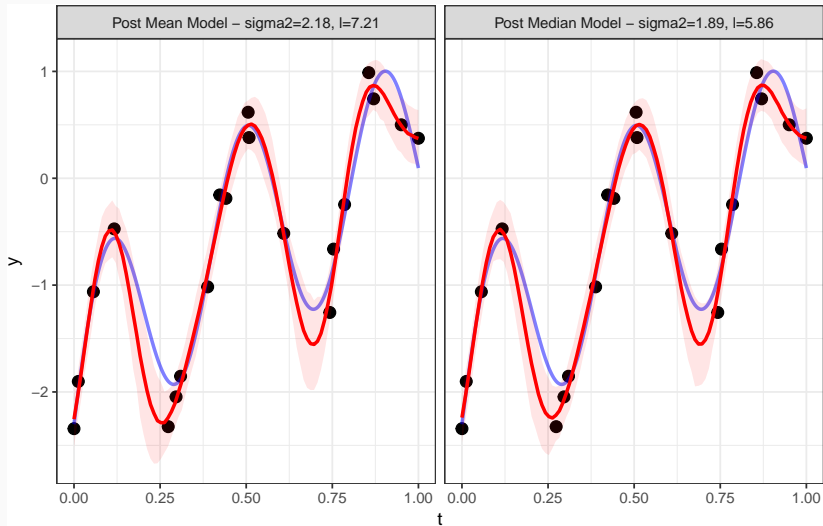
  sigma2 ~ dlnorm(0, 1)
  l ~ dt(0, 2.5, 1) T(0,) # Half-cauchy(0,2.5)
}"
```

Trace plots



param	post_mean	post_med	post_lower	post_upper
l	7.21	6.79	5.23	10.95
sigma2	2.18	1.86	0.84	5.24

Fitted models



Forecasting

