

# Lecture 18

## Fitting CAR and SAR Models

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## Fitting areal models

- Simultaneous Autoregressive (SAR)

$$y(s_i) = \phi \sum_{j=1}^n W_{ij} y(s_j) + \epsilon$$

$$\mathbf{y} \sim \mathcal{N}(0, \sigma^2 ((\mathbf{I} - \phi \mathbf{W})^{-1})((\mathbf{I} - \phi \mathbf{W})^{-1})^t)$$

- Conditional Autoregressive (CAR)

$$y(s_i) | \mathbf{y}_{-s_i} \sim \mathcal{N} \left( \phi \sum_{j=1}^n W_{ij} y(s_j), \sigma^2 \right)$$

$$\mathbf{y} \sim \mathcal{N}(0, \sigma^2 (\mathbf{I} - \phi \mathbf{W})^{-1})$$

## Some generalizations

Generally speaking we will want to work with a scaled / normalized version of the weight matrix,

$$\frac{W_{ij}}{W_i}$$

When  $W$  is an adjacency matrix we can express this as

$$\mathbf{D}^{-1}\mathbf{W}$$

where  $\mathbf{D}^{-1} = \text{diag}(1/|N(s_i)|)$ .

We can also allow  $\sigma^2$  to vary between locations, we can define this as  $\mathbf{D}_\tau = \text{diag}(1/\sigma_i^2)$  and most often we use

$$\mathbf{D}_\sigma^{-1} = \text{diag}\left(\frac{\sigma^2}{|N(s_i)|}\right) = \sigma^2\mathbf{D}^{-1}.$$

- Formula Model

$$y(s_i) = X_i.\beta + \phi \sum_{j=1}^n D_{jj}^{-1} W_{ij} (y(s_j) - X_j.\beta) + \epsilon_i$$

$$\epsilon \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{D}^{-1})$$

- Joint Model

- Formula Model

$$y(s_i) = X_i \beta + \phi \sum_{j=1}^n D_{jj}^{-1} W_{ij} (y(s_j) - X_j \beta) + \epsilon_i$$

$$\epsilon \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{D}^{-1})$$

- Joint Model

$$\begin{aligned} \mathbf{y} &= \mathbf{X}\beta + \phi \mathbf{D}^{-1} \mathbf{W} (\mathbf{y} - \mathbf{X}\beta) + \epsilon \\ (\mathbf{y} - \mathbf{X}\beta) &= \phi \mathbf{D}^{-1} \mathbf{W} (\mathbf{y} - \mathbf{X}\beta) + \epsilon \\ (\mathbf{y} - \mathbf{X}\beta) (\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W})^{-1} &= \epsilon \\ \mathbf{y} &= \mathbf{X}\beta + (\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W})^{-1} \epsilon \end{aligned}$$

- Formula Model

$$y(s_i) = X_i \beta + \phi \sum_{j=1}^n D_{jj}^{-1} W_{ij} (y(s_j) - X_j \beta) + \epsilon_i$$

$$\epsilon \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{D}^{-1})$$

- Joint Model

$$\mathbf{y} = \mathbf{X}\beta + \phi \mathbf{D}^{-1} \mathbf{W} (\mathbf{y} - \mathbf{X}\beta) + \epsilon$$

$$(\mathbf{y} - \mathbf{X}\beta) = \phi \mathbf{D}^{-1} \mathbf{W} (\mathbf{y} - \mathbf{X}\beta) + \epsilon$$

$$(\mathbf{y} - \mathbf{X}\beta)(\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W})^{-1} = \epsilon$$

$$\mathbf{y} = \mathbf{X}\beta + (\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W})^{-1} \epsilon$$

$$\mathbf{y} \sim \mathcal{N} \left( \mathbf{X}\beta, (\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W})^{-1} \sigma^2 \mathbf{D}^{-1} ((\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W})^{-1})^t \right)$$

- Conditional Model

$$y(s_i) | \mathbf{y}_{-s_i} \sim \mathcal{N} \left( X_i \cdot \beta + \phi \sum_{j=1}^n \frac{W_{ij}}{D_{ii}} (y(s_j) - X_j \cdot \beta), \sigma^2 D_{ii}^{-1} \right)$$

- Joint Model



- Conditional Model

$$y(s_i) | \mathbf{y}_{-s_i} \sim \mathcal{N} \left( X_i \cdot \beta + \phi \sum_{j=1}^n \frac{W_{ij}}{D_{ii}} (y(s_j) - X_j \cdot \beta), \sigma^2 D_{ii}^{-1} \right)$$

- Joint Model

$$\mathbf{y} \sim \mathcal{N}(\mathbf{X}\beta, \Sigma_{CAR})$$

- Conditional Model

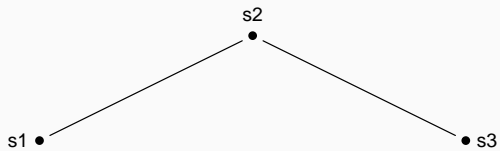
$$y(s_i) | \mathbf{y}_{-s_i} \sim \mathcal{N} \left( X_{i \cdot} \beta + \phi \sum_{j=1}^n \frac{W_{ij}}{D_{ii}} (y(s_j) - X_{j \cdot} \beta), \sigma^2 D_{ii}^{-1} \right)$$

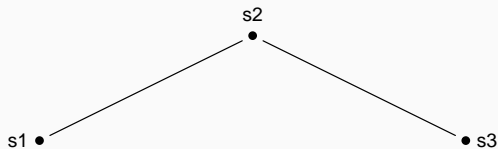
- Joint Model

$$\mathbf{y} \sim \mathcal{N}(\mathbf{X}\beta, \Sigma_{CAR})$$

$$\begin{aligned} \Sigma_{CAR} &= (\mathbf{D}_\sigma (\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W}))^{-1} \\ &= (1/\sigma^2 \mathbf{D} (\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W}))^{-1} \\ &= (1/\sigma^2 (\mathbf{D} - \phi \mathbf{W}))^{-1} \\ &= \sigma^2 (\mathbf{D} - \phi \mathbf{W})^{-1} \end{aligned}$$

## Toy CAR Example





$$\mathbf{W} = \begin{pmatrix} 0 & 1 & 0 \\ 1 & 0 & 1 \\ 0 & 1 & 0 \end{pmatrix} \quad \mathbf{D} = \begin{pmatrix} 1 & 0 & 0 \\ 0 & 2 & 0 \\ 0 & 0 & 1 \end{pmatrix}$$

...

$$\boldsymbol{\Sigma} = \sigma^2 (\mathbf{D} - \phi \mathbf{W}) = \sigma^2 \begin{pmatrix} 1 & -\phi & 0 \\ -\phi & 2 & -\phi \\ 0 & -\phi & 1 \end{pmatrix}^{-1}$$

## When does $\Sigma$ exist?

```
check_sigma = function(phi) {  
  Sigma_inv = matrix(c(1,-phi,0,-phi,2,-phi,0,-phi,1), ncol=3, byrow=TRUE)  
  solve(Sigma_inv)  
}
```

```
check_sigma(phi=0)  
##      [,1] [,2] [,3]  
## [1,]    1  0.0    0  
## [2,]    0  0.5    0  
## [3,]    0  0.0    1
```

```
check_sigma(phi=0.5)  
##      [,1]      [,2]      [,3]  
## [1,] 1.1666667 0.3333333 0.1666667  
## [2,] 0.3333333 0.6666667 0.3333333  
## [3,] 0.1666667 0.3333333 1.1666667
```

```
check_sigma(phi=-0.6)  
##      [,1]      [,2]      [,3]  
## [1,]  1.28125 -0.46875  0.28125  
## [2,] -0.46875  0.78125 -0.46875  
## [3,]  0.28125 -0.46875  1.28125
```

```
check_sigma(phi=1)
```

```
## Error in solve.default(Sigma_inv): Lapack routine dgesv: system is exactl
```

```
check_sigma(phi=-1)
```

```
## Error in solve.default(Sigma_inv): Lapack routine dgesv: system is exactl
```

```
check_sigma(phi=1.2)
```

```
##           [,1]      [,2]      [,3]
```

```
## [1,] -0.6363636 -1.363636 -1.6363636
```

```
## [2,] -1.3636364 -1.136364 -1.3636364
```

```
## [3,] -1.6363636 -1.363636 -0.6363636
```

```
check_sigma(phi=-1.2)
```

```
##           [,1]      [,2]      [,3]
```

```
## [1,] -0.6363636  1.363636 -1.6363636
```

```
## [2,]  1.3636364 -1.136364  1.3636364
```

```
## [3,] -1.6363636  1.363636 -0.6363636
```

## When is $\Sigma$ positive semidefinite?

```
check_sigma_pd = function(phi) {  
  Sigma_inv = matrix(c(1,-phi,0,-phi,2,-phi,0,-phi,1), ncol=3, byrow=TRUE)  
  chol(solve(Sigma_inv))  
}
```

```
check_sigma_pd(phi=0)  
##      [,1]      [,2] [,3]  
## [1,]    1 0.0000000    0  
## [2,]    0 0.7071068    0  
## [3,]    0 0.0000000    1
```

```
check_sigma_pd(phi=0.5)  
##      [,1]      [,2]      [,3]  
## [1,] 1.080123 0.3086067 0.1543033  
## [2,] 0.000000 0.7559289 0.3779645  
## [3,] 0.000000 0.0000000 1.0000000
```

```
check_sigma_pd(phi=-0.6)  
##      [,1]      [,2]      [,3]  
## [1,] 1.131923 -0.4141182 0.2484709  
## [2,] 0.000000 0.7808688 -0.4685213  
## [3,] 0.000000 0.0000000 1.0000000
```

```
check_sigma_pd(phi=1)
```

```
## Error in solve.default(Sigma_inv): Lapack routine dgesv: system is exactl
```

```
check_sigma_pd(phi=-1)
```

```
## Error in solve.default(Sigma_inv): Lapack routine dgesv: system is exactl
```

```
check_sigma_pd(phi=1.2)
```

```
## Error in chol.default(solve(Sigma_inv)): the leading minor of order 1 is
```

```
check_sigma_pd(phi=-1.2)
```

```
## Error in chol.default(solve(Sigma_inv)): the leading minor of order 1 is
```



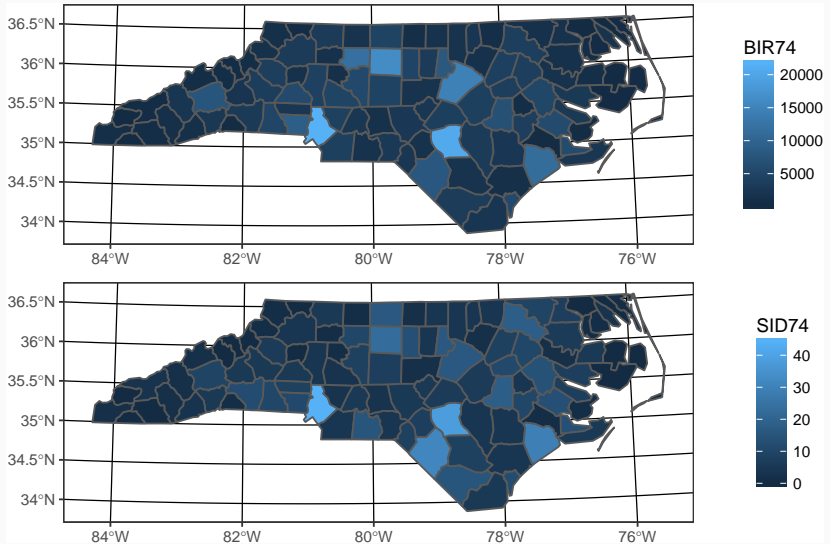
## Conclusions

Generally speaking just like the AR(1) model for time series we require that  $|\phi| < 1$  for the CAR model to be proper.

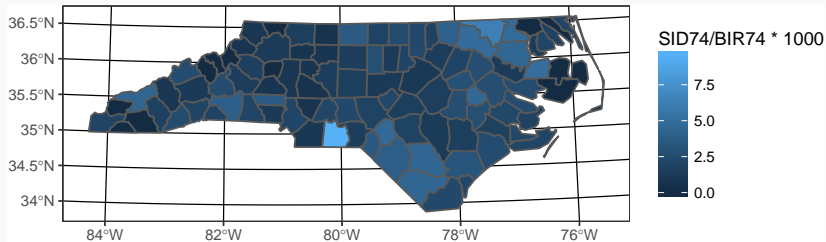
These results for  $\phi$  also apply in the context where  $\sigma_i^2$  is constant across locations (i.e.  $\Sigma = (\sigma^2 (\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W}))^{-1}$ )

As a side note, the special case where  $\phi = 1$  is known as an intrinsic autoregressive (IAR) model and they are popular as an *improper* prior for spatial random effects. An additional sum constraint is necessary for identifiability ( $\sum_{i=1}^n y(s_i) = 0$ ).

# Example - NC SIDS



```
ggplot() + geom_sf(data=nc, aes(fill=SID74/BIR74*1000))
```



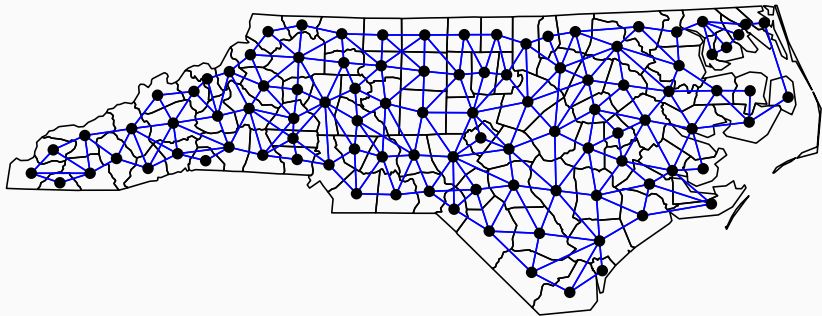
## Using `spautolm` from `spdep`

```
library(spdep)

W = st_touches(nc, sparse=FALSE)
listW = mat2listw(W)

listW
## Characteristics of weights list object:
## Neighbour list object:
## Number of regions: 100
## Number of nonzero links: 490
## Percentage nonzero weights: 4.9
## Average number of links: 4.9
##
## Weights style: M
## Weights constants summary:
##      n    nn  S0  S1   S2
## M 100 10000 490 980 10696
```

```
nc_coords = nc %>% st_centroid() %>% st_coordinates()  
  
plot(st_geometry(nc))  
plot(listW, nc_coords, add=TRUE, col="blue", pch=16)
```



## Moran's I

```
spdep::moran.test(nc$SID74, listW)
##
## Moran I test under randomisation
##
## data: nc$SID74
## weights: listW
##
## Moran I statistic standard deviate = 2.1707, p-value = 0.01498
## alternative hypothesis: greater
## sample estimates:
## Moran I statistic      Expectation      Variance
##      0.119089049      -0.010101010      0.003542176
spdep::moran.test(1000*nc$SID74/nc$BIR74, listW)
##
## Moran I test under randomisation
##
## data: 1000 * nc$SID74/nc$BIR74
## weights: listW
##
## Moran I statistic standard deviate = 3.6355, p-value = 0.0001387
## alternative hypothesis: greater
## sample estimates:
## Moran I statistic      Expectation      Variance
##      0.210046454      -0.010101010      0.003666802
```

```
spdep::geary.test(nc$SID74, listW)
##
## Geary C test under randomisation
##
## data: nc$SID74
## weights: listW
##
## Geary C statistic standard deviate = 0.91949, p-value = 0.1789
## alternative hypothesis: Expectation greater than statistic
## sample estimates:
## Geary C statistic      Expectation      Variance
##      0.88988684      1.00000000      0.01434105
spdep::geary.test(nc$SID74/nc$BIR74, listW)
##
## Geary C test under randomisation
##
## data: nc$SID74/nc$BIR74
## weights: listW
##
## Geary C statistic standard deviate = 3.0989, p-value = 0.0009711
## alternative hypothesis: Expectation greater than statistic
## sample estimates:
## Geary C statistic      Expectation      Variance
##      0.67796679      1.00000000      0.01079878
```

## CAR Model

```
nc_car = spautolm(formula = SID74/BIR74 ~ 1, data = nc,
                  listw = listW, family = "CAR")

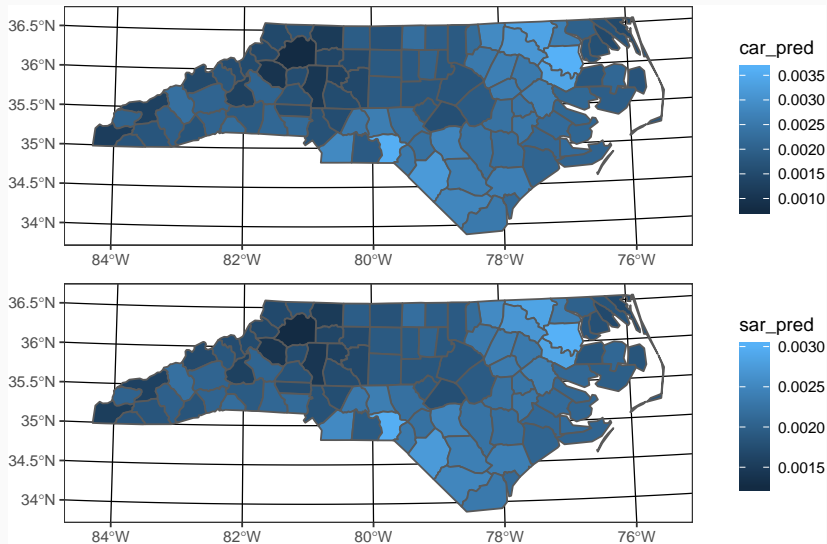
summary(nc_car)
##
## Call: spautolm(formula = SID74/BIR74 ~ 1, data = nc, listw = listW,
##               family = "CAR")
##
## Residuals:
##           Min           1Q       Median           3Q           Max
## -0.00213872 -0.00083535 -0.00022355  0.00055014  0.00768640
##
## Coefficients:
##               Estimate Std. Error z value Pr(>|z|)
## (Intercept)  0.00200203  0.00024272   8.2484 2.22e-16
##
## Lambda: 0.13062 LR test value: 8.6251 p-value: 0.0033157
## Numerical Hessian standard error of lambda: 0.030469
##
## Log likelihood: 508.3767
## ML residual variance (sigma squared): 2.1205e-06, (sigma: 0.0014562)
## Number of observations: 100
## Number of parameters estimated: 3
## AIC: -1010.8
```



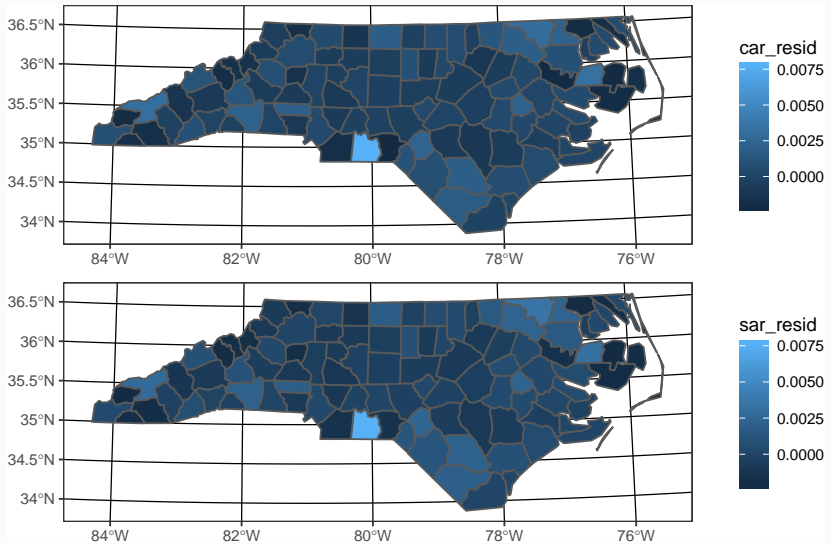
```
nc_sar = spautolm(formula = SID74/BIR74 ~ 1, data = nc,
                  listw = listW, family = "SAR")

summary(nc_sar)
##
## Call: spautolm(formula = SID74/BIR74 ~ 1, data = nc, listw = listW,
##               family = "SAR")
##
## Residuals:
##           Min           1Q       Median           3Q          Max
## -0.00209307 -0.00087039 -0.00020274  0.00051156  0.00762830
##
## Coefficients:
##               Estimate Std. Error z value Pr(>|z|)
## (Intercept)  0.00201084  0.00023622   8.5127 < 2.2e-16
##
## Lambda: 0.079934 LR test value: 8.8911 p-value: 0.0028657
## Numerical Hessian standard error of lambda: 0.0246
##
## Log likelihood: 508.5097
## ML residual variance (sigma squared): 2.1622e-06, (sigma: 0.0014704)
## Number of observations: 100
## Number of parameters estimated: 3
## AIC: -1011
```

# Predictions

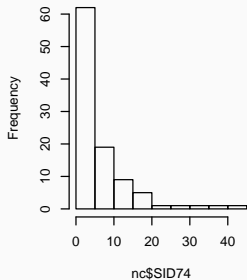


# Residuals

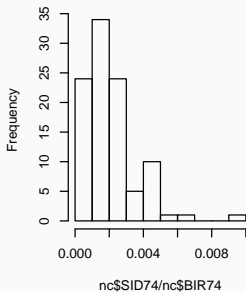


# What's wrong?

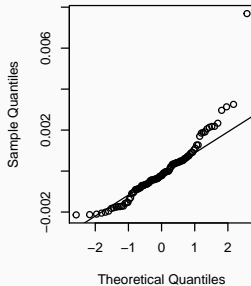
Histogram of nc\$SID74



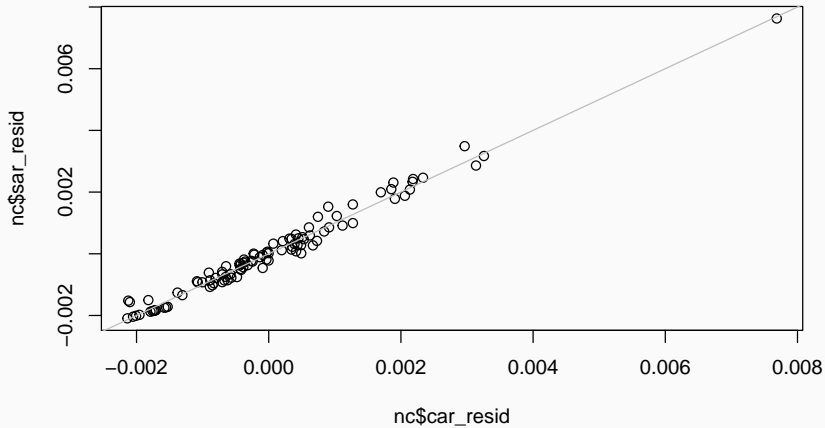
Histogram of nc\$SID74/nc\$BIR74



CAR Residuals



CAR vs SAR Residuals



## Stan CAR Model

```
car_model = "  
data {  
  int<lower=0> N;  
  vector[N] y;  
  matrix[N,N] W;  
  matrix[N,N] D;  
}  
parameters {  
  vector[N] w_s;  
  real beta;  
  real<lower=0> sigma2;  
  real<lower=0> sigma2_w;  
  real<lower=0,upper=1> phi;  
}  
transformed parameters {  
  vector[N] y_pred = beta + w_s;  
}  
model {  
  matrix[N,N] Sigma_inv = (D - phi * W) / sigma2;  
  
  w_s ~ multi_normal_prec(rep_vector(0,N), Sigma_inv);  
  
  beta ~ normal(0,10);  
  sigma2 ~ cauchy(0,5);  
  sigma2_w ~ cauchy(0,5);  
  
  y ~ normal(beta+w_s, sigma2_w);  
}
```

```
data = list(  
  N = nrow(nc),  
  y = nc$SID74 / nc$BIR74,  
  W = W * 1,  
  D = diag(rowSums(W))  
)
```

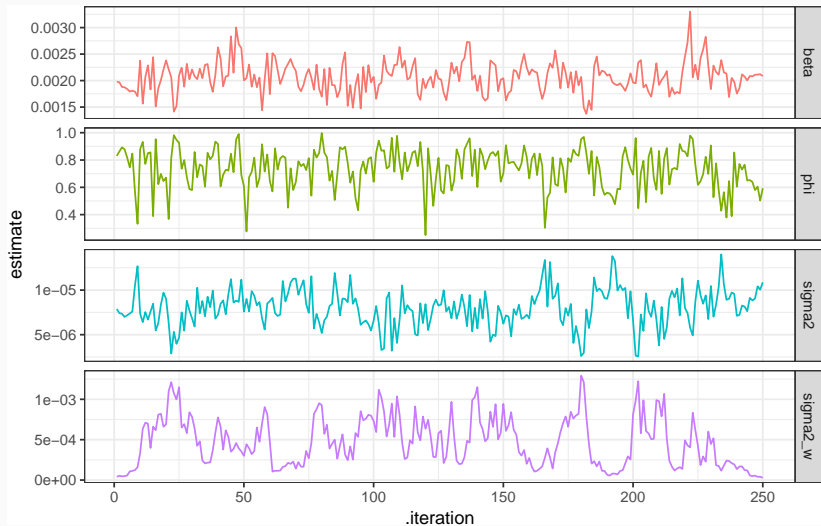
```
car_fit = rstan::stan(  
  model_code = car_model, data = data,  
  iter = 10000, chains = 1, thin=20  
)
```

```
data = list(  
  N = nrow(nc),  
  y = nc$SID74 / nc$BIR74,  
  W = W * 1,  
  D = diag(rowSums(W))  
)  
  
car_fit = rstan::stan(  
  model_code = car_model, data = data,  
  iter = 10000, chains = 1, thin=20  
)
```

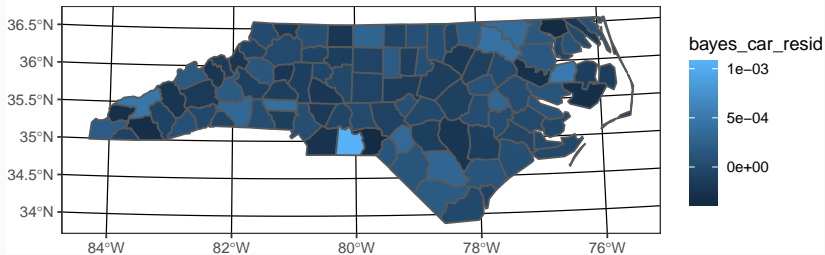
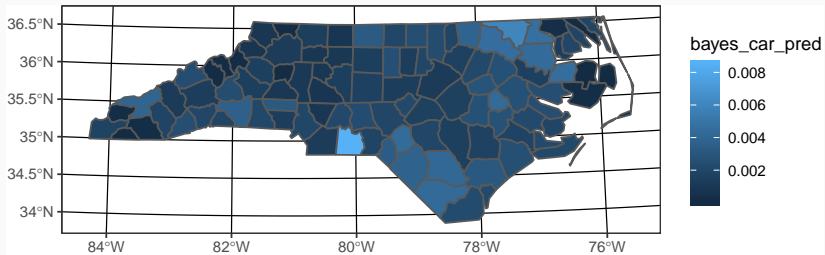
Why don't we use the conditional definition for the  $y$ 's?



# Model Results









$$\Sigma_{SAR} = (\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W})^{-1} \sigma^2 \mathbf{D}^{-1} ((\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W})^{-1})^t$$

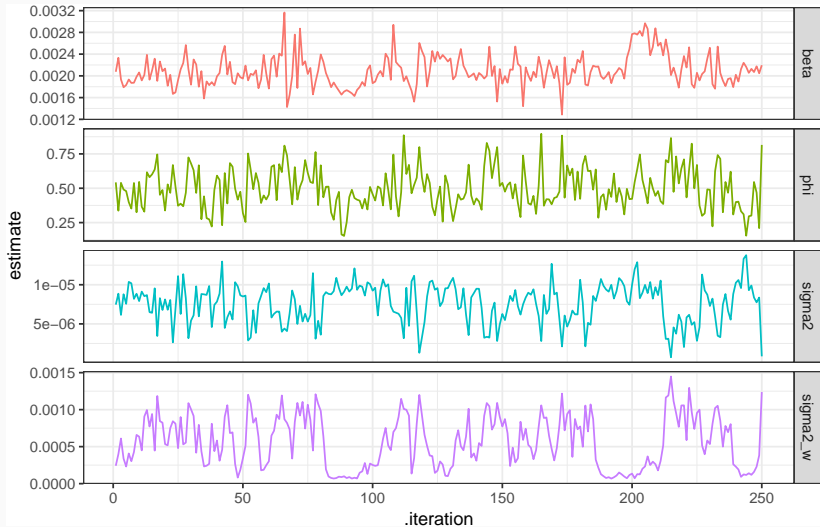
$$\begin{aligned}\Sigma_{SAR}^{-1} &= \left( (\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W})^{-1} \sigma^2 \mathbf{D}^{-1} ((\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W})^{-1})^t \right)^{-1} \\ &= \left( ((\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W})^{-1})^t \right)^{-1} \frac{1}{\sigma^2} \mathbf{D} (\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W}) \\ &= \frac{1}{\sigma^2} (\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W})^t \mathbf{D} (\mathbf{I} - \phi \mathbf{D}^{-1} \mathbf{W})\end{aligned}$$

# Jags SAR Model

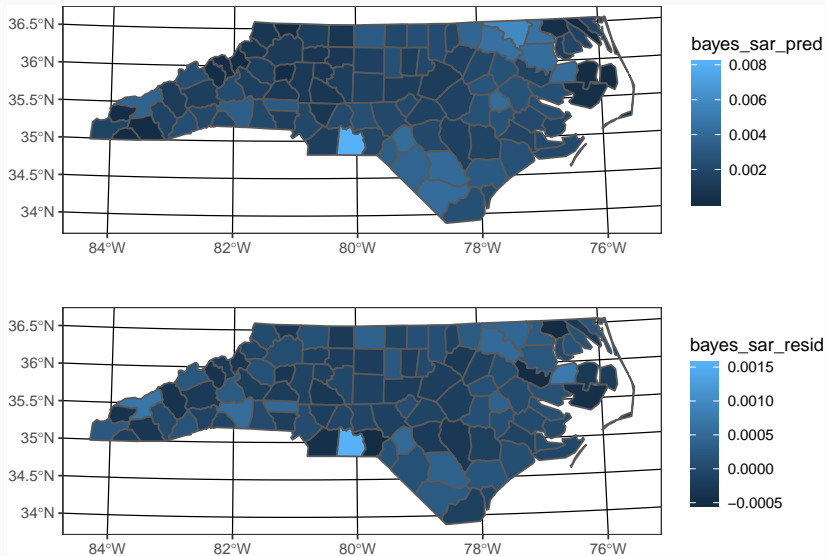
```
sar_model = "  
data {  
  int<lower=0> N;  
  vector[N] y;  
  matrix[N,N] W_tilde;  
  matrix[N,N] D;  
}  
transformed data {  
  matrix[N,N] I = diag_matrix(rep_vector(1, N));  
}  
parameters {  
  vector[N] w_s;  
  real beta;  
  real<lower=0> sigma2;  
  real<lower=0> sigma2_w;  
  real<lower=0,upper=1> phi;  
}  
transformed parameters {  
  vector[N] y_pred = beta + w_s;  
}  
model {  
  matrix[N,N] C = I - phi * W_tilde;  
  matrix[N,N] Sigma_inv = C' * D * C / sigma2;  
  
  w_s ~ multi_normal_prec(rep_vector(0,N), Sigma_inv);  
  
  beta ~ normal(0,10);  
  sigma2 ~ cauchy(0,5);  
  sigma2_w ~ cauchy(0,5);  
  
  y ~ normal(beta + w_s, sigma2_w);  
}  
"
```

```
D = diag(rowSums(W))  
D_inv = diag(1/diag(D))  
data = list(  
  N = nrow(nc),  
  y = nc$SID74 / nc$BIR74,  
  x = rep(1, nrow(nc)),  
  D_inv = D_inv,  
  W_tilde = D_inv %*% W  
)  
  
sar_fit = rstan::stan(  
  model_code = sar_model, data = data,  
  iter = 10000, chains = 1, thin=20  
)
```

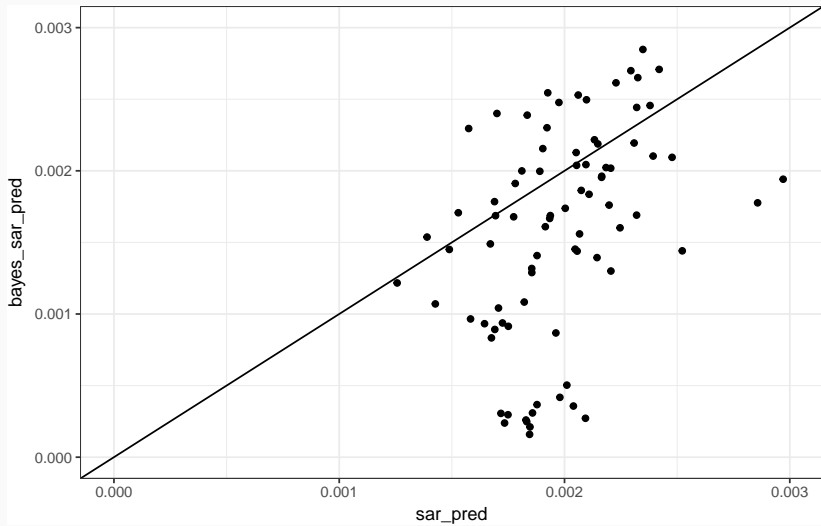
# Model Results



# Predictions







## Comparing Predictions

```
# RMSE
sqrt(mean(nc$bayes_car_resid^2))
## [1] 0.0002092447

sqrt(mean(nc$bayes_sar_resid^2))
## [1] 0.0002983034

sqrt(mean(nc$car_resid^2))
## [1] 0.001448564

sqrt(mean(nc$sar_resid^2))
## [1] 0.001470432
```

# Comparing Parameters

