

Lecture 7

- CCF
- Differencing
- AR(1) Model

~~Autocorrelation~~

2/08/2018

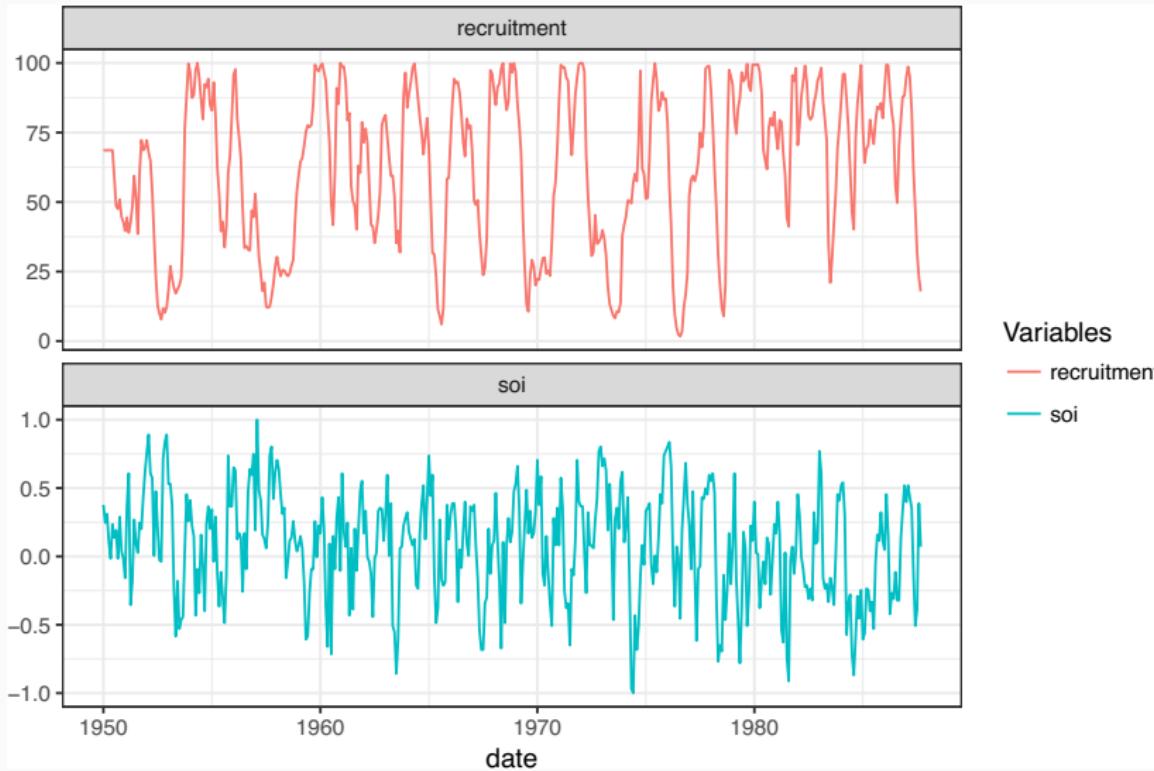
Lagged Predictors and CCFs

Southern Oscillation Index & Recruitment

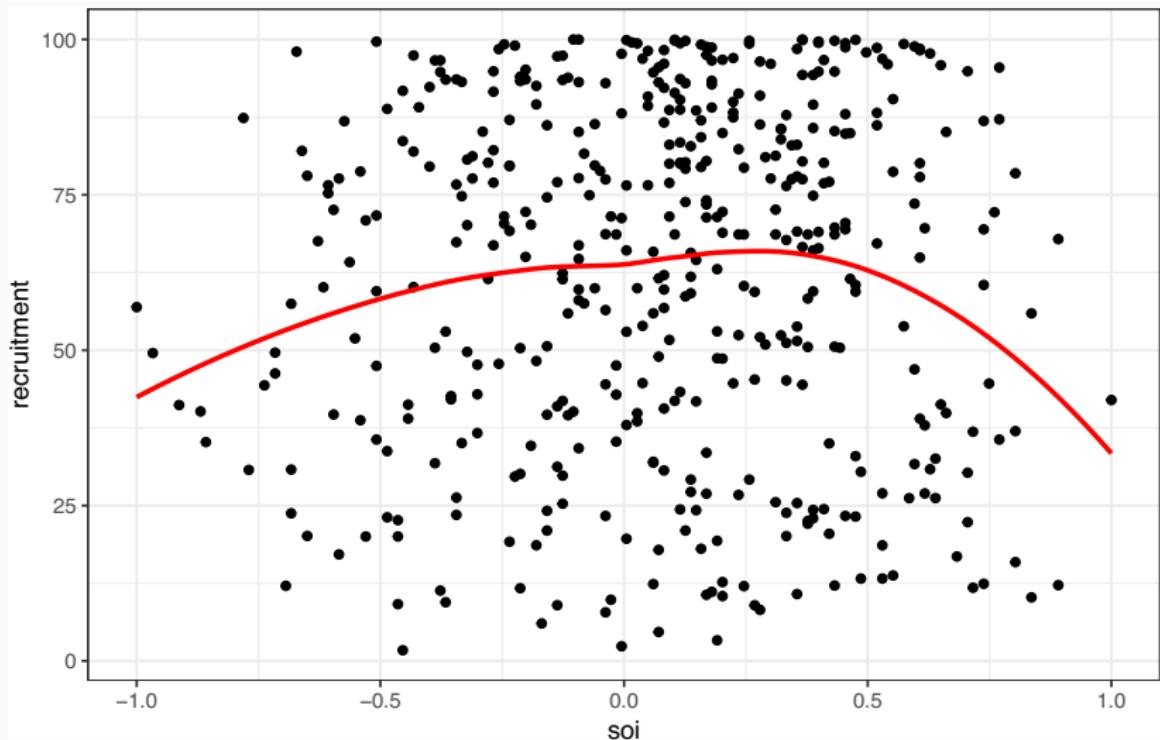
The Southern Oscillation Index (SOI) is an indicator of the development and intensity of El Niño (negative SOI) or La Niña (positive SOI) events in the Pacific Ocean. These data also included the estimate of “recruitment”, which indicate fish population sizes in the southern hemisphere.

```
##  
## Attaching package: 'astsa'  
## The following object is masked from 'package:forecast':  
##  
##     gas  
## # A tibble: 453 x 3  
##   date      soi recruitment  
##   <dbl>    <dbl>      <dbl>  
## 1 1950    0.377      68.6  
## 2 1950    0.246      68.6  
## 3 1950    0.311      68.6  
## 4 1950    0.104      68.6  
## 5 1950   -0.0160     68.6  
## 6 1950    0.235      68.6  
## 7 1950    0.137      59.2  
## 8 1951    0.191      48.7  
## 9 1951   -0.0160     47.5  
## 10 1951   0.290      50.9  
## # ... with 443 more rows
```

Time series



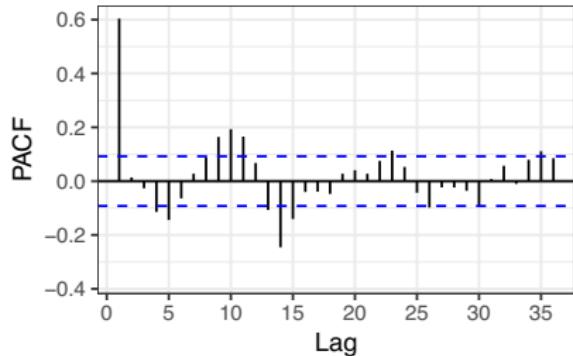
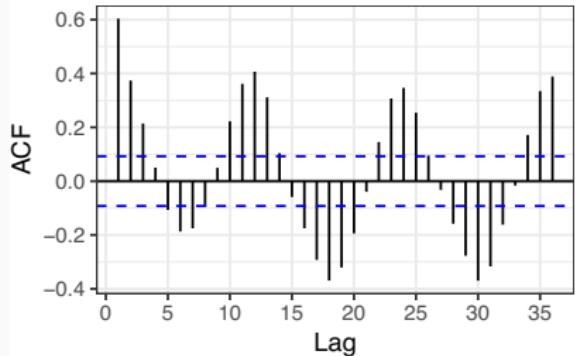
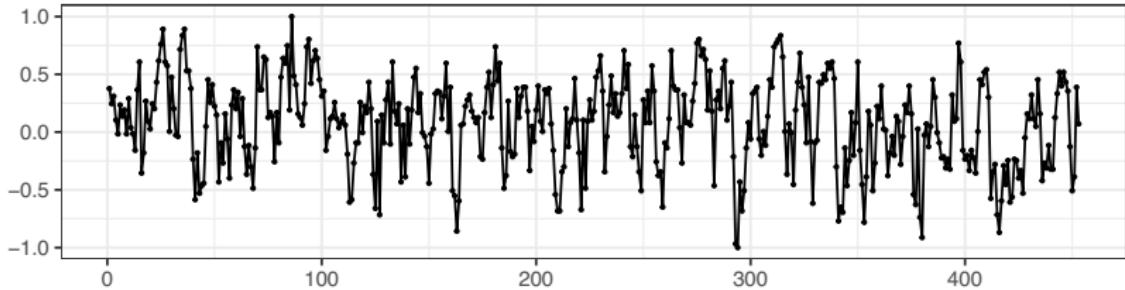
Relationship?



sois ACF & PACF

```
forecast::ggtsdisplay(fish$soi, lag.max = 36)
```

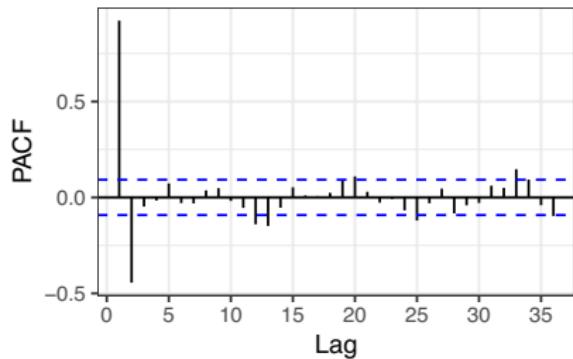
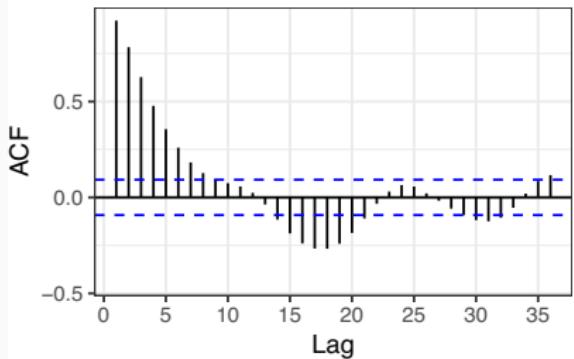
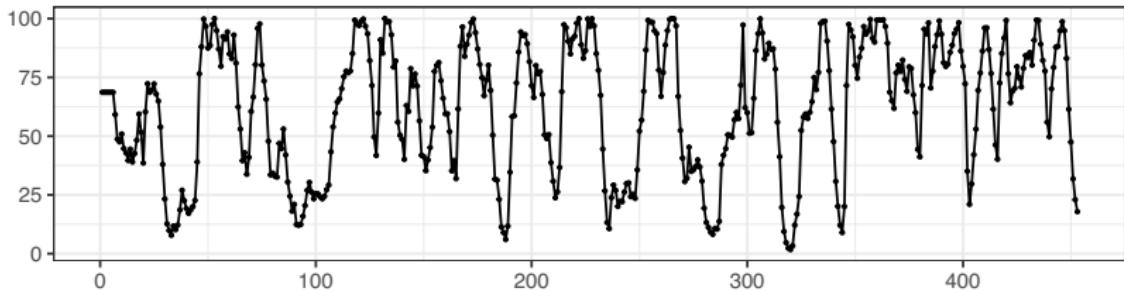
fish\$soi



recruitment

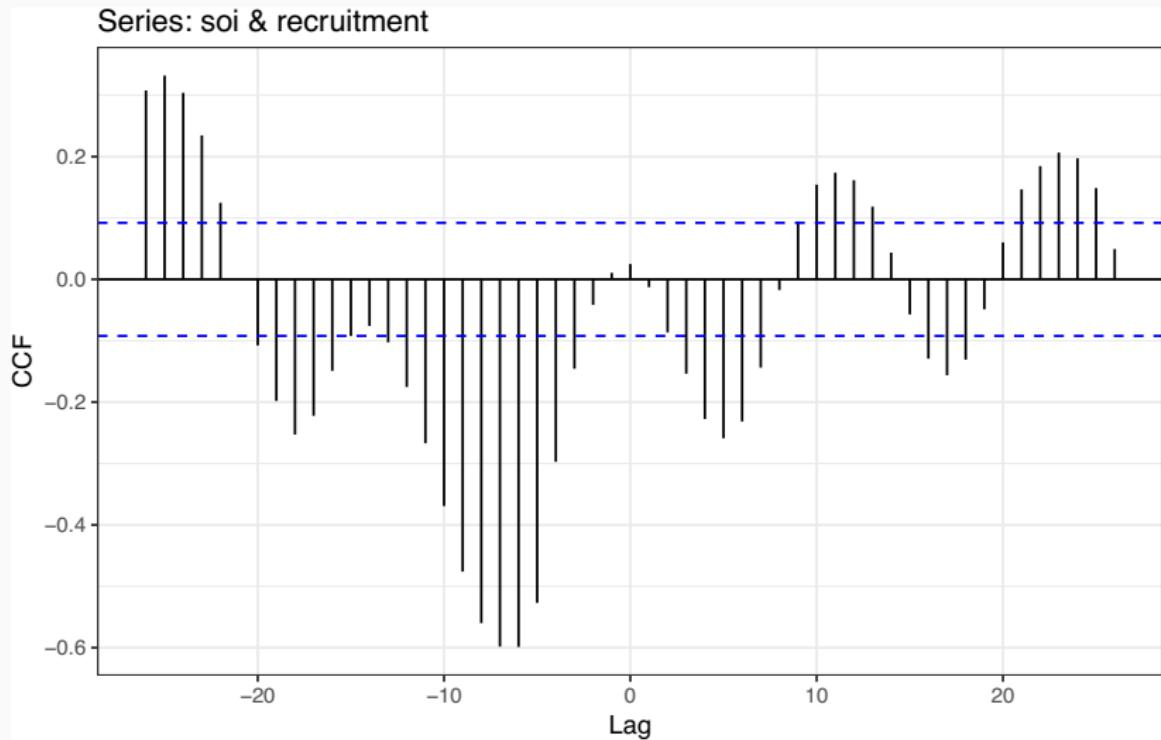
```
forecast::ggtsdisplay(fish$recruitment, lag.max = 36)
```

fish\$recruitment

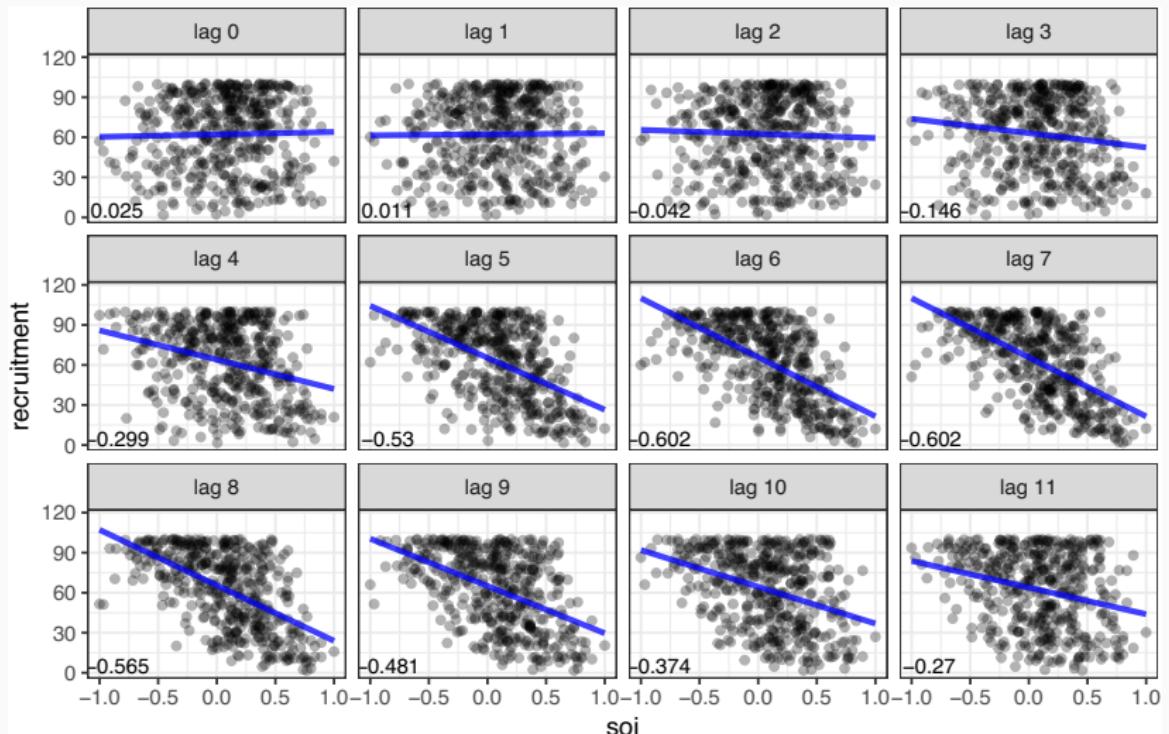


Cross correlation function

```
with(fish, forecast::ggCcf(soi, recruitment))
```



Cross correlation function - Scatter plots

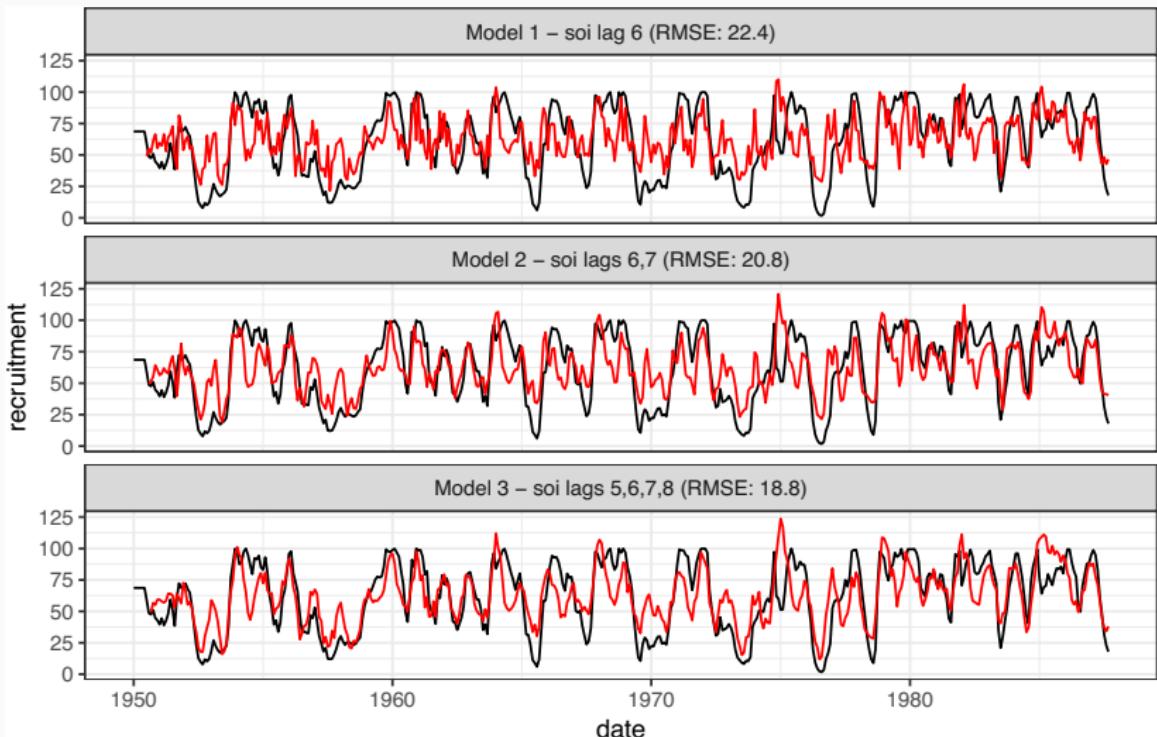


Model

```
model1 = lm(recruitment~lag(soi,6), data=fish)
model2 = lm(recruitment~lag(soi,6)+lag(soi,7), data=fish)
model3 = lm(recruitment~lag(soi,5)+lag(soi,6)+lag(soi,7)+lag(soi,8), data=fish)

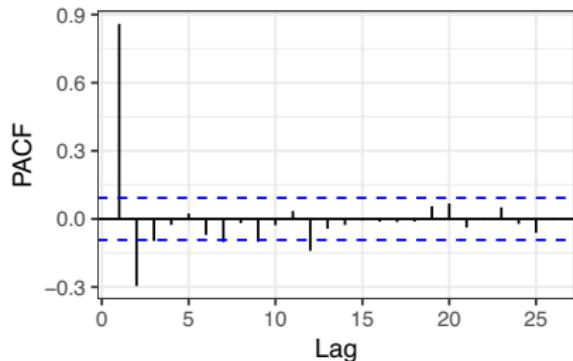
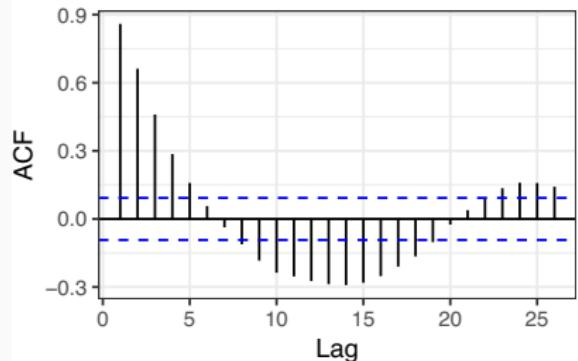
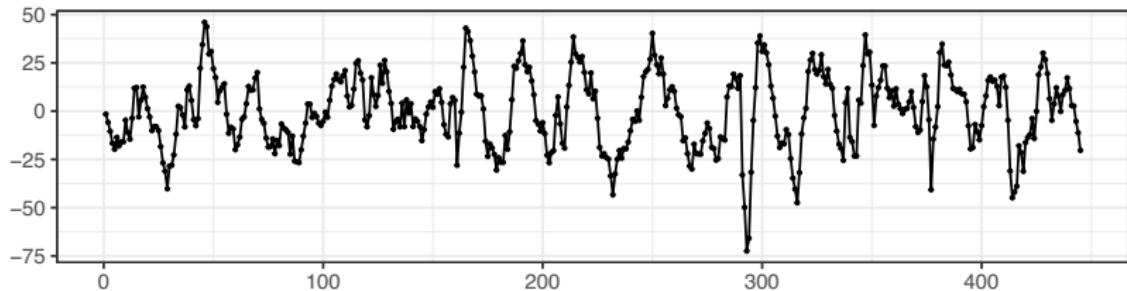
summary(model3)
##
## Call:
## lm(formula = recruitment ~ lag(soi, 5) + lag(soi, 6) + lag(soi,
##     7) + lag(soi, 8), data = fish)
##
## Residuals:
##     Min      1Q  Median      3Q     Max
## -72.409 -13.527   0.191  12.851  46.040
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)
## (Intercept) 67.9438    0.9306  73.007 < 2e-16 ***
## lag(soi, 5) -19.1502    2.9508  -6.490 2.32e-10 ***
## lag(soi, 6) -15.6894    3.4334  -4.570 6.36e-06 ***
## lag(soi, 7) -13.4041    3.4332  -3.904 0.000109 ***
## lag(soi, 8) -23.1480    2.9530  -7.839 3.46e-14 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 18.93 on 440 degrees of freedom
##   (8 observations deleted due to missingness)
## Multiple R-squared:  0.5539, Adjusted R-squared:  0.5498
## F-statistic: 136.6 on 4 and 440 DF,  p-value: < 2.2e-16
```

Prediction



Residual ACF - Model 3

residuals(model3)



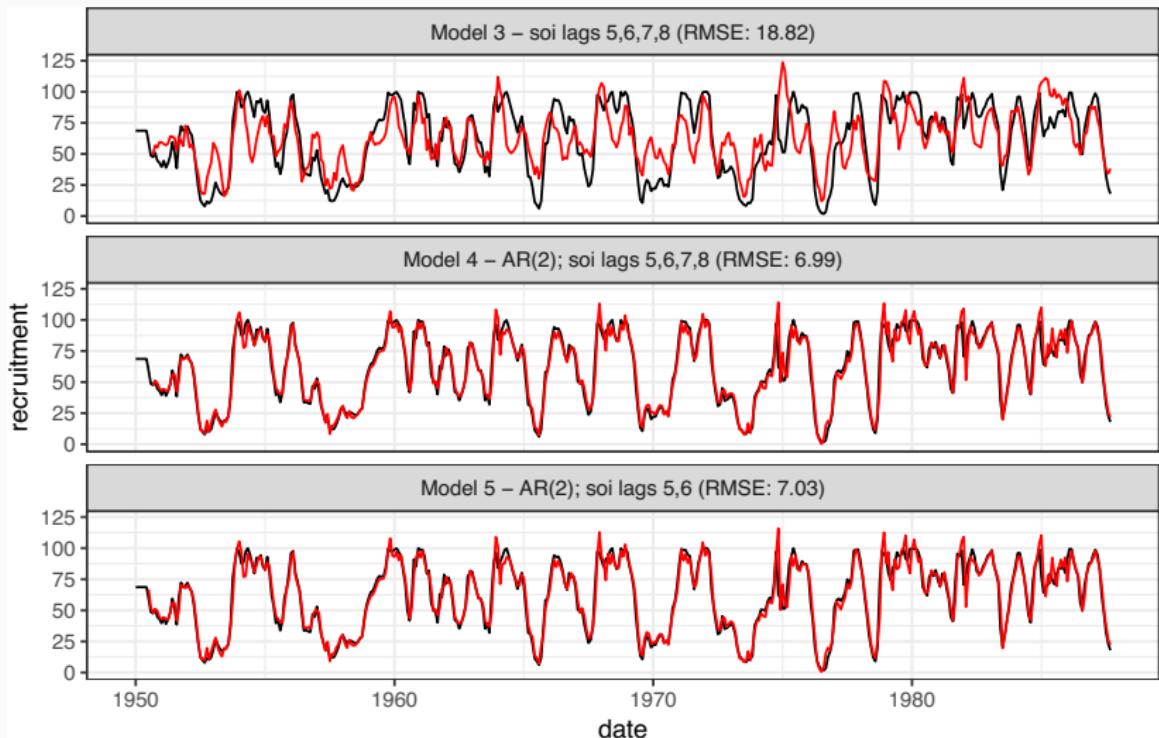
Autoregressive model 1

```
model4 = lm(recruitment~lag(recruitment,1) + lag(recruitment,2) +
            lag(soi,5)+lag(soi,6)+lag(soi,7)+lag(soi,8),
            data=fish)
summary(model4)
##
## Call:
## lm(formula = recruitment ~ lag(recruitment, 1) + lag(recruitment,
##      2) + lag(soi, 5) + lag(soi, 6) + lag(soi, 7) + lag(soi, 8),
##      data = fish)
##
## Residuals:
##     Min      1Q  Median      3Q     Max 
## -51.996   -2.892    0.103   3.117  28.579 
## 
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)    
## (Intercept) 10.25007  1.17081   8.755 < 2e-16 ***
## lag(recruitment, 1) 1.25301  0.04312  29.061 < 2e-16 ***
## lag(recruitment, 2) -0.39961  0.03998 -9.995 < 2e-16 ***
## lag(soi, 5)       -20.76309  1.09906 -18.892 < 2e-16 ***
## lag(soi, 6)        9.71918  1.56265   6.220 1.16e-09 ***
## lag(soi, 7)       -1.01131  1.31912  -0.767  0.4437  
## lag(soi, 8)       -2.29814  1.20730  -1.904  0.0576 .  
## ---                
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## 
## Residual standard error: 7.042 on 438 degrees of freedom
## (8 observations deleted due to missingness)
```

Autoregressive model 2

```
model5 = lm(recruitment~lag(recruitment,1) + lag(recruitment,2) +
            lag(soi,5) + lag(soi,6),
            data=fish)
summary(model5)
##
## Call:
## lm(formula = recruitment ~ lag(recruitment, 1) + lag(recruitment,
##       2) + lag(soi, 5) + lag(soi, 6), data = fish)
##
## Residuals:
##     Min      1Q  Median      3Q     Max 
## -53.786  -2.999  -0.035   3.031  27.669 
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)    
## (Intercept) 8.78498   1.00171   8.770 < 2e-16 ***
## lag(recruitment, 1) 1.24575   0.04314  28.879 < 2e-16 ***
## lag(recruitment, 2) -0.37193   0.03846  -9.670 < 2e-16 ***
## lag(soi, 5)    -20.83776   1.10208 -18.908 < 2e-16 ***
## lag(soi, 6)     8.55600   1.43146   5.977 4.68e-09 ***
## ---      
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 7.069 on 442 degrees of freedom
##   (6 observations deleted due to missingness)
## Multiple R-squared:  0.9375, Adjusted R-squared:  0.937 
## F-statistic: 1658 on 4 and 442 DF,  p-value: < 2.2e-16
```

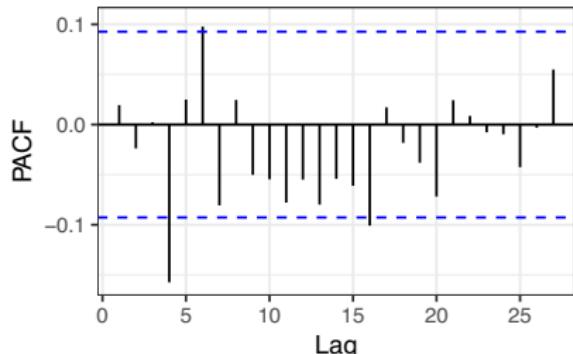
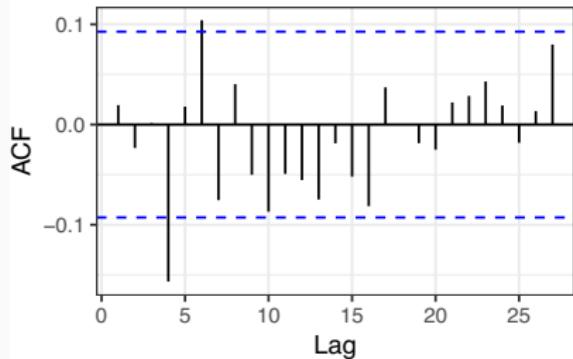
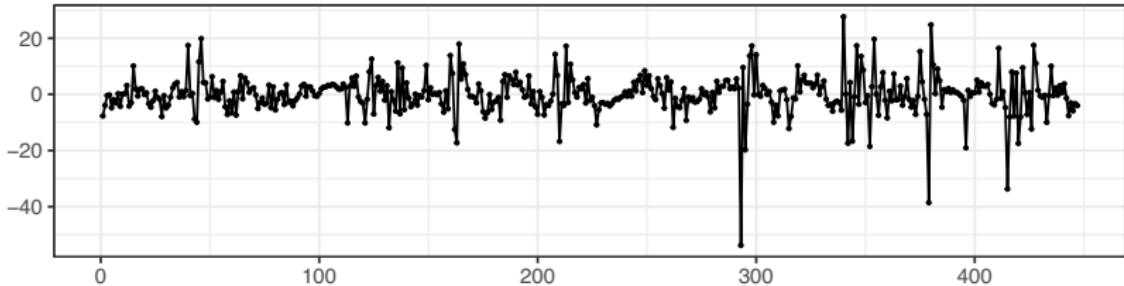
Prediction



Residual ACF - Model 5

```
forecast::ggtstdisplay(residuals(model5))
```

residuals(model5)



Non-stationarity

Non-stationary models

All happy families are alike; each unhappy family is unhappy in its own way.

- Tolstoy, *Anna Karenina*

This applies to time series models as well, just replace happy family with stationary model.

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- Tolstoy, Anna Karenina

This applies to time series models as well, just replace happy family with stationary model.

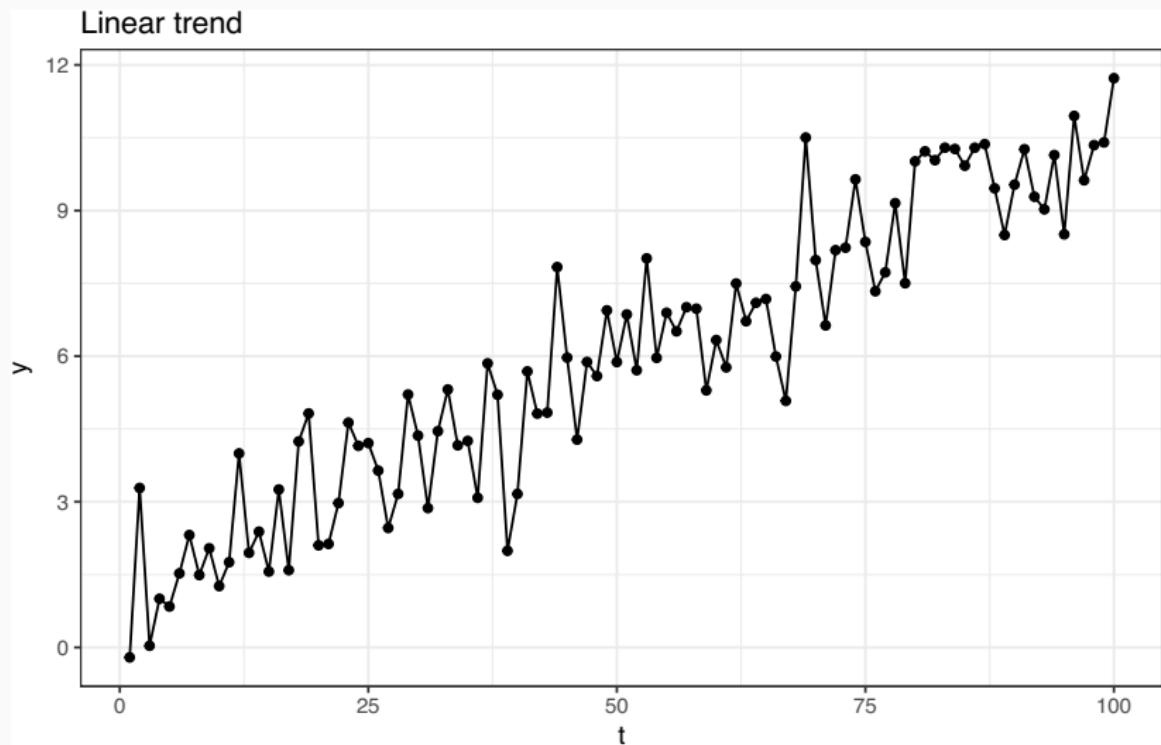
A simple example of a non-stationary time series is a trend stationary model

$$y_t = \mu(t) + w_t$$

where $\mu(t)$ denotes a time dependent trend and w_t is a white noise (stationary) process.

Linear trend model

Lets imagine a simple model where $y_t = \delta + \beta t + x_t$ where δ and β are constants and x_t is a stationary process.

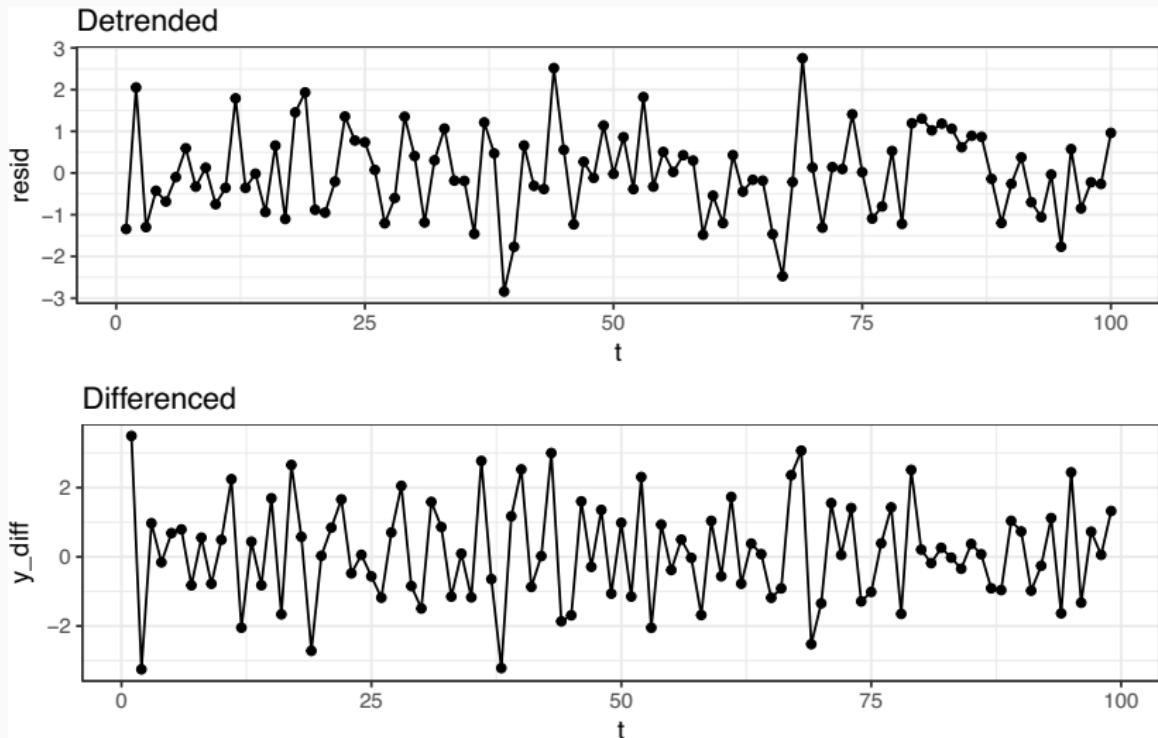


Differencing

An simple approach to remove trend is to difference your response variable, specifically examine $y_t - y_{t-1}$ instead of y_t .

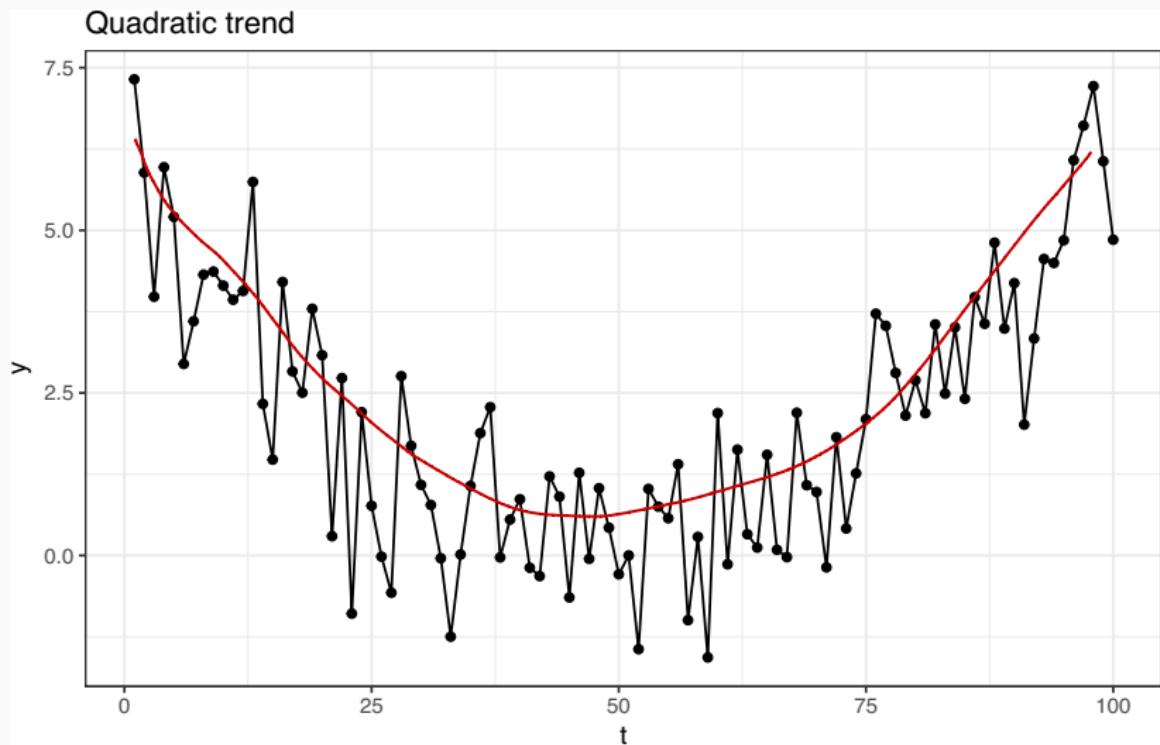
$$\begin{aligned}d_t &= y_t - y_{t-1} \\&= (\delta + \beta t + x_t) - (\delta + \beta(t-1) + x_{t-1}) \\&= \beta + \underbrace{y_t - x_{t-1}}_{\text{Stationary}}\end{aligned}$$

Detrending vs Difference



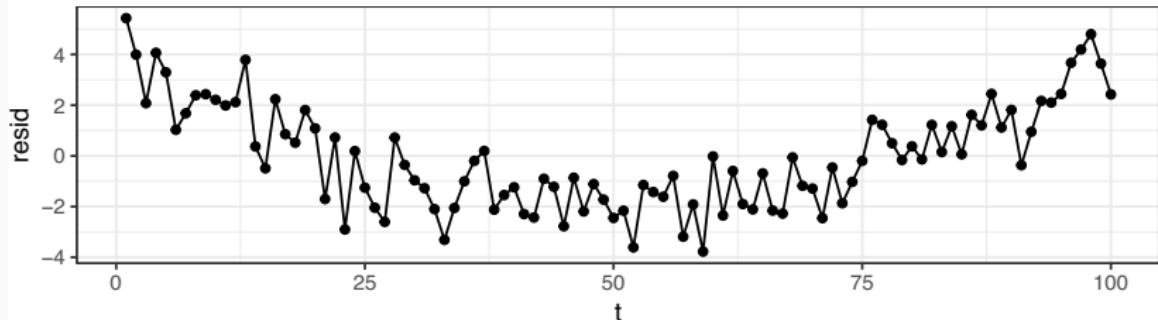
Quadratic trend model

Lets imagine another simple model where $y_t = \delta + \beta t + \gamma t^2 + x_t$ where δ , β , and γ are constants and x_t is a stationary process.

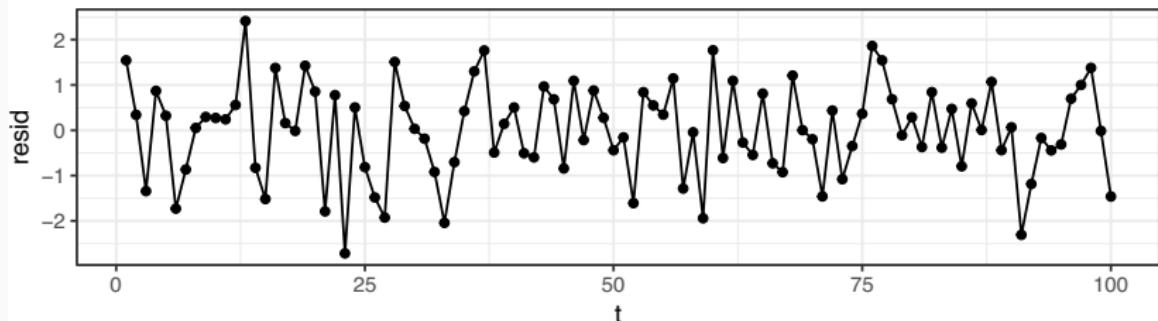


Detrending

Detrended – Linear



Detrended – Quadratic



2nd order differencing

Let $d_t = y_t - y_{t-1}$ be a first order difference then $d_t - d_{t-1}$ is a 2nd order difference.

$$y_t = \delta + \beta t + \gamma t^2 + x_t$$

$$d_t - d_{t-1} = (y_t - y_{t-1}) - (y_{t-1} - y_{t-2})$$

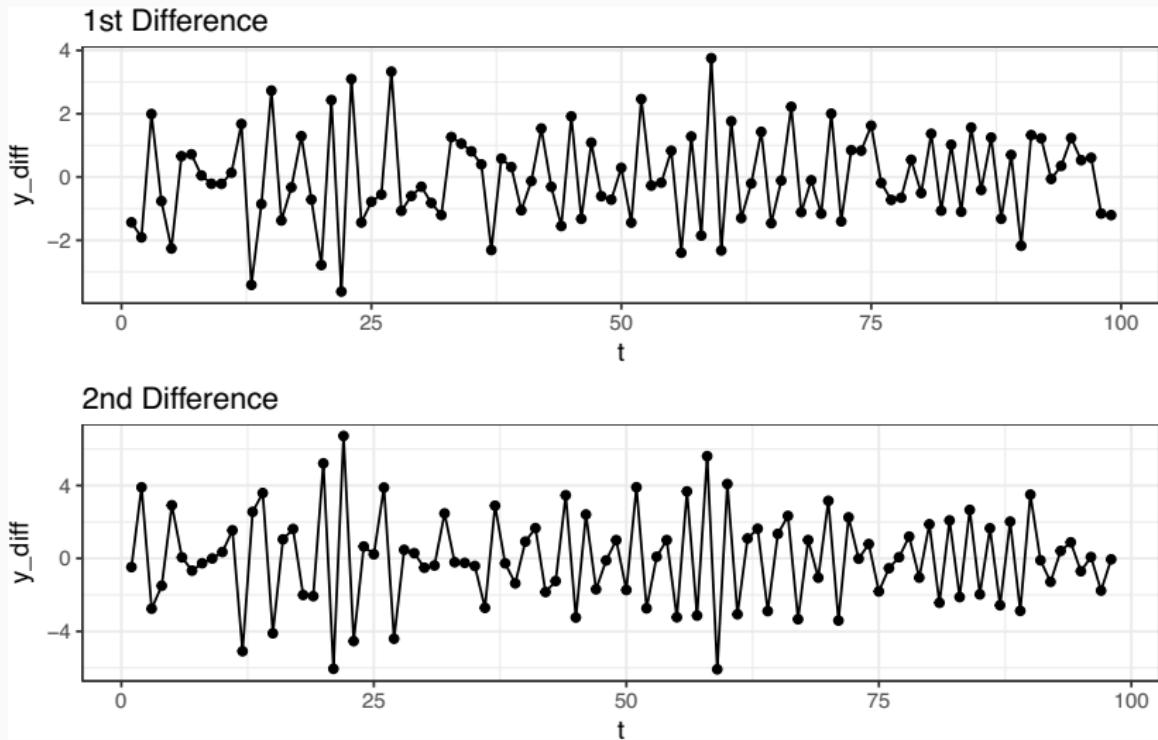
$$d_t = (\delta + \alpha t + \gamma t^2 + x_t) - (\delta + \alpha(t-1) + \gamma(t-1)^2 + x_{t-1})$$

$$= \cancel{\alpha} + \cancel{2\alpha t} - \cancel{\alpha} + (x_t - x_{t-1})$$

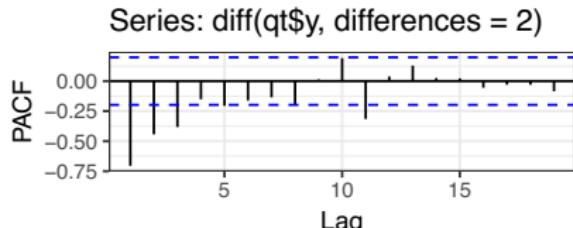
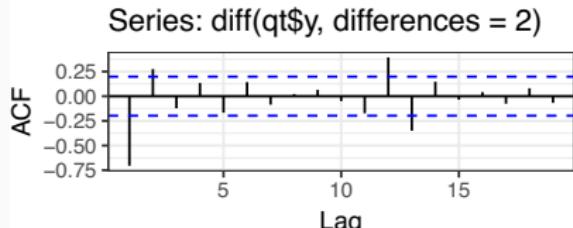
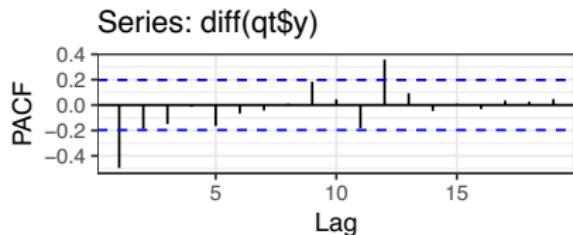
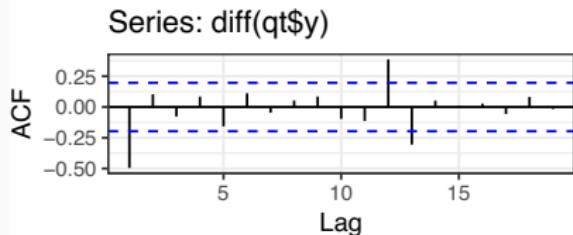
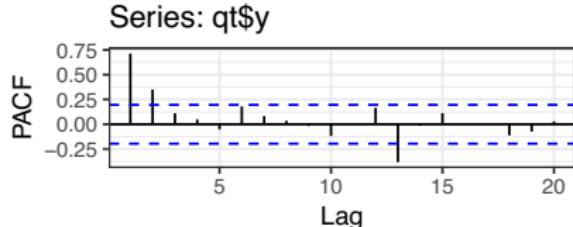
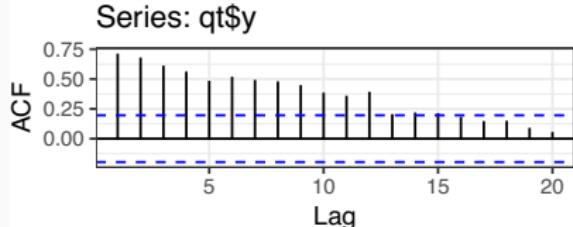
$$d_{t-1} = -\cancel{\alpha} + \cancel{2\alpha(t-1)} - \cancel{\alpha} + (x_{t-1} - x_{t-2})$$

$$= \gamma + \underbrace{x_t - 2x_{t-1} + x_{t-2}}_{\text{Stationary}}$$

Differencing



Differencing - ACF



AR Models

AR(1)

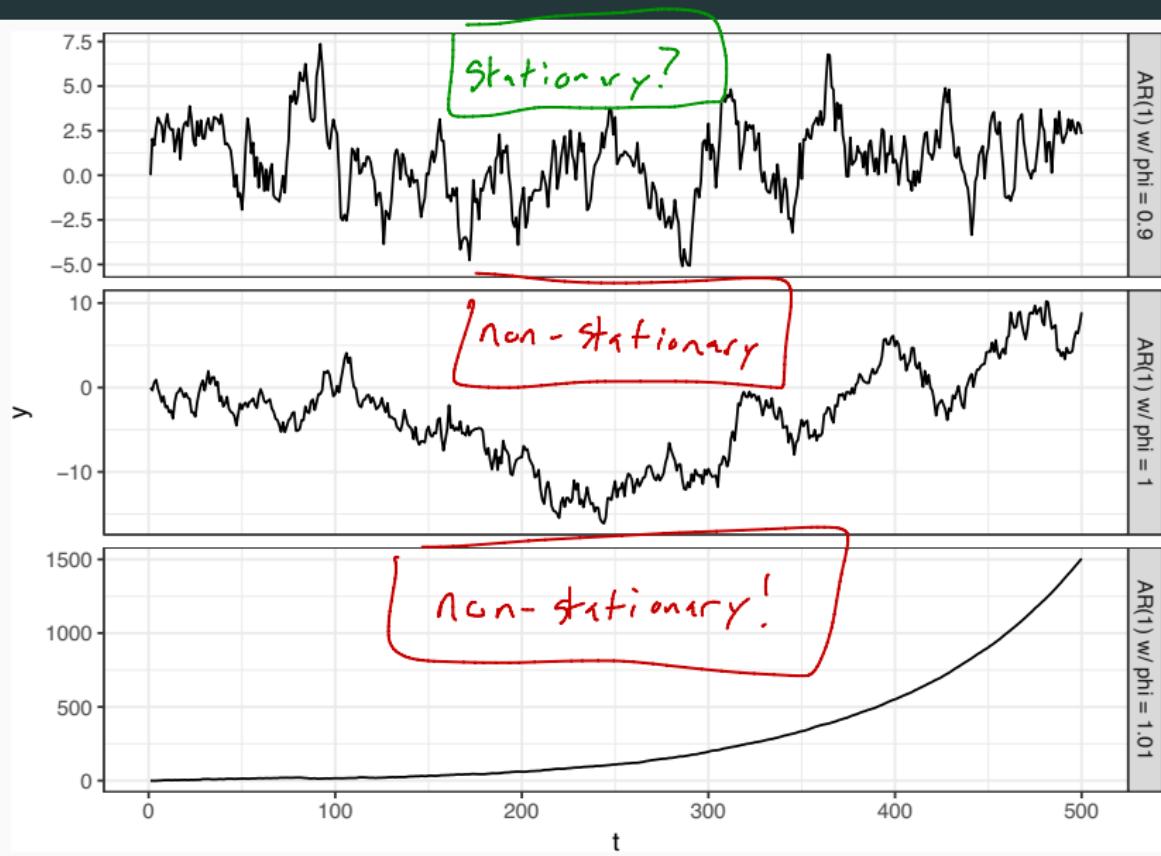
Last time we mentioned a random walk with trend process where

$$y_t = \delta + y_{t-1} + w_t.$$

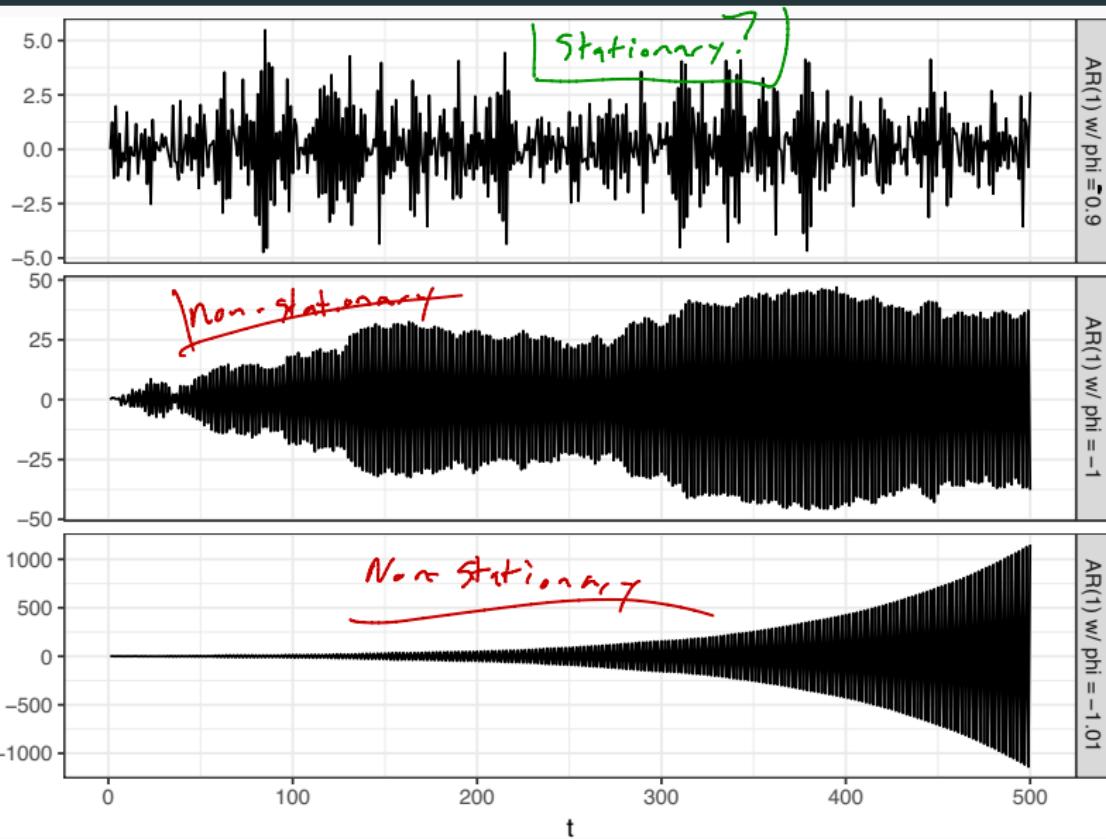
The AR(1) process is a generalization of this where we include a coefficient in front of the y_{t-1} term.

$$AR(1) : \quad y_t = \delta + \phi y_{t-1} + w_t$$

AR(1) - Positive ϕ



AR(1) - Negative ϕ



Stationarity of $AR(1)$ processes

Lets rewrite the $AR(1)$ without any autoregressive terms $(t \rightarrow \infty)$

$$y_t = \delta + \phi y_{t-1} + \epsilon_t$$

$$= \delta + \phi (\delta + \phi y_{t-2} + \epsilon_{t-1}) + \epsilon_t$$

$$= \delta + \phi \delta + \phi^2 y_{t-2} + \epsilon_t + \phi \epsilon_{t-1}$$

$$= \delta + \phi \delta + \phi^2 (\delta + \phi y_{t-3} + \epsilon_{t-2}) + \epsilon_t + \phi \epsilon_{t-1}$$

$$= \delta + \phi \delta + \phi^2 \delta + \phi^3 y_{t-3} + \epsilon_t + \phi \epsilon_{t-1} + \phi^2 \epsilon_{t-2}$$

$$= \delta (1 + \phi + \phi^2 + \phi^3 + \dots) + (\epsilon_t + \phi \epsilon_{t-1} + \phi^2 \epsilon_{t-2} + \phi^3 \epsilon_{t-3})$$

Stationarity of AR(1) processes

Under what conditions will an AR(1) process be stationary?

1. $E(Y_t^2) < \infty$
2. $E(Y_t) = \mu$
3. $\gamma(h) = \text{Cov}(Y_t, Y_{t+h}) = \epsilon(h)$

$$Y_t = \delta(1 + \phi + \phi^2 + \dots) + (\epsilon_t + \phi\epsilon_{t-1} + \phi^2\epsilon_{t-2} + \dots)$$

$$\begin{aligned} E(Y_t) &= \delta(1 + \phi + \phi^2 + \phi^3 + \dots) + \sigma \\ &= \frac{\delta}{1-\phi} \quad \text{if } |\phi| < 1 \end{aligned}$$

$$\begin{aligned} \text{Var}(Y_t) &= \text{Var}(\epsilon_t + \phi\epsilon_{t-1} + \phi^2\epsilon_{t-2} + \dots) \\ &= \sigma^2 + \phi^2\sigma^2 + \phi^4\sigma^2 + \phi^6\sigma^2 + \dots \\ &= \sigma^2 (1 + \phi^2 + \phi^4 + \phi^6 + \dots) = \frac{\sigma^2}{1-\phi^2} \end{aligned}$$

\checkmark

Properties of $AR(1)$ processes

$$\gamma(h) = \text{Cov}(y_t, y_{t-h}) \quad \gamma(0) = \text{Var}(y_t) = \frac{\sigma^2}{1-\phi^2}$$

$$\begin{aligned} \gamma(1) &= \text{Cov}(y_t, y_{t-1}) = E((y_t - \mu)(y_{t-1} - \mu)) \\ &= E((\cancel{\epsilon_t} + \phi \epsilon_{t-1} + \phi^2 \epsilon_{t-2} + \phi^3 \epsilon_{t-3} + \dots \\ &\quad (\uparrow \quad \uparrow \quad \uparrow \quad \uparrow) \quad (\cancel{\epsilon_t} + \phi \epsilon_{t-1} + \phi^2 \epsilon_{t-2} + \phi^3 \epsilon_{t-3} + \dots)) \end{aligned}$$

$$= \phi E(\epsilon_{t-1}^2) + \phi^3 E(\epsilon_{t-2}^2) + \phi^5 E(\epsilon_{t-3}^2) + \dots$$

$$= \phi \sigma^2 + \phi^3 \sigma^2 + \phi^5 \sigma^2 + \dots$$

$$= \phi \sigma^2 (1 + \phi^2 + \phi^4 + \phi^6 + \dots) = \phi \gamma(0)$$

Properties of $AR(1)$ processes

$$\gamma(2) = \phi^2 \gamma(0)$$

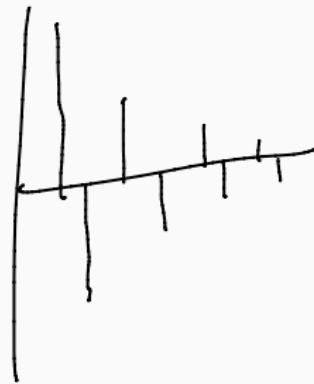
$$\gamma(3) = \phi^3 \gamma(0)$$

⋮

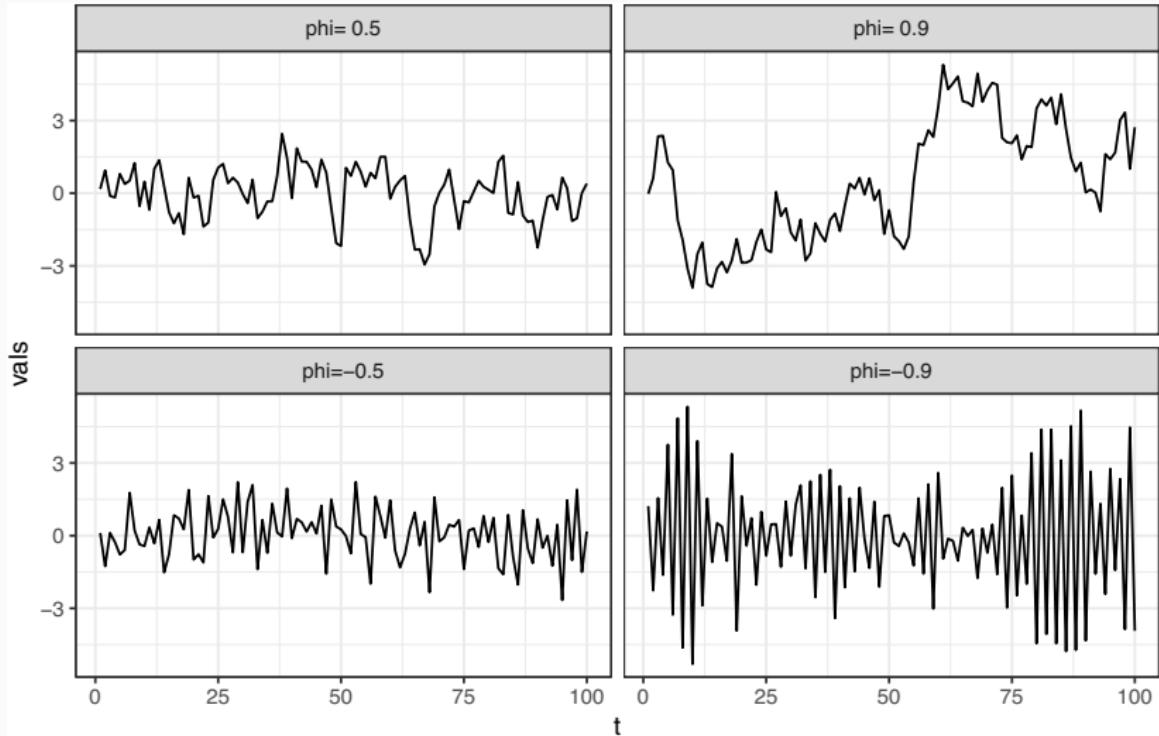
$$\gamma(h) = \phi^h \gamma(0)$$



$$\rho(h) = \frac{\gamma(h)}{\gamma(0)} = \phi^h$$



Identifying AR(1) Processes



Identifying AR(1) Processes - ACFs

