STA 640 — Causal Inference

Chapter 4.1 Treatment Effect Heterogeneity: Causal Subgroup Analysis

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- Acknowledgement: Large part of this lecture is written by Joey Antonelli of University of Florida
- There is huge interest in understanding whether a treatment or policy affects certain individuals more than others
 - Referred to as treatment effect heterogeneity or heterogeneous treatment effects
- Personalized medicine is a huge area of interest
 - What treatment should an individual get
 - Physicians are implicitly considering how treatment effects vary when determining what treatment to assign a patient
 - Given their characteristics, treatment history, etc.

- There are countless other applications for which heterogeneity of the treatment effect is of scientific interest
- Many cancer treatments only work on a subset of the population

Why? What subsets of the population?

- Limited resource settings where not everyone can be assigned treatment
 - Give it to those individuals most likely to benefit
- ► Helps to transport causal effects from one population to another
 - Two populations might have different characteristics and therefore different ATEs

- An additional issue is that sometimes average or marginal treatment effects can mask the effect of a policy
- What if a policy has a positive impact on some individuals and a negative impact on others?
 - ATE will likely be very close to zero
 - Hypothesis tests indicate no treatment effect
 - In truth the treatment is very important
- Looking at heterogeneous treatment effects provides more scientific information than marginal effects alone
 - Immediately recover marginal effects from heterogeneous ones

- There are many questions one can answer in a study of heterogenous treatment effects
 - Which covariates modify the treatment effect?
 - Is there any heterogeneity whatsoever?
 - ► For a given *X*, what is the expected treatment effect (CATE)
 - ► For a given individual, what is their treatment effect (ITE)
- Choice of statistical approach will depend on the goals of the study

The most common target estimand is the conditional average treatment effect

$$CATE = \tau(x) = \mathbb{E}[Y(1) - Y(0)|X = x]$$

Note the ATE is simply the average CATE

ATE =
$$\mathbb{E}[Y(1) - Y(0)] = \int_{x} \tau(x) f_X(x) dx$$

- This shows how the CATE provides additional information over the ATE
 - Once we know the CATE, we immediately know the ATE

- Another relevant estimand refers to subgroup analysis
- Assume we have a subset of the covariate space defined by C, e.g. specific age or gender or medical history
- A subgroup specific estimand is given by

 $\mathbb{E}[Y(1) - Y(0) | X \in C]$

► Commonly we will have non-overlapping regions given by *C*₁,..., *C*_{*G*}, and we estimate

$$\mathbb{E}[Y(1) - Y(0) | X \in C_g]$$
 for $g = 1, ..., G$

And again we can easily recover the ATE by marginalizing over these

Sometimes the CATE is not of interest, but focus is on a subset of predictors given by V ⊂ X:

$$\mathbb{E}[Y(1) - Y(0)|V = v]$$

- Maybe we simply care whether a particular covariate modifies the treatment effect
- This construction is really useful in high-dimensional settings where X is high-dimensional, but we care more about heterogeneity by certain covariates
 - Still need to account for X when adjusting for confounding, but not when estimating heterogeneous treatment effects

Individual treatment effects (ITE) are also of concern

$$\tau_i = Y_i(1) - Y_i(0)$$

- For example, this is the question that personalized medicine looks to address
 - How will the treatment affect this particular individual
- Generally speaking, these are much harder to estimate
 - More uncertainty
 - Prediction intervals are wider than intervals for a mean
 - Stronger assumptions

- ► The literature often conflates the ITE and the CATE
- Clearly, we have that

$$Y_i(1) - Y_i(0) \neq \mathbb{E}[Y(1) - Y(0) | X = X_i]$$

- Related concepts, and certainly the CATE evaluated at X_i is a good point estimate for the ITE of individual i
- Under outcome modeling approach, all estimands are estimated in the same fashion

Identifying assumptions

Estimation of heterogeneous treatment effects (HTE) differs from that of marginal treatment effects, but identification is effectively the same

Easy to see that under SUTVA and unconfoundedness we have

 $\tau(x) = \mathbb{E}[Y(1) - Y(0)|X = x]$ = $\mathbb{E}[Y(1)|X = x] - \mathbb{E}[Y(0)|X = x]$ = $\mathbb{E}[Y(1)|Z = 1, X = x] - \mathbb{E}[Y(0)|Z = 0, X = x]$ = $\mathbb{E}[Y|Z = 1, X = x] - \mathbb{E}[Y|Z = 0, X = x]$

- Unconfoundedness allows us to use data with $Z_i = 0$ to estimate $\mathbb{E}[Y(0)|X = x]$ in the whole population
 - Same for Y(1)

Identifying assumptions

- Overlap is still a fundamental assumption for heterogeneous treatment effects as well
 - With little overlap, causal inference is problematic conceptually and has large uncertainity operationally
- Suppose we have certain regions of the covariate space that are always treated
- ► We have to then extrapolate our estimates of E[Y|Z = 0, X = x] to these individuals with different covariate values
 - Heavily reliant on model specification
 - Difficult to understand the degree of extrapolation
 - Unclear impacts on uncertainty quantification
- We will discuss overlap a bit more in subgroup analysis

Identifying assumptions

In this section, we will mostly cover estimation issues

- There are a lot!
- A lot of other issues inherent to a causal analysis apply here as well
 - Considering plausibility of causal assumptions
 - Sensitivity analysis (to be covered in a couple of weeks)
 - Overlap and balance checks
- When these issues differ in ways unique to heterogeneous treatment effect estimation, we will cover them as they come up

Subgroup analysis

- This sub-chapter focuses on the simplest form of HTE: subgroup analysis
- The simplest form of heterogeneity is subgroup analysis (SGA)
- ► Again suppose we have non-overlapping subsets of the covariate space given by C₁,...,C_G
- Our goal is estimation of

$$\mathbb{E}[Y(1) - Y(0) | X \in C_g]$$
 for $g = 1, ..., G$

We will see that many of the same estimation strategies we've already learned about can be utilized here analogously

Subgroup analysis

- Important that these groups are chosen beforehand, often in a one-variable-at-a-time fashion
 - Might look old-fashioned, by still informative and widely used in practice, e.g. medical research
- There are data-driven approaches for finding the subsets of the population that benefit from treatment
- Generally speaking using the data to find subgroups complicates analyses
 - Valid inference becomes challenging
 - Post selection inference issues
 - Can use data splitting to alleviate these issues
- We will focus for now on situations where these groups are known beforehand

Subgroup analysis (SGA): weighting

- All balancing weights can be directly applied to SGA. Below we will focus on IPW for simplicity
- Recall the original IPW estimator of the ATE

$$\frac{1}{N} \left\{ \sum_{i=1}^{N} \frac{Y_i Z_i}{e(X_i)} - \sum_{i=1}^{N} \frac{Y_i (1-Z_i)}{1-e(X_i)} \right\}$$

which is used as a sample estimate of

$$\mathbb{E}\left[\frac{ZY}{e(X)} - \frac{(1-Z)Y}{1-e(X)}\right]$$

We are using the empirical distribution from the sample to approximate this expectation that is with respect to the overall target population

Subgroup analysis: weighting

▶ Now our target population is the subset of individuals within C_g :

$$\mathbb{E}[Y(1) - Y(0)|X \in C_g] = \mathbb{E}\left[\frac{ZY}{e(X)} - \frac{(1 - Z)Y}{1 - e(X)}\middle|X \in C_g\right]$$

so a natural estimator of this is simply

$$\frac{1}{N_g} \left\{ \sum_{i:X_i \in C_g} \frac{Y_i Z_i}{e(X_i)} - \sum_{i:X_i \in C_g} \frac{Y_i (1 - Z_i)}{1 - e(X_i)} \right\}$$

where $N_g = \sum_{i=1}^n \mathbb{I}(X_i \in C_g)$

Subgroup analysis: weighting

- We simply use the IPW estimator but instead average over just the individuals in the desired subgroup
- The same procedure applies to other balancing weights, e.g. overlap weights, ATT weights
- Can apply this procedure separately within each subgroup to estimate subgroup specific effects
- Remember that balancing weights are intended to construct a weighted population for which the covariates are balanced across treatment groups
 - Does that happen here?

Subgroup analysis: variance-bias tradeoff

- An important question is how the PS is estimated
 - Using the entire sample
 - Using just the individuals in C_g
- Using the full sample aims to ensure balance in the entire target population, not the subgroup specific one
- Nonetheless, if the PS is correctly specified, using the full sample should work well
- ► Using just the individuals in *C_g* will improve balance within the subgroup
 - Less efficient. Bias/variance trade-off
- A simple logistic propensity score model with only main effects of all covariates is not usually adequate in SGA

Subgroup analysis: outcome modeling

- Similar issues occur for outcome modeling or doubly robust estimators
- Recall the outcome modeling estimator

$$\frac{1}{N}\left\{\sum_{i=1}^n \hat{\mu}_1(X_i) - \hat{\mu}_0(X_i)\right\}$$

Can similarly replace this with

$$\frac{1}{N_g} \left\{ \sum_{i:X_i \in C_g} \hat{\mu}_1(X_i) - \hat{\mu}_0(X_i) \right\}$$

in order to estimate the subgroup effect

Subgroup analysis

- Similar decisions need to be made here
- ► Fitting the outcome model only on individuals with X_i ∈ C_g is more flexible, but also less efficient
- One alternative is to fit a model on the full sample, but include interactions between covariates and indicators of subgroup index
 - Similar to fitting separate regression models
 - Can use penalization on interaction terms to shrink/regularize towards the standard model fit on the full data
 - Balance bias and variance concerns
- One solution is the subgroup balancing propensity score (Dong et al. 2020): estimating PS that reaches a comprise between global and subgroup balance
- ► In general, cumbersome to implement

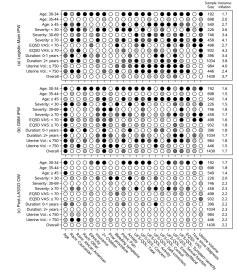
Post-LASSO Algorithm to Balance Bias-Variance Tradeoff

- Yang et al. 2021 (SIM) proposed to use post-LASSO estimate PS. The procedure is:
 - S1. Fit a logistic PS model with all pre-specified covariates and subgroup variables along with pairwise covariate-subgroup interactions, and perform LASSO to select covariate-subgroup interactions (without penalizing the main effects in the model).
 - S2. Estimate PS by refitting the logistic regression with all main effects and selected covariate-subgroup interactions from S1.
 - S3. Calculate a chosen type of weights (e.g. IPW or OW) based on the PS estimated from S2, and check subgroup balance before and after weighting.
 - S4. Estimate the causal effects for all prespecified subgroups using the Hajek estimator within subgroup with the weights from S3.

Visualizing Subgroup Balance: Connect-S plot

- Difficult to visualize subgroup balance. For K subgroups and p covariate, there are Kp standardized differences
- One can draw K love plots, each for p covariates, but still cumbersome
- Connect-S plot (Yang et al. 2021) visualizes *Kp* balance statistics all at once
 - each row represents a subgroup variable, (e.g. a race group)
 - each column represents a confounder/covariate that we want to balance (e.g. age).
 - Each dot corresponds to a specific subgroup and confounder, and the shade of the dot is coded based on the corresponding balance statistics, with darker color meaning more severe imbalance.

Connect S plot: example of COMPARE UF Yang et al. 2021



Confounder

Absolute Standardized Mean Difference ○ ASMD ≤ 0.10 ◎ ASMD 0.10 - 0.15 ● ASMD 0.15 - 0.20 ● ASMD > 0.20

Case study of subgroup analysis - COMPARE UF Yang et al. 2021

- Goal: determine whether certain patient subgroups should receive myomectomy versus hysterectomy (two treatments)
- Pre-specified 35 subgroups: categories of 16 variables including race, age, and baseline symptom severity
- 20 covariates/confounders: including demographics, disease history, quality of life and symptoms
- Total sample size: 1430, 567 in the myomectomy group and 863 patients in the hysterectomy group
- Outcome: quality of life score after 1 year
- Connect-S plot shows imbalance in many subgroup-confounder combinations

Case study of subgroup analysis - COMPARE UF Yang et al. 2021

			UFS-QoL Total at 1 Yea			
Subgroups	Myom. Mean	Hyst. Mean	Mean Difference (95% CI) Myom Hyst.	1		
Age						
30-34	86.6	81.9	4.71 (-10.6, 20.03)		-	
35-44	86.2	94.3	-8.13 (-11.5, -4.72)			
≥ 45	92.1	93.6	-1.49 (-4.91, 1.94)			
30-34	84.5	87.4	-2.84 (-14.6, 8.97)			
35-44	86.5	94.8	-8.29 (-11.8, -4.84)			
≥ 45	90.5	94.8	-4.35 (-7.95, -0.75)			
Severity						
< 30	95.1	97.4	-2.25 (-4.87, 0.38)	-8-		
30-69	87.6	93.4	-5.82 (-9.40, -2.24)			
≥ 70	87.6	90.5	-2.96 (-8.88, 2.97)		-	
< 30	95.0	96.7	-1.65 (-4.13, 0.84)			
30-69	86.5	94.9	-8.40 (-12.0, -4.76)			
≥ 70	84.5	91.3	-6.75 (-12.5, -1.04)			
EQ5D VAS						
< 70	85.4	91.3	-5.97 (-11.1, -0.84)			
≥ 70	90.3	94.3	-4.04 (-7.14, -0.94)			
< 70	82.6	91.0	-8.33 (-14.1, -2.58)			
≥ 70	89.7	95.9	-6.27 (-8.89, -3.65)			
Duration						
0-1 years	90.5	95.1	-4.62 (-9.09, -0.15)			
2+ years	88.2	92.6	-4.45 (-7.69, -1.22)			
0-1 years	89.6	95.5	-5.86 (-10.0, -1.68)			
2+ years	86.5	93.9	-7.36 (-10.6, -4.15)			
Uterine Volume						
≤ 750	88.4	94.0	-5.50 (-8.62, -2.39)			
> 750	89.7	91.9	-2.22 (-7.00, 2.57)		-	
≤ 750	85.8	95.3	-9.54 (-12.7, -6.43)			
> 750	91.1	92.3	-1.13 (-5.57, 3.31)		_	
			-	-10 0	10	20
				-10 0	10	20
				← Hyst. Better	Myom. Better	•
				_	-	

Logistic-Main IPW Post-LASSO OW

UFS-QoL Total at 1 Year