

①

The F-test for nested Model evaluation

$$E[y | \underline{x}_1, \underline{x}_2] = \underline{\beta}_1^T \underline{x}_1 + \underline{\beta}_2^T \underline{x}_2 \quad \begin{array}{l} \underline{x}_1 \in \mathbb{R}^{p_1} \\ \underline{x}_2 \in \mathbb{R}^{p_2} \end{array}$$

Examples:

Old faithful data: $\underline{x}_1 = (1, \text{eruptions})$
 $\underline{x}_2 = (\text{eruptions}^2, \text{eruptions}^3)$

Salary data: $\underline{x}_1 = (1, \text{years}, \text{assoc}, \text{prof})$
 $\underline{x}_2 = (\text{years} * \text{assoc}, \text{years} * \text{prof})$

Goal: Evaluate evidence of an effect of $\underline{x}_2$ (=)
 " that $\underline{\beta}_2 = 0$

$$H: \underline{\beta}_2 = 0$$

$$K: \underline{\beta}_2 \neq 0$$

=

F-test procedure

- 1) fit reduced model $y \sim \underline{x}_1$, get RSS_r
- 2) fit full model $y \sim \underline{x}_1 + \underline{x}_2$, get RSS_f
- 3) compute $F_{\text{OBS}} = \frac{(RSS_r - RSS_f) / p_2}{RSS_f / (n - p_1 - p_2)}$
- 4) Reject $H: \underline{\beta}_2 = 0$ if F_{OBS} is large.

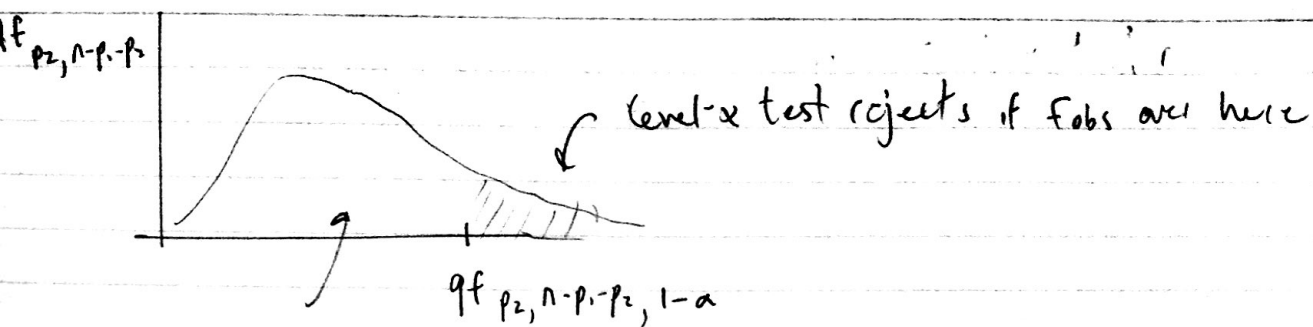
How large is large?

$$\text{If } 1) y_i = \underline{\beta}_1^T \underline{x}_{1i} + \epsilon_i$$

then the dist of F_{OBS} is $F_{p_2, n-p_1-p_2}$

$$2) \epsilon_1, \dots, \epsilon_n \sim \text{i.i.d. } N(0, \sigma^2),$$

(2)



level α test doesn't reject if F_{obs} over here.

What is the F-dist? Where does it come from?

χ^2 dist

$$\text{If } z_1, \dots, z_n \sim \text{iid } N(0, 1) \Rightarrow \sum z_i^2 \sim \chi_n^2$$

$$y_1, \dots, y_n \sim \text{iid } N(\mu, \sigma^2) \Rightarrow \frac{1}{\sigma^2} \sum (y_i - \bar{y})^2 \sim \chi_{n-1}^2$$

$$y \sim N(X\beta, \sigma^2 I) \Rightarrow \frac{1}{\sigma^2} \sum (y_i - \hat{y}_i)^2 \sim \chi_{n-p}^2$$

$$\Rightarrow \text{RSS}/\sigma^2 \sim \chi_{n-p}^2$$

Also! if $\left. \begin{array}{l} X_1 \sim \chi_{k_1}^2 \\ X_2 \sim \chi_{k_2}^2 \\ X_1, X_2 \text{ indep} \end{array} \right\} \Rightarrow X_1 + X_2 \sim \chi_{k_1+k_2}^2$

F-dist

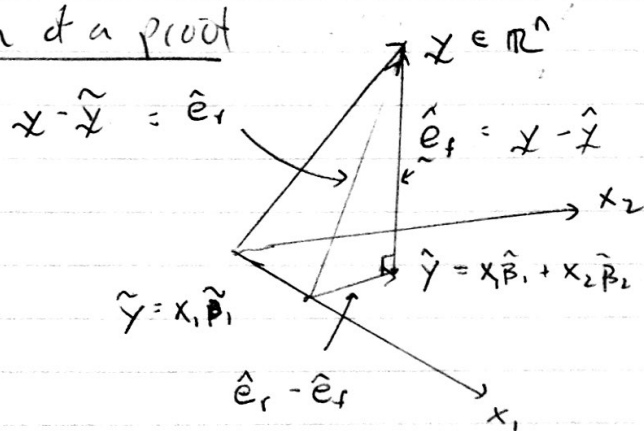
$$\text{If } \left. \begin{array}{l} X_1 \sim \chi_{k_1}^2 \\ X_2 \sim \chi_{k_2}^2 \\ X_1 \perp X_2 \end{array} \right\} \Rightarrow \frac{X_1/k_1}{X_2/k_2} \sim F_{k_1, k_2}$$

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Claim: Under H_0 and normality assumptions

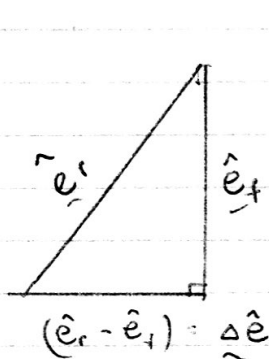
$$\frac{(RSS_r - RSS_f) / p_2}{RSS_f / (n - p_1 - p_2)} \sim F_{p_2, n - p_1 - p_2}$$

Sketch of a proof



$\hat{e}_f$ = resid under full model

$\hat{e}_r$ = resid under reduced model



$$\|\hat{e}_f\|^2 = RSS_f$$

$$\|\hat{e}_r\|^2 = RSS_r$$

$$\|\hat{e}_r - \hat{e}_f\|^2 = ? \quad (\text{Pythagoras: } a^2 + b^2 = c^2) \\ = RSS_r - RSS_f$$

$\hat{e}_f$ is n dimensional, but lives in $n - p_1 - p_2$ dim space
 $\frac{1}{\sigma^2} RSS_f \sim \chi^2_{n - p_1 - p_2}$

$\Delta \hat{e}$ is p_2 dimensional
 $\frac{1}{\sigma^2} (RSS_r - RSS_f) \sim \chi^2_{p_2}$

Also, $\Delta \hat{e}$, $\hat{e}_f$ are orthogonal $\Rightarrow$ statistically independent

$\Rightarrow RSS_f$ indep of $RSS_r - RSS_f$

$$\Rightarrow \frac{\frac{1}{\sigma^2} (RSS_r - RSS_f) / p_2}{\frac{1}{\sigma^2} (RSS_f) / (n - p_1 - p_2)} \stackrel{\chi}{=} \frac{\chi^2_{p_2} / p_2}{\chi^2_{n - p_1 - p_2} / (n - p_1 - p_2)} \sim F_{p_2, n - p_1 - p_2}$$