

Inference for non-Stationary, non-Gaussian, Irregularly-Sampled Processes

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- 1 Introduction
Easy Time Series Examples
- 2 Extension 1: Irregular Sampling
- 3 Extension 2: Non-Gaussian
- 4 Extension 3: Non-Stationary
- 5 Wrap-up

Outline

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Basic Ideas

The entire talk in one slide:

- Basic idea 1: Even if we only observe events Y_i at discrete times $\{t_i\}$, we can *model* these as observations $Y(t_i)$ of a *process* $Y(t)$ that is well-defined (if un-observed) in **continuous time**. This overcomes the “regularly-spaced observations” limitation.

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- Basic idea 2: Almost anything you can do with Gaussian distributions can also be done with other **Infinitely-Divisible (ID) distributions**, such as Poisson, Negative Binomial, Gamma, α -Stable, Cauchy, Beta Process. Some of these feature discrete (integer) values, and some feature heavy tails, offering a wider range of behavior than is possible with Gaussian time series or processes.

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- Basic idea 3: **Start with Stationary** processes (like the Ornstein-Uhlenbeck process) or processes with Stationary increments (like Brownian Motion), **then extend** to *non-stationary* processes (like Diffusions).

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Easy TS Example: AR(1)

Fix $\mu \in \mathbb{R}$, $\sigma^2 > 0$, and $\rho \in (0, 1)$.

Draw $X_0 \sim \text{No}(\mu, \sigma^2)$.

For times $t = 1, 2, \dots$, set

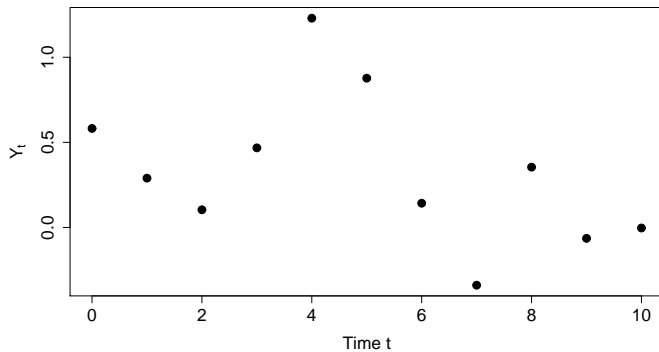
$$\begin{aligned} X_t &:= \rho X_{t-1} + \zeta_t \\ &= \rho^t X_0 + \sum_{s=0}^{t-1} \zeta_{t-s} \rho^s = \sum_{s=0}^{\infty} \zeta_{t-s} \rho^s \end{aligned}$$

with Normal innovations $\{\zeta_t\} \stackrel{\text{iid}}{\sim} \text{No}((1 - \rho)\mu, (1 - \rho^2)\sigma^2)$

Like most commonly-studied Time Series models, this features:

- **Regularly-spaced** observation times $t = 0, 1, 2, 3, \dots$;
- **Gaussian** marginal distributions $X_t \sim \text{No}(\mu, \sigma^2)$;
- **Stationary** distributions, with autocorrelation $\text{Corr}(X_s, X_t) = \rho^{|t-s|}$.

Gaussian AR(1) Time Series: Equally-Spaced



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Extension 1: Continuous Time

Set $\lambda := -\log \rho > 0$ and let $\zeta(ds)$ be a random measure assigning to disjoint intervals $(a_i, b_i]$ independent random variables

$$\zeta((a, b]) \sim \text{No}((b - a)\lambda\mu, (b - a)2\lambda\sigma^2).$$

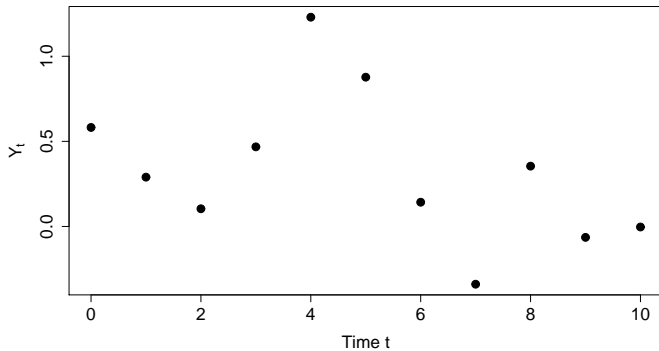
Now take $X_0 \sim \text{No}(\mu, \sigma^2)$ and for $t > 0$ set

$$X_t := X_0 e^{-\lambda t} + \int_0^t e^{-\lambda(t-s)} \zeta(ds) = \int_{-\infty}^t e^{-\lambda(t-s)} \zeta(ds)$$

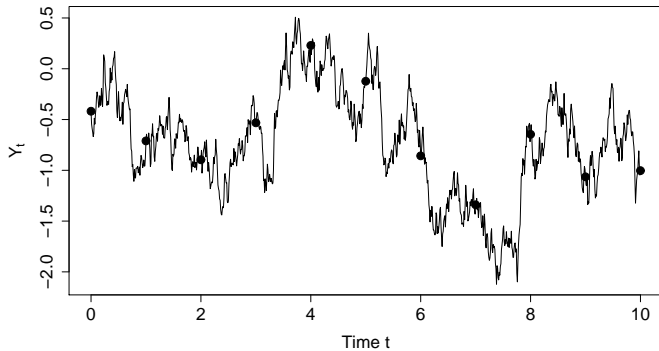
to find a process *for all* $t > 0$ (or all $t \in \mathbb{R}$) with the exact same joint distribution at $t \in \mathbb{N}_0 = \{0, 1, 2, \dots\}$, with

- **Irregularly-spaced** observation times $\{t_i\} \subset \mathbb{R}_+$;
- **Gaussian** marginal distributions $X_t \sim \text{No}(\mu, \sigma^2)$;
- **Stationary** distributions, with autocorrelation $\text{Corr}(X_s, X_t) = \rho^{|t-s|}$.

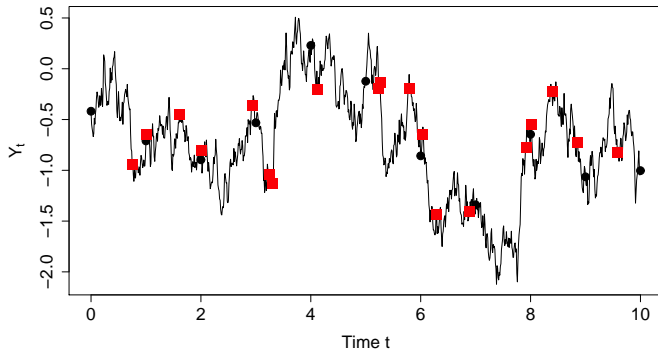
Gaussian AR(1) Time Series: Equally-Spaced



O-U Process with Regularly-Spaced Observations



O-U Process with Irregularly-Spaced Observations



Inference 1: Continuous Time

Inference is easy— MLEs or conjugate Bayes estimators are readily available, as is LH for arbitrary observation pairs $\{(X_{t_i}, t_i)\}$ because $s < t \Rightarrow$

$$X_t \mid \{X_u : u \leq s\} \sim \text{No}(\mu + (X_s - \mu)\rho, \sigma^2(1 - \rho^2)), \quad \rho := e^{-\lambda|t-s|}$$

so with Metropolis-Hastings we can sample posterior for any prior $\pi(\mu, \sigma^2, \lambda)$.

Easy extension to AR(p):

$$X_t + a_1 X_{t-1} + \dots + a_p X_{t-p} := \zeta_t$$

For example, by expressing as a **vector** AR(1) for

$$\mathbf{X}_t = [X_t, X_{t-1}, \dots, X_{t-p+1}]'$$

$$\mathbf{X}_t := R\mathbf{X}_{t-1} + \zeta_t$$

for some $a \in \mathbb{R}^p$

Extension 1: Continuous Time, More Broadly

Any stationary $AR(p)$ has a $MA(\infty)$ representation:

$$\rho(L)X(t) = \zeta_t$$

for the left-shift operator $LX(t) := X(t - 1)$ and the polynomial

$$\rho(z) = 1 + \sum_{j=1}^p a_j z^j = 1 + a_1 z + \cdots + a_p z^p.$$

If $\rho(z)$ has all its roots outside the unit circle, then

$$\frac{1}{\rho(z)} = 1 + \sum_{i=1}^{\infty} b_i z^i = 1 + b_1 z + b_2 z^2 + \cdots$$

$$X(t) = \frac{1}{\rho}(L)(\zeta_t) = \zeta_t + b_1 \zeta_{t-1} + b_2 \zeta_{t-2} + \cdots$$

Under **suitable conditions**, this has a continuous version

$$X(t) = \int_{-\infty}^t b(t-s)\zeta(ds)$$

for all $t \geq 0$, with the same distribution at times $t \in \mathbb{N}_0$.

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Extension 2: Non-Gaussian

Just as **AR(ρ) Time Series**

$$X_t = \rho^t X_0 + \sum_{s=0}^{t-1} \zeta_{t-s} \rho^s$$

can be constructed with *any* iid innovations $\{\zeta_s\}$, so too for the continuous-time version

$$X_t := X_0 e^{-\lambda t} + \int_0^t e^{-\lambda(t-s)} \zeta(ds).$$

The **random measure** $\zeta(ds)$ can be anything that is “iid” in the sense that:

- For disjoint sets $A_i \subset \mathbb{R}$, the RVs $\{\zeta(A_i)\}$ are indep;
- The $\{\zeta(a, b]\}$ distributions only depend on $(b - a)$.

For example, can have

$$\zeta(a, b] \sim \text{No}(\mu(b - a), \sigma^2(b - a))$$

$$\zeta(a, b] \sim \text{Ga}(\alpha(b - a), \beta)$$

$$\zeta(a, b] \sim \text{St}_A(\alpha, \beta, \gamma(b - a), \delta(b - a))$$

Ornstein-Uhlenbeck

Positive, Expo. Tails

Heavy (Pareto) Tails

Extension 2: Non-Gaussian

The two conditions

- For disjoint sets $A_i \subset \mathbb{R}$, the RVs $\{\zeta(A_i)\}$ are indep;
- The $\{\zeta(a, b)\}$ distributions only depend on $(b - a)$.

require that each $\zeta(A)$ should be **Infinitely Divisible**, or **ID**.

Examples:

ID Continuous	ID Discrete	Not ID
Normal	Poisson	Binomial
Gamma	Negative Binomial	Beta
α -Stable	$p_i \propto \frac{1}{(i+a)(i+b)}$	Uniform

Extension 2: Non-Gaussian

But what about **Autocorrelated Count Data**?

- Binned photon counts in satellite Gamma Ray detectors?
- Binned photon counts in particle accelerator detectors?
- Seismic event counts?
- Pyroclastic flow counts?
- Rockfall counts near active volcano?
- Failures of complex systems?
- Rare disease case counts?

AR(ρ) and its continuous versions wouldn't respect **Integer Nature** of data. If $X_t \in \mathbb{Z}$, and $|\rho| < 1$, then

$$X_{t+1} = \rho X_t + \zeta_{t+1} \notin \mathbb{Z}.$$

Alternatives?

Extension 2: Non-Gaussian

Here's a way to construct, and make inference in,

- **stationary** (for now) process X_t with
- **continuous time** $t \in \mathbb{R}_+$ for *any*
- **non-Gaussian** marginal **Infinitely Divisible (ID)** dist'ns including both continuous dist'ns ($\text{Ga}(\theta, \beta)$, $\text{St}_A(\alpha, \beta, \theta, \delta)$, *etc.*) and discrete **count distributions** ($\text{Po}(\theta)$, $\text{NB}(\theta, p)$, *etc.*)

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In fact, except for the special cases of Gaussian and Poisson, we can do this in **many ways**.

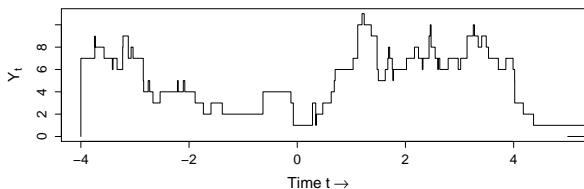
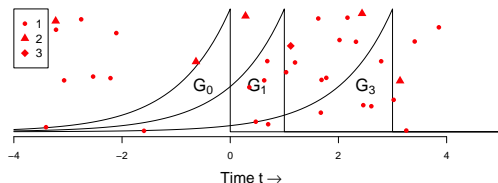
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In fact, except for the special cases of Gaussian and Poisson, we can do this in **many ways**. Let's look at an example.

Ex 2: Negative Binomial Example



AR(1)-like Negative Binomial Process

Based on **Random Measure** $\mathcal{N}(dx dy) \sim \text{NB}(\alpha dx dy, \beta)$ on \mathbb{R}^2

AR(1)-like Negative Binomial Process

Properties:

- $X_t \sim \text{NB}(\alpha, \beta)$ for all t ;
- $\text{Corr}(X_s, X_t) = \exp(-\lambda|t - s|)$.

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Does this characterize the joint distribution of $\{X_{t_i}\}$?

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Properties:

- $X_t \sim \text{NB}(\alpha, \beta)$ for all t ;
- $\text{Corr}(X_s, X_t) = \exp(-\lambda|t - s|)$.

Does this characterize the joint distribution of $\{X_t\}$?

Nope. Here's a different process, the *Branching* NB:

- **Immigration** at rate ι ;
- **Birth** at rate β ;
- **Death** at rate δ ;

$$X_{t+\epsilon} = \begin{cases} X_t + 1 & \text{with probability } o(\epsilon) + \epsilon(\iota + \beta X_t) \\ X_t & \text{with probability } o(\epsilon) + 1 - \epsilon(\iota + (\beta + \delta)X_t) \\ X_t - 1 & \text{with probability } o(\epsilon) + \epsilon\delta X_t \end{cases}$$

$\sim \text{NB}(\alpha, \rho)$, with autocovariance $\exp(-\lambda|t - s|)$.

AR(1)-like Negative Binomial Process

Are these two processes the same?

AR(1)-like Negative Binomial Process

Are these two processes the same?

Nope.

- The **Random Measure NB** process has jumps of all possible non-zero magnitudes $\Delta X_t := [X_t - X_{t-}] \in \mathbb{Z} \setminus \{0\}$, while the **Branching NB** process has only jumps of $\Delta X_t = \pm 1$;
- The **Branching NB** process is Markov, so for $s < t$

$$P[X_t \in A \mid \mathcal{F}_s] = P[X_t \in A \mid X_s],$$

while the **Random Measure NB** process isn't: if it has a jump $\Delta X_s = 7$, for example, then sooner or later there must follow a jump $\Delta X_t = -7$, so $P[X_t = 3 \mid X_s = 10]$ depends on the history \mathcal{F}_s , not just the value X_s .

- The **Random Measure NB** process is **multivariate** ID, while **Branching NB** process is not.

AR(1)-like Negative Binomial Process

Are these the only two?

AR(1)-like Negative Binomial Process

Are these the only two?

Oh no.

- A discrete time AR(1)-like Markov process exists for **any ID distribution** and any “auto-correlation” ρ , based on **Thinning**:

$$Z \sim f(z | \lambda, \phi) = X + Y,$$

$$X \sim f(x | \rho\lambda, \phi) \perp\!\!\!\perp Y \sim f(y | \bar{\rho}\lambda, \phi) \quad [\bar{\rho} := (1 - \rho)]$$

$$X | Z \sim f(x | z, \rho, \lambda, \phi) = f(x | \rho\lambda, \phi) f(z - x | \bar{\rho}\lambda, \phi) / f(z | \lambda, \phi)$$

- Given $\{X_m : m \leq n\} \sim f(x | \lambda, \phi)$, draw

$$\xi_{n+1} \sim f(\xi | X_n, \rho, \lambda, \phi) \perp\!\!\!\perp \eta_{n+1} \sim f(\eta | \bar{\rho}\lambda, \phi)$$

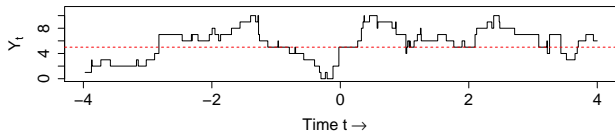
and set

$$X_{n+1} := \xi_{n+1} + \eta_{n+1} \sim f(x | \lambda, \phi).$$

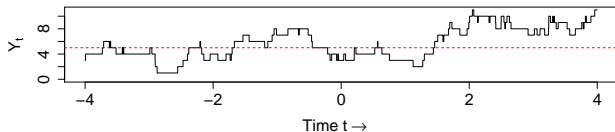
- A continuous-time version exists too.

AR(1)-like Negative Binomial Process

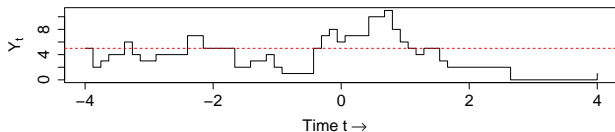
Random Measure NB($\alpha = 10$, $p = 0.67$, $\lambda = 1$)



Branching NB($\alpha = 10$, $p = 0.67$, $\lambda = 1$)



Continuous Thin NB($\alpha = 10$, $p = 0.67$, $\lambda = 1$)



AR(1)-like Negative Binomial Process

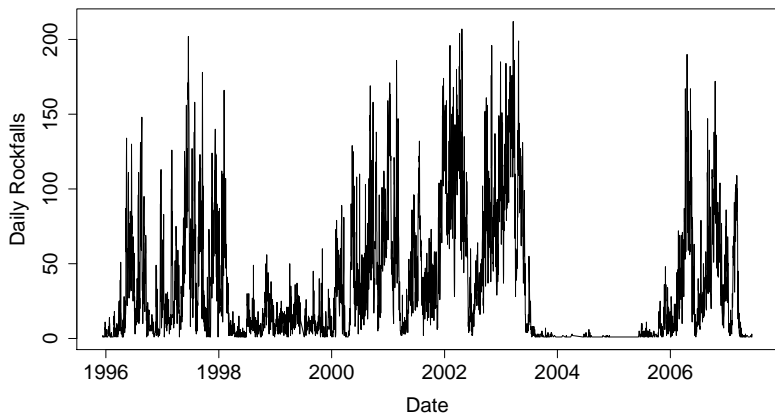
Larry Brown (U Penn) & I found six different AR(1)-like (also AR(p)-like) processes for each ID distribution, with subtly different properties:

- Markov?
- Are *all* finite-dimensional marginals ID?
- In continuous time, are paths continuous? Increments ± 1 ? Or bigger?
- Time-reversible?

In Wolpert & L. D. Brown (2016+) we present a complete class theorem characterizing all *Markov Infinitely-Divisible Stationary Time-Reversible Integer-Valued* (“MISTI”) Processes.

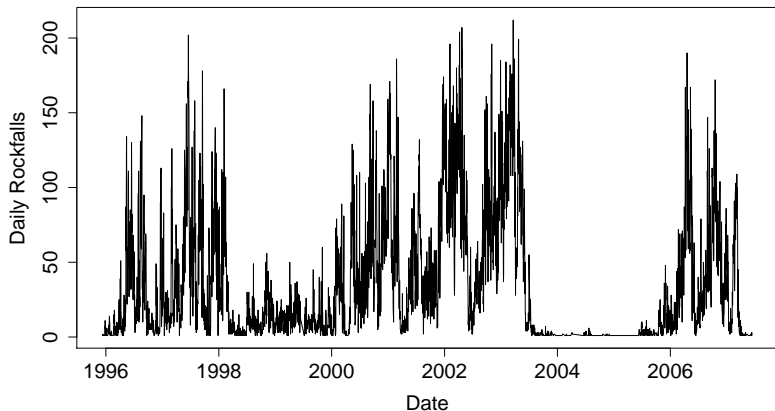
Let's look at a motivating example (not Astro, sorry!).

Motivation for a Negative Binomial series...



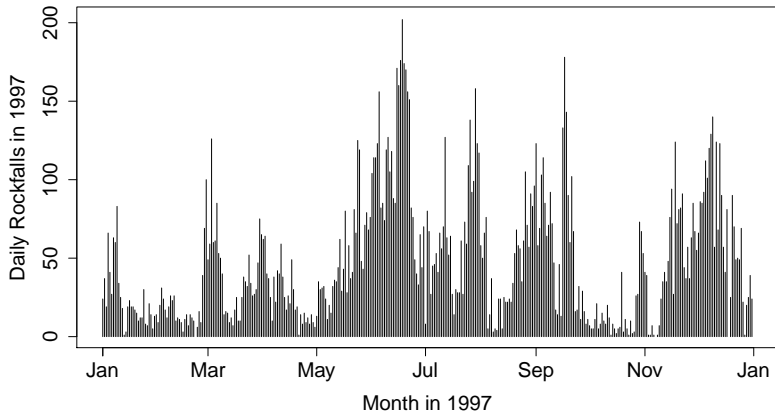
Rockfall counts at Soufrière Hills Volcano, Montserrat

Stationary?



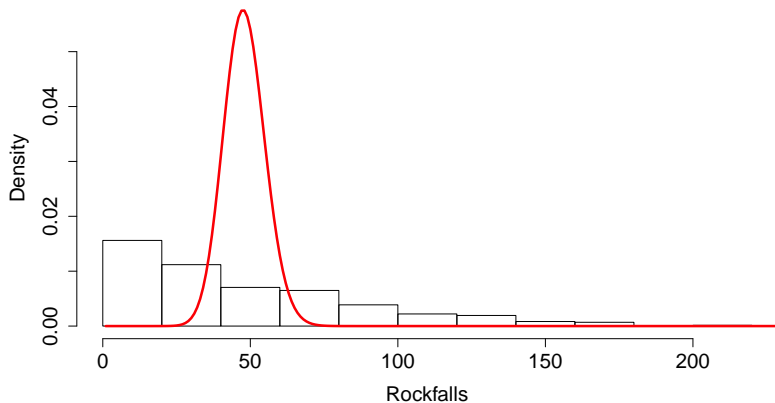
Looks a little patchy; maybe we can find a homogeneous subset. Looks like 1997 might be a less patchy year...

Subset, just the 1997 Rockfalls



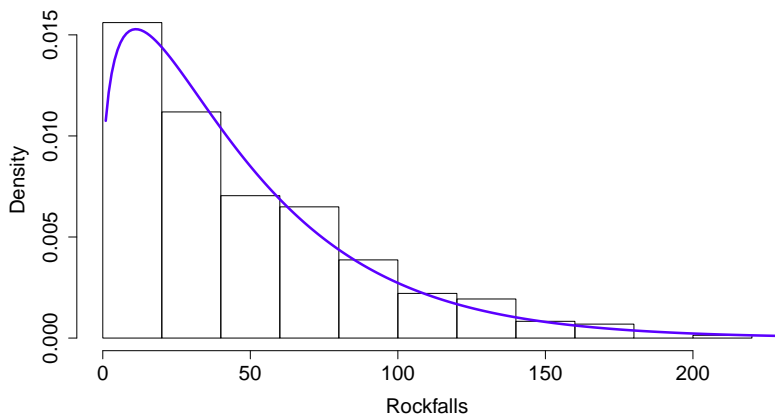
Poisson fit is horrible...

Histogram of 1997 Rockfall Counts with $Po(\lambda = 48.01)$ Overlay

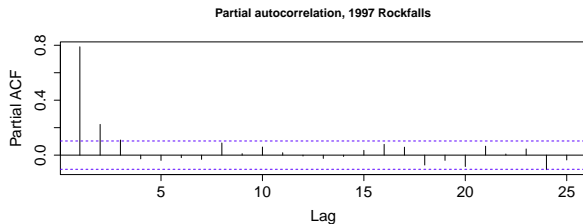
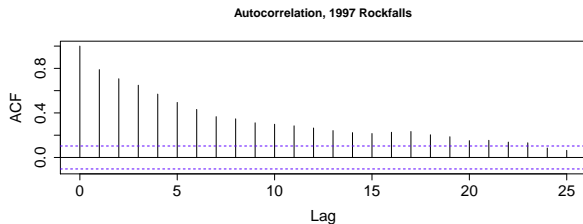


Negative Binomial fit is great.

Histogram of 1997 Rockfall Counts with NB($\alpha = 1.32$, $p = 0.03$) Overlay



But not IID, need autocorrelation



SO,

We (esp. **Jianyu Wang** '13) built a regression model based on one of the AR(1)-like NB models above, to relate

→ (easily counted) rockfall counts to

→ (hard to measure) subsurface volcanic magma flows,

to improve volcanic hazard forecasting.

Related issues were addressed by **Mary Beth Broadbent** ('14) in constructing NPB light curve models for Gamma Ray Bursts, with the help of astronomers **Tom Loredo** and **Jon Hakkila** (Thanks!).

A shout out about **Another cool modeling idea:**

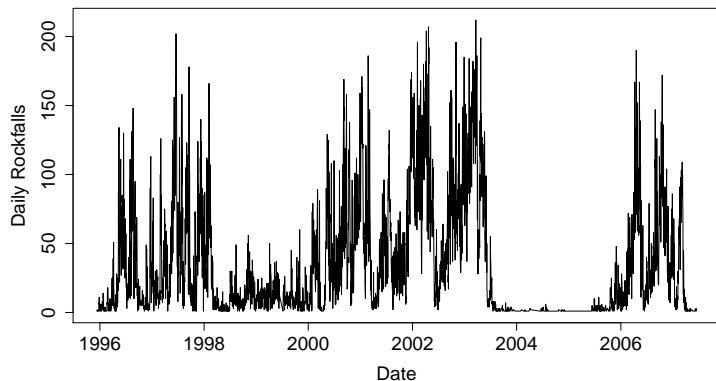
Robert Lund (Clemson) and his student Yunwei Cui (2009) found an interesting new way to model stationary integer-valued time series and processes using [Renewal Theory](#). In discrete time their method (**unlike ours**) is able to model [negative autocorrelation](#) and cyclic behavior. If you have negatively-autocorrelated integer data, look into it.

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Extension 3: Non-Stationary

You might have noticed that the rockfall counts don't look stationary



So, we are implementing **random time-change**.

Extension 3: Non-Stationary

Begin with stationary process $X(t)$...

Then, construct **random time-change** model $t \rightarrow R_t$ and set

$$Y_t := X(R_t)$$

We used:

$$R_t = \int_0^t \lambda(s) ds,$$

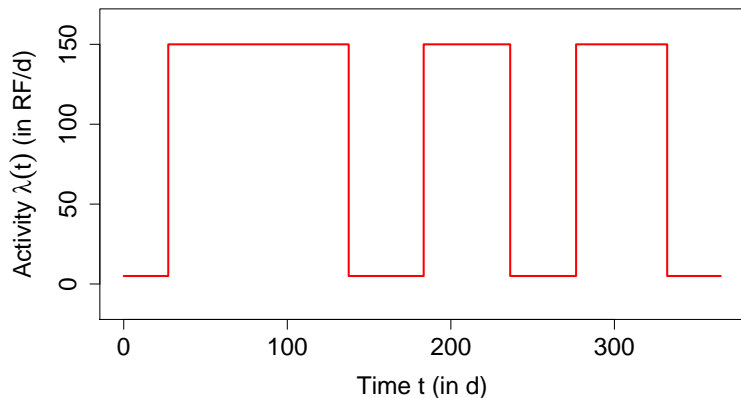
with two-level

$$\lambda(s) = \begin{cases} \lambda_+ & s_i < s \leq t_i \\ \lambda_- & t_i < s \leq s_{i+1} \end{cases}$$

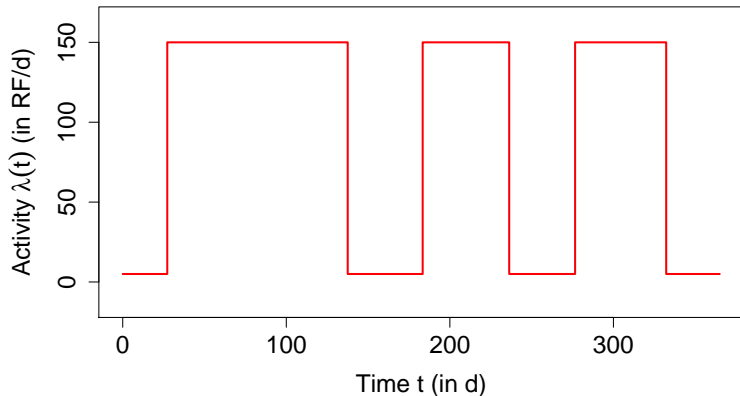
for uncertain levels $0 < \lambda_- < \lambda_+ < \infty$ and transition times

$s_1 < t_1 < s_2 < t_2 < \dots$

Extension 3: Non-Stationary



Extension 3: Non-Stationary



A work in progress, wish us luck.

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Wrap-up

Conclusions:

- Statistical methods are available for data that are **non-Stationary**, or **non-Gaussian**, or **irregularly sampled**, or all three.
- They're not built into SAS.
- BUT, together, Astronomers and Statisticians can build problem-specific tools to support *estimation* and *prediction* and (especially Bayesian) *inference*,
- using routine simulation-based computational methods.

Thanks!

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Statistical, Mathematical and Computational Methods for Astronomy (ASTRO)

Glad to see you here!

