

Answer 1

Probability of fair, six-sided dice. Let $A, B \in \{1, 2, \dots, 6\}$ represent the events from the first and second coin flip respectively.

We want to find: $\mathbf{p}(A = 5 \text{ or } B = 5)$

We know that $\mathbf{p}(A \text{ or } B) = \mathbf{p}(A) + \mathbf{p}(B) - \mathbf{p}(A \text{ and } B)$

So,

$$\mathbf{p}(A = 5 \text{ or } B = 5) = \mathbf{p}(A = 5) + \mathbf{p}(B = 5) - \mathbf{p}(A = 5 \text{ and } B = 5) \quad (1)$$

$$\mathbf{p}(A = 5) = \mathbf{p}(B = 5) = \frac{1}{6} \quad (2)$$

$$\mathbf{p}(A = 5 \text{ and } B = 5) = \mathbf{p}(A = 5) \mathbf{p}(B = 5) = \frac{1}{36} \text{ (by independence)} \quad (3)$$

Substituting eqns (2) and (3) into (1)

$$\Rightarrow \mathbf{p}(A = 5 \text{ or } B = 5) = \frac{1}{6} + \frac{1}{6} - \frac{1}{36} = \frac{11}{36}$$

Answer 2

Bayes Rule. Let $X = \{0, 1\}$ represent if I have the disease or not where $X = 1$ means I have the disease and $X = 0$ means I do not have the disease. Also let $Y = \{\text{pos}, \text{neg}\}$ where $Y = \text{pos}$ means the test is positive and $Y = \text{neg}$ means the test is negative. We know:

$$\mathbf{p}(X = 1) = 0.001$$

$$\mathbf{p}(X = 0) = 0.999$$

$$\mathbf{p}(Y = \text{pos} \mid X = 1) = 0.9$$

$$\mathbf{p}(Y = \text{neg} \mid X = 1) = 0.1$$

$$\mathbf{p}(Y = \text{pos} \mid X = 0) = 0.2$$

$$\mathbf{p}(Y = \text{neg} \mid X = 0) = 0.8$$

We want to find: $\mathbf{p}(X = 1 \mid Y = \text{pos})$. Using Bayes Rule:

$$\mathbf{p}(X = 1 \mid Y = \text{pos}) = \frac{\mathbf{p}(Y = \text{pos} \mid X = 1) \mathbf{p}(X = 1)}{\mathbf{p}(Y = \text{pos})} \quad (4)$$

Where,

$$\begin{aligned} \mathbf{p}(Y = \text{pos}) &= \mathbf{p}(Y = \text{pos} \mid X = 0) \mathbf{p}(X = 0) + \mathbf{p}(Y = \text{pos} \mid X = 1) \mathbf{p}(X = 1) \\ &= 0.2 \times 0.999 + 0.9 \times 0.001 \approx 0.2007 \end{aligned} \quad (5)$$

Now, eqn (4) is

$$\begin{aligned} \mathbf{p}(X = 1 \mid Y = \text{pos}) &= \frac{0.9 \times 0.001}{0.2007} \\ &\approx 0.0045 = 0.45\% \end{aligned}$$

Answer 3

Uniform Distribution. X is a collection of N i.i.d $\text{Uniform}(a, b)$ random variables, where $a = 0$ and $b = \frac{1}{2}$.

(a) $X \sim \text{Uniform}(x|a, b) = \frac{1}{b-a} = \frac{1}{0.5-0} = 2$

So, $\mathbf{p}(X = x|a, b) = 2 \quad \forall x \in (0, 0.5)$

(b) This is asking for the probability of $X = 0.00027$ according to the pdf. So, $\mathbf{p}(X = 0.00027 \mid a, b) = 2$, since $0.00027 \in (0, 0.5)$

(c) This is asking for the probability of attaining a single value, which is zero since an integral over a single point evaluates to 0. $\Pr(X = 0.00027|a, b) = 0$.

Answer 4

Poisson Distribution. $X \sim \text{Pois}(\lambda)$, where $\lambda > 0$.

$$\text{Exponential family reference:} \quad h(x) \exp\{\eta T(x) - A(\eta)\} \quad (6)$$

$h(x)$: scaling constant

η : natural parameter

$T(x)$: sufficient statistics

$A(\eta)$: log partition function

(a) Exponential family form:

$$\exp\{x \ln(\lambda) - \lambda - \ln(x!)\}$$

$$\frac{1}{x!} \exp\{x \ln(\lambda) - \lambda\}$$

(b) $T(x) \approx x$

(c) $A(\eta) = \exp\{\eta\}$

(d) Response function: $\lambda = \exp\{\eta\}$, also Link function: $\eta = \ln(\lambda)$

(e) $\mathbb{E}(X) = \frac{\partial A(\eta)}{\partial \eta} = \lambda$

$$\text{Var}(X) = \frac{\partial^2 A(\eta)}{\partial \eta^2} = \lambda$$

(f)

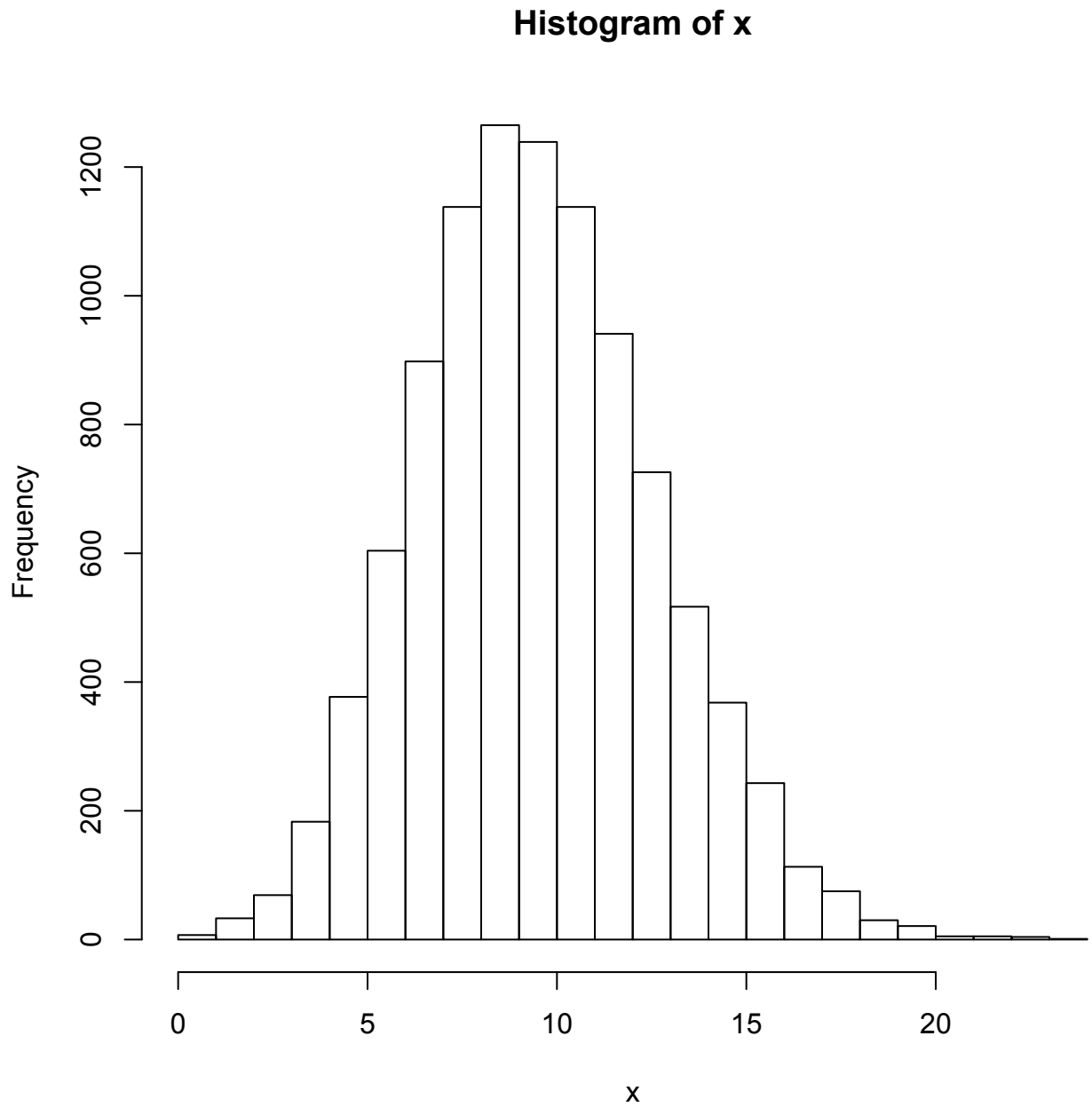
$$\begin{aligned}
\mathbf{p}(X | \lambda) &= \prod_{i=1}^n \frac{\lambda^{X_i} e^{-\lambda}}{X_i!} \\
\ln \mathbf{p}(X | \lambda) &= -n\lambda + \ln \lambda \sum_{i=1}^n X_i - \ln \prod_{i=1}^n X_i! \\
\frac{\partial \ln \mathbf{p}(X | \lambda)}{\partial \lambda} &= -n + \frac{\sum_{i=1}^n X_i}{\lambda} = 0 \\
\Rightarrow \hat{\lambda}_{MLE} &= \frac{\sum_{i=1}^n X_i}{n}
\end{aligned}$$

(g)

$$\begin{aligned}
\mathbf{p}(\lambda | X, \alpha, \beta) &= \mathbf{p}(X | \lambda, \alpha, \beta) \mathbf{p}(\lambda | \alpha, \beta) \\
&= \prod_{i=1}^n \frac{\lambda^{X_i} e^{-\lambda}}{X_i!} \frac{\lambda^{\alpha-1}}{e^{\beta\lambda}} \\
&= \left(\lambda^{\sum_{i=1}^n X_i + \alpha - 1} e^{-(\beta+n)\lambda} \right) / \prod_{i=1}^n X_i! \\
\ln \mathbf{p}(\lambda | X, \alpha, \beta) &= \left(\sum_{i=1}^n X_i + \alpha - 1 \right) \ln \lambda - (\beta + n)\lambda - \ln \prod_{i=1}^n X_i! \\
\frac{\partial \ln \mathbf{p}(\lambda | X, \alpha, \beta)}{\partial \lambda} &= \frac{\sum_{i=1}^n X_i + \alpha - 1}{\lambda} - (\beta + n) = 0 \\
\Rightarrow \hat{\lambda}_{MAP} &= \frac{\sum_{i=1}^n X_i + \alpha - 1}{\beta + n}
\end{aligned}$$

Answer 5

Extension to Problem 4, MAP and MLE simulation.



- (a)
- (b) 10.0171
- (c) 10.0161

(d) 10.0260

(e) 10.0170

(f) We are trying to determine the best approximation of λ given the data X . Since we assumed our data is from the Poisson distribution, whose expectation is λ , and we know that the average of our data is 10.0171 (from part a), we would expect $\lambda \simeq 10.0171$. We assume the prior distribution of λ is the Gamma distribution, whose expectation given parameters α and β is $\frac{\alpha}{\beta}$. Therefore, the simulation that has $\alpha = 10$ and $\beta = 1$ should be a good approximation since $\frac{\alpha}{\beta}$ is the closest one to 10.0171. Also another thing to note is that we have a large sample size of our data ($n = 10,000$ time points), which according to our MAP estimate and initializations for both α and β , decreases the emphasis on our prior distribution. This is why we do not see a drastic difference in the MLE and MAP estimates of λ in parts b-e. In fact, the data X was created from a λ value of 10; so the “uniformed” prior estimate with $\alpha = 1$ and $\beta = 1$ achieved the best result.

Answer 6

Will post the code for this after the next assignment.

Answer 7

MAP Estimator for Logistic Regression.

Since $P(y = 1|x)$ is a Bernoulli random variable (say $y = \{-1, 1\}$) rewrite the following generalized linear model

$$\begin{aligned} f(x) = \beta^T x &= \log \left(\frac{P(y = 1|x)}{P(y = -1|x)} \right) \\ &= \log \left(\frac{P(y = 1|x)}{1 - P(y = 1|x)} \right) \end{aligned}$$

which implies

$$\begin{aligned} P(y = 1|x) &= \frac{1}{1 + \exp(\beta^T x)} \\ P(y = -1|x) &= \frac{1}{1 + \exp(-\beta^T x)} \\ P(y = \pm 1|x) &= \frac{1}{1 + \exp(y \beta^T x)}. \end{aligned}$$

So the likelihood is

$$\prod_{i=1}^n \frac{1}{1 + \exp(y_i \beta^T x_i)}$$

if we use the same prior as in the ridge regression setting we obtain the following MAP estimator

$$\hat{\beta}_{\text{MAP}} = \arg \min_{\beta} \left[n^{-1} \sum_{i=1}^n \log(1 + \exp(-y_i \beta^T x_i)) + \lambda \|\beta\|^2 \right],$$