

## Back to estimators...

So far, we have:

- Identified estimators for common parameters
- Discussed the sampling distributions of estimators
- Introduced ways to judge the “goodness” of an estimator (bias, MSE, etc.)
- Used estimators in confidence intervals, hypothesis testing, etc.

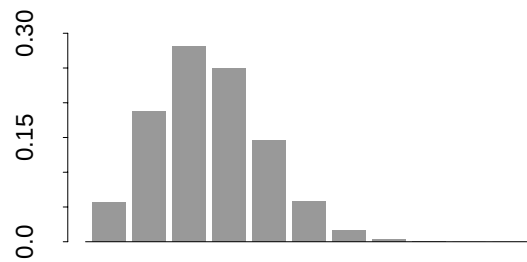
We haven't spent a lot of time on developing/thinking of good estimators for a parameter. What can we do to estimate a parameter that we've never considered before?

## Binomial example

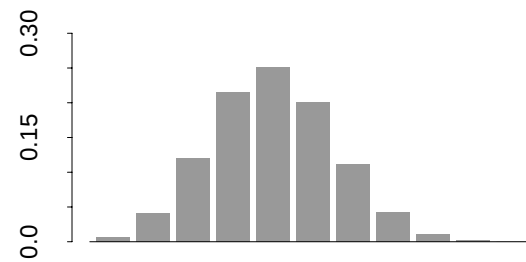
Let's begin with a familiar case.

- We have a RV that we know has a binomial distribution.
- Familiar example: We have 10 people receiving a new treatment for a serious disease? What is the probability of recovery (success)  $p$ ?
- Another familiar example: What proportion of those casting ballots voted for Gore? Try to estimate this based on a random sample of 1000 voters.
- The problem is that we don't know what  $p$  is, now that we have implemented a new treatment regimen.

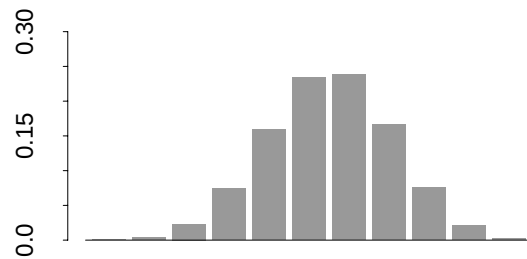
# Binomial probability histograms



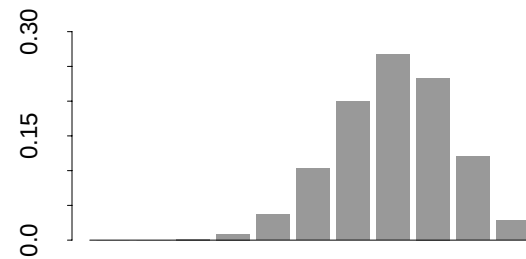
Binomial prob. hist. with  $n=10$  and  $p=0.25$



Binomial prob. hist. with  $n=10$  and  $p=0.4$



Binomial prob. hist. with  $n=10$  and  $p=0.55$



Binomial prob. hist. with  $n=10$  and  $p=0.7$

## Estimating $p$

Pretend for a few moments that we haven't previously discussed using  $\hat{p}$  as an estimator for  $p$ .

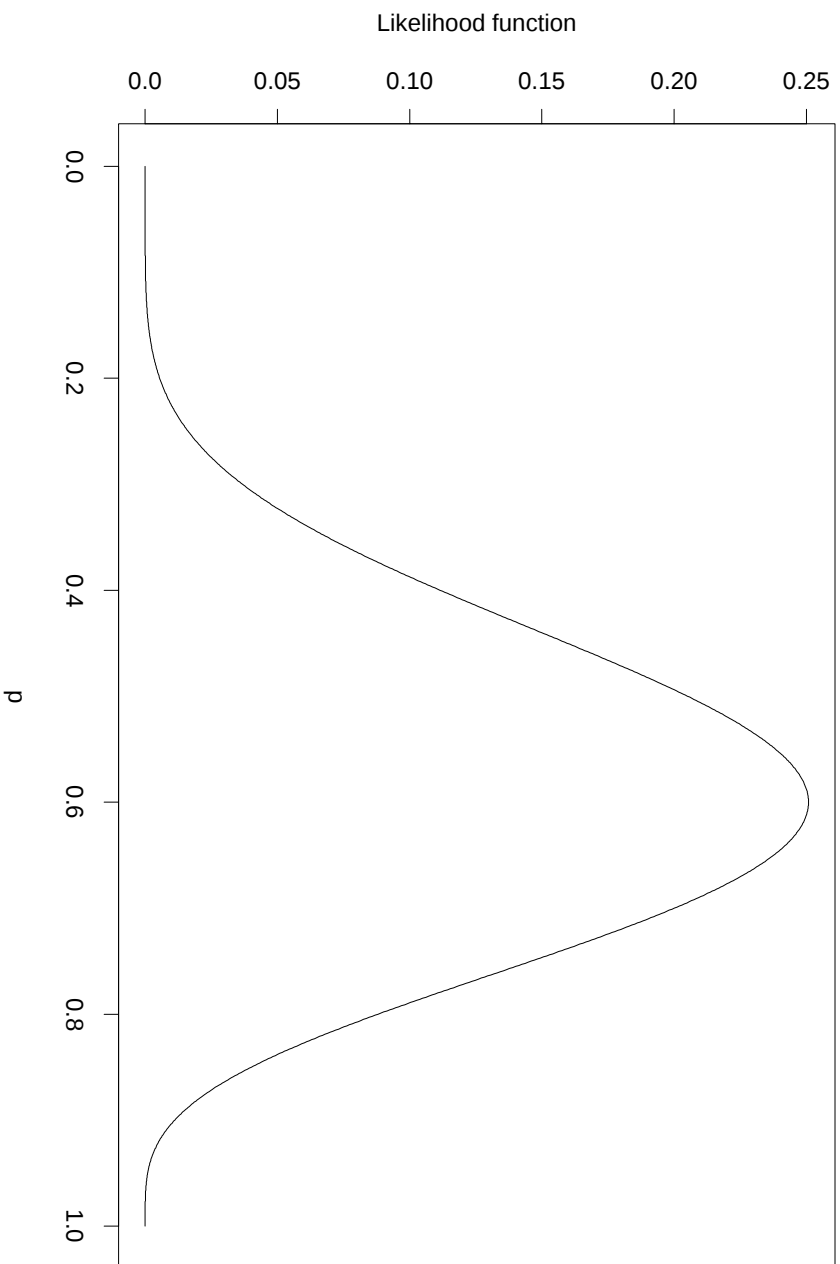
- Our intuition would tell us that the best guess for  $p$  was the proportion of successes in our sample ( $\hat{p}$ ).
- There is also a more mathematical way to approach the problem.
  1. Think about the probability distribution for  $Y$ , the number of successes in your sample of size  $n$ .
  2. Although the parameter  $p$  is unknown, we can substitute into the probability distribution the information we do know:  $Y$  and  $n$ .
  3. Determine what value of  $p$  would give the highest probability of obtaining that sample with  $y$  successes in  $n$  trials.

## Maximum likelihood methodology

Assume that 6 of the 10 patients receiving the new treatment regimen recover.

- Write down the probability distribution for the number of successes in  $n$  trials. (This is the likelihood function.)
- Substitute in the information from your sample.

## Likelihood function for $p$



## Maximum likelihood methodology

- Find the value of  $p$  that maximizes  $L(p)$ .
- This is your *maximum likelihood estimate* for  $p$ .

# Maximum likelihood methodology

Let's work through the more general case and find a formula for the maximum likelihood estimate for  $p$  - a formula which depends just on observed data.

- As before, write down likelihood function.
- How to take the derivative, set it equal to zero, and solve for  $p$ ? Looks messy...

## Maximum likelihood estimate of $p$

- Find the value of  $p$  that maximizes  $L(p)$ .
- This is your *maximum likelihood estimate* for  $p$ .

## Invariance property of MLEs

- Say we have the MLE for parameter 1, but we want to know the MLE for parameter 2, which is a one-to-one function of parameter 1.
- To find the MLE for parameter 2, substitute the MLE for parameter 1 into the function that gives parameter 2.
- If  $t(\theta)$  is a function of  $\theta$  and  $\hat{\theta}$  is the MLE for  $\theta$ , then the MLE for  $t(\theta)$  is given by  $t(\hat{\theta})$ .

## Using the invariance property

We know that for a binomial proportion of successes  $p$ , the MLE is given by  $\hat{p}$ . What is the MLE for the variance of  $Y$ ?

## More about MLEs

- Can be used to help you find an estimator for a parameter that is new to you
- Invariance property of MLEs means that for many functions, you can build a MLE for the function using MLEs that are already known.
- MLEs are not always unbiased. However, in some cases adjusting the MLE by a constant can yield an unbiased estimator with minimum variance.

The method of maximum likelihood is a commonly used procedure in statistics.

## MLE for mean of the normal

Assume that we have a random sample  $x_1, x_2, \dots, x_n$  from a normally distributed population for which the variance  $\sigma^2$  is known. What is the MLE for the mean  $\mu$ ? The probability density function for the normal distribution is:  $f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{1}{2\sigma^2}(x - \mu)^2\right]$ ,  $-\infty < x < \infty$ ,  $-\infty < \mu < \infty$

## MLE for variance of the normal

Assume that we have a random sample  $x_1, x_2, \dots, x_n$  from a normally distributed population for which the mean  $\mu$  is known. What is the MLE for the variance  $\sigma^2$ ? The probability density function for the normal distribution is:

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{1}{2\sigma^2}(x - \mu)^2\right], \quad -\infty < x < \infty, \sigma^2 > 0$$

## Large sample properties of MLEs

- Already know from the invariance property that if  $\hat{\theta}$  is the MLE for  $\theta$ , then for a function of  $\theta$ ,  $t(\theta)$ , the MLE  $\widehat{t(\theta)}$  is  $t(\hat{\theta})$
- Under certain regularity conditions (that apply for the distributions that we will consider),  $\widehat{t(\theta)}$  is a consistent estimator for  $t(\theta)$
- This means that as the sample size  $n$  grows,  $\widehat{t(\theta)}$  tends to get closer to  $t(\theta)$
- For large sample sizes, we also know that  $t(\hat{\theta})$  is normally distributed, so that the following has an approximately standard normal distribution

$$Z = \frac{t(\hat{\theta}) - t(\theta)}{\sqrt{\frac{[\frac{dt(\theta)}{d\theta}]^2}{nE[-\frac{d^2 \ln f(Y|\theta)}{d\theta^2}]}}}$$

## Confidence intervals using MLEs

For large sample sizes, we can obtain a confidence interval for  $t(\theta)$  using the MLE estimator and its large sample properties:

$$t(\hat{\theta}) \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\left[\frac{dt(\theta)}{d\theta}\right]^2}{nE\left[-\frac{d^2 \ln f(Y|\theta)}{d\theta^2}\right]}}$$

However, this formula can still depend on  $\theta$ , so in practice, we can substitute the MLE  $\hat{\theta}$  for  $\theta$ :

$$t(\hat{\theta}) \pm z_{\frac{\alpha}{2}} \sqrt{\left. \frac{\left[\frac{dt(\theta)}{d\theta}\right]^2}{nE\left[-\frac{d^2 \ln f(Y|\theta)}{d\theta^2}\right]} \right|_{\theta=\hat{\theta}}}$$

## CI for normal mean

As before, assume that we have a random sample  $x_1, x_2, \dots, x_n$  from a normally distributed population for which the variance  $\sigma^2$  is known. Show that a  $100(1 - \alpha)\%$  confidence interval for  $\mu$ , when  $\sigma$  is known, is given by  $\bar{x} \pm z_{\frac{\alpha}{2}} \frac{\sigma}{\sqrt{n}}$ , as we learned before.

## Example

Suppose that  $X_1, X_2, \dots, X_n$  form a random sample from a distribution for which the p.d.f.  $f(x|\theta)$  is given below. Also, suppose  $\theta$  is unknown ( $\theta > 0$ ). Find the MLE of  $\theta$ .

$$f(x|\theta) = \begin{cases} \theta x^{\theta-1} & \text{for } 0 < x < 1 \\ 0 & \text{otherwise} \end{cases}$$

## Example continued

Find a  $100(1 - \alpha)\%$  confidence interval for the  $\theta^2$  (reference the distribution given in the last slide).

$$t(\hat{\theta}) \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\left[\frac{dt(\theta)}{d\theta}\right]^2}{nE\left[-\frac{d^2 \ln f(Y|\theta)}{d\theta^2}\right] \Big|_{\theta=\hat{\theta}}}}$$