

INTRODUCING LINEAR REGRESSION MODELS

- **Response** or **Dependent** variable y
- **Predictor** or **Independent** variable x
- Model with error: for $i = 1, \dots, n$,

$$y_i = \alpha + \beta x_i + \varepsilon_i$$

- ε_i : independent errors (sampling, measurement, lack of fit)
- Typically $\varepsilon \sim N(0, \sigma^2)$
- **Analysis and inference:**
 - Estimate parameters $(\alpha, \beta, \sigma^2)$
 - Assess model fit — adequate? good? if inadequate, how?
 - Explore implications: $\beta, \beta x$
 - Predict new (“future”) responses at new x_{n+1}, \dots

BIG PICTURE:

- Understanding variability in y as a function of x
- Exploring $p(y|x)$ as a function of x
- One aspect: Regression function $E(y|x)$ as x varies
- Special case: normal, linear in mean
 - Other cases: binomial y , success prob depends on x
 - e.g., logistic regression, dose-response models
- How much variability does x explain?
- Normal models: Variance measures “variability”

- Observational studies versus Designed studies
 - “Random” x versus “Controlled” x
- Bivariate data (y_i, x_i) , but take x_i fixed
- “Special” status of response variable
- Several or many predictor variables

SAMPLE SUMMARY STATISTICS

- Sample means \bar{x}, \bar{y}
- Sample variances s_x^2, s_y^2

$$s_y^2 = S_{yy}/(n - 1), \quad s_x^2 = S_{xx}/(n - 1)$$

- sample covariance

$$s_{xy} = S_{xy}/(n - 1)$$

where the “Sums of Squares” are:

- $S_{yy} = \sum_{i=1}^n (y_i - \bar{y})^2$ – “Total Variation in response”
- $S_{xx} = \sum_{i=1}^n (x_i - \bar{x})^2$
- $S_{xy} = \sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})$

Standardized scale for covariance:

SAMPLE CORRELATION:

$$r = \frac{s_{xy}}{s_x s_y}$$

$-1 < r < 1$, measure of dependence

for a single predictor, $r^2 = R^2$

SQUARED ERRORS AND “FIT” OF CHOSEN LINES

Measurement error version of model: $y_i = \alpha + \beta x_i + \varepsilon_i$

For any chosen α, β ,

$$Q(\alpha, \beta) = \sum_{i=1}^n \varepsilon_i^2 = \sum_{i=1}^n (y_i - \alpha - \beta x_i)^2$$

measures “fit” of chosen line $\alpha + \beta x$ to response data

LEAST SQUARES LINE:

- Choose $\hat{\alpha}, \hat{\beta}$ to *minimise* $Q(\alpha, \beta)$
- Geometric interpretation
- *Least squares estimates* (LSE) $\hat{\alpha}, \hat{\beta}$
- (Venerable/ad-hoc) “principal” of **least squares estimation**
- Least squares fit is also the MLE

LEAST SQUARES ESTIMATES

FACTS:

$$\hat{\beta} = \frac{s_{xy}}{s_x^2}, \quad \hat{\alpha} = \bar{y} - \hat{\beta}\bar{x}$$

Or

$$\hat{\beta} = r \left(\frac{s_y}{s_x} \right)$$

$\hat{\beta}$ is correlation coefficient r , corrected for relative scales of $y : x$
so that the units of the “fitted values” $\hat{\beta}x$ are on scale of y

Note also

$$\hat{\beta} = \frac{S_{xy}}{S_{xx}}$$

of use in theoretical derivations

R^2 measure of model fit:

Simplest model: $\beta = \hat{\beta} = 0$ so y_i are a normal random sample

$$\hat{\alpha} = \bar{y}, \quad Q(\bar{y}, 0) = S_{yy} = \text{total sum of squares}$$

Any other model fit: **Residual Sum of Squares** $Q(\hat{\alpha}, \hat{\beta})$

DEFINE: $R^2 = 1 - Q(\hat{\alpha}, \hat{\beta})/S_{yy}$

– proportion of variation “explained” by model –

FACT: $R^2 = r^2$

- “Multiple regression” generalisation later
- Higher %variation explained is better: Higher correlation
- Measures linear correlation
 - not general dependence
 - not causation

EXAMINING MODEL FIT

- **Fitted values** $\hat{y}_i = \hat{\alpha} + \hat{\beta}x_i$
- **Residuals** $\hat{\varepsilon}_i = y_i - \hat{y}_i$ estimates of ε_i
- **Residual sum of squares** $Q(\hat{\alpha}, \hat{\beta}) = \sum_{i=1}^n \hat{\varepsilon}_i^2$
 - measures remaining/residual variation in response data –
- **Residual sample variance:**

$$s_{Y|X}^2 = \frac{RSS}{n-2} = \sum_{i=1}^n \frac{\hat{\varepsilon}_i^2}{n-2}$$

- $s_{Y|X}^2$ is a point estimate of σ^2 from fitted model
 - note: $n - 2$ degrees of freedom, not $n - 1$
 - “lose” 2 degrees of freedom for estimation of α, β

Conjugate Priors for Regression

- Normal is conjugate for α and β
- Inverse-Gamma is conjugate for σ^2

Common re-parameterization: Precision $\phi = \frac{1}{\sigma^2}$

- Gamma is conjugate for ϕ

A Reference Prior for Regression

Take limits of conjugate prior as the prior variance goes to infinity (information goes to zero)

- $\lim_{t \rightarrow \infty} N(0, t)$ is proportional to a **constant**
- $\lim_{\substack{a \rightarrow 0, \\ b \rightarrow 0}} \Gamma(a, b)$ or $\Gamma^{-1}(a, b)$ is proportional to the **inverse**

$$p(\alpha, \beta, \sigma^2) \propto \frac{1}{\sigma^2}$$

$$p(\alpha, \beta, \phi) \propto \frac{1}{\phi}$$

THEORY FOR INFERENCE: REFERENCE POSTERIOR

Some key aspects of the reference posterior for $(\alpha, \beta, \sigma^2)$:

- (marginal) posterior for β is t distribution with $n - 2$ df.

$$t_{n-2}(\hat{\beta}, s_{Y|X}^2 v_{\beta}^2)$$

where $v_{\beta}^2 = 1/S_{xx}$

- $s_{Y|X}^2$ is a posterior estimate of σ^2 – residual variance

Key to assessing *significance* of regression fit and measuring the “explanatory power” of chosen predictor x

Intervals (HPD or equal-tailed):

$$\hat{\beta} \pm (s v_{\beta}) t_{p/2}$$

where $t_{p/2}$ is $100(p/2)\%$ quantile of standard t_{n-2}

“TESTING” SIGNIFICANCE OF THE REGRESSION FIT

Question: How probable is $\beta = 0$ under the posterior?

Answer:

- Compute posterior probability on β values with lower posterior density than $\beta = 0$
- “Measures” probability of β “less likely” than $\beta = 0$
- Informal “test” of significance –
Probability in tails = **significance level** = (Bayesian) ***p*–value**
- Symmetric posterior density: double one tail area
- Classical testing terminology:
“The regression on x is significant at the 5% level (or 1%, etc) if the *p*–value is smaller than 0.05 (or 0.01, etc)”

R/S-Plus: $2*(1-pt(abs(t), n-2))$ where $t = \hat{\beta} / s_{Y|X} v_{\beta}$ –
standardized T Statistic

F TESTS, ANOVA AND DEVIANCES

F test of regression fit:

Theory: If $t \sim T_k(0, 1)$ then $F = t^2 \sim F_{1, n-2}$

So

- $p\text{-value} = Pr(F \geq f_{obs})$
- $f_{obs} = \hat{\beta}^2 / s_{Y|X}^2 v_{\beta}^2$
- T and F tests are equivalent: same $p\text{-value}$
- S-Plus output: quotes T values, $p\text{-values}$ in coefficient table
and F test result

F TESTS, ANOVA AND DEVIANCES

Deviances = Sums of squares:

Deviance decomposition:

$$S_{yy} = Q(\hat{\alpha}, \hat{\beta}) + \hat{\beta}^2 / v_{\beta}^2$$

- Total deviance $S_{yy} = \sum_{i=1}^n (y_i - \bar{y})^2$
- Residual deviance $Q(\hat{\alpha}, \hat{\beta}) = \sum_{i=1}^n (y_i - \hat{y}_i)^2$
- *Fitted or explained deviance:* $\hat{\beta}^2 / v_{\beta}^2$
– here equal to $s_{Y|X}^2 f_{obs}$ –
- Large deviance explained \equiv large $F \equiv$ significant regression
- ANOVA: analysis of variance (deviance)

PREDICTION FROM FITTED MODEL

Question: What is the posterior predictive distribution for a new case,

$$y_{n+1} = \alpha + \beta x_{n+1} + \varepsilon_{n+1}$$

Answer: Also a Student t distribution with $n - 2$ df.

$$y_{n+1} \sim T_{n-2}(\hat{y}, s_{Y|X}^2 v_y^2)$$

- Mean is $\hat{y} = \hat{\alpha} + \hat{\beta}x_{n+1}$
- Spread: $s_{Y|X}^2 v_y^2 = s_{Y|X}^2 + s_{Y|X}^2 w^2 \dots$
 - $s_{Y|X}^2 w^2$ – posterior uncertainty about $\alpha + \beta x_{n+1}$
depends on x_{n+1} , spread is higher for x_{n+1} far from \bar{x}
 - additional variability $+s_{Y|X}^2$ due to ε_{n+1} , estimating σ^2 by $s_{Y|X}^2$
- Can use S-Plus function `predict.lm()`

Model fit assessment/implications: Explore predictive distributions

Residual analysis: Graphical exploration of fitted residuals $\hat{\varepsilon}_i$

- Standardize: $r_i = \hat{\varepsilon}_i / \sqrt{\text{var}(\hat{\varepsilon}_i)}$
- Check normality assumption
- Treat $\hat{\varepsilon}_i$ as “new data” – look at structure, other predictors

Other predictors?