

Midterm Examination II

STA 711: Probability & Measure Theory

Wednesday, 2016 Nov 16, 1:25 – 2:40pm

This is a closed-book exam. You may use a single sheet of prepared notes, if you wish, but you may not share materials.

If a question seems ambiguous or confusing, *please* ask me to clarify it. Unless a problem states otherwise, you must **show** your **work**. There are blank worksheets at the end of the test if you need more room for this, and also a pdf/pmf sheet.

It is to your advantage to write your solutions as clearly as possible, and to box answers I might not find.

For full credit, answers must be given in **closed form** with no unevaluated sums, integrals, maxima, unreduced fractions. Wherever possible, **Simplify**.

Good luck!

Print Name Clearly: _____

1.	/20
2.	/20
3.	/20
4.	/20
5.	/20
Total:	/100

Problem 1: Let $\{X_n\}_{n \in \mathbb{N}} \stackrel{\text{ind}}{\sim} \text{Ex}(n^2)$ be independent exponentially-distributed random variables that satisfy $\mathbb{P}[X_n > x] = \exp(-n^2x)$ for $x > 0$, and set $S_n := \sum_{m=1}^n X_m$.

a) (5) Does S_n converge in L_1 ? Yes No Why?

b) (5) Does S_n converge in L_2 ? Yes No Why?

c) (5) Does X_n converge in L_∞ ? Yes No Why?

d) (5) Does X_n converge almost-surely? Yes No Why?

e) (XC) Does S_n converge almost-surely? Yes No Why?

Problem 2: Let $(\Omega, \mathcal{F}, \mathbf{P})$ be $\Omega = (0, 1]$ with the Borel sets and Lebesgue measure, and let $\{U_n\} \stackrel{\text{iid}}{\sim} \text{Un}(0, 1)$ be iid standard uniform random variables on $(\Omega, \mathcal{F}, \mathbf{P})$. In each part below, indicate in which (if any) sense(s) the sequence $\{X_n\}$ converges to zero. No explanations are necessary.

a) (5) $X_n := 2^n \mathbf{1}_{\{0 < \omega \leq 4^{-n}\}}$: *a.s.* *pr.* L_1 L_2 L_∞

b) (5) $X_n := \sqrt{n} \mathbf{1}_{\{U_n < 1/n^2\}}$: *a.s.* *pr.* L_1 L_2 L_∞

c) (5) $X_n := U_n^{-1} \mathbf{1}_{\{U_n < 1/n\}}$: *a.s.* *pr.* L_1 L_2 L_∞

d) (5) $X_n := 1/(nU_n)$: *a.s.* *pr.* L_1 L_2 L_∞

Problem 3: Let $\{U_i\} \stackrel{\text{iid}}{\sim} \text{Un}(0, 1)$ be iid standard Uniform random variables, and set $X_i := U_i^{-1/4}$. Show your work in finding:

a) (4) Show that the pdf for X_i is $f(x) = 4x^{-5}\mathbf{1}_{\{x>1\}}$:

b) (4) Find $\|X_i\|_p$ for every $0 < p \leq \infty$.
 $\|X_i\|_p =$

c) (4) Set $S_n := \sum_{i \leq n} X_i$. Find sequences a_n and b_n , if possible, so that

$$\frac{S_n - a_n}{b_n} \Rightarrow \text{No}(0, 1)$$

has approximately a standard Normal distribution for large n . Justify.

$$a_n = \qquad b_n =$$

Problem 3 (cont'd): Still $X_i := U_i^{-1/4}$ for $\{U_i\} \stackrel{\text{iid}}{\sim} \text{Un}(0, 1)$.

d) (4) Set $T_n := \sum_{i \leq n} X_i^2$. Find sequences c_n and d_n , if possible, so that

$$\frac{T_n - c_n}{d_n} \Rightarrow \text{No}(0, 1)$$

has approximately a standard Normal distribution for large n . Justify.

$c_n =$

$d_n =$

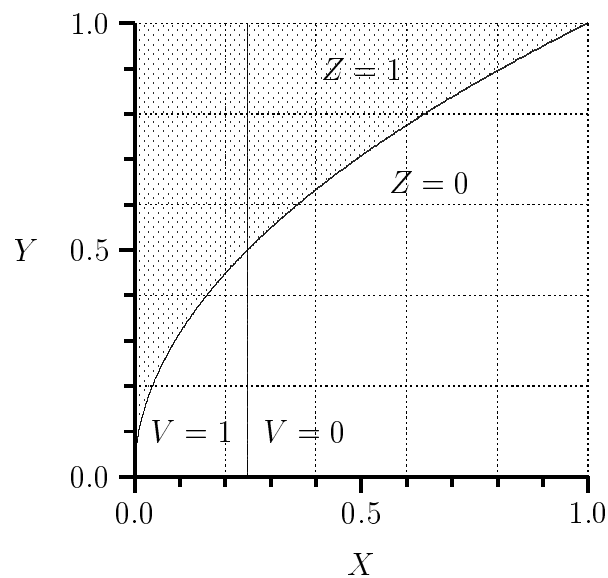
e) (4) Find and justify the indicated limits:

$$\lim_{n \rightarrow \infty} \frac{S_n}{n} =$$

$$\lim_{n \rightarrow \infty} \frac{T_n}{n} =$$

Problem 4: The random variables $X, Y \stackrel{\text{iid}}{\sim} \text{Un}(0, 1]$, while

$$V := \mathbf{1}_{\{X \leq 0.25\}} \quad Z := \mathbf{1}_{\{X \leq Y^2\}}$$



(and **simplify!**):

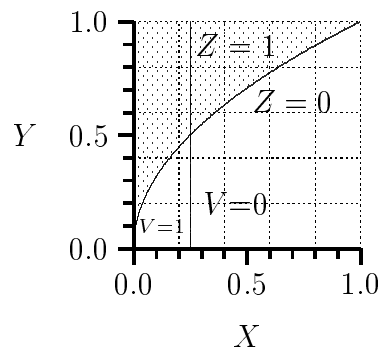
a) (4) Fill in¹ the following table of probabilities:

	$V = 0$	$V = 1$
$Z = 0$		
$Z = 1$		

b) (6) $E[Z | V] =$

¹Remember: Simplify, with no unreduced fractions

Problem 4 (cont'd): Still $X, Y \stackrel{\text{iid}}{\sim} \text{Un}(0, 1]$, $V := \mathbf{1}_{\{X \leq 0.25\}}$, $Z := \mathbf{1}_{\{X \leq Y^2\}}$.



Find:

c) (5) $E[X | V] =$

d) (5) $E[Z | X] =$

Problem 5: True or false? Circle one; each answer is worth 2 points. No explanations are needed, but you can give one if you think a question seems ambiguous or tricky. All random variables are real on some $(\Omega, \mathcal{F}, \mathbf{P})$.

- a) T F If $\{X, Y, Z\}$ are iid & positive then $W := \frac{Z}{X+Y+Z} \in L_1$ and $\mathbf{E}W = 1/3$.
- b) T F If X has a continuous pdf with $f(0) > 0$ then $\frac{1}{X} \notin L_1$.
- c) T F If $\{X_n\}$ are iid and L_∞ with mean $\mu = \mathbf{E}X_n$ then for some $c > 0$ and all $\epsilon > 0$ and $n \in \mathbb{N}$, $\mathbf{P}[(\bar{X}_n - \mu) > \epsilon] \leq \exp(-cn\epsilon^2)$.
- d) T F If $\|X_n\|_2 \rightarrow 0$ then also $\mathbf{E}\sqrt{|X_n|} \rightarrow 0$.
- e) T F For the Cauchy distribution, $\mathbf{E}[\exp(i\omega X)]$ is infinite for all $\omega \in \mathbb{R}$ except for $\omega = 0$ because the Cauchy pdf has heavy tails.
- f) T F If $\mathcal{G} \subset \mathcal{F}$ and $Y = \mathbf{E}[X | \mathcal{G}]$ with $X \in L_1$, then $\mathbf{E}[X] = \mathbf{E}[Y]$.
- g) T F If $\mathcal{G} \subset \mathcal{F}$ and $Y = \mathbf{E}[X | \mathcal{G}]$ with $X \in L_2$, then $\mathbf{V}[X] = \mathbf{V}[Y]$.
- h) T F Every ch.f. $\phi(\omega) = \mathbf{E}[e^{i\omega X}]$ is continuously differentiable.
- i) T F X and $(-X)$ have the same distribution if and only if $(\forall \omega) \phi_X(\omega) = \phi_X(-\omega)$.
- j) T F If $\{X_i\}$ are iid with ch.f. $\phi(\omega)$, then $\prod_{i=1}^n X_i$ has ch.f. $\phi(\omega)^n$.

Name: _____ STA 711: Prob & Meas Theory

Blank Worksheet

Name: _____ STA 711: Prob & Meas Theory

Another Blank Worksheet

Name	Notation	pdf/pmf	Range	Mean μ	Variance σ^2
Beta	$\text{Be}(\alpha, \beta)$	$f(x) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}$	$x \in (0, 1)$	$\frac{\alpha}{\alpha+\beta}$	$\frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$
Binomial	$\text{Bi}(n, p)$	$f(x) = \binom{n}{x} p^x q^{(n-x)}$	$x \in 0, \dots, n$	np	$npq \quad (q = 1 - p)$
Exponential	$\text{Ex}(\lambda)$	$f(x) = \lambda e^{-\lambda x}$	$x \in \mathbb{R}_+$	$1/\lambda$	$1/\lambda^2$
Gamma	$\text{Ga}(\alpha, \lambda)$	$f(x) = \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x}$	$x \in \mathbb{R}_+$	α/λ	α/λ^2
Geometric	$\text{Ge}(p)$	$f(x) = p q^x$ $f(y) = p q^{y-1}$	$x \in \mathbb{Z}_+$ $y \in \{1, \dots\}$	q/p $1/p$	$q/p^2 \quad (q = 1 - p)$ $q/p^2 \quad (y = x + 1)$
HyperGeo.	$\text{HG}(n, A, B)$	$f(x) = \frac{\binom{A}{x} \binom{B}{n-x}}{\binom{A+B}{n}}$	$x \in 0, \dots, n$	nP	$nP(1-P) \frac{N-n}{N-1} \quad (P = \frac{A}{A+B})$
Logistic	$\text{Lo}(\mu, \beta)$	$f(x) = \frac{e^{-(x-\mu)/\beta}}{\beta[1+e^{-(x-\mu)/\beta}]^2}$	$x \in \mathbb{R}$	μ	$\pi^2 \beta^2 / 3$
Log Normal	$\text{LN}(\mu, \sigma^2)$	$f(x) = \frac{1}{x\sqrt{2\pi\sigma^2}} e^{-(\log x - \mu)^2 / 2\sigma^2}$	$x \in \mathbb{R}_+$	$e^{\mu + \sigma^2/2}$	$e^{2\mu + \sigma^2} (e^{\sigma^2} - 1)$
Neg. Binom.	$\text{NB}(\alpha, p)$	$f(x) = \binom{x+\alpha-1}{x} p^\alpha q^x$ $f(y) = \binom{y-1}{y-\alpha} p^\alpha q^{y-\alpha}$	$x \in \mathbb{Z}_+$ $y \in \{\alpha, \dots\}$	$\alpha q/p$ α/p	$\alpha q/p^2 \quad (q = 1 - p)$ $\alpha q/p^2 \quad (y = x + \alpha)$
Normal	$\text{No}(\mu, \sigma^2)$	$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2 / 2\sigma^2}$	$x \in \mathbb{R}$	μ	σ^2
Pareto	$\text{Pa}(\alpha, \epsilon)$	$f(x) = (\alpha/\epsilon)(1+x/\epsilon)^{-\alpha-1}$ $f(y) = \alpha \epsilon^\alpha / y^{\alpha+1}$	$x \in \mathbb{R}_+$ $y \in (\epsilon, \infty)$	$\frac{\epsilon}{\alpha-1}^*$ $\frac{\epsilon \alpha}{\alpha-1}^*$	$\frac{\epsilon^2 \alpha}{(\alpha-1)^2(\alpha-2)}^*$ $\frac{\epsilon^2 \alpha}{(\alpha-1)^2(\alpha-2)}^* \quad (y = x + \epsilon)$
Poisson	$\text{Po}(\lambda)$	$f(x) = \frac{\lambda^x}{x!} e^{-\lambda}$	$x \in \mathbb{Z}_+$	λ	λ
Snedecor F	$F(\nu_1, \nu_2)$	$f(x) = \frac{\Gamma(\frac{\nu_1+\nu_2}{2}) (\nu_1/\nu_2)^{\nu_1/2}}{\Gamma(\frac{\nu_1}{2})\Gamma(\frac{\nu_2}{2})} \times x^{\frac{\nu_1-2}{2}} \left[1 + \frac{\nu_1}{\nu_2} x\right]^{-\frac{\nu_1+\nu_2}{2}}$	$x \in \mathbb{R}_+$	$\frac{\nu_2}{\nu_2-2}^*$	$\left(\frac{\nu_2}{\nu_2-2}\right)^2 \frac{2(\nu_1+\nu_2-2)^*}{\nu_1(\nu_2-4)}$
Student t	$t(\nu)$	$f(x) = \frac{\Gamma(\frac{\nu+1}{2})}{\Gamma(\frac{\nu}{2})\sqrt{\pi\nu}} [1+x^2/\nu]^{-(\nu+1)/2}$	$x \in \mathbb{R}$	0^*	$\nu/(\nu-2)^*$
Uniform	$\text{Un}(a, b)$	$f(x) = \frac{1}{b-a}$	$x \in (a, b)$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$
Weibull	$\text{We}(\alpha, \beta)$	$f(x) = \alpha\beta x^{\alpha-1} e^{-\beta x^\alpha}$	$x \in \mathbb{R}_+$	$\frac{\Gamma(1+\alpha^{-1})}{\beta^{1/\alpha}}$	$\frac{\Gamma(1+2/\alpha) - \Gamma^2(1+1/\alpha)}{\beta^{2/\alpha}}$