

Example : Chris, an Environmental Engineer, wants to know the Hg concentration of fish in the Eno river. She takes a rod and reel to a spot in Duke Forest to try to catch some fish. From past experience she knows that the average number of fishes she catches per hour is $\lambda = 6$. Let y denote the number of fish she catches in one hour of fishing. What is the distribution for y ?

Divide the time interval (an hour) into n subintervals (small enough, say, a second) and assume

- the probability of catching more than one fish in a subinterval is (almost) zero
- the probability of catching a fish is the same for all subintervals
- the number of fish caught in each subinterval is independent of other subintervals

Then y has *approximately* a binomial distribution with mean $\mu = np = \lambda$ (**Why?**), so the probability of catching a fish in one subinterval must be $p = \lambda/n$ and the

probability distribution for y must be approximately

$$\begin{aligned} p(y) &\approx \frac{n!}{y!(n-y)!} (\lambda/n)^y (1 - \lambda/n)^{n-y} \\ &= \frac{n(n-1)(n-2)\dots(n-y+1)}{y!} \frac{\lambda^y}{n^y} (1 - \lambda/n)^{n-y} \\ &= \frac{\lambda^y}{y!} \left(\frac{n}{n}\right) \left(\frac{n-1}{n}\right) \dots \left(\frac{n-y+1}{n}\right) \frac{(1 - \lambda/n)^n}{(1 - \lambda/n)^y} \\ &\rightarrow \frac{\lambda^y}{y!} e^{-\lambda} \quad \text{for } y = 0, 1, \dots \text{ as } n \rightarrow \infty, \end{aligned}$$

the **Poisson distribution** with mean parameter $\lambda = 6$.

Let y_t denote the number of fishes she catches in t hours.

Q: What's the distribution for y_t ?

A: Poisson with mean parameter λt .

Q: What's the probability she catches exactly one fish in twenty minutes?

*A: $2e^{-2} \approx 0.2707$ ($\lambda t = 6 * 1/3 = 2$; remember units!).*

Let z_1 denote the waiting time (in hours) before Chris catches the first fish.

What is the distribution of z_1 ?

- Note that the waiting time is shorter than any number t **if and only if** Chris catches at least one fish in the first t hours... so

$$\begin{aligned} F(t) &= P(z_1 \leq t) \\ &= P(y_t \geq 1) \\ &= 1 - P(y_t = 0) \\ &= 1 - \frac{(\lambda t)^0}{0!} e^{-\lambda t} \\ &= 1 - e^{-\lambda t}, \quad t \geq 0. \end{aligned}$$

The density function is:

$$\begin{aligned} f(t) &= F'(dt) \\ &= \lambda e^{-\lambda t}, \quad t \geq 0. \end{aligned}$$

This is called the **exponential distribution**.

Let z_α denote the waiting time (in hours) before Chris catches the α^{th} fish, for $\alpha = 1, 2, \dots$

What is the distribution of z_α ?

- Note that the waiting time does not exceed any number t **if and only if** Chris catches **at least** α fish in the first t hours... so

$$\begin{aligned} F(t) &= P(y_t \geq \alpha) \\ &= \sum_{j=\alpha}^{\infty} \frac{(\lambda t)^j}{j!} e^{-\lambda t}, \quad t \geq 0 \\ &= \text{gammainc}(1 * t, a) \text{ in MatLab.} \end{aligned}$$

Differentiating term-by-term leads to some cancellation (try it!), with:

$$\begin{aligned} f(t) &= F'(t) \\ &= \frac{\lambda^\alpha t^{\alpha-1}}{\Gamma(\alpha)} e^{-\lambda t}, \quad t \geq 0. \end{aligned}$$

This is called the **gamma distribution**.

Aside on the Gamma Function

The Gamma distribution is named after the function $\Gamma(\alpha)$ in the numerator,

$$\Gamma(\alpha) = \int_0^{\infty} t^{\alpha-1} e^{-t} dt$$

$$= \text{gamma}(a) \text{ in MatLab}$$

$$\Gamma(\alpha) = \int_0^{\infty} (\alpha-1)t^{\alpha-2} e^{-t} dt \text{ (integrate by pts)}$$

$$= (\alpha-1)\Gamma(\alpha-1)$$

$$= (\alpha-1)(\alpha-2)\Gamma(\alpha-2)$$

$$= (\alpha-1)(\alpha-2)\cdots 3 \cdot 2 \cdot 1 \cdot \Gamma(1)$$

$$= (\alpha-1)! \quad \text{for positive integers } \alpha$$

$$\Gamma(1) = 0! = 1 = 10^0,$$

$$\Gamma(10) = 9! = 362880 \approx 3.6 \times 10^5,$$

$$\Gamma(100) \approx 10^{156},$$

$$\Gamma(1000) \approx 10^{2565},$$

$$\Gamma(\alpha) \approx \alpha^{\alpha} e^{-\alpha} \sqrt{2\pi/\alpha} \text{ (Stirling's Approx)}$$

$$\ln \Gamma(\alpha) = \text{gamma}(\ln(a)) \text{ in MatLab}$$

(otherwise it gets HUGE!)

Exponential and Gamma Distributions

The Exponential distribution is the special case of the Gamma distribution, with $\alpha = 1$. Both distributions are sometimes parametrized by $\beta = 1/\lambda$, the average time-per-event, and sometimes by $\lambda = 1/\beta$, the average rate of events.

The mean and variance for the **exponential distribution** are

$$\begin{aligned}\mu &= \int_0^{\infty} t \lambda e^{-\lambda t} dt \\ &= \lambda^{-1} = \beta \quad (\text{Why?}) \\ \sigma^2 &= \int_0^{\infty} (t - \lambda^{-1})^2 \lambda e^{-\lambda t} dt \\ &= \lambda^{-2} = \beta^2.\end{aligned}$$

and, for the **gamma distribution**,

$$\begin{aligned}\mu &= \int_0^\infty t \frac{\lambda^\alpha t^{\alpha-1}}{\Gamma(\alpha)} e^{-\lambda t} dt \\ &= \alpha \lambda^{-1} = \alpha \beta \quad (\text{Why?}) \\ \sigma^2 &= \int_0^\infty (t - \lambda^{-1})^2 \frac{\lambda^\alpha t^{\alpha-1}}{\Gamma(\alpha)} e^{-\lambda t} dt \\ &= \alpha \lambda^{-2} = \alpha \beta^2.\end{aligned}$$

Recall that the **Standard Normal Distribution** has mean $\mu = 0$, variance $\sigma^2 = 1$, and density function

$$f(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2}$$

It turns out (we'll see why later) that the *square* of a standard normal $y = z^2$ has a gamma distribution with $\alpha = 1/2$ and $\beta = 2$. This turns out to be important in statistics when we consider the sum of squared errors in regression.

Chi-Square Distribution

- A **chi-square (χ^2) random variable** is a gamma-type random variable with $\alpha = \nu/2$ and $\beta = 2$ (or $\lambda = 1/2$)

$$f(\chi^2) = c (\chi^2)^{(\nu/2)-1} e^{-\chi^2/2} \quad \chi^2 \geq 0$$

where

$$c = \frac{1}{2^{\nu/2} \Gamma(\frac{\nu}{2})}$$

- Mean and Variance

$$\mu = \nu \quad \sigma^2 = 2\nu$$

- The parameter ν is called the **number of degrees of freedom** for the chi-square distribution.
- Important in statistics: $\chi^2 = z_1^2 + \dots + z_\nu^2$, each z_i a standard normal distribution.

Failure Time Distributions

- The **reliability** of a product is the probability that the product will meet a set of specifications for a given period of time.
- The **failure time** of a product is the length of time that the product performs according to specifications.
- The **failure time distribution** for a product is the density function of the failure time t , denoted by $f(t)$.
- Called **survival time** in medical applications (e.g. clinical trials).

- The probability that the product (or subject) will fail before any fixed time t_0 is

$$F(t_0) = \int_0^{t_0} f(t) dt$$

- Call a product *reliable* if it survives until time t_0 . Then the **reliability** of the product (i.e., the probability that it will survive at least time t_0) is

$$R(t_0) = 1 - F(t_0),$$

also called the **Survival function** $S(t_0)$.

Hazard Rates

Given that a product has lasted at least time t , what is the probability that it will fail in the next short time period of length dt ? Does this failure rate increase over time, decrease, or stay the same? What does it look like for human lifetimes? For computer chips? For automobile bearings?

- Consider the two events

A : Item fails in the interval $(t, t + dt]$

B : Item survives until t

$$\begin{aligned} P(A|B) &= \frac{P(A \cap B)}{P(B)} = \frac{P(A)}{P(B)} \\ &\approx \frac{f(t)dt}{1 - F(t)} = \frac{f(t)dt}{R(t)} \end{aligned}$$

- The **hazard rate** for a product is defined to be

$$h(t) = \frac{f(t)}{1 - F(t)} = \frac{f(t)}{R(t)}$$

where $f(t)$ is the density function of the product's failure time distribution.

- Note $h(t) = -[\ln(1 - F(t))]'$ (chain rule), so

$$F(t) = 1 - e^{-\int_0^t h(s) ds}, \quad t > 0$$

and we may recover the distribution function from the hazard function.

- Example: If $h(t) \equiv \lambda$ for all $t > 0$, then

$$F(t) = 1 - e^{-\int_0^t \lambda ds} = 1 - e^{-\lambda t}, \quad t > 0$$

the Exponential Distribution with rate λ (or mean $\beta = 1/\lambda$).

- **Example 17.1** The exponential distribution is often used in industry to model the failure time distribution of a product.

Find the hazard rate for the exponential distribution.

$$F(t) = \int_{-\infty}^t f(y) dy = \int_0^t \lambda e^{-\lambda y} dy = 1 - e^{-\lambda t}, \quad t > 0.$$

Then the hazard rate is

$$h(t) = \frac{f(t)}{1 - F(t)} = \frac{\lambda e^{-\lambda t}}{1 - (1 - e^{-\lambda t})} = \lambda.$$

- Constant hazard rate implies that the product does not wear out, that is, it is just as likely to survive one more hour at any age (**Lack of Memory Property**)
- In some applications the failure rate may **increase** (why?) or **decrease** (how is *that* possible?).
- The **Weibull distribution** allows $h(t) \propto t^{\alpha-1}$ for $\alpha > 1$ (increasing), $\alpha < 1$ (decreasing), or $\alpha = 1$ (exponential).

Weibull Distribution

- Density Function

$$f(y) = \alpha \lambda y^{\alpha-1} e^{-\lambda y^\alpha}, \quad y > 0$$

Two parameters: $\alpha, \lambda > 0$

- Mean and Variance

$$\mu = \lambda^{-1/\alpha} \Gamma\left(\frac{\alpha+1}{\alpha}\right)$$
$$\sigma^2 = \lambda^{-2/\alpha} \left[\Gamma\left(\frac{\alpha+2}{\alpha}\right) - \Gamma\left(\frac{\alpha+1}{\alpha}\right)^2 \right]$$

- Reliability and Hazard Rate

$$\begin{aligned} \text{CDF } F(t) &= \int_0^t \alpha \lambda y^{\alpha-1} e^{-\lambda y^\alpha} dy \\ &= 1 - e^{-\lambda t^\alpha}, \quad t > 0 \end{aligned}$$

$$\text{Reliability } R(t) = 1 - F(t) = e^{-\lambda t^\alpha}$$

$$\text{Hazard Rate } h(t) = \frac{f(t)}{R(t)} = \alpha \lambda t^{\alpha-1}$$

Hazard Rate for Weibull Distribution

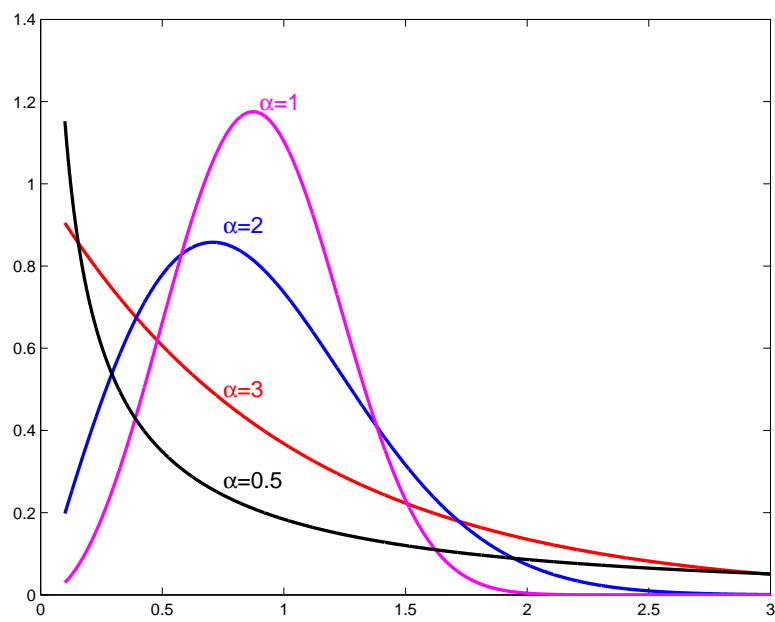
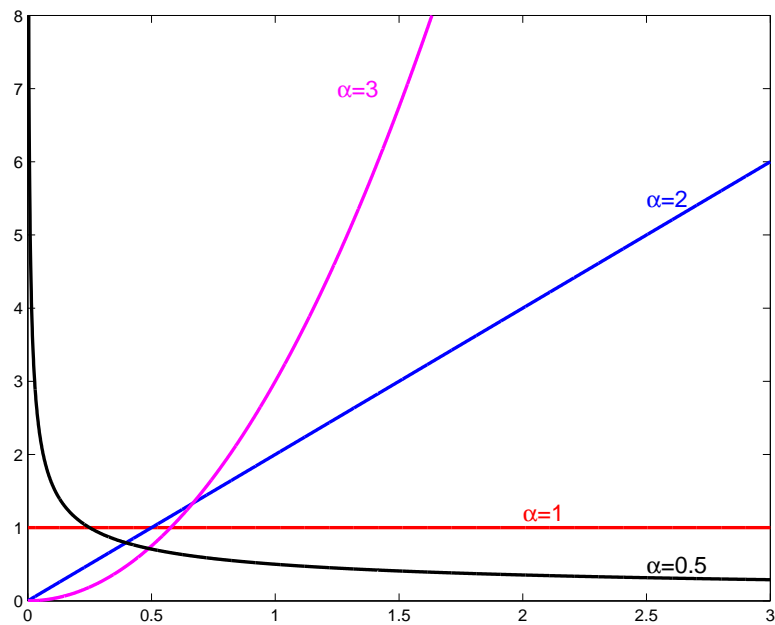
$$h(t) = \frac{f(t)}{R(t)} = \alpha \lambda t^{\alpha-1}, \quad t > 0$$

Weibull distribution provides a great deal of flexibility to model the system when the hazard rate

- increases with time (bearing wear): $\alpha > 1$
- decreases with time (semiconductors): $\alpha < 1$
- constant with time (failures caused by external shocks): $\alpha = 1$

Note x has the Weibull distribution if and only if x^α has the exponential distribution with rate λ .

Plotting of Weibull distribution's hazard rates and density functions with $\beta = 1$ and various α values.



Beta Distribution

- The probability density function for a **beta-type random variable** is given by

$$f(y) = \frac{y^{\alpha-1}(1-y)^{\beta-1}}{B(\alpha, \beta)}, \quad 0 < y < 1$$

for parameters $\alpha, \beta > 0$ where

$$\begin{aligned} B(\alpha, \beta) &= \int_0^1 y^{\alpha-1}(1-y)^{\beta-1} dy = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha + \beta)} \\ &= \text{beta}(a, b) \text{ in MatLab} \end{aligned}$$

- Recall that

$$\Gamma(\alpha) = \int_0^\infty y^{\alpha-1} e^{-y} dy$$

and $\Gamma(\alpha) = (\alpha - 1)!$ when α is a positive integer.

- Mean and Variance are

$$\mu = \frac{\alpha}{\alpha + \beta} \quad \sigma^2 = \frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)}$$

The cumulative distribution function (CDF) of the beta distribution is called **incomplete beta function**.

$$\begin{aligned} F(p) &= \int_0^p \frac{y^{\alpha-1}(1-y)^{\beta-1}}{B(\alpha, \beta)} dy \\ &= \text{betainc}(p, a, b) \text{ in MatLab} \\ &= \sum_{j=\alpha}^n p(j) \text{ if } \alpha \text{ and } \beta \text{ are integers,} \end{aligned}$$

where $p(j)$ is a binomial probability distribution with parameters p and $n = (\alpha + \beta - 1)$.

Discrete Distributions

Bi(p): Bernoulli: Independent zero-one values with same probability p of success.

Bi(n, p): Binomial: Number of successes in a fixed number n of independent trials with the same probability p of success; also number of successes when sampling a population with replacement.

G(n, A, B): Hypergeometric: Number of successes in a fixed number n of samples without replacement from a finite population of A successes, B failures.

MN(n, \vec{p}): Multinomial: Numbers of each of k possible outcomes in a fixed number n of independent trials with the same outcome probability vector $\vec{p} = \{p_i\}$.

Ge(p): **Geometric**: Number of **independent** trials needed for one success.

NB(p, α): **Negative Binomial**: Number of **independent** trials needed for α successes.

Po(λ): **Poisson**: Number of **events** in a **fixed period** if events in different periods are **independent** with constant rate λ .

Un(n): **Uniform**: Finite number n of equally-likely outcomes.

Continuous Distributions

Un(S): **Uniform**: Density function **constant** on some set S .

No(μ, σ^2): **Normal**: Sum or average of **large number** of **independent** quantities.

Ex(β): **Exponential**: Failure time if hazard is **constant** $1/\beta$; time-to-first-event distribution for Poisson with rate $\lambda = 1/\beta$.

We(α, λ): **Weibull**: Failure time if hazard is **power** $\alpha - 1$.

Ga(α, β): **Gamma**: Time-to- α^{th} -event distribution for Poisson with rate $\lambda = 1/\beta$.

Be(α, β): **Beta**: Order statistics: α^{th} largest ($= \beta^{th}$ smallest) of $n = \alpha + \beta - 1$ indep. uniforms.