## Final Examination

## STA 215: Statistical Inference

Monday, 2002 Apr 29, 2:00 – 5:00 pm

This is a closed-book examination; please put all your books and notes on the floor. A normal distribution table, a PMF/PDF handout, and a blank worksheet are attached to the exam. If a questions seems ambiguous or confusing *please* ask me instead of guessing. You may use a calculator but not a PDA, cell phone, or laptop computer.

You must **show** your **work** to get partial credit. Unsupported answers are not acceptable, even if they are correct. It is to your advantage to write your solutions as clearly as possible, since I cannot give credit for solutions I do not understand. Good luck.

Print Name:

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**Problem 1:** Let  $X_j \sim \mathsf{Ge}(\theta)$  be independent geometrically-distributed random variables with probability mass function

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$$f(x \mid \theta) = \theta (1 - \theta)^x, \qquad x = 0, 1, \dots$$

and hence mean  $\mu = \frac{1-\theta}{\theta}$  and variance  $\sigma^2 = \frac{1-\theta}{\theta^2}$  for some  $\theta \in \Theta = (0, 1)$ .

a. (5) Find the "characteristic function"  $\phi(\omega) \equiv \mathsf{E}[e^{i\omega X_j}], \quad \omega \in \mathbb{R}$ 

 $\phi(\omega) =$ and verify the mean given above by computing

 $\phi'(0) = i\mu = \_$ 

b. (5) Find the Fisher Information  $I(\theta)$  for a single observation:

 $I(\theta) =$ \_\_\_\_\_

**Problem 1 (cont'd):** Recall  $X_j \stackrel{\text{ind}}{\sim} \mathsf{Ge}(\theta)$  with p.m.f.

 $f(x \mid \theta) = \theta (1 - \theta)^x, \qquad x = 0, 1, \dots$ 

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c. (5) Find the indicated probability, for integers  $0 \le a < b < \infty$ . Simplify as much as possible.

 $\mathsf{P}[a \le X_j \le b] = \_$ 

d. (5) Find an upper 90% confidence interval  $[L_X, 1]$  for for  $\theta$  from a single observation  $X_1 = 3$ , so that  $0.90 = \mathsf{P}[L_X \le \theta \le 1]$ :

 $L_3 =$  \_\_\_\_\_

**Problem 2:** Again let  $X_j \stackrel{\text{ind}}{\sim} \mathsf{Ge}(\theta)$  with p.m.f.

 $f(x \mid \theta) = \theta (1 - \theta)^x, \qquad x = 0, 1, \dots$ 

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and let  $\vec{x} = (x_1, ..., x_n) \in \mathcal{X} = \mathbb{N}^n$  be a random sample of size n:

a. (5) Find the likelihood function for  $\theta$  upon observing  $\vec{x} \in \mathcal{X}$ .

 $L(\theta \mid \vec{x}) = \_$ 

b. (15) If this is an Exponential Family, find the dimension k, the canonical parameter  $\eta(\theta) \in \mathbb{R}^k$ , the canonical sufficient statistic  $T(\vec{x}) = T_n(\vec{x}) \in \mathbb{R}^k$ , and the normalizing function  $B(\theta) \in \mathbb{R}$  in the Exponential Family representation

$$f(\vec{x} \mid \theta) = h_n(\vec{x}) e^{\eta(\theta) \cdot T_n(\vec{x}) - n B(\theta)}$$

for a random sample of size n (you need not give  $h_n(\vec{x}) = \prod h(x_j)$ ). If it is *not* an exponential family, explain why.

$$k = \underline{\qquad} T_n(\vec{x}) = \underline{\qquad} B(\theta) = \underline{\qquad} \eta(\theta) = \underline{\qquad}$$

**Problem 3:** With the same geometric  $Ge(\theta)$  model as above, and with  $\vec{x} = (x_1, ..., x_n) \in \mathcal{X} = \mathbb{N}^n$  a random sample of size n,

a. (10) Show that the Beta  $\theta \sim \mathsf{Be}(\alpha, \beta)$  distribution with density

$$\pi(\theta) \equiv \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \ \theta^{\alpha - 1} \ (1 - \theta)^{\beta - 1}, \qquad 0 < \theta < 1$$

is conjugate for  $\vec{x}$  and find  $\alpha^*$ ,  $\beta^*$  (which may depend on  $\alpha$ ,  $\beta$ , n, and  $\vec{x}$ ) such that  $\pi(\theta \mid \vec{x}) \sim \mathsf{Be}(\alpha^*, \beta^*)$ :

$$\alpha^{\star} = \underline{\qquad} \qquad \beta^{\star} = \underline{\qquad}$$

b. (10) Find the Jeffreys prior density function  $\pi_J(\theta)$  and give the Jeffreys posterior distribution (upon observing  $\vec{x} \in \mathcal{X}$ ), either by giving its pdf or (better) by giving the distribution's name and parameters:

 $\pi_J(\theta) \propto \_$   $\pi_J(\theta \mid \vec{x}) \sim \_$ 

**Problem 4:** With the same geometric  $Ge(\theta)$  model as above, and again with  $\vec{x} = (x_1, ..., x_n) \in \mathcal{X} = \mathbb{N}^n$  a random sample of size n,

a. (5) Find the MLE  $\hat{\theta}$ :

 $\hat{\theta} =$  \_\_\_\_\_\_

b. (5) Find a Method of Moments (MOM) estimator  $\tilde{\theta}$  (Hint: Solve for  $\theta$  as a function of  $\mu$ , then substitute  $\bar{x}$  for  $\mu$ ):

 $\tilde{\theta} =$  \_\_\_\_\_\_

c. (5) Find the posterior mean  $\bar{\theta}$  for a Bayesian analysis with Beta prior distribution  $\theta \sim \mathsf{Be}(\alpha, \beta)$ :

 $\bar{\theta} =$ 

d. (5) Find the posterior mean  $\bar{\theta}_J$  for a Bayesian analysis with the Jeffreys prior distribution  $\pi_J(\theta)$ :

 $\bar{\theta}_J =$ 

**Problem 5:** With the same geometric  $Ge(\theta)$  model as above, and with n = 10 observations  $\vec{x} = (x_1, ..., x_{10})$  satisfying

$$S_{1} \equiv \sum_{j=1}^{n} x_{j} = 30 | S_{2} \equiv \sum_{j=1}^{n} x_{j}^{2} = 166$$
  

$$S_{3} \equiv \sum_{j=1}^{n} \log(1+x_{j}) = 11.75 | S_{4} \equiv \sum_{j=1}^{n} \frac{1}{1+x_{j}} = 3.79$$
  

$$S_{5} \equiv \min_{1 \le j \le n} x_{j} = 0 | S_{6} \equiv \max_{1 \le j \le n} x_{j} = 10$$

a. (10) Find an expression for the Jeffreys posterior probability of the hypothesis  $H_0$ : [ $\theta \leq 0.10$ ]. Give the answer as your choice of Maple, Matlab, Mathematica, or S-plus commands or as an explicit one-dimensional integral expression:

$$\pi_J(H_0 \mid \vec{x}) = \_$$

b. (10) Suppose we know only  $S_6 = \max x_j = 10$  for n = 10 observations. Perform a significance test of the hypothesis  $H_0 : [\theta \le 0.10]$ against the alternative  $H_1 : [\theta > 0.10]$ . Give the *p*-value to four decimals. Do you accept or reject  $H_0$  at level  $\alpha = .05$ ?

*P* = \_\_\_\_\_

**Problem 6:** Ron thinks that the numbers of times  $X_j$  that pieces of chalk can be used without breaking have a Ge(1/4) distribution, with mean three, while Tom thinks they have a Ge(1/10) distribution, with mean nine. Consider the following statistics, each based on an independent sample  $\vec{x} = \{x_j\}$  of size n:

$$S_{1} = \sum_{j=1}^{n} x_{j} \\ S_{4} = \sum_{j=1}^{n} \frac{1}{1+x_{j}} \begin{vmatrix} S_{2} = \sum_{j=1}^{n} x_{j}^{2} \\ S_{5} = \min_{1 \le j \le n} x_{j} \end{vmatrix} \begin{vmatrix} S_{3} = \sum_{j=1}^{n} \log(1+x_{j}) \\ S_{6} = \max_{1 \le j \le n} x_{j} \end{vmatrix}$$

(20) How should Ron and Tom resolve their dispute? Suggest and defend a **statistic** based on one or more of the statistics  $S_k = S_k(\vec{x}_n)$  above and a procedure for deciding whether  $x_j \sim \text{Ge}(1/4)$  or  $x_j \sim \text{Ge}(1/10)$ . Describe briefly what computation would be required. Please present **both** a (10) Frequentist and a (10) Bayesian solution. Be "fair" to both participants!

Spring 2002

## Problem X: Extra Credit:

a) (+5) (Problem 6, revisited). Use a normal approximation if necessary to complete the analysis (either one) you begain in Problem 6. Decide whether Ron or Tom is correct, and defend your choice, based on the following data for n = 100 observations:

$$S_{1} \equiv \sum_{j=1}^{n} x_{j} = 344 | S_{2} \equiv \sum_{j=1}^{n} x_{j}^{2} = 2196 S_{3} \equiv \sum_{j=1}^{n} \ln(1+x_{j}) = 110.07 | S_{4} \equiv \sum_{j=1}^{n} \frac{1}{1+x_{j}} = 42.58 S_{5} \equiv \min_{1 \le j \le n} t_{j} = 0 | S_{6} \equiv \max_{1 \le j \le n} t_{j} = 13$$

b) (+5) (Problem 1, revisited). Use a normal approximation to give a 90% symmetric confidence interval for  $\theta$ , on the basis of the data above:

 $[L_X, R_X] = [ \_ ___, \_ \_]$ 

c) (+0) What are Ron and Tom's *last* names?

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Name: \_

Extra worksheet, if needed:

$$\Phi(x) = \int_{-\infty}^{x} \frac{1}{\sqrt{2\pi}} e^{-t^{2}/2} dt;$$

Name:

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x	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
.0	.5000	.5040	.5080	.5120	.5160	.5199	.5239	.5279	.5319	.5359
.1	.5398	.5438	.5478	.5517	.5557	.5596	.5636	.5675	.5714	.5753
.2	.5793	.5832	.5871	.5910	.5948	.5987	.6026	.6064	.6103	.6141
.3	.6179	.6217	.6255	.6293	.6331	.6368	.6406	.6443	.6480	.6517
.4	.6554	.6591	.6628	.6664	.6700	.6736	.6772	.6808	.6844	.6879
.5	.6915	.6950	.6985	.7019	.7054	.7088	.7123	.7157	.7190	.7224
.6	.7257	.7291	.7324	.7357	.7389	.7422	.7454	.7486	.7517	.7549
.7	.7580	.7611	.7642	.7673	.7704	.7734	.7764	.7794	.7823	.7852
.8	.7881	.7910	.7939	.7967	.7995	.8023	.8051	.8078	.8106	.8133
.9	.8159	.8186	.8212	.8238	.8264	.8289	.8315	.8340	.8365	.8389
1.0	.8413	.8438	.8461	.8485	.8508	.8531	.8554	.8577	.8599	.8621
1.1	.8643	.8665	.8686	.8708	.8729	.8749	.8770	.8790	.8810	.8830
1.2	.8849	.8869	.8888	.8907	.8925	.8944	.8962	.8980	.8997	.9015
1.3	.9032	.9049	.9066	.9082	.9099	.9115	.9131	.9147	.9162	.9177
1.4	.9192	.9207	.9222	.9236	.9251	.9265	.9279	.9292	.9306	.9319
1.5	.9332	.9345	.9357	.9370	.9382	.9394	.9406	.9418	.9429	.9441
1.6	.9452	.9463	.9474	.9484	.9495	.9505	.9515	.9525	.9535	.9545
1.7	.9554	.9564	.9573	.9582	.9591	.9599	.9608	.9616	.9625	.9633
1.8	.9641	.9649	.9656	.9664	.9671	.9678	.9686	.9693	.9699	.9706
1.9	.9713	.9719	.9726	.9732	.9738	.9744	.9750	.9756	.9761	.9767
2.0	.9772	.9778	.9783	.9788	.9793	.9798	.9803	.9808	.9812	.9817
2.1	.9821	.9826	.9830	.9834	.9838	.9842	.9846	.9850	.9854	.9857
2.2	.9861	.9864	.9868	.9871	.9875	.9878	.9881	.9884	.9887	.9890
2.3	.9893	.9896	.9898	.9901	.9904	.9906	.9909	.9911	.9913	.9916
2.4	.9918	.9920	.9922	.9925	.9927	.9929	.9931	.9932	.9934	.9936
2.5	.9938	.9940	.9941	.9943	.9945	.9946	.9948	.9949	.9951	.9952
2.6	.9953	.9955	.9956	.9957	.9959	.9960	.9961	.9962	.9963	.9964
2.7	.9965	.9966	.9967	.9968	.9969	.9970	.9971	.9972	.9973	.9974
2.8	.9974	.9975	.9976	.9977	.9977	.9978	.9979	.9979	.9980	.9981
2.9	.9981	.9982	.9982	.9983	.9984	.9984	.9985	.9985	.9986	.9986
3.0	.9987	.9987	.9987	.9988	.9988	.9989	.9989	.9989	.9990	.9990
3.1	.9990	.9991	.9991	.9991	.9992	.9992	.9992	.9992	.9993	.9993

 $3.2 \ .9993 \ .9993 \ .9994 \ .9994 \ .9994 \ .9994 \ .9994 \ .9995 \ .9995$ 

3.3 .9995 .9995 .9995 .9996 .9996 .9996 .9996 .9996 .9996

3.4 .9997 .9997 .9997 .9997 .9997 .9997 .9997 .9997 .9997

**Table 5.1**Area  $\Phi(x)$  under the Standard Normal Curve to the left of x.

.9995

.9997

.9998